### Check for updates

#### OPEN ACCESS

EDITED BY Valiollah Palangi, Ege University, Türkiye

REVIEWED BY Zoey Durmic, University of Western Australia, Australia Babak Darabighane, Semnan University, Iran

\*CORRESPONDENCE Xuezhao Sun [xuezhaos@hotmail.com](mailto:xuezhaos@hotmail.com)

RECEIVED 14 August 2024 ACCEPTED 05 November 2024 PUBLISHED 02 December 2024

#### CITATION

Malik MI, Li J, Capucchio MT, Hassan T and Sun X (2024) Effects of distiller's dried grains with solubles on enteric methane emissions in dairy and beef cattle: a meta-analysis. *Front. Vet. Sci.* 11:1480682. [doi: 10.3389/fvets.2024.1480682](https://doi.org/10.3389/fvets.2024.1480682)

#### COPYPICHT

© 2024 Malik, Li, Capucchio, Hassan and Sun. This is an open-access article distributed under the terms of the [Creative Commons](http://creativecommons.org/licenses/by/4.0/)  [Attribution License \(CC BY\).](http://creativecommons.org/licenses/by/4.0/) The use, distribution or reproduction in other forums is permitted, provided the original author(s) and the copyright owner(s) are credited and that the original publication in this journal is cited, in accordance with accepted academic practice. No use, distribution or reproduction is permitted which does not comply with these terms.

# [Effects of distiller's dried grains](https://www.frontiersin.org/articles/10.3389/fvets.2024.1480682/full)  [with solubles on enteric methane](https://www.frontiersin.org/articles/10.3389/fvets.2024.1480682/full)  [emissions in dairy and beef cattle:](https://www.frontiersin.org/articles/10.3389/fvets.2024.1480682/full)  [a meta-analysis](https://www.frontiersin.org/articles/10.3389/fvets.2024.1480682/full)

# Muhammad Irfan Malik<sup>1,2</sup>, Jianping Li<sup>1,3</sup>, Maria Teresa Capucchio<sup>2</sup>, Talal Hassan<sup>2</sup> and Xuezhao Sun<sup>1,3,4</sup>\*

1 The Innovation Centre of Ruminant Precision Nutrition and Smart Farming, Jilin Agricultural Science and Technology University, Jilin, China, <sup>2</sup>Department of Veterinary Sciences, University of Turin, Turin, Italy, 3Jilin Inter-Regional Cooperation Centre for the Scientific and Technological Innovation of Ruminant Precision Nutrition and Smart and Ecological Farming, Jilin, China, <sup>4</sup>Grasslands Research Centre, AgResearch Limited, Palmerston North, New Zealand

Introduction: Distiller's dried grains with solubles (DDGS), a by-product of grain fermentation for ethanol production, are extensively used in livestock feed. Given their nutrient composition, DDGS could potentially influence methane (CH4) emissions, a significant greenhouse gas concern in ruminant production systems. This study utilized a multilevel random-effects meta-analysis to assess the impact of DDGS inclusion in cattle diets on  $CH<sub>4</sub>$  production and yield.

Methods: The literature search was conducted on 23 July 2024. Studies reporting  $CH<sub>4</sub>$  emissions and dry matter intake (DMI) in cattle fed DDGS-based diets were identified, and data extraction was performed. The meta-analysis calculated the mean difference (MD) for DMI and  $CH_4$  yield and the relative mean difference (RMD) for CH<sub>4</sub> production across the selected studies.

**Results:** A total of  $k = 25$  effect sizes from 10 studies were included in the DMI meta-analysis. DDGS had no significant effect on DMI in dairy or beef cattle ( $p$  = 0.770, MD = 0.070, 95% confidence interval [CI] from −0.420 to 0.561). For  $CH<sub>4</sub>$  production,  $k = 24$  effect sizes from 10 studies were analyzed, revealing no significant effect (*p* = 0.759, RMD = −1.045, 95% CI: from −8.025 to 5.935). Similarly, the meta-regression model indicated that the diet's ether extract (EE) had no significant influence (*p* = 0.815, 95% CI from −1.121 to 1.409) on CH<sub>4</sub> production. For CH<sub>4</sub> yield,  $k = 23$  effect sizes from 10 studies were included, with results showing no significant effect (*p* = 0.475, MD = −0.434 g/kg DMI, 95% CI: from −1.673 to 0.805). The regression model for the EE content of the diet also showed no significant impact on CH<sub>4</sub> yield ( $p = 0.311$ , 95% CI: from −0.366 to 0.122).

Discussion: The findings suggest that the inclusion of DDGS does not significantly affect DMI, enteric  $CH_4$  production, or  $CH_4$  yield in cattle. Moreover, the EE content in DDGS-containing diets does not significantly influence CH<sub>4</sub> outcomes. These results indicate that DDGS can be incorporated into cattle diets without exacerbating  $CH_4$  emissions, contributing to sustainable livestock feeding practices.

#### **KEYWORDS**

distillers dried grains with solubles, methane, dairy cows, cattle, meta-analysis

# 1 Introduction

Distiller's dried grains with solubles (DDGS) are widely utilized as a feed ingredient in livestock systems due to their abundant availability and robust nutritional profile. As a by-product of ethanol production through grain fermentation, DDGS is produced when two-thirds of the corn starch is converted to ethanol, leaving behind nutrients concentrated in the stillage [\(1\)](#page-13-0). These nutrients are then recovered and processed into DDGS, resulting in a product with significantly enhanced nutritional content compared to the original grain. Specifically, the fermentation process triples the concentrations of protein, fiber, fat, and phosphorus in DDGS relative to corn, with typical DDGS compositions including 10–30% crude protein (CP), 4–12% fat, 12–36% neutral detergent fiber (NDF), and 0.3–0.9% phosphorus on a dry matter (DM) basis ([2](#page-13-1)). The growing demand for bioethanol has led to increased production of DDGS, making it an increasingly important component of livestock feed. For instance, in 2023 alone, the United States exported 10.8 million metric tons of DDGS [\(3\)](#page-13-2). The widespread adoption of DDGS in feed not only reduces reliance on imported soybean meal and cereals but also contributes to lowering the carbon footprint and enhancing food security ([4\)](#page-13-3). Corn DDGS is particularly well-established in dairy cattle diets, with inclusion levels of up to 300g/kg of diet DM reported without adverse effects on milk yield [\(5\)](#page-13-4). Due to its high protein content, DDGS is primarily used as a protein source for ruminants [\(6\)](#page-13-5). However, there is limited research exploring the impact of DDGS on enteric methane  $(CH<sub>4</sub>)$  emissions in dairy and beef cattle.

Methane emissions are a critical issue in livestock production due to their significant contribution to greenhouse gases and their impact on climate change [\(7\)](#page-13-6). Studies have shown mixed effects of DDGS inclusion on CH<sub>4</sub> emissions. In dairy cows, for instance, DDGS has been shown to reduce enteric CH<sub>4</sub> emissions without negatively impacting feed intake or milk production ([8\)](#page-13-7). However, DDGS inclusion has also been associated with increased manure  $CH<sub>4</sub>$  emissions by up to 15% ([9](#page-13-8)). In beef cattle, high levels of DDGS supplementation (40% on a DM basis) can reduce  $CH_4$ emissions but may simultaneously increase nitrous oxide emissions, highlighting a trade-off between different greenhouse gases [\(10](#page-13-9)).

Several studies have reported reductions in  $CH<sub>4</sub>$  emissions when feeding DDGS to beef ([11](#page-13-10), [12](#page-13-11)) and dairy cattle ([8](#page-13-7)). Hünerberg et al. ([10](#page-13-9)) also reviewed that DDGS consistently resulted in lower  $CH<sub>4</sub>$ emissions. The potential mechanism behind this reduction could be attributed to the higher fat content in DDGS [\(2](#page-13-1)), which can negatively affect ruminal fiber degradation, alter the acetate-topropionate ratio, and reduce protozoa numbers, thereby decreasing  $CH<sub>4</sub>$  production [\(8\)](#page-13-7).

Due to inconsistencies in the literature, with some studies indicating that  $CH<sub>4</sub>$  emissions are unaffected by varying levels of DDGS inclusion ([13\)](#page-13-12), animal nutritionists, policymakers, and farmers struggled to make informed decisions regarding the inclusion of DDGS as a CH<sub>4</sub>-mitigating feed ingredient in dairy and beef ration. Therefore, this meta-analysis was conducted to quantify the effect of DDGS inclusion in the diet on  $CH<sub>4</sub>$ production and yield. Additionally, this study aimed to evaluate whether any reductions observed in  $CH_4$  emissions in dairy or beef cattle-fed DDGS are associated with the fat content of the diet.

# 2 Materials and methods

# 2.1 Search strategy and data processing

The literature search was conducted on 23 July 2024, with no time restrictions applied. We selected two databases, PubMed<sup>[1](#page-1-0)</sup> and Scopus,<sup>2</sup> along with Google Scholar, for our search. For PubMed and Scopus, we used the following keywords: DDGS OR dried distiller's grains with solubles AND methane OR CH<sub>4</sub> AND cattle OR cows OR beef OR steer OR cow OR heifer. For Google Scholar, the keywords were dried distiller's grains with solubles OR DDGS AND methane. The detailed information on the search strategy is presented in the PRISMA flowchart [\(Figure 1](#page-2-0)) [\(14\)](#page-13-13).

Only English-language, peer-reviewed articles were included, and studies reporting enteric CH<sub>4</sub> emissions were selected. Articles that reported CH<sub>4</sub> emissions from *in vitro* studies were excluded. Eligible studies had to involve dairy cattle, heifers, or beef cattle (either steers or heifers) and provide  $CH_4$  emission data. Data for  $CH_4$ emissions (g/day) were considered as  $CH_4$  production, and  $CH_4$  yield was reported as grams per kilogram of dry matter intake (DMI). We extracted data for  $CH_4$  production,  $CH_4$  yield, and DMI, acetate, propionate, and butyrate, along with sample size, standard deviation (SD), or standard error of the mean (SEM). For studies providing variance as SED, we used the RevMan calculator (Version 5.4, [15](#page-13-14)) to compute the SEM. Study characteristics, such as experimental design, diet composition (including % of forage in the diet, % of concentrate in the diet, NDF, EE, CP, starch, % of DDGS in the diet, and types of DDGS: wheat or corn), and types of animals were extracted ([Table 1\)](#page-3-0). For the study by Bernier et al. ([16\)](#page-13-15), where EE of the diet was not reported, it was calculated using the nutritional dynamic system (NDS) Professional Software. Methane production and yield reported in liters were converted to g/day and g/kg DMI, respectively. Liters per day were converted to grams per day, assuming that a mole of CH<sub>4</sub>, weighing 16.0 g, has a volume of 22.4 L  $(17)$ .

# 2.2 Data analysis

The analysis utilized mean difference (MD) as the outcome measure for DMI and CH<sub>4</sub> yield (treatment mean – control mean). Methane production was calculated as relative mean difference  $(RMD) = [(treatment mean - control mean)/(control mean)] \times 100.$ The RMD, a dimensionless variable, was used to account for large variations and is particularly useful for expressing percentage changes in methane production, which is of greater interest to readers ([18](#page-13-17)). The standardized mean difference (SMD) for volatile fatty acids is a statistical technique commonly employed in meta-analyses to compare and synthesize findings from different studies that use varying measurement scales ([19](#page-13-18), [20](#page-13-19)). To calculate the SMD, the mean of the control group is subtracted from the mean of the treatment group, and the result is divided by the pooled standard deviation [\(19\)](#page-13-18). A positive SMD indicates that the treatment group had a higher mean than the control group, while a negative SMD suggests the opposite.

<span id="page-1-0"></span><sup>1</sup> [https://pubmed.ncbi.nlm.nih.gov](https://pubmed.ncbi.nlm.nih.gov/)

<span id="page-1-1"></span><sup>2</sup> <https://www.scopus.com/home.uri>

<span id="page-2-0"></span>

We applied a multilevel random-effects model to address the dependency of effect sizes from the same study. This three-level metaanalytical model is appropriate for handling dependence and heterogeneity among studies. In this model, effect sizes extracted from the same study are considered nested within higher levels, making it suitable for scenarios with varying degrees of variation both within and between studies. The multilevel meta-analysis technique provides more precise effect sizes of treatment effects and helps identify sources of heterogeneity. The variance distribution in the model is as follows: level  $1$  = sampling variance, level  $2$  = effect sizes extracted from the same study, and level 3=variance between studies. By accounting for the varying levels of variation within and between studies, the multilevel meta-analysis technique can provide more precise effect sizes of treatment effects and aid in identifying the sources of heterogeneity [\(21,](#page-13-20) [22](#page-13-21)). We applied an equal effect model for acetate, as the limited number of studies prevented the multilevel model from converging. Convergence refers to the optimizer's ability to identify the best-fitting parameters for the applied model. Successful convergence occurs when the algorithm effectively minimizes or maximizes the target function. Conversely, failure to converge can result from issues, such as poorly specified models, insufficient data, or constraints, that hinder the optimizer's ability to find an optimal solution. Additionally, a subgroup analysis was conducted based on the types of DDGS fed to the animals, with subgroups created for wheat and corn DDGS.

Heterogeneity  $(\tau^2)$  was estimated using the DerSimonian-Laird estimator  $(23)$ , and the  $I<sup>2</sup>$  statistic  $(24)$  $(24)$  was reported and calculated as follows:



#### <span id="page-3-0"></span>TABLE 1 Database characteristics of primary studies included in the meta-analysis.

DDGS, distiller's dried grains with solubles; DOE, design of experiment; EE, ether extract; NDF, neutral detergent fiber; CP, crude protein of the diet; LSD, Latin square design; RCBD, randomized control block design; SF<sub>6</sub>, sulfur hexafluoride.

$$
I^2 = \frac{Q - (k-1)}{Q} \times 100
$$

where Q is the  $\chi^2$  statistic and *k* is the number of studies included in the meta-analysis.

A prediction interval for the true outcomes was also provided ([25](#page-13-24)). The Knapp and Hartung adjustment method was used for the tests and confidence intervals ([26](#page-13-25)). Potential outliers and influential studies were assessed using studentized residuals and Cook's distances [\(27\)](#page-13-26). Meta-regression was performed to test the hypothesis that CH<sub>4</sub> emissions decreased with increased EE contents in the diet. with EE included as a continuous variable in the multilevel randomeffects meta-regression model. Studies with studentized residuals larger than the  $100 \times [1-0.05/(2 \times k)]$  percentile of a standard normal distribution were considered potential outliers (Bonferroni correction with two-sided  $\alpha$  = 0.05 for *k* studies). Studies with Cook's distances larger than the median plus 6 times the interquartile range were deemed influential. Sensitivity analyses assessed the robustness of the results by removing statistical outliers with 95% confidence intervals lying outside the pooled effect size [\(27\)](#page-13-26). Funnel plot asymmetry was checked using the rank correlation test [\(28\)](#page-13-32) and the regression test by Sterne and Egger [\(29\)](#page-13-33), with the standard error of observed outcomes as the predictor. Data analysis was performed using R (version 4.4.0) ([30\)](#page-13-34) and the metafor package (version  $4.6.0)$   $(31)$  $(31)$ .

# 3 Results

# 3.1 Database characteristics

The data analysis included 6 studies on beef cattle and 4 studies on dairy cattle, yielding 11 effect sizes for beef and 14 effect sizes for dairy. The experimental design was a randomized complete block design (RCBD) in two studies and a Latin square design (LSD) in the remaining eight. Two types of DDGS were used: wheat-based DDGS in three studies and corn-based DDGS in seven. Methane quantification methods varied, with the sulfur hexafluoride ( $SF<sub>6</sub>$ ) trace gas technique used in one study, an infrared analyzer in another, indirect calorimetry in two, and a respiratory chamber in five ([Table 1\)](#page-3-0).

On average, the inclusion rate of DDGS was 29.74% for beef cattle and 19.54% for dairy cattle, with concentrate levels at 64.54 and 41.28%, respectively ([Table 2](#page-4-0)). For dairy cattle, forage averaged 58.71% of the diet, with CP at 17.79%, NDF at 36.47%, and starch at 19.41%. In contrast, beef cattle diets had a higher DDGS content (29.74%) and more variable forage levels (35.45%), with CP averaging at 17.49%, NDF lower at 32.69%, and starch higher at 35.18% ([Table 2](#page-4-0)).

A summary of the multilevel random-effects meta-analysis and meta-regression for DMI and methane production and yield is provided in [Table 3](#page-5-0), offering a concise overview of the statistical results.

# 3.2 Dry matter intake

A total of  $k = 25$  effect sizes from 10 studies were included in the analysis. The observed mean differences ranged from −0.92 to 4.60, with 48% of the effect sizes being negative. The multilevel random-effects meta-analysis indicated that DDGS had no significant effect on DMI in dairy or beef cattle  $(p=0.770,$ MD = 0.070, 95% CI: from −0.420 to 0.561). An orchard plot illustrating the observed outcomes and the effect size from the multilevel random-effects model is presented in [Figure 2.](#page-5-1) The subgroup analysis for the different types of DDGS was also non-significant  $(p > 0.05)$  for corn, wheat, or a mixture of both. The effect sizes were as follows: corn DDGS  $(p=0.529,$  $MD = 0.146$ , 95% CI = from  $-0.328$  to 0.612), wheat DDGS (*p* = 0.135, MD =−0.509, 95% CI = from −1.191 to 0.172), and a mixture of corn and wheat DDGS ( $p = 0.189$ , MD = 1.327, 95% CI = from −0.704 to 3.358). The Q-test revealed heterogeneity among the true outcomes  $(Q = 39.56, p = 0.023, \tau^2 = 0.148,$  $I^2$  = 39.34%). Since the heterogeneity ( $I^2$ ) was below 40% and the primary outcome was non-significant, meta-regression was not conducted, as adding covariates would be meaningless. An

<span id="page-4-0"></span>TABLE 2 Descriptive statistics for the dietary characteristics of the studies included in the meta-analysis.



All units are in %, otherwise mentioned. SD, standard deviation; DDGS, distiller's dried grains with solubles; EE, ether extract; NDF, neutral detergent fiber; CP, crude protein.



<span id="page-5-0"></span>TABLE 3 Summary statistics of the multilevel random effects meta-analysis and meta-regression for dry matter intake and methane production and yield.

SE, standard error; DF, degree of freedom (number of effect size); 95% CI, 95% confidence interval; DDGS, distiller's dried grains with solubles; DMI, dry matter intake.

<span id="page-5-1"></span>

dotted vertical line marks the line of no effect.

examination of studentized residuals showed no outliers, with no values exceeding ±3.09. Additionally, Cook's distances indicated that none of the studies were overly influential. The funnel plot of

the effect sizes, shown in [Figure 3,](#page-6-0) indicated potential asymmetry, supported by the rank correlation and Egger's regression tests (*p* = 0.009 and *p* = 0.001, respectively).

<span id="page-6-0"></span>

# 3.3 Methane production

A total of *k*=24 effect sizes from 10 studies were analyzed. Sensitivity analysis identified a treatment with 45% distiller grains as an outlier and overly influential, leading to its exclusion from the final analysis [\(32](#page-13-27)). The observed RMD was −1.045%, with 52% of the effect sizes being positive. Methane production was found to be non-significant (*p*=0.759, RMD=−1.045, 95% CI: from −8.025 to 5.935). An orchard plot showing the observed outcomes and the prediction interval is presented in [Figure 4](#page-7-0). The subgroup analysis of different types of DDGS showed no significant impact on methane production. For corn-based DDGS, the effect size was non-significant (*p*=0.496, RMD=−3.502, 95% CI=from −14.02 to 7.019). Similarly, wheat-based DDGS had no notable effect (*p*=0.937, RMD=−0.243, 95% CI=from −15.17 to 14.68). The combination of corn and wheat DDGS also showed no significant influence (*p*=0.726, RMD=4.347, 95% CI=from −21.14 to 29.84). The regression model indicated that the EE of the diet had no significant effect on  $CH_4$  production (*p*=0.815, 95% CI: from −1.121 to 1.409), with an increase of  $0.144\%$  in CH<sub>4</sub> production per unit increase in EE. The Q-test suggested homogeneity among the true outcomes  $(Q=21.5,$  $p = 0.550$ ,  $\tau^2 = 0$ ,  $I^2 = 0$ ), indicating no heterogeneity. The funnel plot in [Figure 5](#page-7-1) showed no asymmetry, as confirmed by the rank correlation and Egger's regression tests ( $p=0.549$  and  $p=0.469$ , respectively).

## 3.4 Methane yield

A total of *k*=23 effect sizes from 10 studies were included in the analysis. Sensitivity analysis identified two treatments with 30 and 45% distiller grains as outliers, which were subsequently removed from the final analysis [\(32\)](#page-13-27). The observed mean differences for  $CH_4$ yield ranged from −3.90 to 3.63, with 57% of the effect sizes being

negative. Methane yield was found to be non-significant (*p*=0.475, MD=−0.434g/kg DMI, 95% CI: from −1.673 to 0.805). An orchard plot depicting the observed outcomes and the prediction interval is shown in [Figure 6](#page-8-0). The regression model for EE indicated no significant effect on CH4 yield (*p*=0.311, 95% CI: from −0.366 to 0.122), with a  $-0.122$  g/kg DMI increase in CH<sub>4</sub> yield per unit increase in EE. The subgroup analysis for types of DDGS suggests that DDGS types have no significant effect on methane yield. The effect sizes were as follows: corn DDGS (*p*=0.330, MD=−0.835, 95% CI=from −2.580 to 0.910), wheat DDGS (*p*=0.498, MD=0.903, 95% CI=from −1.826 to 3.632), and a mixture of corn and wheat DDGS  $(p=0.882,$ MD=−0.359, 95% CI=from −5.364 to 4.646). The Q-test indicated heterogeneity among the true outcomes ( $Q = 48$ ,  $p = 0.001$ ,  $\tau^2 = 0.55$ ,  $I^2$ =54.16%). The funnel plot in Figure 7 showed no significant asymmetry, supported by the rank correlation and Egger's regression tests ( $p = 0.183$  and  $p = 0.161$ , respectively) ([Figure 8](#page-9-0)).

# 3.5 Acetate

A total of *k*=10 effect sizes from four studies were included in the analysis. The observed SMD for acetate was found to be significant (*p*=0.005, SMD=−0.463, 95% CI: from −0.749 to −0.176). Subgroup analysis by DDGS type indicated that corn DDGS significantly decreased rumen acetate production, with an effect size of  $(p=0.001,$ SMD=−1.048, 95% CI: from−1.526 to −0.570) [\(Table 4](#page-9-1)). In contrast, wheat DDGS showed no significant difference (*p*=0.176, SMD=−0.313, 95% CI: from −0.801 to 0.173). Due to the substantially reduced acetate production, a meta-regression was conducted to identify potential moderators influencing acetate levels. We found that increasing the inclusion level of DDGS in dairy cattle diets significantly reduced acetate (*p*=0.005, SMD=−0.024, 95% CI: from −0.040 to −0.009). Similarly, the inclusion of EE had a significant effect on rumen acetate production (*p*=0.002, SMD=−0.102, 95% CI: −0.159

<span id="page-7-0"></span>

Orchard plot for methane production % (relative mean difference): The overall effect size from a multilevel random-effects meta-analysis of 24 effect sizes is centered on zero, with a 95% confidence interval (CI) spanning the line of no effect (dotted line). The thick black horizontal line represents the prediction interval, while the dotted vertical line marks the line of no effect.

<span id="page-7-1"></span>

to 0.046). The Q-test indicated no significant heterogeneity among the true outcomes (Q=8.28,  $p=0.506$ ,  $\tau^2=0$ ,  $I^2=0$ %). Funnel plot asymmetry was also non-significant, as supported by both the rank correlation and Egger's regression tests  $(p=0.216$  and  $p=0.461$ , respectively).

# 3.6 Butyrate

A total of *k*=10 effect sizes from four studies were analyzed. Rumen butyrate production was found to be non-significant (*p*=0.159, SMD=0.569, 95% CI: from −0.270 to 1.409) ([Figure 9](#page-10-0)).

<span id="page-8-0"></span>

#### FIGURE 6

Orchard plot for methane yield: The overall effect size from a multilevel random-effects meta-analysis of 23 effect sizes is centered on zero, with a 95% confidence interval (CI) spanning the line of no effect (dotted line). The thick black horizontal line represents the prediction interval, while the dotted vertical line marks the line of no effect. SE, standard error.

<span id="page-8-1"></span>

Subgroup analysis by DDGS type showed no significant impact on butyrate production [\(Table 5\)](#page-10-1). For corn-based DDGS, the effect size was non-significant (*p*=0.102, SMD=0.784, 95% CI: from −0.198 to 1.766), and wheat-based DDGS also showed no notable effect (*p*=0.384, SMD=0.389, 95% CI: from −0.586 to 1.365). The

regression model indicated that the EE of the diet had no significant effect on CH<sub>4</sub> production ( $p=0.067$ , SMD=0.131, 95% CI: from −0.011 to 0.274). The Q-test suggested significant heterogeneity among the true outcomes  $(Q=32.33, p=0.0002, \tau^2=0.563,$  $I^2$  = 76.58%). The funnel plot showed asymmetry; the rank correlation

<span id="page-9-0"></span>

FIGURE 8

Orchard plot for acetate: The overall effect size from a multilevel random-effects meta-analysis of 10 effect sizes is centered on zero, with a 95% confidence interval (CI) spanning the line of no effect (dotted line). The thick black horizontal line represents the prediction interval, while the dotted vertical line marks the line of no effect. SE, standard error.

<span id="page-9-1"></span>TABLE 4 Summary statistics for the equal effect meta-analysis and meta-regression for rumen acetate production.



SE, standard error; DF, degree of freedom (number of effect size); 95% CI, 95% confidence interval; DDGS, distiller's dried grains with solubles. All effect sizes are expressed as standardized mean difference (SMD).

test was non-significant ( $p=0.216$ ), while Egger's regression test was significant ( $p = 0.006$ ).

# 3.7 Propionate

A total of *k*=10 effect sizes from four studies were included in the analysis. The observed SMD for acetate was found to be non-significant (*p*=0.508, SMD=−0.125, 95% CI: from −0.538 to −0.286) [\(Figure 10\)](#page-11-0). Subgroup analysis by DDGS type indicated that corn DDGS had no effect on rumen propionate production (*p*=0.622, SMD=0.116, 95% CI: from −0.408 to 0.641). Similarly, wheat DDGS showed no significant difference (*p*=0.139, SMD=−0.382, 95% CI: from −0.920

to 0.155) [\(Table 6](#page-10-2)). EE also had no significant effect on rumen propionate production (*p*=0.913, SMD=−0.004, 95% CI: from −0.087 to 0.079). The Q-test indicated no significant heterogeneity among the true outcomes ( $Q = 9.99$ ,  $p = 0.350$ ,  $\tau^2 = 0.029$ ,  $I^2 = 30.73\%$ ). Funnel plot asymmetry was non-significant, as confirmed by the rank correlation and Egger's regression tests  $(p=1.0$  and  $p=0.417$ , respectively).

# 4 Discussion

Methane emissions from livestock production are a significant contributor to climate change, posing a major challenge among

test



<span id="page-10-1"></span>TABLE 5 Summary statistics for the multilevel random effects meta-analysis and meta-regression for rumen butyrate production.



SE, standard error; DF, degree of freedom (number of effect size); 95% CI, 95% confidence interval; DDGS, distiller's dried grains with solubles. All effect sizes are expressed as standardized mean difference (SMD).

<span id="page-10-0"></span>

<span id="page-10-2"></span>



SE, standard error; DF, degree of freedom (number of effect size); 95% CI, 95% confidence interval; DDGS, distiller's dried grains with solubles. All effect sizes are expressed as standardized mean difference (SMD).

<span id="page-11-0"></span>

atmospheric pollutants. It is well established that  $CH<sub>4</sub>$  emissions are influenced by both the quantity and type of nutrients fermented in the rumen. Typically, CH<sub>4</sub> production increases with DMI, and the specific nutrients fermented play a crucial role in rumen methanogenesis.

Our findings suggest that DDGS have no significant effect on DMI in both dairy and beef cattle, consistent with previous studies in beef cattle ([11](#page-13-10), [33\)](#page-13-28) and dairy cattle [\(16](#page-13-15)). However, this is in contrast to other studies that observed increased DMI when DDGS replaced soybean meal and corn in dairy cattle diets [\(34\)](#page-13-36). The discrepancy may be attributed to our study's focus on literature that specifically evaluates enteric  $CH<sub>4</sub>$  emissions, potentially excluding studies that might have reported positive effects on DMI without examining  $CH_4$  outcomes [\(8\)](#page-13-7). This limitation highlights the need for a more comprehensive analysis that includes a broader range of studies.

Our meta-analysis found that both CH4 yield and production were non-significant, indicating that DDGS does not influence CH<sub>4</sub> emissions in dairy or beef cattle. This finding contrasts with some studies that suggest DDGS can impact these emissions ([11](#page-13-10), [33\)](#page-13-28). For instance, research has shown that feeding DDGS to dairy cows can mitigate enteric CH<sub>4</sub> emissions without negatively affecting intake and milk production [\(8](#page-13-7)). In beef cattle, DDGS inclusion has also been associated with reduced CH<sub>4</sub> emissions  $(11, 33)$  $(11, 33)$  $(11, 33)$  $(11, 33)$  $(11, 33)$ , attributed to the high EE (EE) content of DDGS (12.7% of DM), which can range from 2.0 to 5.1% of DM [\(11\)](#page-13-10).

The reduction in  $CH_4$  production in these studies is often linked to increased EE supply from DDGS, which affects ruminal fiber degradation, the ratio of acetate to propionate, and protozoa numbers. These factors collectively contribute to lower  $CH<sub>4</sub>$  production. However, our meta-regression analysis did not find a significant influence of EE on CH<sub>4</sub> production or yield, contradicting the hypothesis that higher fat content from DDGS would reduce CH4 emissions. This suggests that the relationship between dietary fat content in DDGS-supplemented cows and  $CH<sub>4</sub>$  emissions may be more complex than previously thought and warrants further investigation. Another potential mechanism could be related to sulfur concentration. Buckner et al. [\(35\)](#page-13-37) analyzed 1,200 DDGS (corn =400 and wheat=800) samples from six ethanol processing facilities over 10 months, reporting an average sulfur content of 0.78%. Higher sulfur content may reduce  $CH_4$  emissions by redirecting ruminal  $H_2$ from methanogenesis for  $CH_4$  production [\(36](#page-13-38)) toward hydrogen sulfide (H<sub>2</sub>S) production. Hydrogen sulfide has been shown to inhibit methanogenic archaea directly [\(37\)](#page-13-39). The activity of sulfate-reducing bacteria (SRB) depends on the availability of  $H_2$  and sulfate levels, as these bacteria use sulfate as a terminal electron acceptor in anaerobic respiration [\(38,](#page-13-40) [39\)](#page-13-41). By increasing the sulfate level in the rumen, the capacity of SRB to outcompete methanogens as an  $H_2$  sink is enhanced, which could further reduce  $CH_4$  emissions ([40](#page-13-42)). This suggests that sulfur and sulfate levels in DDGS may influence the microbial dynamics, favoring pathways that reduce  $CH<sub>4</sub>$  production. Our findings suggest that dietary fat is not responsible for  $CH_4$  reduction

in DDGS-supplemented cows. The reduction in  $CH<sub>4</sub>$  emissions reported in some studies might be associated with higher sulfur contents, and variations in sulfur levels due to regional and processing differences could explain the differing results across studies.

The findings of the current meta-analysis suggest that acetate production decreases significantly in cows supplemented with DDGS, which aligns with previous studies in dairy cattle ([8](#page-13-7), [41,](#page-13-43) [42](#page-13-44)). This reduction in acetate may be linked to a decline in ruminal fiber digestion and a decrease in the ruminal degradability of hay as the proportion of DDGS in the diet increases ([8](#page-13-7)). These results are further supported by the meta-regression model, which shows a linear decrease in acetate production with increasing DDGS inclusion. In contrast, butyrate and propionate production were not influenced by the percentage or type of DDGS. The literature shows inconsistencies regarding butyrate and propionate production. Leupp et al. ([43\)](#page-13-45) reported a decrease in acetate molar proportion alongside an increase in propionate, with no effect on butyrate in beef cattle. Meanwhile, Anderson et al. [\(44](#page-13-46)) observed a numerical decrease in acetate and increases in both propionate and butyrate molar proportions in dairy cows fed DDGS diets. There was evidence of publication bias in both DMI and butyrate. This bias may be linked to the unilaterally skewed effect sizes observed in the meta-analysis. Additionally, meta-analyses with a smaller number of studies are more susceptible to publication bias than those with a larger number of studies, which can affect the reliability and representativeness of the findings ([45\)](#page-13-47). The implications of our findings for livestock management and  $CH<sub>4</sub>$  mitigation are significant. Although DDGS may not consistently reduce CH<sub>4</sub> emissions, their diet inclusion offers other nutritional benefits, such as improved nitrogen utilization. However, it is essential to consider the environmental impact of increased nitrogen excretion when evaluating the overall sustainability of DDGS in cattle diets ([8](#page-13-7)). Future research should focus on identifying the conditions under which DDGS can effectively reduce CH<sub>4</sub> emissions and exploring the underlying mechanisms in greater detail. Studies should also investigate the potential relationship between  $CH<sub>4</sub>$  reduction and sulfur content in cattle diets, particularly when supplemented with DDGS. Given that DDGS is rich in both fats and sulfur, it is important to distinguish the individual effects of these components on CH4 emissions and overall cow health. Understanding how sulfur and fats interact within the rumen and their combined impact on CH4 reduction will be crucial in developing more sustainable cattle diets that mitigate environmental impact while ensuring animal health. Additionally, addressing the significant variability observed in the literature could provide clearer insights into the role of DDGS in CH<sub>4</sub> mitigation.

# 5 Conclusion

Our meta-analysis indicates that the inclusion of DDGS has no significant impact on DMI in dairy or beef cattle in studies that evaluated enteric CH<sub>4</sub> emissions. Similarly, DDGS supplementation in cattle diets does not influence enteric CH<sub>4</sub> production or yield. Furthermore, the EE content of diets containing DDGS does not significantly affect CH<sub>4</sub> production or yield in these cattle.

These findings have important implications for livestock producers and policymakers seeking to balance the nutritional benefits of DDGS with the need for effective  $CH<sub>4</sub>$  mitigation strategies. Continued research is essential to refine our understanding of DDGS's role in  $CH<sub>4</sub>$  emissions and to explore alternative dietary strategies that can contribute to more sustainable livestock production systems.

A key limitation of this meta-analysis is the relatively small number of studies available on enteric CH<sub>4</sub> emissions and rumen volatile fatty acids in dairy and beef cattle supplemented with DDGS. This limited dataset may reduce the statistical power and generalizability of the results, as fewer studies can increase variability and the potential for bias. Future research involving a larger body of studies would be valuable in validating and expanding upon the conclusions drawn here.

# Data availability statement

The data analyzed in this study is subject to the following licenses/ restrictions: Inquiries regarding the original data for the meta-analysis can be directed to Muhammad Irfan Malik, [dr.irfan279@gmail.com.](mailto:dr.irfan279@gmail.com)

# Author contributions

MM: Conceptualization, Data curation, Formal analysis, Investigation, Methodology, Project administration, Validation, Visualization, Writing – original draft, Writing – review & editing. JL: Funding acquisition, Writing – review & editing. MC: Writing – review & editing. TH: Writing – review & editing. XS: Funding acquisition, Writing – original draft, Writing – review & editing.

# Funding

The author(s) declare that financial support was received for the research, authorship, and/or publication of this article. This study was funded by the Department of Science and Technology of Jilin Province, China (grant no.: 20220202052NC).

# Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

The author(s) declared that they were an editorial board member of Frontiers, at the time of submission. This had no impact on the peer review process and the final decision.

# Publisher's note

All claims expressed in this article are solely those of the authors and do not necessarily represent those of their affiliated organizations, or those of the publisher, the editors and the reviewers. Any product that may be evaluated in this article, or claim that may be made by its manufacturer, is not guaranteed or endorsed by the publisher.

# References

<span id="page-13-0"></span>1. Belyea RL, Rausch KD, Tumbleson ME. Composition of corn and distillers dried grains with solubles from dry grind ethanol processing. *Bioresour Technol*. (2004) 94:293–8. doi: [10.1016/j.biortech.2004.01.001](https://doi.org/10.1016/j.biortech.2004.01.001)

<span id="page-13-1"></span>2. Klopfenstein TJ, Erickson GE, Bremer VR. Board-invited review: Use of distillers by-products in the beef cattle feeding industry. *J Anim Sci*. (2008) 86:1223–31. doi: [10.2527/jas.2007-0550](https://doi.org/10.2527/jas.2007-0550)

<span id="page-13-2"></span>3. USDA. (2023). Ethanol and distillers' dried grains with solubles transportation dashboard. Available at: [https://agtransport.usda.gov/stories/s/Ethanol-and-DDGS](https://agtransport.usda.gov/stories/s/Ethanol-and-DDGS-exports/dvd5-axr4/)[exports/dvd5-axr4/](https://agtransport.usda.gov/stories/s/Ethanol-and-DDGS-exports/dvd5-axr4/) (Accessed July 23, 2024).

<span id="page-13-3"></span>4. Crawshaw R. Co-product feeds in Europe: Animal feeds derived from industrial processing. Morrisville, NC: Lulu. Com (2019).

<span id="page-13-4"></span>5. Schingoethe D, Kalscheur K, Hippen A, Garcia A. Invited review: The use of distillers products in dairy cattle diets. *J Dairy Sci*. (2009) 92:5802–13. doi: [10.3168/](https://doi.org/10.3168/jds.2009-2549) [jds.2009-2549](https://doi.org/10.3168/jds.2009-2549)

<span id="page-13-5"></span>6. Klopfenstein T, Waller J, Merchen N, Petersen L. Distillers grains as a naturally protected protein for ruminants. *Dist. Feed Res. Council Proc.* (1978) 33:38–45.

<span id="page-13-6"></span>7. Gerber PJ, Hristov AN, Henderson B, Makkar H, Oh J, Lee C, et al. Technical options for the mitigation of direct methane and nitrous oxide emissions from livestock: A review. *Animal*. (2013) 7:220–34. doi: [10.1017/s1751731113000876](https://doi.org/10.1017/s1751731113000876)

<span id="page-13-7"></span>8. Benchaar C, Hassanat F, Gervais R, Chouinard P, Julien C, Petit H, et al. Effects of increasing amounts of corn dried distillers grains with solubles in dairy cow diets on methane production, ruminal fermentation, digestion, N balance, and milk production. *J Dairy Sci*. (2013) 96:2413–27. doi: [10.3168/jds.2012-6037](https://doi.org/10.3168/jds.2012-6037)

<span id="page-13-8"></span>9. Massé DI, Jarret G, Benchaar C, Hassanat F. Effects of adding corn dried distiller grains with solubles (DDGS) to the dairy cow diet and effects of bedding in dairy cow slurry on fugitive methane emissions. *Animals*. (2014) 4:767–78. doi: [10.3390/](https://doi.org/10.3390/ani4040767) [ani4040767](https://doi.org/10.3390/ani4040767)

<span id="page-13-9"></span>10. Hünerberg M, Little SM, Beauchemin KA, McGinn SM, O'Connor D, Okine EK, et al. Feeding high concentrations of corn dried distillers' grains decreases methane, but increases nitrous oxide emissions from beef cattle production. *Agric Syst*. (2014) 127:19–27. doi: [10.1016/j.agsy.2014.01.005](https://doi.org/10.1016/j.agsy.2014.01.005)

<span id="page-13-10"></span>11. McGinn S, Chung Y-H, Beauchemin K, Iwaasa A, Grainger C. Use of corn distillers' dried grains to reduce enteric methane loss from beef cattle. *Can J Anim Sci*. (2009) 89:409–13. doi: [10.4141/CJAS08133](https://doi.org/10.4141/CJAS08133)

<span id="page-13-11"></span>12. Hünerberg M, McGinn S, Beauchemin K, Okine E, Harstad O, McAllister T. Effect of dried distillers grains plus solubles on enteric methane emissions and nitrogen excretion from growing beef cattle. *J Anim Sci*. (2013) 91:2846–57. doi: [10.2527/](https://doi.org/10.2527/jas.2012-5564) [jas.2012-5564](https://doi.org/10.2527/jas.2012-5564)

<span id="page-13-12"></span>13. Garnsworthy PC, Marsden M, Goodman JR, Saunders N. Inclusion of wheat dried distillers' grains with solubles from bioethanol plants in diets for dairy cows. *Animals*. (2021) 11:70. doi: [10.3390/ani11010070](https://doi.org/10.3390/ani11010070)

<span id="page-13-13"></span>14. Liberati A, Altman DG, Tetzlaff J, Mulrow C, Gøtzsche PC, Ioannidis JP, et al. The PRISMA statement for reporting systematic reviews and meta-analyses of studies that evaluate health care interventions: explanation and elaboration. *Ann Intern Med*. (2009) 151:136. doi: [10.7326/0003-4819-151-4-200908180-00136](https://doi.org/10.7326/0003-4819-151-4-200908180-00136)

<span id="page-13-14"></span>15. The Cochrane Collaboration. (2020). The nutritional dynamic system (NDS) Professional Software version 3.9.9.05 (RUM&N Sas, Reggio Emilia, Italy).

<span id="page-13-15"></span>16. Bernier J, Undi M, Plaizier J, Wittenberg K, Donohoe G, Ominski K. Impact of prolonged cold exposure on dry matter intake and enteric methane emissions of beef cows overwintered on low-quality forage diets with and without supplemented wheat and corn dried distillers' grain with solubles. *Can J Anim Sci*. (2012) 92:493–500. doi: [10.4141/cjas2012-040](https://doi.org/10.4141/cjas2012-040)

<span id="page-13-16"></span>17. Appuhamy JRN, Strathe A, Jayasundara S, Wagner-Riddle C, Dijkstra J, France J, et al. Anti-methanogenic effects of Monensin in dairy and beef cattle: A meta-analysis. *J Dairy Sci*. (2013) 96:5161–73. doi: [10.3168/jds.2012-5923](https://doi.org/10.3168/jds.2012-5923)

<span id="page-13-17"></span>18. Feng X, Dijkstra J, Bannink A, Van Gastelen S, France J, Kebreab E. Antimethanogenic effects of nitrate supplementation in cattle: A meta-analysis. *J Dairy Sci*. (2020) 103:11375–85. doi: [10.3168/jds.2020-18541](https://doi.org/10.3168/jds.2020-18541)

<span id="page-13-18"></span>19. Andrade C. Mean difference, standardized mean difference (SMD), and their use in meta-analysis: As simple as it gets. *J Clin Psychiatr*. (2020) 81:11349. doi: [10.4088/](https://doi.org/10.4088/JCP.20f13681) [JCP.20f13681](https://doi.org/10.4088/JCP.20f13681)

<span id="page-13-19"></span>20. Cohen J. Statistical power analysis for the behavioral, Sciences. New York, NY: Academic Press (2013).

<span id="page-13-20"></span>21. Cheung MW-L. Modeling dependent effect sizes with three-level meta-analyses: A structural equation modeling approach. *Psychol Methods*. (2014) 19:211–29. doi: [10.1037/a0032968](https://doi.org/10.1037/a0032968)

<span id="page-13-21"></span>22. Assink M, Wibbelink CJ. Fitting three-level meta-analytic models in R: A step-bystep tutorial. *Quant Methods Psychol*. (2016) 12:154–74. doi: [10.20982/tqmp.](https://doi.org/10.20982/tqmp.12.3.p154) [12.3.p154](https://doi.org/10.20982/tqmp.12.3.p154)

<span id="page-13-22"></span>23. DerSimonian R, Laird N. Meta-analysis in clinical trials. *Control Clin Trials*. (1986) 7:177–88. doi: [10.1016/0197-2456\(86\)90046-2](https://doi.org/10.1016/0197-2456(86)90046-2)

<span id="page-13-23"></span>24. Higgins JP, Thompson SG. Quantifying heterogeneity in a meta-analysis. *Stat Med*. (2002) 21:1539–58. doi: [10.1002/sim.1186](https://doi.org/10.1002/sim.1186)

<span id="page-13-24"></span>25. Riley RD, Higgins JP, Deeks JJ. Interpretation of random effects meta-analyses. *BMJ*. (2011) 342:d549. doi: [10.1136/bmj.d549](https://doi.org/10.1136/bmj.d549)

<span id="page-13-25"></span>26. Knapp G, Hartung J. Improved tests for a random effects meta-regression with a single covariate. *Stat Med*. (2003) 22:2693–710. doi: [10.1002/sim.1482](https://doi.org/10.1002/sim.1482)

<span id="page-13-26"></span>27. Viechtbauer W, Cheung MWL. Outlier and influence diagnostics for metaanalysis. *Res Synth Methods*. (2010) 1:112–25. doi: [10.1002/jrsm.11](https://doi.org/10.1002/jrsm.11)

<span id="page-13-32"></span>28. Begg CB, Mazumdar M. Operating characteristics of a rank correlation test for publication bias. *Biometrics*. (1994) 50:1088–101. doi: [10.2307/2533446](https://doi.org/10.2307/2533446)

<span id="page-13-33"></span>29. Sterne JA, Egger M. Regression methods to detect publication and other bias in meta-analysis. *Pub Bias Meta-Anal Prev Assessment Adjust*. (2005) 22:99–110. doi: [10.1002/0470870168.ch6](https://doi.org/10.1002/0470870168.ch6)

<span id="page-13-34"></span>30. R Core Team. (2020). *R: A language and environment for statistical computing*. Vienna, Austria: R Foundation for Statistical Computing. Available at: [https://www.R](https://www.R-project.org/)[project.org/](https://www.R-project.org/)

<span id="page-13-35"></span>31. Viechtbauer W. Conducting meta-analyses in R with the metafor package. *J Stat Softw*. (2010) 36:1–48. doi: [10.18637/jss.v036.i03](https://doi.org/10.18637/jss.v036.i03)

<span id="page-13-27"></span>32. Hales K, Cole N, MacDonald J. Effects of increasing concentrations of wet distillers grains with solubles in steam-flaked, corn-based diets on energy metabolism, carbonnitrogen balance, and methane emissions of cattle. *J Anim Sci*. (2013) 91:819–28. doi: [10.2527/jas.2012-5418](https://doi.org/10.2527/jas.2012-5418)

<span id="page-13-28"></span>33. Hales K, Cole N, Varel V. Effects of corn processing method and dietary inclusion of corn wet distillers grains with solubles on odor and gas production in cattle manure. *J Anim Sci*. (2012) 90:3988–4000. doi: [10.2527/jas.2011-4679](https://doi.org/10.2527/jas.2011-4679)

<span id="page-13-36"></span>34. Kleinschmit D, Schingoethe D, Kalscheur K, Hippen A. Evaluation of various sources of corn dried distillers grains plus solubles for lactating dairy cattle. *J Dairy Sci*. (2006) 89:4784–94. doi: [10.3168/jds.S0022-0302\(06\)72528-0](https://doi.org/10.3168/jds.S0022-0302(06)72528-0)

<span id="page-13-37"></span>35. Buckner C, Vanness S, Erickson G, Klopfenstein T, Benton J. Nutrient composition and variation among wet and modified distillers grains plus solubles. *J Anim Sci*. (2008) 86:84

<span id="page-13-38"></span>36. Zhao Y, Zhao G. Decreasing ruminal methane production through enhancing the sulfate reduction pathway. *Anim Nutr*. (2022) 9:320–6. doi: [10.1016/j.aninu.2022.01.006](https://doi.org/10.1016/j.aninu.2022.01.006)

<span id="page-13-39"></span>37. Van Zijderveld S, Gerrits W, Apajalahti J, Newbold J, Dijkstra J, Leng R, et al. Nitrate and sulfate: Effective alternative hydrogen sinks for mitigation of ruminal methane production in sheep. *J Dairy Sci*. (2010) 93:5856–66. doi: [10.3168/jds.2010-3281](https://doi.org/10.3168/jds.2010-3281)

<span id="page-13-40"></span>38. Jiang G, Gutierrez O, Sharma KR, Yuan Z. Effects of nitrite concentration and exposure time on sulfide and methane production in sewer systems. *Water Res*. (2010) 44:4241–51. doi: [10.1016/j.watres.2010.05.030](https://doi.org/10.1016/j.watres.2010.05.030)

<span id="page-13-41"></span>39. Singh SB, Lin HC. Hydrogen sulfide in physiology and diseases of the digestive tract. *Microorganisms*. (2015) 3:866–89. doi: [10.3390/microorganisms3040866](https://doi.org/10.3390/microorganisms3040866)

<span id="page-13-42"></span>40. Bryant M, Campbell LL, Reddy C, Crabill M. Growth of desulfovibrio in lactate or ethanol media low in sulfate in association with H2-utilizing methanogenic bacteria. *Appl Environ Microbiol*. (1977) 33:1162–9. doi: [10.1128/aem.33.5.1162-1169.1977](https://doi.org/10.1128/aem.33.5.1162-1169.1977)

<span id="page-13-43"></span>41. Kelzer JM, Kononoff PJ, Gehman A, Tedeschi L, Karges K, Gibson M. Effects of feeding three types of corn-milling coproducts on milk production and ruminal fermentation of lactating Holstein cattle. *J Dairy Sci*. (2009) 92:5120–32. doi: [10.3168/](https://doi.org/10.3168/jds.2009-2208) [jds.2009-2208](https://doi.org/10.3168/jds.2009-2208)

<span id="page-13-44"></span>42. Sasikala-Appukuttan A, Schingoethe D, Hippen A, Kalscheur K, Karges K, Gibson M. The feeding value of corn distillers solubles for lactating dairy cows. *J Dairy Sci*. (2008) 91:279–87. doi: [10.3168/jds.2007-0250](https://doi.org/10.3168/jds.2007-0250)

<span id="page-13-45"></span>43. Leupp J, Lardy G, Karges K, Gibson M, Caton J. Effects of increasing level of corn distillers dried grains with solubles on intake, digestion, and ruminal fermentation in steers fed seventy percent concentrate diets. *J Anim Sci*. (2009) 87:2906–12. doi: [10.2527/](https://doi.org/10.2527/jas.2008-1712) [jas.2008-1712](https://doi.org/10.2527/jas.2008-1712)

<span id="page-13-46"></span>44. Anderson J, Schingoethe D, Kalscheur K, Hippen A. Evaluation of dried and wet distillers grains included at two concentrations in the diets of lactating dairy cows. *J Dairy Sci*. (2006) 89:3133–42. doi: [10.3168/jds.S0022-0302\(06\)72587-5](https://doi.org/10.3168/jds.S0022-0302(06)72587-5)

<span id="page-13-47"></span>45. Borenstein M, Hedges LV, Higgins JP, Rothstein HR. Introduction to Meta-Analysis. New York, NY: John Wiley & Sons (2021). 276–292.

<span id="page-13-29"></span>46. Hünerberg M, McGinn S, Beauchemin KA, Okine E, Harstad OM, McAllister TA. Effect of dried distillers' grains with solubles on enteric methane emissions and nitrogen excretion from finishing beef cattle. *Can J Anim Sci*. (2013) 93:373–85. doi: [10.4141/](https://doi.org/10.4141/cjas2012-151) [cjas2012-151](https://doi.org/10.4141/cjas2012-151)

<span id="page-13-30"></span>47. Castillo-Lopez E, Jenkins C, Aluthge N, Tom W, Kononoff P, Fernando SC. The effect of regular or reduced-fat distillers grains with solubles on rumen methanogenesis and the rumen bacterial community. *J Appl Microbiol*. (2017) 123:1381–95. doi: [10.1111/jam.13583](https://doi.org/10.1111/jam.13583)

<span id="page-13-31"></span>48. Judy J, Bachman G, Brown-Brandl T, Fernando SC, Hales K, Miller PS, et al. Reducing methane production with corn oil and calcium sulfate: Responses on wholeanimal energy and nitrogen balance in dairy cattle. *J Dairy Sci*. (2019) 102:2054–67. doi: [10.3168/jds.2018-14567](https://doi.org/10.3168/jds.2018-14567)