



# Value of imaging examinations in diagnosing lumbar disc herniation: A systematic review and meta-analysis

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**Purpose:** To systematically review the clinical value of three imaging examinations (Magnetic Resonance Imaging, Computed Tomography, and myelography) in the diagnosis of Lumbar Disc Herniation.

**Methods:** Databases including PubMed, Embase, The Cochrane Library, Web of Science, CBM, CNKI, WanFang Data, and VIP were electronically searched to collect relevant studies on three imaging examinations in the diagnosis of Lumbar Disc Herniation from inception to July 1, 2021. Two reviewers using the Quality Assessment of Diagnostic Accuracy Studies-2 tool independently screened the literature, extracted the data, and assessed the risk of bias of included studies. Then, meta-analysis was performed by using Meta-DiSc 1.4 software and Stata 15.0 software.

**Results:** A total of 38 studies from 19 articles were included, involving 1,875 patients. The results showed that the pooled Sensitivity, pooled Specificity, pooled Positive Likelihood Ratio, pooled Negative Likelihood Ratio, pooled Diagnostic Odds Ratio, Area Under the Curve of Summary Receiver Operating Characteristic, and Q\* were 0.89 (95%CI: 0.87–0.91), 0.83 (95%CI: 0.78–0.87), 4.57 (95%CI: 2.95–7.08), 0.14 (95%CI: 0.09–0.22), 39.80 (95%CI: 18.35–86.32), 0.934, and 0.870, respectively, for Magnetic Resonance Imaging. The pooled Sensitivity, pooled Specificity, pooled Positive Likelihood Ratio, pooled Negative Likelihood Ratio, pooled Diagnostic Odds Ratio, Area Under the Curve of Summary Receiver Operating Characteristic, and Q\* were 0.82 (95%CI: 0.79–0.85), 0.78 (95%CI: 0.73–0.82), 3.54 (95%CI: 2.86–4.39), 0.19 (95%CI: 0.12–0.30), 20.47 (95%CI: 10.31–40.65), 0.835, and 0.792, respectively, for Computed Tomography. The pooled Sensitivity, pooled Specificity, pooled Positive Likelihood Ratio, pooled Negative Likelihood Ratio, pooled Diagnostic Odds Ratio, Area Under the Curve of Summary Receiver Operating Characteristic, and Q\* were 0.79 (95%CI: 0.75–0.82), 0.75 (95%CI: 0.70–0.80), 2.94 (95%CI: 2.43–3.56), 0.29 (95%CI: 0.21–0.42), 9.59 (95%CI: 7.05–13.04), 0.834, and 0.767 respectively, for myelography.

**Conclusion:** Three imaging examinations had high diagnostic value. In addition, compared with myelography, Magnetic Resonance Imaging had a higher diagnostic value.

### KEYWORDS

Lumbar Disc Herniation, Magnetic Resonance Imaging, Computed Tomography, myelography, meta-analysis

## Introduction

Lumbar Disc Herniation (LDH) is defined as a localized displacement of disc material (nucleus, cartilage, fragmented apophyseal bone, annular tissue, or any combination thereof) (1). When the displaced disc material compresses the local nerve, there will be a series of symptoms such as low back pain, radiating pain on one or both sides of the lower limb, numbness, intermittent claudication, difficulty walking, and even muscle atrophy, which will seriously affect the daily life of patients. It is estimated that approximately 2%–3% of the population may be affected, with a prevalence of 4.8% among men and 2.5% among women older than 35 (2). Approximately 95% of herniated discs occur at the low lumbar spine (L4/5 and L5/S1 level) in people aged 25 to 55 (3). With the rapid development of the economy and society, people's way of life and work has changed. Patients with LDH show an increasing trend and tend to be younger. The treatment cycle of the disease is long, and the cost is high, resulting in a heavy burden on the family and society. Timely and accurate diagnosis plays an important role in the later treatment and rehabilitation of LDH.

Imaging examinations are often used in patients with low back pain and/or leg pain to assess the compression of a nerve root caused by disc herniation or spinal and cauda equina syndrome (4–7). Furthermore, imaging examinations can also be used to identify the clinical symptoms of affected disc levels before surgery (8). However, there are different reports on the accuracy of imaging examinations in the diagnosis of LDH, and there is a lack of multicenter and large-scale research. The purpose of this meta-analysis is to systematically review the published literature on the diagnosis of LDH by Magnetic Resonance Imaging (MRI), Computed Tomography (CT), and myelography through meta-analysis so as to provide a basis for clarifying the accuracy of imaging examination in the diagnosis of LDH.

## Materials and methods

This review followed the meta-analysis of Standards for Reporting of Diagnostic Accuracy Studies (STARD) 2015 guidelines (9) and was reported in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) 2020 (10). The review was registered in International Prospective Register of Systematic Reviews (PROSPERO) (registration number CRD42021269796).

## Literature search strategy

All relevant literature from eight databases, namely, PubMed, Embase, The Cochrane Library, Web of Science,

CBM, CNKI, WanFang Data, and VIP, were explored from inception to July 1, 2021. To minimize the missing literature, the references in the included studies were traced to [supplement data](#).

## Eligibility criteria

The inclusion criteria were as follows: participants with suspected LDH who underwent MRI, CT, or myelography before reference standard examinations (not limited by age, race, and nationality); prospective or retrospective study design; direct or indirect availability of the results—True Positive (TP), False Positive (FP), False Negative (FN), and True Negative (TN). The exclusion criteria were as follows: duplicate articles; reviews, conference abstracts, animal studies, and case reports; studies that did not describe specific diagnostic reference standards of LDH; studies with unclear measurement indicators, inappropriate statistical methods adopted, or important outcome indicators not fully explained; studies that were unable to obtain full text directly or indirectly; and studies not in English or Chinese.

## Literature screening, data extraction

Two reviewers independently screened the literature and extracted and cross-checked the data. In case of disagreements, a third party was consulted to assist in the judgment. During literature screening, first, the title and abstract were read. Then, after the exclusion of irrelevant literature, the full text was read to determine whether it was finally included. Data extraction mainly included the basic characteristics of the included studies, such as author, publication year, country, design type, sample size, diagnostic method, and reference standard. Results considered, such as TP, FP, FN, and TN.

## Risk of bias assessment of included studies

Two reviewers independently used the Quality Assessment of Diagnostic Accuracy Studies-2 (QUADAS-2) tool to evaluate the risk of bias of included studies (11). In case of disagreements, a third party was consulted to assist in the judgment. Each item was assessed as “yes” (low bias or good suitability), “no” (high bias or poor suitability), or “unclear” (lack of relevant information or uncertainty for the bias).

## Outcome indicators

These include pooled Sensitivity (Sen), pooled Specificity (Spe), pooled Positive Likelihood Ratio (+LR),

pooled Negative Likelihood Ratio ( $-LR$ ), pooled Diagnostic Odds Ratio (DOR), Summary Receiver Operating Characteristic (SROC), Area Under the Curve (AUC) of SROC, and  $Q^*$ .

## Statistical analysis

Review Manager 5.3 software was used to evaluate the risk of bias of the included studies, Meta-DiSc 1.4 was used for meta-analysis, Stata 15.0 was used for sensitivity analysis, and publication bias test. First, Spearman's correlation coefficient between the logarithm of Sen and the logarithm of  $(1-Spe)$  was calculated to analyze the heterogeneity caused by the threshold effect: if the  $P$ -value of Spearman's correlation coefficient was less than 0.05, it indicated that there was heterogeneity caused by threshold effect. It was necessary to conduct a meta-analysis after adjusting and combining the confounding factors between studies and considering the interaction between Sen and Spe. If the  $P$ -value of Spearman's correlation coefficient was more than 0.05, it indicated that there was no heterogeneity caused by the threshold effect. The next step was to test the heterogeneity of the no-threshold effect (12, 13).  $I^2$  was used to analyze the no-threshold heterogeneity (14): if  $I^2 < 50\%$ , it indicated that there was little heterogeneity between studies, and the fixed-effects model was used for pooling. If  $I^2 \geq 50\%$ , it indicated that there was great heterogeneity between studies. Meta-regression was used to find the potential factors causing heterogeneity (15), and then subgroup analysis was performed (16). If the source of heterogeneity could not be found, the random effects model was used for pooling. According to the corresponding model, calculated  $Sen_{(pooled)}$ ,  $Spe_{(pooled)}$ ,  $+LR_{(pooled)}$ ,  $-LR_{(pooled)}$ , and  $DOR_{(pooled)}$ ; draw SROC; and calculated AUC and  $Q^*$  of the included studies. Among them, the higher the Sen, Spe, DOR,  $Q^*$ , and  $+LR$ , the lower the  $-LR$ , and the closer was AUC to 1, indicating the higher value of imaging examinations in diagnosing LDH; otherwise, the value was lower (13, 17–19). The stability of the research results was analyzed by sensitivity analysis. The included literature was excluded one by one, and then meta-analysis was performed again. The results were compared with those before exclusion. If the change was small, it indicated that the stability of the included literature was good and the results were credible. If there were significant changes, it indicated that the results were not credible (20). Finally, the publication bias was tested by Deek's funnel plot (21). If the  $P$ -value of the slope coefficient was more than 0.05, it indicated that there was no publication bias. On the contrary, it indicated that there was publication bias.

## Results

### The result of the literature search

A total of 8,034 relevant articles were obtained. After the layer-by-layer screening, 19 articles (22–40) were finally included. The process of the literature search is shown in Figure 1 and Supplementary Method S1. Detailed information on the included literature is shown in Table 1.

### Risk of bias assessment of the included studies

The results of the QUADAS-2 tool showed that the implementation of diagnostic tests and the rationality of the reference standard included in this meta-analysis were of good quality, suggesting that the included studies had high quality and were less likely to cause selection bias (41). However, we were not satisfied with the reference standard. The main reason is that most studies regarded surgical findings as the reference standard, and all patients need imaging examinations before surgery. The detailed information is shown in Figures 2, 3 and Supplementary Table S1.

### Meta-analysis of MRI

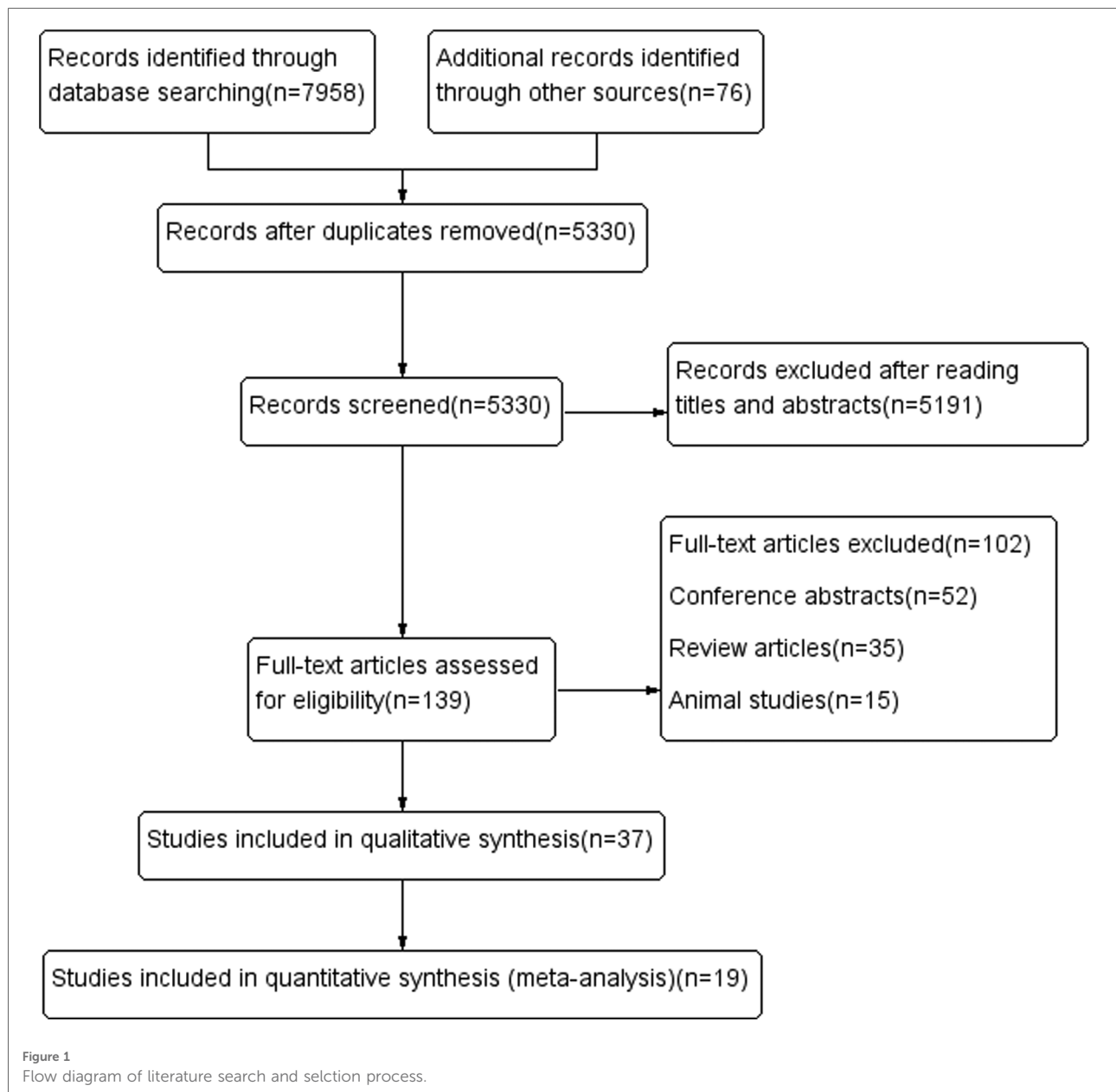
Thirteen articles with 19 studies were included (Table 2).

### Heterogeneity test

By Spearman's correlation analysis, the correlation coefficient between the logarithm of Sen and the logarithm of  $(1-Spe)$  was  $-0.394$ ,  $P = 0.095$ , indicating that there was no threshold effect in this meta-analysis. The heterogeneity test results showed that the heterogeneity of Sen ( $\chi^2 = 77.02$ ,  $P = 0.000$ ,  $I^2 = 76.6\%$ ), Spe ( $\chi^2 = 52.88$ ,  $P = 0.000$ ,  $I^2 = 66.0\%$ ),  $+LR$  (Cochran-Q = 52.29,  $P = 0.000$ ,  $I^2 = 65.6\%$ ),  $-LR$  (Cochran-Q = 69.73,  $P = 0.000$ ,  $I^2 = 74.2\%$ ), and DOR (Cochran-Q = 57.49,  $P = 0.001$ ,  $I^2 = 68.7\%$ ) among the studies were high (Figure 4). The cause of heterogeneity was not found through meta-regression or subgroup analysis. Therefore, the effect sizes were pooled using a random effects model.

### Evaluation index of diagnostic test

The effect sizes of  $Sen_{(pooled)}$ ,  $Spe_{(pooled)}$ ,  $+LR_{(pooled)}$ ,  $-LR_{(pooled)}$ ,  $DOR_{(pooled)}$ , AUC of SROC, and  $Q^*$  were 0.89 (95%CI: 0.87–0.91), 0.83 (95%CI: 0.78–0.87), 4.57 (95%CI:



2.95–7.08), 0.14 (95%CI: 0.09–0.22), 39.80 (95%CI: 18.35–86.32), 0.934, and 0.870, respectively (**Figure 4**).

### Sensitivity analysis and publication bias analysis

After the exclusion of individual studies one by one, the remaining studies were pooled and analyzed again. The results showed that each excluded study had a minor impact on the amount of pooling effect, indicating that the

results of this meta-analysis were stable and reliable (**Figure 5**). Funnel plot was drawn with the inverse of the square root of the effective sample size (ESS) as the ordinate and DOR as the abscissa. The results the slope coefficient was 1.00, suggesting that there was no publication bias (**Figure 6**).

### Meta-analysis of CT

Ten articles with 10 studies were included (**Table 3**).



TABLE 1 Characteristics of included studies.

Study	Country	Type of study	Reported outcome	Reference standard	Age (range)	Sample size (male/female)
Aejmelaeus 1984 (22)	Finland	Prospective	Myelography	Surgical findings	43.8 (14–82)	200 (109/91)
Bernard 1994 (23)	America	Prospective	MRI	Surgical findings	50 (23–74)	33 (20/13)
Birney 1992 (24)	America	Retrospective	MRI	Surgical findings	39 (20–71)	90 (48/49)
Bischoff 1993 (25)	America	Retrospective	Myelography;MRI	Surgical findings	20–79	57 (29/28)
Chawalparit 2006 (26)	Thailand	Prospective	MRI	Surgical findings	42.9 (21–60)	123 (61/62)
Firooznia 1984 (27)	America	Retrospective	CT	Surgical findings	49 (19–76)	100 (61/39)
Forristall 1988 (28)	America	Prospective	CT;MRI	Surgical findings	45 (22–74)	32 (25/7)
Gillström 1986 (29)	Sweden	Prospective	Myelography;CT	Surgical findings	23–74	37 (22/15)
Haughton 1982 (30)	America	Prospective	Myelography;CT	Surgical findings	13–72	107 (58/49)
Huang 2020 (31)	China	Prospective	MRI	Clinical data	51.9 (37–65)	161 (93/68)
Jackson 1989 (32)	America	Prospective	Myelography;CT	Surgical findings	42.7 (21–76)	124 (87/37)
Jackson 1989 (33)	America	Prospective	Myelography;CT;MRI	Surgical findings	39.6 (18–70)	59 (33/26)
Janssen 1994 (34)	America	Retrospective	Myelography;MRI	Surgical findings	46 (27–73)	60 (23/37)
Kamal 2009 (35)	Bangladesh	Prospective	MRI	Surgical findings	NA	40 (28/12)
Masaryk 1987 (36)	America	Prospective	MRI	Surgical findings	26–66	20 (13/7)
Modic 1986 (37)	America	Prospective	Myelography;CT;MRI	Surgical findings	46 (19–73)	48 (NA)
Mullin 2000 (38)	America	Prospective	MRI	Surgical findings	NA	28 (NA)
Schipper 1987 (39)	The Netherlands	Prospective	Myelography; CT	Surgical findings	43 (NA)	461 (NA)
Thornbury 1993 (40)	America	Retrospective	CT;MRI	Surgical findings	39.6 (21–72)	95 (61/34)

### Heterogeneity test

By Spearman’s correlation analysis, the correlation coefficient between the logarithm of Sen and the logarithm of (1–Spe) was 0.539,  $P = 0.108$ , indicating that there was no threshold effect in this meta-analysis. The heterogeneity test results showed that the heterogeneity of Sen ( $\chi^2 = 103.22$ ,  $P = 0.000$ ,  $I^2 = 91.3\%$ ), –LR (Cochran- $Q = 51.60$ ,  $P = 0.000$ ,  $I^2 = 82.6\%$ ), and DOR (Cochran- $Q = 24.59$ ,  $P = 0.004$ ,  $I^2 = 63.4\%$ ) among the studies were high. The cause of heterogeneity was not found by meta-regression or subgroup analysis, so the random effects model were used for pooling. The heterogeneity of Spe ( $\chi^2 = 8.92$ ,  $P = 0.444$ ,  $I^2 = 0.0\%$ ) and

+LR (Cochran- $Q = 5.84$ ,  $P = 0.756$ ,  $I^2 = 0.0\%$ ) among the studies were low (Figure 7). Therefore, the effect sizes were pooled using a fixed effects model.

### Evaluation index of diagnostic test

The effect sizes of  $Sen_{(pooled)}$ ,  $Spe_{(pooled)}$ ,  $+LR_{(pooled)}$ ,  $-LR_{(pooled)}$ ,  $DOR_{(pooled)}$ , AUC of SROC, and  $Q^*$  were 0.82 (95%CI: 0.79–0.85), 0.78 (95%CI: 0.73–0.82), 3.54 (95%CI: 2.86–4.39), 0.19 (95%CI: 0.12–0.30), 20.47 (95%CI: 10.31–40.65), 0.835, and 0.792, respectively (Figure 7).

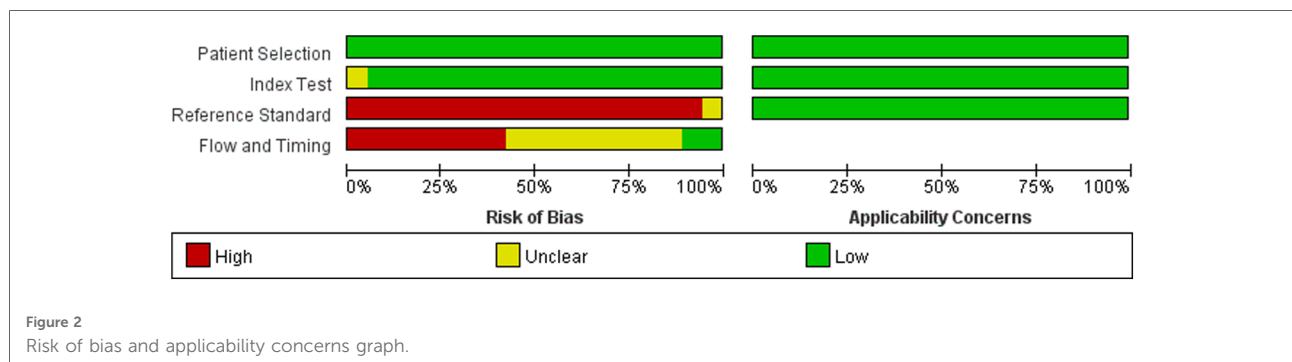


Figure 2 Risk of bias and applicability concerns graph.

	Risk of Bias				Applicability Concerns		
	Patient Selection	Index Test	Reference Standard	Flow and Timing	Patient Selection	Index Test	Reference Standard
Aejmelaeus 1984	+	+	-	-	+	+	+
Bernard 1994	+	+	-	?	+	+	+
Birney 1992	+	+	-	-	+	+	+
Bischoff 1993	+	+	-	-	+	+	+
Chawalparit 2006	+	+	-	-	+	+	+
Firooznia 1984	+	+	-	?	+	+	+
Forristall 1988	+	+	-	-	+	+	+
Gillström 1986	+	+	-	?	+	+	+
Haughton 1982	+	+	-	-	+	+	+
Huang 2020	+	?	?	?	+	+	+
Jackson 1989	+	+	-	+	+	+	+
Jackson 1989	+	+	-	?	+	+	+
Janssen 1994	+	+	-	?	+	+	+
Kamal 2009	+	+	-	+	+	+	+
Masaryk 1988	+	+	-	?	+	+	+
Modic 1986	+	+	-	?	+	+	+
Mullin 2000	+	+	-	?	+	+	+
Schipper 1987	+	+	-	-	+	+	+
Thornbury 1993	+	+	-	-	+	+	+

 High	 Unclear	 Low
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Figure 3  
Risk of bias and applicability concerns summary.

TABLE 2 Characteristics of MRI diagnostic tests.

Study name	TP	FP	FN	TN
Bernard 1994 (23)	33	9	13	11
Birney 1992 (24)	70	0	5	1
Bischoff 1993 (25)	25	10	10	27
Chawalparit 2006 (26)	19	2	4	8
Chawalparit 2006 (26)	19	3	4	7
Forrestall 1988 (28)	22	0	2	7
Huang 2020 (31)	158	3	3	32
Jackson 1989 (33)	38	8	21	53
Janssen 1994 (34)	65	1	3	33
Kamal 2009 (35)	33	2	2	3
Masaryk 1987 (36)	8	2	1	9
Modic 1986 (37)	28	7	4	23
Mullin 2000 (38)	20	0	1	10
Mullin 2000 (38)	20	0	1	10
Mullin 2000 (38)	20	1	1	9
Mullin 2000 (38)	20	1	1	9
Mullin 2000 (38)	19	0	2	10
Mullin 2000 (38)	18	0	3	10
Thornbury 1993 (40)	68	10	6	11

### Sensitivity analysis and publication bias analysis

After the exclusion of individual studies one by one, the remaining studies were pooled and analyzed again. The results showed that each excluded study had a minor impact on the amount of pooling effect, indicating that the results of this meta-analysis were stable and reliable (Figure 8). Funnel plot was drawn with 1/root (ESS) as the ordinate and DOR as the abscissa. The results showed that the P-value of the slope coefficient was 0.31, suggesting that there was no publication bias (Figure 9).

### Meta-analysis of myelography

Nine articles with nine studies were included (Table 4).

### Heterogeneity test

By Spearman's correlation analysis, the correlation coefficient between the logarithm of Sen and the logarithm of (1-Spe) was 0.583,  $P=0.099$ , indicating that there was no threshold effect in this meta-analysis. The heterogeneity test results showed that the heterogeneity of Sen ( $\chi^2=52.12$ ,  $P=0.000$ ,  $I^2=84.7\%$ ), Spe ( $\chi^2=18.98$ ,  $P=0.015$ ,  $I^2=57.9\%$ ), and

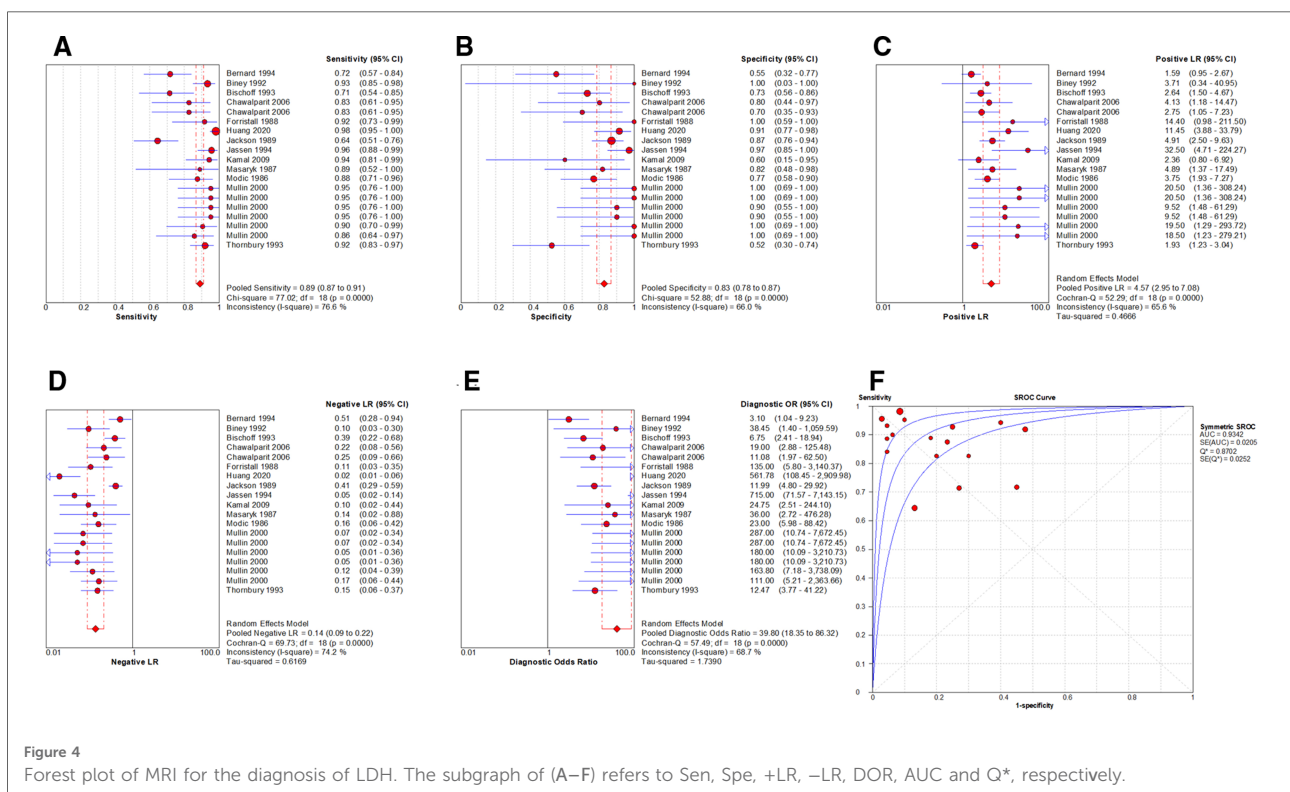


Figure 4 Forest plot of MRI for the diagnosis of LDH. The subgraph of (A–F) refers to Sen, Spe, +LR, –LR, DOR, AUC and Q\*, respectively.

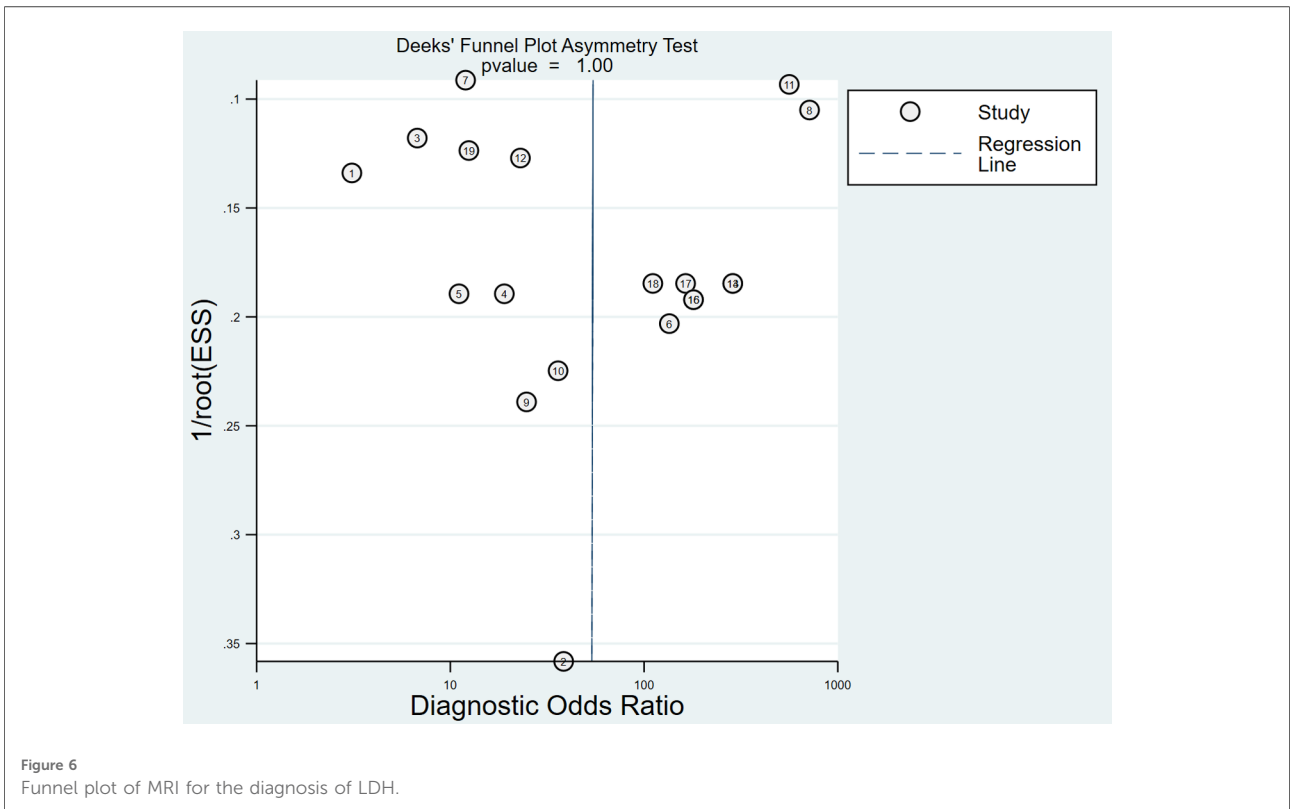
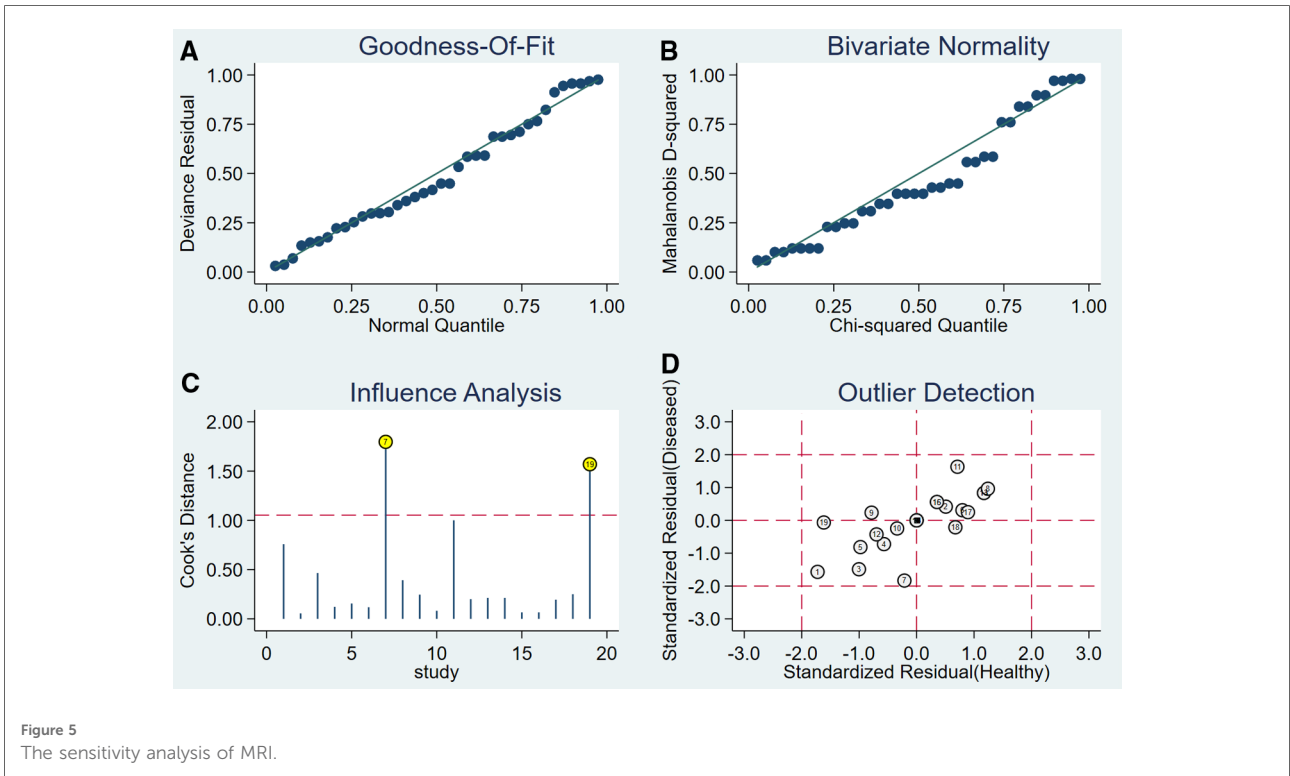


TABLE 3 Characteristics of CT diagnostic tests.

Study name	TP	FP	FN	TN
Firooznia 1984 (27)	97	4	8	7
Forristall 1988 (28)	20	2	4	5
Gillström 1986 (29)	28	1	0	2
Haughton 1982 (30)	29	8	1	17
Huang 2020 (31)	156	5	5	30
Jackson 1989 (32)	89	25	36	81
Jackson 1989 (33)	35	8	24	53
Modic 1986 (37)	25	5	4	19
Schipper 1987 (39)	140	8	57	30
Thornbury 1993 (40)	17	5	1	9

-LR (Cochran-Q = 33.82,  $P = 0.000$ ,  $I^2 = 76.3\%$ ) among the studies were high. The cause of heterogeneity was not found by meta-regression or subgroup analysis, so the effect sizes were pooled using a random effects model. The heterogeneity of +LR (Cochran-Q = 10.65,  $P = 0.222$ ,  $I^2 = 24.9\%$ ) and DOR (Cochran-Q = 13.11,  $P = 0.108$ ,  $I^2 = 39.0\%$ ) among the studies were low (Figure 10). Therefore, the effect sizes were pooled using a fixed effects model.

### Evaluation index of diagnostic test

The effect sizes of  $Sen_{(pooled)}$ ,  $Spe_{(pooled)}$ ,  $+LR_{(pooled)}$ ,  $-LR_{(pooled)}$ ,  $DOR_{(pooled)}$ , AUC of SROC, and  $Q^*$  were 0.79

(95%CI: 0.75–0.82), 0.75 (95%CI: 0.70–0.80), 2.94 (95%CI: 2.43–3.56), 0.29 (95%CI: 0.21–0.42), 9.59 (95%CI: 7.05–13.04), 0.834, and 0.767, respectively (Figure 10).

### Sensitivity analysis and publication bias analysis

After the exclusion of individual studies one by one, the remaining studies were pooled and analyzed again. The results showed that each excluded study had a minor impact on the amount of pooling effect, indicating that the results of this meta-analysis were stable and reliable (Figure 11). Funnel plot was drawn with  $1/\sqrt{ESS}$  as the ordinate and DOR as the abscissa. The results showed that the  $P$ -value of the slope coefficient was 0.30, suggesting that there was no publication bias (Figure 12).

### Discussion

Imaging examinations have important clinical significance for the diagnosis and treatment of LDH. They can provide not only a basis for diagnosis but also a basis for choosing conservative treatment or surgical treatment and surgical methods (42) so as to improve the treatment level. At present, the commonly used imaging examinations include MRI, CT, myelography, and x-ray. MRI is the most established of the imaging examinations, as it has the advantage of not using

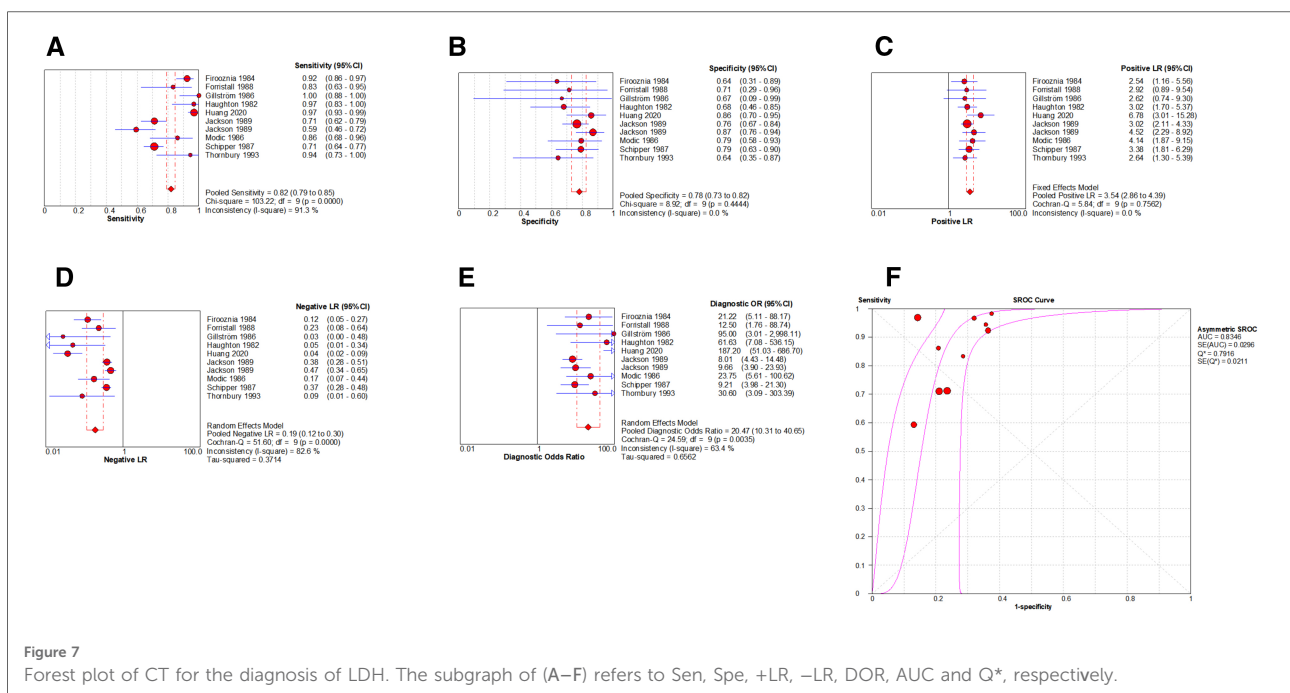


Figure 7 Forest plot of CT for the diagnosis of LDH. The subgraph of (A–F) refers to Sen, Spe, +LR, -LR, DOR, AUC and  $Q^*$ , respectively.



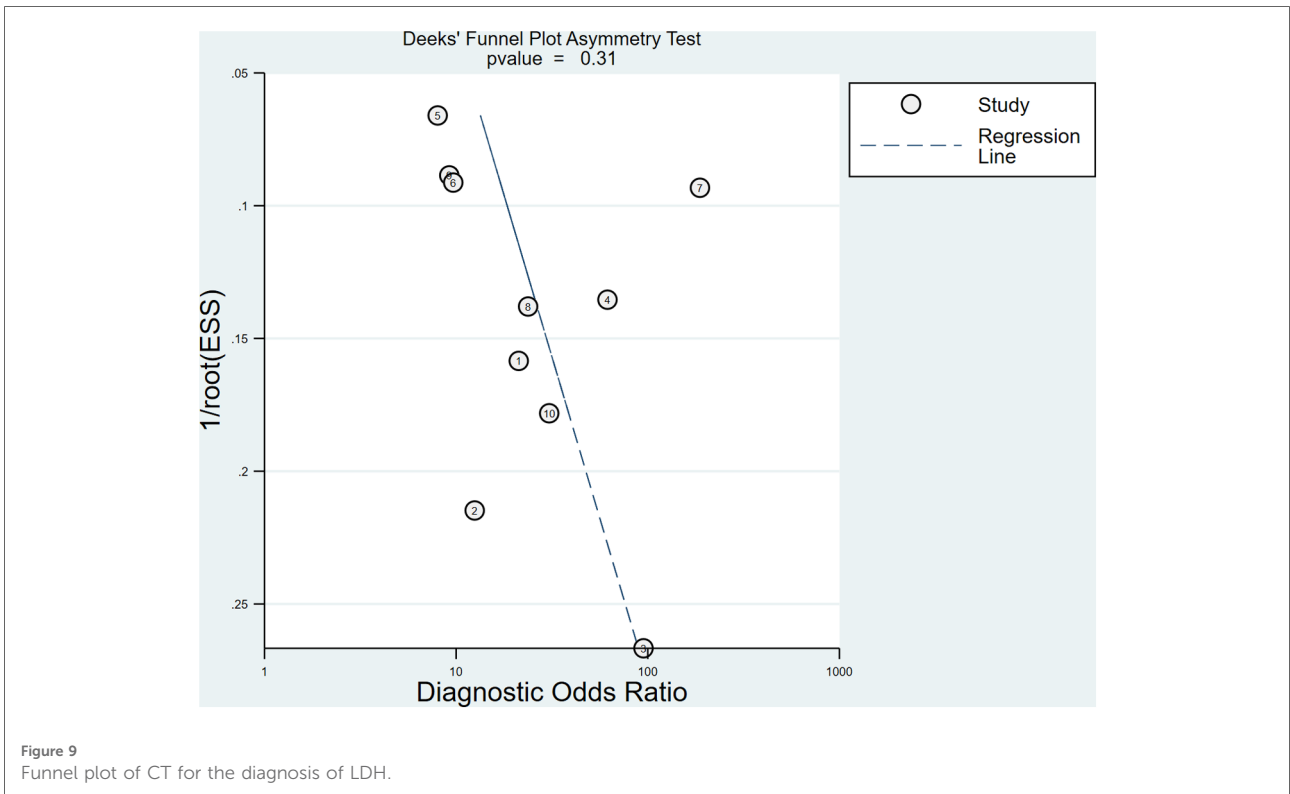
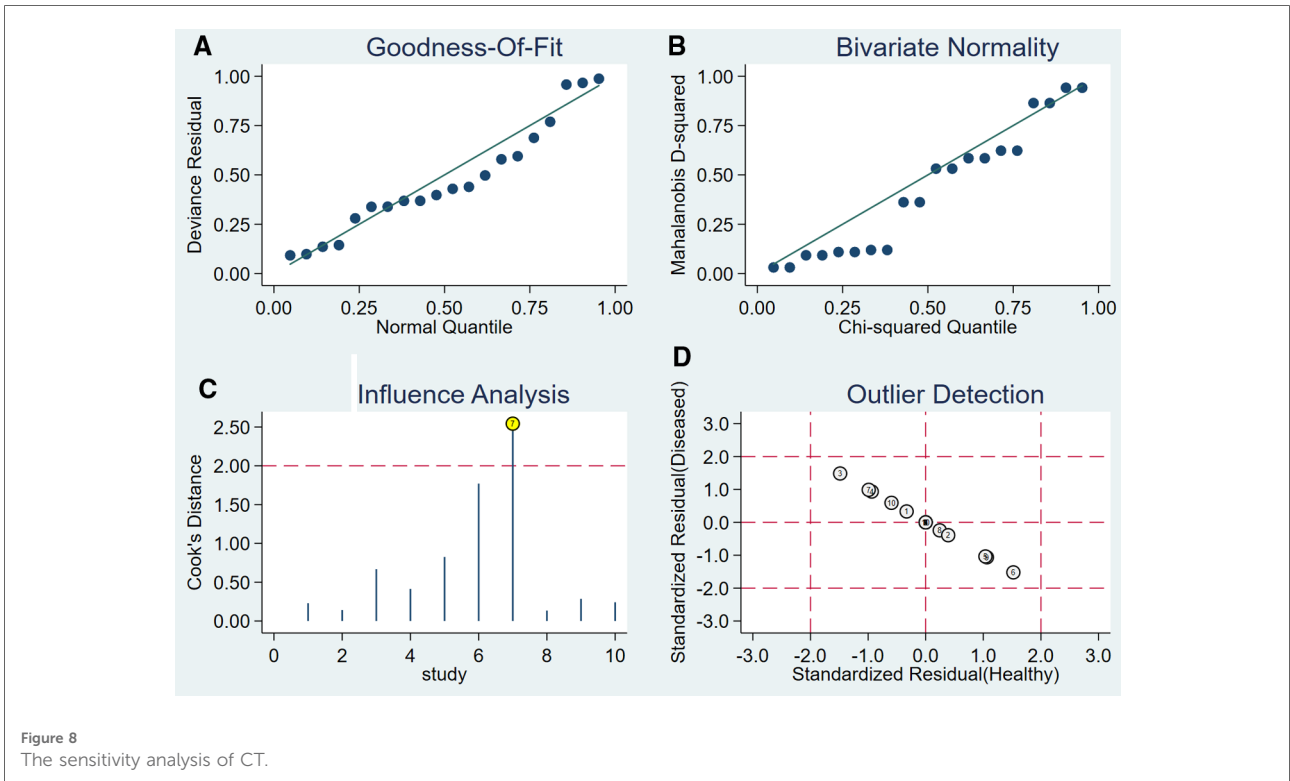


TABLE 4 Characteristics of myelography diagnostic tests.

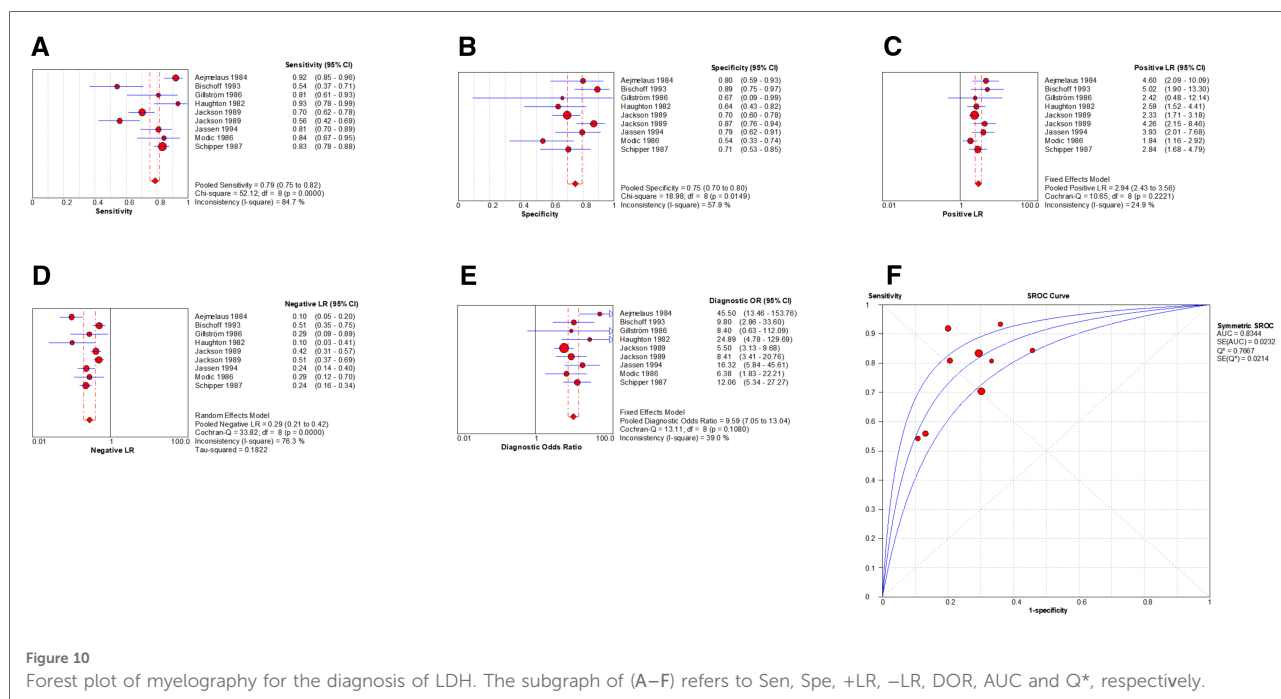
Study name	TP	FP	FN	TN
Aejmelaeus 1984 (22)	91	5	8	20
Bischoff 1993 (25)	19	4	16	33
Gillström 1986 (29)	21	1	5	2
Haughton 1982 (30)	28	9	2	16
Jackson 1989 (32)	88	32	37	74
Jackson 1989 (33)	33	8	26	53
Janssen 1994 (34)	55	7	13	27
Modic 1986 (37)	27	11	5	13
Schipper 1987 (39)	191	10	38	24

ionizing radiation and has good visualizing capacities, especially for soft tissue (43). MRI can also comprehensively observe whether each lumbar intervertebral disc has lesions, identify the degree and location of nucleus pulposus herniation on the sagittal plane, and distinguish whether there are other space-occupying lesions in the spinal canal. CT can show the shape of the bony spinal canal and the size and direction of intervertebral disc herniation. It has great diagnostic value for this disease. At present, CT is being commonly used (44). Compared with MRI, CT has the advantages of low cost, shorter total testing time, and larger availability of CT scanners in hospital settings but has the disadvantage of exposure to ionizing radiation. Myelography requires an injection of a contrast medium when testing, under specific circumstances (e.g., metal implant or malalignment of the spine). Myelography can replace MRI as the imaging

examination (45). An x-ray cannot directly identify the existence of LDH. Scoliosis, vertebral marginal hyperplasia, and narrowing of intervertebral space on the film all suggest degenerative changes. If the lumbosacral structure is abnormal (e.g., transitional spine, spondylolisthesis, and spondylolysis), it indicates that the adjacent intervertebral discs will accelerate the degeneration and increase the chance of protrusion owing to the increase of stress. With the development of technology in recent times, an x-ray examination is rarely used at present (46).

The comparison of effect sizes showed that the pooled Sen of MRI [0.89 (95%CI: 0.87–0.91)] was higher than that of myelography [0.79 (95%CI: 0.75–0.82)]. The pooled DOR of MRI [39.80 (95%CI: 18.35–86.32)] was also higher than that of myelography [9.59 (95%CI: 7.05–13.04)].

To improve the stability and reliability of the research results, during the implementation of this meta-analysis, two reviewers independently extracted data and evaluated the risk of bias in the included studies. Strict inclusion criteria and exclusion criteria were formulated during literature screening. Considering the differences between studies, the effect sizes with high heterogeneity were analyzed by meta-regression and subgroup analysis. But the source of heterogeneity was not found, the random effects model was used for pooling. Sensitivity and publication bias analyses were performed to make the final results more reliable. However, because of the differences in the condition of patients, medical equipment, and the doctors' proficiency in imaging examination, the heterogeneity of some effect sizes could be high, and it was difficult to find



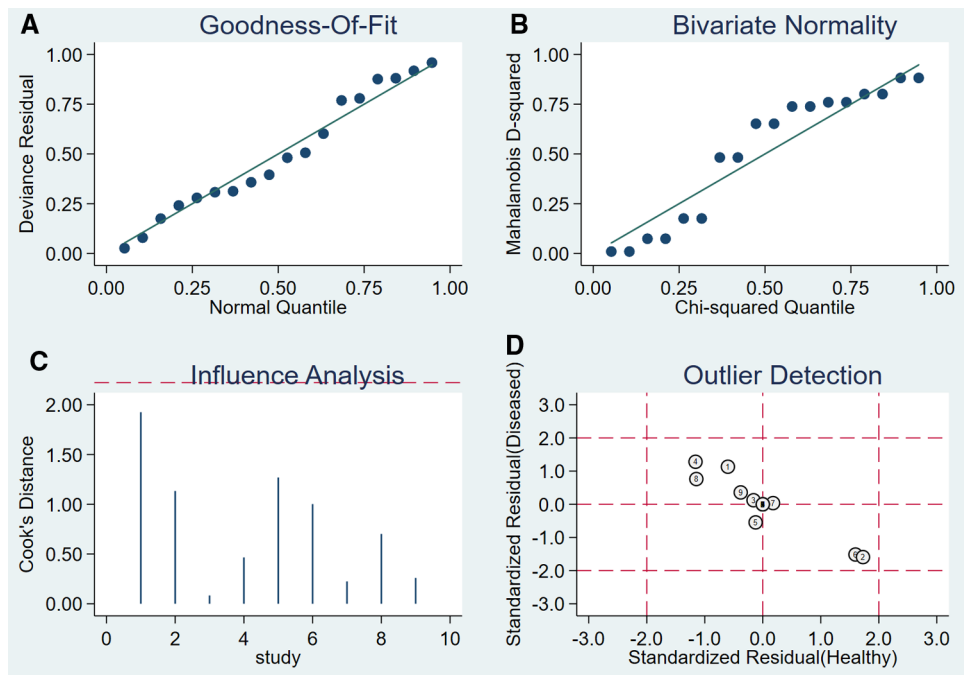


Figure 11 The sensitivity analysis of myelography.

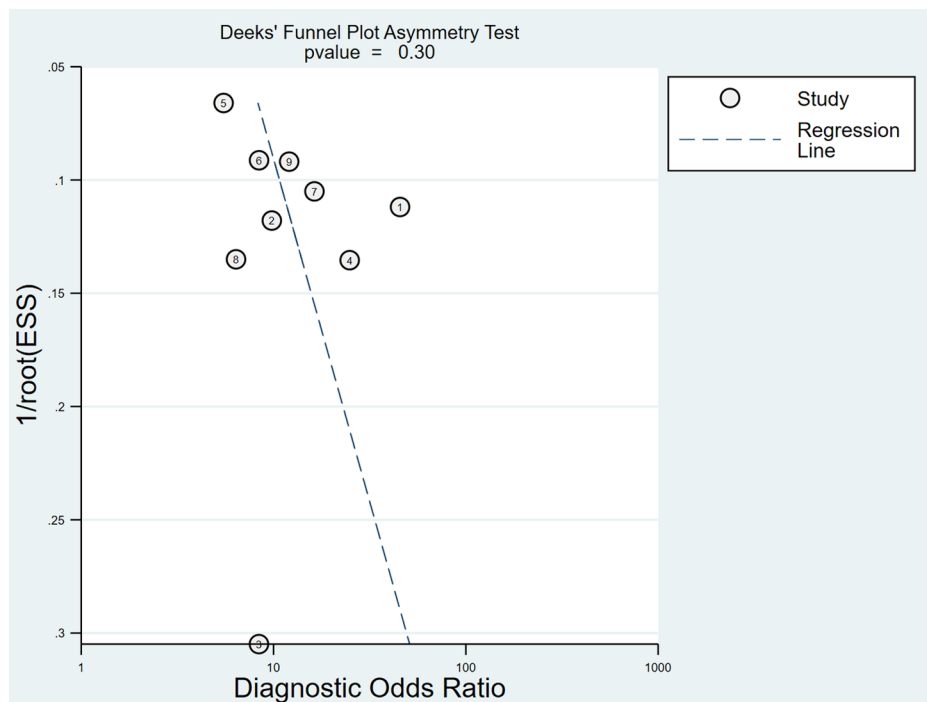


Figure 12 Funnel plot of myelography for the diagnosis of LDH.

the source of heterogeneity. Finally, some studies regarded patients as research objects, while others regarded lumbar discs as research objects, which also affect the results of this meta-analysis.

## Conclusion

MRI, CT, and myelography have a high value in the diagnosis of LDH; however, the diagnostic value of MRI is higher than that of myelography. Therefore, reasonable selection should be made in combination with the patients' condition.

## Data availability statement

The original contributions presented in the study are included in the article/**Supplementary Materials**, further inquiries can be directed to the corresponding author/s.

## Author contributions

ZH, PZ, CZ, and JW devised the project, the main conceptual ideas, and planned the research. ZH, PZ, and RL worked out the methodology. ZH, PZ, and CZ performed the data collection. ZH and PZ also organized and maintained research data for analysis. ZH performed analytic calculations.

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ZH, PZ, and JW validated reproducibility of the results. ZH, CZ, and JW wrote the manuscript with input from all authors. ZH, CZ, and RL extensively reviewed the work and further edited the manuscript. All authors contributed to the article and approved the submitted version.

## Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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## Supplementary material

The Supplementary Material for this article can be found online at: <https://www.frontiersin.org/articles/10.3389/fsurg.2022.1020766/full#supplementary-material>.

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