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Corrigendum: Degenerate Beta autoregressive model for proportion time-series with zeros or ones: an application to antimicrobial resistance rate using R shiny app

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KEYWORDS

Beta distribution, time-series model, mixture distribution, rates, proportions, inflated distribution, AMR, resistance

A corrigendum on
[Degenerate Beta autoregressive model for proportion time-series with zeros or ones: an application to antimicrobial resistance rate using R shiny app](#)
 by Lobo, J., Kamath, A., and Kalwaje Eshwara, V. (2023). *Front. Public Health* 10:969777.
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In the published article reference 14 was not cited in the article and an additional citation for reference 11 was missed. The citations have now been inserted in **Material and Methods, Degenerate Beta Autoregressive (DeβAR) model, Parameter estimation** and should read:

“Here, let $x_t^* = \log(\frac{x_t}{1-x_t})$ if $x_t \in (0, 1)$ else $x_t^* = 0$ (11, 14) and $\psi(\mu_t \zeta) - \psi((1 - \mu_t)\zeta) = \mu_t^*$, where $\psi(\cdot)$ is a digamma function.”

In the published article, there was an error. Inbetween steps of likelihood derivation was missed.

A correction has been made to **Material and Methods, Degenerate Beta Autoregressive (DeβAR) model, Parameter estimation**. “This sentence previously stated:”

$$L(\theta; x) = \prod_{t=p+1}^n f_{X_t}(x_t | \mathcal{F}_{t-1})$$

The likelihood function for the parameters of Degenerate Beta AR model is given by,

$$L(\theta; x) = \prod_{t=p+1}^n \{ \omega \mathcal{I}_{x_t=c} + \mathcal{I}_{x_t \in (0,1)} (1 - \omega) \frac{\Gamma(\zeta)}{\Gamma(\mu_t \zeta) \Gamma((1 - \mu_t)\zeta)} x_t^{\mu_t \zeta - 1} (1 - x_t)^{(1 - \mu_t)\zeta - 1} \}$$

“The corrected sentence appears below:”

$$L(\theta; x_t) = \prod_{t=p+1}^n f_{X_t}(x_t | \mathcal{F}_{t-1})$$

$$L(\theta; x_t) = \prod_{t=p+1}^n b_{i_c}(z_t; \omega, \mu_t, \zeta) = L_1(\omega)L_2(\mu_t, \zeta)$$

where,

$$L_1(\omega) = \prod_{t=p+1}^n \omega^{\mathcal{I}_{x_t=c}}(1 - \omega)^{\mathcal{I}_{x_t \in (0,1)}}$$

$$L_2(\mu_t, \zeta) = \prod_{t=p+1}^n \left\{ \frac{\Gamma(\zeta)}{\Gamma(\mu_t \zeta) \Gamma((1 - \mu_t) \zeta)} x_t^{\mu_t \zeta - 1} (1 - x_t)^{(1 - \mu_t) \zeta - 1} \right\}^{\mathcal{I}_{x_t \in (0,1)}}$$

The likelihood function for the parameters of Degenerate Beta AR model is given by,

$$L(\theta; x_t) = \prod_{t=p+1}^n \left\{ \omega^{\mathcal{I}_{x_t=c}} + \mathcal{I}_{x_t \in (0,1)}(1 - \omega) \frac{\Gamma(\zeta)}{\Gamma(\mu_t \zeta) \Gamma((1 - \mu_t) \zeta)} x_t^{\mu_t \zeta - 1} (1 - x_t)^{(1 - \mu_t) \zeta - 1} \right\}$$

In the published article, there was an error. Limitation of the model has been added and reference 19 has been added.

A correction has been made to **Material and Methods, Degenerate Beta Autoregressive (DeβAR) model, Parameter estimation**. “This sentence previously stated:”

Large sample inference: If the model specified by Equation (5) follows the regularity condition of maximum likelihood estimation (MLE) then, MLEs of θ and $J(\theta)$ (Fisher information matrix) are consistent. Assuming that $I(\theta) = \lim_{n \rightarrow \infty} \{n^{-1}J(\theta)\}$ exists and is non-singular, we have $\sqrt{n}(\hat{\theta} - \theta)$ converges in distribution to $N(0, I(\theta)^{-1})$.

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Note: The proposed DeβAR model is applicable when x_t^* is converted to 0 as mentioned above. To overcome with this limitation Bayer et al. (19) proposed Inflated beta autoregressive moving average models, which are more suitable when interval data includes 0 or 1.

The authors apologize for this error and state that this does not change the scientific conclusions of the article in any way. The original article has been updated.

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