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Associations between blood essential metal mixture and serum uric acid: a cross-sectional study

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Introduction: Although several studies have explored the associations between single essential metals and serum uric acid (SUA), the study about the essential metal mixture and the interactions of metals for hyperuricemia remains unclear.

Methods: We performed a cross-sectional study to explore the association of the SUA levels with the blood essential metal mixture, including magnesium (Mg), calcium (Ca), iron (Fe), copper (Cu), zinc (Zn), manganese (Mn) in Chinese community-dwelling adults (n=1039). The multivariable linear regression, the weighted quantile sum (WQS) regression and Bayesian kernel machine regression (BKMR) were conducted to estimate the associations of blood essential metals with SUA levels and the BKMR model was also conducted to estimate the interactions of the essential metals on SUA.

Results: In the multivariable linear regression, the association of blood Mg, Mn, and Cu with SUA was statistically significant, both in considering multiple metals and a single metal. In WQS regression [β =13.59 (95%CI: 5.57, 21.60)] and BKMR models, a positive association was found between the mixture of essential metals in blood and SUA. Specifically, blood Mg and Cu showed a positive association with SUA, while blood Mn showed a negative association. Additionally, no interactions between individual metals on SUA were observed.

Discussion: In conclusion, further attention should be paid to the relationship between the mixture of essential metals in blood and SUA. However, more studies are needed to confirm these findings.

KEYWORDS

manganese, magnesium, copper, essential metal mixture, interactions, serum uric acid

Introduction

Serum uric acid (SUA) is the end-product of purine metabolism, and hyperuricemia can result from either increased production or decreased excretion of SUA (1). In recent years, the prevalence of hyperuricemia has sharply increased. Epidemiological investigations have revealed that the prevalence of hyperuricemia in America increased from 19.1% (1988–1994) to 21.5% (2007–2008) (2). A previous meta-analysis reported that the prevalence of hyperuricemia in Chinese individuals was 13.3% (3). It is widely recognized that high SUA levels are closely associated with various diseases, including gout (4), diabetes, hypertension (5), chronic kidney disease (6), non-alcoholic fatty liver disease, liver fibrosis (7), and adverse cardiovascular outcomes (8). Therefore, it is crucial to identify the risk factors for hyperuricemia.

In recent years, there has been an increasing amount of epidemiological research investigating the potential relationship between essential metals and hyperuricemia. A cross-sectional study on middle-aged Chinese indicated that dietary zinc (Zn) intake was correlated with hyperuricemia negatively (9). Another cross-sectional study suggested the positive relationship between copper (Cu) and hyperuricemia in middle-aged and older adults population (10). One recent study found magnesium (Mg) inversely correlated with the risk of new-onset hyperuricemia (11). However, these studies only assessed the individual association between hyperuricemia and essential metals, rather than evaluating the mixture of the metals and the possible interactions between them.

In China, an increasing number of people are taking multivitamin-mineral tablets for wellness without measuring the concentrations of the metals in the blood (12). People are always exposed to multiple metals daily, possibly interacting with each other. Previous studies only focused on the associations between co-exposure to heavy metals (lead, cadmium, mercury) and SUA (13–15). A recent study found the potential interactions between manganese (Mn) and cobalt (Co), Mn and chromium (Cr), Mn and Zn, and Cr and Co on hyperuricemia (9, 16). However, the associations between blood essential metal mixture and SUA have yet to be studied so far, as well as the interactions of the essential metals.

In the current study, we chose a metal mixture including Mg, calcium (Ca), iron (Fe), Cu, Zn, and Mn, which are commonly tested clinically. We aimed to explore the associations between SUA and the blood essential metal mixture and the potential interactions of the essential metals, including Mg, Ca, Fe, Cu, Zn, and Mn, on SUA levels in Chinese community-dwelling adults.

Materials and methods

Study population

This cross-sectional study was conducted at www.chictr.org.cn (ChiCTR2100054130) using data from participants who were enrolled in 2021 from Lecong, Shunde District, Foshan, China, using a stratified cluster sampling method (17, 18). Initially, 1,111 potential participants aged ≥ 18 , who had lived in Shunde for more than 6 months and were not pregnant, were identified. Four individuals who had not undergone blood examination and 68 individuals who had taken SUA lowering drugs were excluded, resulting in a final sample of 1,039 participants (Supplementary Figure S1). The study protocol (20211103) was approved by the Ethics Committee of Shunde Hospital of Southern Medical University and was conducted in accordance with the ethical guidelines of the 1975 Declaration of Helsinki. Prior to participants.

Measurements and definitions

Data collection

A team of well-versed researchers conducted a standardized survey to gather data, encompassing demographics, lifestyle, medical history, and medication usage (19). Anthropometric characteristics, such as weight, height, and blood pressure, were measured in accordance with established protocols. Body Mass Index (BMI) was computed as weight in kilograms divided by height in meters squared (kg/m²). Blood pressure readings were obtained twice using an automated electronic device (HEM-752 FUZZY, Omron, China), and the mean of these measurements was selected (20).

Laboratory assays

All samples were collected from participants who had fasted for at least 10h overnight between 08:00 am and 10:00 am. The samples were transported under cold chain management to a central laboratory certified by the College of American Pathologists and were centrifuged within 4h. The SUA concentration was detected using a Beckman Coulter AU 5800 (Beckman Coulter Inc., Brea, CA, United States). The blood concentrations of six metals, including Mg, Ca, Fe, Cu, Zn, and Mn, were measured using inductively coupled plasma mass spectrometry (ICP-MS) (Thermo Fisher Scientific, iCAP RQ, Waltham, United States). Fasting plasma glucose (FPG) was assessed using a Hitachi LABOSPECT 008AS (Tokyo, JAPAN), and 2h postprandial plasma glucose (PPG) was measured using a Hitachi 7,600 automatic biochemical analyzer (Hitachi, Tokyo, Japan). Glycated hemoglobin (HbA1c) was assessed using highperformance liquid chromatography (TOSOH, HLC-723G8, Tokyo, JAPAN). Additionally, the serum creatinine (Scr) and blood lipid profiles, including total cholesterol (TC), triglyceride (TG), high-density lipoprotein (HDL), and low-density lipoprotein (LDL), were measured using a BS800 (Mindray, Shenzhen, China).

Covariates

All the models were adjusted for categorical covariates: age categories (\leq 40, 40–60 and >60 years), sex (men, women), educational levels(<high school, high school, and>high school), smoking status (never, ever, and current) (21), alcohol abuse (yes, no), BMI categories $(\langle 24, 24-28, \geq 28)$ (22), diabetes (yes, no), hypertension (yes, no), dyslipidemia (yes, no) and eGFR (<60, 60–89, ≥90). It was considered alcohol abuse if it was>30 g/day for men and>20 g/day for women (23). Diabetes was defined as FPG level≥7.0 mmol/L, PPG ≥11.1 mmol/L, or HbA1c ≥6.5% and/or self-reported history of disease diagnosis (24). Hypertension was defined if systolic blood pressure \geq 140 mmHg or diastolic blood pressure \geq 90 mmHg, and/or self-reported history of disease diagnosis (25). Dyslipidemia was defined as a higher level of TC (≥6.22 mmol/L), TG (≥2.26 mmol/L), LDL-C (≥4.14 mmol/L), or HDL-C (<1.04 mmol/L) and/or previous diagnosis of dyslipidemia (26). The estimated glomerular filtration rate (eGFR) was calculated using the modified 4-variable Modification of Diet in Renal Disease (MDRD) equation:

$$eGFR\left(\frac{ml}{min}/m^{2}\right) = 175 \times (Scr)^{-1.154} \times age^{-0.203} \times 0.742 (if female) \times 1.212 (if black),$$

where Scr was expressed as mg/dl and age was expressed in years (27).

Statistical analysis

The data were analyzed using IBM SPSS Statistics (version 24) and R (version 4.2.2). Continuous variables were summarized using mean and standard deviations or median and interquartile range (IQR) whether the variable distribution was symmetric or not, respectively.

The Spearman correlation method was applied to assess the correlation among the different essential metal elements. To achieve normal distribution, the levels of studied metal elements were ln-transformed. The association between blood essential metals and SUA was evaluated using multivariable linear regression, weighted quantile sum (WQS) regression, and Bayesian kernel machine regression (BKMR) models. The statistical level of significance was set at 0.05 and all tests were two-tailed.

The multivariable linear regression: We performed the multivariable linear regression analysis to examine the relationship between blood essential metals and SUA. Corrected β regression coefficients and 95% confidence intervals (CIs) were calculated. Initially, a multivariable linear regression model was fitted for each essential metal, controlling for other metals and potential confounders such as age, sex, education, smoking status, alcohol abuse, BMI, diabetes, hypertension, dyslipidemia, and eGFR. We observed no significant collinearity among these variables (all VIF<3) (Supplementary Table S1). Finally, we investigated the associations between each essential metal and SUA after adjusting for potential confounders.

BKMR model: We applied the nonparametric BKMR method to estimate the overall, single-exposure, and interactive effects of each metal element with SUA, while considering possible interactions and nonlinear associations (28). Specifically, for a continuous outcome, the model can be expressed as follows:

$$Y_i = h(z_{i1}, \dots, z_{iM}) + x_i\beta + \epsilon_i$$

Where Y_i represents the response for individual i (i=1,...,n), z_{iM} is the *m*th exposed variable, h() is the unknown exposure-response function, β is the effect of the covariates x_i , and the residuals $\epsilon_i \sim N(0,\sigma^2)$ are assumed to be independent and identically normally distributed with a common variance (29). The posterior distributions for variables were provided using 50,000 iterations of the Markov Chain Monte Carlo (MCMC) (30–33). We utilized the R package BKMR for this analysis.

WQS regression model: WQS is an approach that can evaluate the mixed effect of metals and estimate the weight of individual metals (34). The study participants were randomly allocated to a training dataset (40%, n=416) and a validation dataset (60%, n=623) (35). We used 1,000 bootstrap samples (B=1,000) from the training dataset to calculate the weights that maximize the likelihood for the model in this analysis (36, 37). The fitted model of WQS can be expressed as follows:

$$g(\mu) = \beta_0 + \beta_1 \sum_{i=1}^{c} w_i q_i + \varphi z'$$

$$WQS = \sum_{i=1}^{c} \overline{w_i} q_i$$
where $\overline{w_i} = \frac{1}{B} \sum_{b=1}^{B} w_i(b) f\left(\hat{\beta}_{1(b)}\right)$ and $f\left(\hat{\beta}_{1(b)}\right)$ is a

pre-specified "signal function" of the estimated slope parameter associated with WQS from the *b*th bootstrap sample to measure the signal strength (30, 34, 38). In the above equation, the term $\sum_{i=1}^{c} w_i q_i$ represents the

overall weighted index for the set of 6 studied metals. In the second equation above, the c=6 essential metal concentrations were categorized into quartiles (*qi*) and assigned a score of 0, 1, 2, or 3 for i=1 to *c*. The weight for the *i*th chemical component *qi* is represented by *wi*. In the first equation, *z* refers to the covariate vector determined before the weights were estimated, and φ represents the coefficients for the covariates in *z* (34, 39). For this analysis, metals with estimated weights greater than 0.167 (1/6) were considered significant contributors to the WQS score (40). To improve the robustness of the regression parameters, we apply the repeated holdout extension of WQS with 100 repetitions. That is, we performed 100 random partitions on the dataset and repeated WQS regressions in each partition. Based on the simulated distribution, the β coefficient of the WQS index was averaged as the final estimate (41–43). The gWQS package in R was employed for this purpose.

Results

General characteristics of participants

Table 1 presents an overview of the participants' general characteristics, which includes age, sex, SUA, FPG, HbA1c, education level, BMI, smoking status, alcohol abuse, hypertension, dyslipidemia, diabetes, eGFR categories, and the investigated metallic elements. Of the 1,039 participants, 438 (42.2%) were male and 601 (57.8%) were female, with a mean age of 49.18 ± 14.27 years. The average SUA level was 378.42 ± 96.80 µmol/L. In terms of BMI, the prevalence of obesity was 13.1%. 98.3% participants reported no alcohol abuse, and 81.1% had never smoked cigarettes. Among all participants, 68.0% had normal eGFR (\geq 90 mL/min/m²). Furthermore, the blood Mn, Ca, Cu and Zn were slightly left skewed, while the blood Mg and Fe were slightly right skewed.

The associations between blood essential metal and SUA using the multivariable linear regression

Supplementary Figure S2 displays the Spearman correlation coefficients among essential metals, which suggested the correlations ranging from weak to moderate (r: 0.01 to 0.65) (44). Figure 1 presents that the associations of blood Mg [β =81.30 (95%CI: 17.11, 145.49), p=0.013], Mn [β =-18.50 (95%CI: -35.73, -1.26), p=0.035], and Cu [β =43.74 (95%CI: 1.22, 86.26), p=0.044] with SUA was statistically significant, considering multiple metals. Furthermore, the associations of blood Mg [β =76.02 (95%CI: 28.00, 124.04), p=0.002], Mn [β =-17.49 (95%CI: -34.52, -0.47), p=0.044], Cu [β =55.83 (95%CI: 16.26, 95.41), p=0.006], and Zn [β =46.69 (95%CI: 11.40, 81.97), p=0.010] with SUA were statistically significant, considering a single metal.

The associations between blood essential metal mixture and SUA using the BKMR model

The BKMR model revealed that higher levels of SUA were associated with increasing levels of common blood metal elements

TABLE 1 General characteristics of	participants in the study.
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	Total (<i>n</i> = 1,039)	
Age, years	49.18±14.27	
Male, %	42.2	
SUA, µmol/L	378.42±96.80	
FPG, mmol/L	5.03 (4.70,5.41)	
HbA1c, %	5.70 (5.40,6.00)	
Education, %		
< High school	28.8	
High school	34.1	
> High school	37.2	
BMI, %		
Normal weight	55.0	
Overweight	32.0	
Obese	13.1	
Smoking status, %		
No	81.1	
Ever	4.7	
Current	14.1	
Alcohol abuse, %	1.7	
Hypertension, %	33.8	
Dyslipidemia, %	42.3	
Diabetes, %	11.4	
eGFR, %		
<60 mL/min/m ²	2.5	
60-90 mL/min/m ²	29.5	
$\geq 90 \text{mL/min/m}^2$	68.0	
Mg, mg/L	41.80 (38.60,45.00)	
Mn, µg/L	12.60 (10.30,15.20)	
Ca, mg/L	62.80 (58.80,67.00)	
Fe, mg/L	506.00 (468.50,543.90)	
Cu, µg/L	875.40 (799.20,963.60)	
Zn, mg/L	6.30 (5.70,6.90)	

The general characteristics of participants were summarized as mean±standard deviation or median and interquartile range (IQR) for continuous variables and frequencies for categorical variables (%).

SUA, serum uric acid; BMI, body mass index; eGFR, estimated glomerular filtration rate; FPG, fasting plasma glucose; HbA1c, glycated hemoglobin; Mg, magnesium; Mn, manganese; Ca, calcium; Fe, iron; Cu, copper; Zn, zinc.

when the concentrations of blood essential metals were above the 55th percentile (Figure 2). To examine the contribution of individual exposures to the overall effect, we fixed the other elements at their 25th, 50th, or 75th percentile and identified significant relationships between Mg, Mn, and Cu with SUA. Our findings indicated that increased exposure to Mg and Cu was associated with higher levels of SUA, while increased exposure to Mn was associated with lower levels of SUA (Figure 3). Additionally, Supplementary Figure S3 illustrates no interactions between individual metals on SUA.

The associations between blood essential metal mixture and SUA using the WQS model

Since there is no evidence of non-linear relationship between the exposures and the outcome and that there is no interaction among the elements in the mixture, we applied the WQS regression. The WQS model demonstrated the association between the mixture of essential metals and SUA. The WQS index showed a positive and significant association using repeated holdout validation [β =13.59 (95%CI: 5.57, 21.60)] (Table 2), indicating that the combined effect of mixed metal elements was positively linked to SUA (42). Table 3 shows the mean weights obtained by the repeated holdout. Specifically, blood Zn (29.3%), Mg (21.4%), Cu (19.7%) and Fe (17.4%) made significant contributions in positive direction. Conversely, blood Mn (65.3%) had the highest weight in the negative direction, and blood Ca (20.9%) also made a considerable contribution. What's more, Figure 4 displays the distribution of all metal weights obtained with 100 repetitions.

Discussion

The current study examined the associations between blood essential metals, including Mg, Ca, Fe, Cu, Zn, and Mn, and SUA levels in Chinese community-dwelling adults using three statistical approaches. In reality, different models serve different purposes, and each approach may have its own advantages and disadvantages, leading to distinct outcomes. The multivariable linear regression is a commonly used and straightforward method for analysis and interpretation of data. And the current study demonstrated a statistically significant association between blood Mg, Mn, and Cu with SUA when considering multiple metals. However, we conducted multivariable linear regression and found a positive relationship between blood Zn and SUA in the analysis of single metal. The moderate correlation between some metals might lead to result distortion (45). Also, it is not enough when considering mixed exposures and their complex interactions in a simple linear regression model (34, 45). The BKMR model, on one hand, can estimate both the interactive effects of exposures and reveal the nonlinear relationships between exposures and outcomes (30, 46, 47). In our analysis, blood Mg and Cu were positively associated with SUA, while blood Mn showed a negative relationship. WQS regression, on the other hand, is capable of assessing the overall effect and contribution of each exposure by estimating the weight of individual metals (48, 49). For instance, the combined effect of mixed metal elements was positively linked to SUA [β=13.59 (95%CI: 5.57, 21.60)]. However, it assumes that all exposures have the same direction of associations with the outcome, which requires making assumptions in two different directions (48, 50, 51). In the current study, we found that the blood Zn (29.3%) had the highest mean weight in positive direction while the blood Mn (65.3%) had the highest mean weight in the negative direction. Overall, the results of both models were consistent and reliable, suggesting that essential metal supplementation should not be blindly undertaken. These findings provide a possible approach to preventing hyperuricemia.

The relationship between Zn and uric acid (UA) has been controversial in previous research. A study from NHANES



FIGURE 1

The associations between blood essential metals and SUA using the multivariable linear regression. (A) Linear regression considering multiple metals (B) linear regression considering a single metal. The levels of studied metal elements were ln-transformed to improve normal distribution. The model was adjusted for age, sex, educational levels, smoking status, alcohol abuse, BMI categories, diabetes, hypertension, dyslipidemia, and eGFR. BMI, body mass index; eGFR, estimated glomerular filtration rate; Mg, magnesium; Mn, manganese; Ca, calcium; Fe, iron; Cu, copper; Zn, zinc.

(2001–2014, n=37,215) suggested that dietary Zn intake (median: 11.82 mg/d for men; 8.45 mg/d for women) is inversely associated with hyperuricemia using 24h dietary recall method (52). Another crosssectional study in China (n=5,168) also indicated that dietary Zn intake (median: 20.54 mg/d for men; 17.20 mg/d for women) was inversely associated with hyperuricemia in men using the food frequency questionnaire (FFQ) (9). Additionally, an animal study inferred that Zn can alleviate hyperuricemia by decreasing the activities of adenosine deaminase and xanthine oxidase (XO), and promoting UA excretion by altering intestinal flora composition (53). However, these studies had focused the dietary Zn intake instead of the blood Zn. A cross-sectional study in China (n = 1,046) shown that higher plasma levels of Zn might increase hyperuricemia risk (54). Another study (n=6,508) also found that the urinary levels of Zn positively associated with hyperuricemia in China (55). The current study indicated a positive association between blood Zn and SUA. Our bodies have a capacity to maintain Zn homeostasis while dietary Zn intake varies (56). So the change of dietary Zn intake would not consistent with other biological samples (blood or urinary). Additionally, the positive associations between Zn and UA had been found in plasma, urinary and blood. And more studies on mechanisms were needed.

There are some researches have explored the relationship between blood Mn and SUA. One study (n=3,926) reported negative associations between Mn and SUA concentration (57), consistent with our findings. Conversely, another study based on NHANES from 2011 to 2018 (n=14,871) indicated that the association between blood Mn and SUA was not statistically significant (15). In accordance of a previous study, people with long-term and low-level of occupational Mn exposure showed a lower level of UA, but high-level of Mn exposure would cause nerve damage (58) which performed as parkinsonian symptoms (59). More and more studies had found that UA has a protective effect on Parkinson's disease (PD) (59–61), and the PD case had the lower levels of UA (62). Above all, there is a hypothesis that excessive Mn exposure leads to nerve damage that ultimately reduce the levels of UA, but further research is needed to find the mechanism between UA, Mn and the nervous system.

The current study has identified a positive association between blood Cu and SUA levels. In a cross-sectional analysis of Italian adults (n=1,197), it was observed that dietary Cu intake was inversely correlated with SUA levels, while serum Cu levels did not demonstrate statistical significance in relation to SUA (63). This finding is consistent with a previous cross-sectional study conducted in Chinese adults (n = 6,212), which suggested a positive relationship between serum Cu levels and hyperuricemia prevalence (10). The inconsistency in results may be attributed to differences in Cu exposure. The physiological mechanisms underlying the association between Cu and SUA remain to be established. Adenosine monophosphate (AMP) is converted to purine base hypoxanthine by deaminase, nucleotidase and purine nucleoside phosphorylase (PNP). Hypoxanthine is then oxidized by XO to form xanthine, which is again oxidized by XO to form the final product, UA (64). A vitro study has proposed that Cu can either inhibit or activate XO, depending on its concentration. Notably, a sharp inhibition of XO activity was observed when the Cu concentration reached 0.7 mM (65). It is plausible that different populations may have varying levels of Cu exposure, resulting in differing outcomes.

A previous study suggested an inverse relation of plasma Mg with the risk of new-onset hyperuricemia (11). Another study from American adults in 2001–2014 (n=26,796) indicated that Mg intake could prevent hyperuricemia (66). However, a meta-analysis showed that exposure to Mg was not associated with hyperuricemia risk (67). The results of these studies were inconsistent with our current study, which found that the relationship of Mg with SUA was positive. The previous studies suggested that Mg is positively associated with all lipoprotein species due to the affinity of specific phospholipids head groups as a divalent cation (68). A cohort investigation demonstrated a positive relationship between hyperuricemia and dyslipidemia in the



older adults population (69). Furthermore, the administration of statin therapy resulted in a significant reduction in SUA levels (70). The observed positive association between Mg and SUA may be attributed to the reported positive relationship between Mg and TG. Additionally, the discrepancy in measurement methods between our study (blood Mg) and previous studies (serum Mg) may be a potential contributing factor.

Despite the various models utilized and comprehensive adoption of results in the current study, there remain several limitations that should be noted. Firstly, given the cross-sectional design of the study, only association analysis could be performed, precluding any causal inferences. Secondly, while our sample size possessed sufficient power, the number of participants was relatively small. Lastly, it is important to note that metal ions in different valence states have varying functions, and further analyses would benefit from the detection of valence states.

In conclusion, our findings show a positive relationship between blood levels of essential metals and SUA. Specifically, we observed positive associations between SUA and blood Mg and Cu, while a negative association was found between SUA and blood Mn. Additionally, no interactions between individual metals on SUA were observed. These results highlight the need to further investigate the relationship between blood essential metals and SUA. Nevertheless, further studies are warranted to confirm our conclusions.

Data availability statement

The raw data supporting the conclusions of this article will be made available by the authors, without undue reservation.

Ethics statement

The studies involving humans were approved by the Ethics Committee of Shunde Hospital of Southern Medical University. The studies were conducted in accordance with the local legislation and institutional requirements. The participants provided their written informed consent to participate in this study.

Author contributions

DW and HD conducted the data analysis. LiL, YH, XC, YJ, QM, GY, SL, NY, YoL, XL, and LaL conducted the data acquisition.



The association between the individual blood essential metal and SUA using the BKMR model. The levels of studied metal elements were Intransformed to improve normal distribution. The model was adjusted for age, sex, educational levels, smoking status, alcohol abuse, BMI categories, diabetes, hypertension, dyslipidemia, and eGFR. The 75th percentile of the individual metal was compared to its 25th percentile, with the other essential metals at their 25th, 50th, or 75th percentile. We considered the results statistically significant if the 95% confidence interval (CI) of the estimate does not include 0. And the position of the individual metal suggested the positive or negative effect on the outcome. BMI, body mass index; eGFR, estimated glomerular filtration rate; Mg, magnesium; Mn, manganese; Ca, calcium; Fe, iron; Cu, copper; Zn, zinc.

TABLE 2	WQS index β	coefficients and	CIs by validation	technique & estimation	ation type.
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Validation technique	Estimation type	β coefficient	Lower limit	Upper limit
None: train/test full dataset	Mean & SE-based 95% CI	20.30	9.50	31.10
Repeated holdout	Mean & SD-based 95% CI	13.59	5.57	21.60

TABLE 3 The mean weights obtained by the repeated holdout.

	Positive	Negative
lnZn	0.293	0.027
lnMg	0.214	0.038
lnCu	0.197	0.041
lnFe	0.174	0.032
lnCa	0.109	0.209
lnMn	0.012	0.653

Mg, magnesium; Mn, manganese; Ca, calcium; Fe, iron; Cu, copper; Zn, zinc.

DW drafted the manuscript. YuL, SZ, HW, and JS revised the manuscript. JS and HW performed the conceptualization. All authors contributed to the article and approved the submitted version.

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FIGURE 4

Chemical of concern identification & uncertainty for blood essential metals and SUA. (A) The positive direction (B) the negative direction. The levels of studied metal elements were In-transformed to improve normal distribution. The hollow points represent the weight value of the 100 repetitions, and the box plots show the 25th, 50th, and 75th percentiles, and the diamonds represent the mean weights of 100 repetitions. The horizontal line represents the threshold value of the weight, and metals with estimated weights greater than threshold 1/6 (16.7%) were considered significant contributors to the WQS score. The model was adjusted for age, sex, educational levels, smoking status, alcohol abuse, BMI categories, diabetes, hypertension, dyslipidemia, and eGFR. BMI, body mass index; eGFR, estimated glomerular filtration rate; Mg, magnesium; Mn, manganese; Ca, calcium; Fe, iron; Cu, copper; Zn, zinc.

Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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References

1. Sun H, Wang N, Chen C, Nie X, Han B, Li Q, et al. Cadmium exposure and its association with serum uric acid and hyperuricemia. *Sci Rep.* (2017) 7:550. doi: 10.1038/ s41598-017-00661-3

 Tan Y, Fu Y, Huang F, Wen L, Weng X, Yao H, et al. Association between blood metal exposures and hyperuricemia in the U.S. general adult: a subgroup analysis from NHANES. *Chemosphere*. (2023) 318:137873. doi: 10.1016/j.chemosphere.2023. 137873

3. Liu R, Han C, Wu D, Xia X, Gu J, Guan H, et al. Prevalence of hyperuricemia and gout in mainland China from 2000 to 2014: a systematic review and Meta-analysis. *Biomed Res Int.* (2015) 2015:762820:1–12. doi: 10.1155/2015/762820

4. Rock KL, Kataoka H, Lai JJ. Uric acid as a danger signal in gout and its comorbidities. *Nat Rev Rheumatol.* (2013) 9:13–23. doi: 10.1038/nrrheum.2012.143

5. Nakanishi N, Okamoto M, Yoshida H, Matsuo Y, Suzuki K, Tatara K. Serum uric acid and risk for development of hypertension and impaired fasting glucose or type II diabetes in Japanese male office workers. *Eur J Epidemiol.* (2003) 18:523–30. doi: 10.1023/A:1024600905574

6. Ruilope LM, Pontremoli R. Serum uric acid and cardio-renal diseases. *Curr Med Res Opin*. (2013) 29:25–31. doi: 10.1185/03007995.2013.790803

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Supplementary material

The Supplementary material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fpubh.2023.1182127/ full#supplementary-material

7. Duan H, Zhang R, Chen X, Yu G, Song C, Jiang Y, et al. Associations of uric acid with liver steatosis and fibrosis applying vibration controlled transient elastography in the United States: a nationwide cross-section study. *Front Endocrinol (Lausanne)*. (2022) 13:930224. doi: 10.3389/fendo.2022.930224

8. Kleber ME, Delgado G, Grammer TB, Silbernagel G, Huang J, Kramer BK, et al. Uric acid and cardiovascular events: a Mendelian randomization study. J Am Soc Nephrol. (2015) 26:2831–8. doi: 10.1681/ASN.2014070660

9. Xie DX, Xiong YL, Zeng C, Wei J, Yang T, Li H, et al. Association between low dietary zinc and hyperuricaemia in middle-aged and older males in China: a cross-sectional study. *BMJ Open.* (2015) 5:e008637. doi: 10.1136/bmjopen-2015-008637

10. Jiang T, Xie D, Wu J, He H, Wang H, Wang N, et al. Association between serum copper levels and prevalence of hyperuricemia: a cross-sectional study. *Sci Rep.* (2020) 10:8687. doi: 10.1038/s41598-020-65639-0

11. Cao J, Zhang J, Zhang Y, Li H, Jiang C, Lin T, et al. Plasma magnesium and the risk of new-onset hyperuricaemia in hypertensive patients. *Br J Nutr.* (2020) 124:1–8. doi: 10.1017/S0007114520001099

12. Force USPST, Mangione CM, Barry MJ, Nicholson WK, Cabana M, Chelmow D, et al. Vitamin, mineral, and multivitamin supplementation to prevent cardiovascular

disease and Cancer: US preventive services task Force recommendation statement. JAMA. (2022) 327:2326–33. doi: 10.1001/jama.2022.8970

13. Park J, Kim Y. Associations of blood heavy metals with uric acid in the Korean general population: analysis of data from the 2016–2017 Korean National Health and nutrition examination survey. *Biol Trace Elem Res.* (2021) 199:102–12. doi: 10.1007/s12011-020-02152-5

14. Sanders AP, Mazzella MJ, Malin AJ, Hair GM, Busgang SA, Saland JM, et al. Combined exposure to lead, cadmium, mercury, and arsenic and kidney health in adolescents age 12–19 in NHANES 2009–2014. *Environ Int.* (2019) 131:104993. doi: 10.1016/j.envint.2019.104993

15. Xu J, Zhu X, Hui R, Xing Y, Wang J, Shi S, et al. Associations of metal exposure with hyperuricemia and gout in general adults. *Front Endocrinol (Lausanne)*. (2022) 13:1052784. doi: 10.3389/fendo.2022.1052784

16. Wu S, Huang H, Ji G, Li L, Xing X, Dong M, et al. Joint effect of multiple metals on hyperuricemia and their interaction with obesity: a community-based cross-sectional study in China. *Nutrients*. (2023) 15:552. doi: 10.3390/nu15030552

17. Wan H, Yu G, Xu S, Chen X, Jiang Y, Duan H, et al. Central sensitivity to free triiodothyronine with MAFLD and its progression to liver fibrosis in euthyroid adults. *J Clin Endocrinol Metab.* (2023):dgad18. doi: 10.1210/clinem/dgad186

18. Luo Y, Zeng H, Ye Y, Yu G, Song C, Liu S, et al. Associations of metal profiles in blood with thyroiditis: a cross-sectional study. *Environ Sci Pollut Res Int.* (2023) 30:21072–80. doi: 10.1007/s11356-022-23625-1

19. Wan H, Wang Y, Xiang Q, Fang S, Chen Y, Chen C, et al. Associations between abdominal obesity indices and diabetic complications: Chinese visceral adiposity index and neck circumference. *Cardiovasc Diabetol.* (2020) 19:118. doi: 10.1186/ s12933-020-01095-4

20. Chia YMF, Teng TK, Tan ESJ, Tay WT, Richards AM, Chin CWL, et al. Disparity between indications for and utilization of implantable cardioverter defibrillators in Asian patients with heart failure. *Circ Cardiovasc Qual Outcomes.* (2017) 10:e003651. doi: 10.1161/CIRCOUTCOMES.116.003651

21. Ha K, Nam K, Song Y. A moderate-carbohydrate diet with plant protein is inversely associated with cardiovascular risk factors: the Korea National Health and Nutrition Examination Survey 2013–2017. *Nutr J.* (2020) 19:84. doi: 10.1186/s12937-020-00603-2

22. Goldstein RF, Abell SK, Ranasinha S, Misso ML, Boyle JA, Harrison CL, et al. Gestational weight gain across continents and ethnicity: systematic review and metaanalysis of maternal and infant outcomes in more than one million women. *BMC Med.* (2018) 16:153. doi: 10.1186/s12916-018-1128-1

23. Ciardullo S, Perseghin G. Statin use is associated with lower prevalence of advanced liver fibrosis in patients with type 2 diabetes. *Metabolism*. (2021) 121:154752. doi: 10.1016/j.metabol.2021.154752

24. Li Y, Teng D, Shi X, Qin G, Qin Y, Quan H, et al. Prevalence of diabetes recorded in mainland China using 2018 diagnostic criteria from the American Diabetes Association: national cross sectional study. *BMJ*. (2020) 369:m997. doi: 10.1136/bmj.m997

25. Wan H, Wang Y, Zhang H, Zhang K, Chen Y, Chen C, et al. Chronic lead exposure induces fatty liver disease associated with the variations of gut microbiota. *Ecotoxicol Environ Saf.* (2022) 232:113257. doi: 10.1016/j.ecoenv.2022.113257

26. Zhu X, Fan Y, Sheng J, Gu L, Tao Q, Huang R, et al. Association between blood heavy metal concentrations and dyslipidemia in the elderly. *Biol Trace Elem Res.* (2021) 199:1280–90. doi: 10.1007/s12011-020-02270-0

27. Levey AS, Coresh J, Greene T, Stevens LA, Zhang YL, Hendriksen S, et al. Using standardized serum creatinine values in the modification of diet in renal disease study equation for estimating glomerular filtration rate. *Ann Intern Med.* (2006) 145:247–54. doi: 10.7326/0003-4819-145-4-200608150-00004

28. Shih YH, Howe CG, Scannell Bryan M, Shahriar M, Kibriya MG, Jasmine F, et al. Exposure to metal mixtures in relation to blood pressure among children 5–7 years old: an observational study in Bangladesh. *Environ Epidemiol*. (2021) 5:e135. doi: 10.1097/ EE9.0000000000000135

29. Bobb JF, Claus Henn B, Valeri L, Coull BA. Statistical software for analyzing the health effects of multiple concurrent exposures via Bayesian kernel machine regression. *Environ Health.* (2018) 17:67. doi: 10.1186/s12940-018-0413-y

30. Zhang Y, Dong T, Hu W, Wang X, Xu B, Lin Z, et al. Association between exposure to a mixture of phenols, pesticides, and phthalates and obesity: comparison of three statistical models. *Environ Int.* (2019) 123:325–36. doi: 10.1016/j.envint.2018.11.076

31. Pavuk M, Rosenbaum PF, Lewin MD, Serio TC, Rago P, Cave MC, et al. Polychlorinated biphenyls, polychlorinated dibenzo-p-dioxins, polychlorinated dibenzofurans, pesticides, and diabetes in the Anniston community health survey follow-up (ACHS II): single exposure and mixture analysis approaches. *Sci Total Environ.* (2023) 877:162920. doi: 10.1016/j.scitotenv.2023.162920

32. Luo L, Tian K, Chen Y, Liu Y, Dai M, Gong L, et al. Single and joint associations of exposure to polycyclic aromatic hydrocarbons with blood coagulation function during pregnancy: a cross-sectional study. *Sci Total Environ.* (2023) 885:163949. doi: 10.1016/j.scitotenv.2023.163949

33. Che Z, Jia H, Chen R, Pan K, Fan Z, Su C, et al. Associations between exposure to brominated flame retardants and metabolic syndrome and its components in U.S. adults. *Sci Total Environ.* (2023) 858:159935. doi: 10.1016/j.scitotenv.2022.159935

34. Carrico C, Gennings C, Wheeler DC, Factor-Litvak P. Characterization of weighted quantile sum regression for highly correlated data in a risk analysis setting. J Agric Biol Environ Stat. (2015) 20:100–20. doi: 10.1007/s13253-014-0180-3

35. Chen Y, Pan Z, Shen J, Wu Y, Fang L, Xu S, et al. Associations of exposure to blood and urinary heavy metal mixtures with psoriasis risk among U.S. adults: a cross-sectional study. *Sci Total Environ.* (2023) 887:164133. doi: 10.1016/j. scitotenv.2023.164133

36. Duc HN, Oh H, Kim MS. The effect of mixture of heavy metals on obesity in individuals ≥50 years of age. *Biol Trace Elem Res.* (2022) 200:3554–71. doi: 10.1007/ s12011-021-02972-z

37. King L, Wang Q, Xia L, Wang P, Jiang G, Li W, et al. Environmental exposure to perchlorate, nitrate and thiocyanate, and thyroid function in Chinese adults: a community-based cross-sectional study. *Environ Int.* (2023) 171:107713. doi: 10.1016/j. envint.2022.107713

38. Czarnota J, Gennings C, Wheeler DC. Assessment of weighted quantile sum regression for modeling chemical mixtures and cancer risk. *Cancer Inform.* (2015) 14:159–71. doi: 10.4137/CIN.S17295

39. Czarnota J, Gennings C, Colt JS, De Roos AJ, Cerhan JR, Severson RK, et al. Analysis of environmental chemical mixtures and non-Hodgkin lymphoma risk in the NCI-SEER NHL study. *Environ Health Perspect.* (2015) 123:965–70. doi: 10.1289/ehp.1408630

40. Nguyen HD, Kim MS. Cadmium, lead, and mercury mixtures interact with nonalcoholic fatty liver diseases. *Environ Pollut.* (2022) 309:119780. doi: 10.1016/j. envpol.2022.119780

41. Busgang SA, Spear EA, Andra SS, Narasimhan S, Bragg JB, Renzetti S, et al. Application of growth modeling to assess the impact of hospital-based phthalate exposure on preterm infant growth parameters during the neonatal intensive care unit hospitalization. *Sci Total Environ.* (2022) 850:157830. doi: 10.1016/j. scitotenv.2022.157830

42. Tanner EM, Bornehag CG, Gennings C. Repeated holdout validation for weighted quantile sum regression. *MethodsX*. (2019) 6:2855–60. doi: 10.1016/j.mex.2019.11.008

43. Invernizzi A, Rechtman E, Oluyemi K, Renzetti S, Curtin P, Colicino E, et al. Topological network properties of resting-state functional connectivity patterns are associated with metal mixture exposure in adolescents. *Front Neurosci.* (2023) 17:1098441. doi: 10.3389/fnins.2023.1098441

44. Diks AM, Khatri I, Oosten LEM, de Mooij B, Groenland RJ, Teodosio C, et al. Highly sensitive flow cytometry allows monitoring of changes in circulating immune cells in blood after Tdap booster vaccination. *Front Immunol.* (2021) 12:666953. doi: 10.3389/fimmu.2021.666953

45. Marill KA. Advanced statistics: linear regression, part II: multiple linear regression. *Acad Emerg Med.* (2004) 11:94–102. doi: 10.1111/j.1553-2712.2004.tb01379.x

46. Bobb JF, Valeri L, Claus Henn B, Christiani DC, Wright RO, Mazumdar M, et al. Bayesian kernel machine regression for estimating the health effects of multi-pollutant mixtures. *Biostatistics*. (2015) 16:493–508. doi: 10.1093/biostatistics/kxu058

47. Zhang Y, Xing H, Hu Z, Xu W, Tang Y, Zhang J, et al. Independent and combined associations of urinary arsenic exposure and serum sex steroid hormones among 6–19-year old children and adolescents in NHANES 2013–2016. *Sci Total Environ.* (2023) 863:160883. doi: 10.1016/j.scitotenv.2022.160883

48. Keil AP, Buckley JP, O'Brien KM, Ferguson KK, Zhao S, White AJ. A quantilebased g-computation approach to addressing the effects of exposure mixtures. *Environ Health Perspect*. (2020) 128:47004. doi: 10.1289/EHP5838

49. Guo X, Wu B, Hu W, Wang X, Su W, Meng J, et al. Associations of perchlorate, nitrate, and thiocyanate with metabolic syndrome and its components among US adults: a cross-sectional study from NHANES. *Sci Total Environ.* (2023) 879:163083. doi: 10.1016/j.scitotenv.2023.163083

50. Wan H, Jiang Y, Yang J, Ma Q, Liu L, Peng L, et al. Sex-specific associations of the urinary fourteen-metal mixture with NAFLD and liver fibrosis among US adults: a nationally representative study. *Ecotoxicol Environ Saf.* (2022) 248:114306. doi: 10.1016/j. ecoenv.2022.114306

51. Shi Y, Wang H, Zhu Z, Ye Q, Lin F, Cai G. Association between exposure to phenols and parabens and cognitive function in older adults in the United States: a cross-sectional study. *Sci Total Environ.* (2023) 858:160129. doi: 10.1016/j. scitotenv.2022.160129

52. Zhang Y, Liu Y, Qiu H. Association between dietary zinc intake and hyperuricemia among adults in the United States. *Nutrients*. (2018) 10:568. doi: 10.3390/nu10050568

53. Sun X, Wen J, Guan B, Li J, Luo J, Li J, et al. Folic acid and zinc improve hyperuricemia by altering the gut microbiota of rats with high-purine diet-induced hyperuricemia. *Front Microbiol.* (2022) 13:907952. doi: 10.3389/fmicb.2022.907952

54. Wang T, Lv Z, Wen Y, Zou X, Zhou G, Cheng J, et al. Associations of plasma multiple metals with risk of hyperuricemia: a cross-sectional study in a mid-aged and older population of China. *Chemosphere*. (2022) 287:132305. doi: 10.1016/j. chemosphere.2021.132305

55. Huang C, Gao E, Xiao F, Wu Q, Liu W, Luo Y, et al. The relative and interactive effects of urinary multiple metals exposure on hyperuricemia among urban elderly in China. *Front Public Health.* (2023) 11:1015202. doi: 10.3389/fpubh.2023.1015202

56. King JC, Shames DM, Woodhouse LR. Zinc homeostasis in humans. J Nutr. (2000) 130:1360S–6S. doi: 10.1093/jn/130.5.1360S

57. Ma Y, Hu Q, Yang D, Zhao Y, Bai J, Mubarik S, et al. Combined exposure to multiple metals on serum uric acid in NHANES under three statistical models. *Chemosphere*. (2022) 301:134416. doi: 10.1016/j.chemosphere.2022.134416

58. Sun X, Liu Y, Wang Z, Li H, Liu Y, Di D. Effect of occupational manganese exposure on uric acid levels in human urine. *Biomed Environ Sci.* (2011) 24:222–7. doi: 10.3967/0895-3988.2011.03.003

59. Roth JA. Homeostatic and toxic mechanisms regulating manganese uptake, retention, and elimination. *Biol Res.* (2006) 39:45–57. doi: 10.4067/S0716-97602006000100006

60. Wirdefeldt K, Adami HO, Cole P, Trichopoulos D, Mandel J. Epidemiology and etiology of Parkinson's disease: a review of the evidence. *Eur J Epidemiol*. (2011) 26:1–58. doi: 10.1007/s10654-011-9581-6

61. Alonso A, Rodriguez LA, Logroscino G, Hernan MA. Gout and risk of Parkinson disease: a prospective study. *Neurology*. (2007) 69:1696–700. doi: 10.1212/01. wnl.0000279518.10072.df

62. Annanmaki T, Muuronen A, Murros K. Low plasma uric acid level in Parkinson's disease. *Mov Disord*. (2007) 22:1133–7. doi: 10.1002/mds.21502

63. Bo S, Durazzo M, Gambino R, Berutti C, Milanesio N, Caropreso A, et al. Associations of dietary and serum copper with inflammation, oxidative stress, and metabolic variables in adults. *J Nutr.* (2008) 138:305–10. doi: 10.1093/jn/138.2.305

64. Maiuolo J, Oppedisano F, Gratteri S, Muscoli C, Mollace V. Regulation of uric acid metabolism and excretion. *Int J Cardiol.* (2016) 213:8–14. doi: 10.1016/j. ijcard.2015.08.109

65. Hadizadeh M, Keyhani E, Keyhani J, Khodadadi C. Functional and structural alterations induced by copper in xanthine oxidase. *Acta Biochim Biophys Sin Shanghai*. (2009) 41:603–17. doi: 10.1093/abbs/gmp048

66. Zhang Y, Qiu H. Dietary magnesium intake and hyperuricemia among US adults. *Nutrients*. (2018) 10:296. doi: 10.3390/nu10030296

67. Gu T, Cao G, Luo M, Zhang N, Xue T, Hou R, et al. A systematic review and metaanalysis of the hyperuricemia risk from certain metals. *Clin Rheumatol.* (2022) 41:3641–60. doi: 10.1007/s10067-022-06362-1

68. Cao Y, Wang C, Guan K, Xu Y, Su YX, Chen YM. Association of magnesium in serum and urine with carotid intima-media thickness and serum lipids in middle-aged and elderly Chinese: a community-based cross-sectional study. *Eur J Nutr.* (2016) 55:219–26. doi: 10.1007/s00394-015-0839-8

69. Xu Y, Dong H, Zhang B, Zhang J, Ma Q, Sun H. Association between dyslipidaemia and the risk of hyperuricaemia: a six-year longitudinal cohort study of elderly individuals in China. *Ann Med.* (2022) 54:2402–10. doi: 10.1080/07853890.2022.2118368

70. Akbari A, Razmi M, Rafiee M, Watts GF, Sahebkar A. The effect of statin therapy on serum uric acid levels: a systematic review and meta-analysis. *Curr Med Chem.* (2023). doi: 10.2174/0929867330666230207124516