



# Subjective Impact of the COVID-19 Pandemic on Schizotypy and General Mental Health in Germany and the United Kingdom, for Independent Samples in May and in October 2020

Sarah Daimer<sup>1</sup>, Lorenz Mihatsch<sup>2</sup>, Lisa Ronan<sup>3</sup>, Graham K. Murray<sup>3,4</sup> and Franziska Knolle<sup>1,3\*</sup>

<sup>1</sup> Department of Diagnostic and Interventional Neuroradiology, School of Medicine, Technical University of Munich, Munich, Germany, <sup>2</sup> Institute for Medical Microbiology, Immunology and Hygiene, Technical University of Munich, Munich, Germany, <sup>3</sup> Department of Psychiatry, University of Cambridge, Cambridge, United Kingdom, <sup>4</sup> Cambridgeshire and Peterborough NHS Foundation Trust, Cambridge, United Kingdom

## OPEN ACCESS

### Edited by:

Ana Tajadura-Jiménez,  
Universidad Carlos III de Madrid,  
Spain

### Reviewed by:

Phillip Grant,  
Fresenius University of Applied  
Sciences, Germany  
Sara Marelli,  
San Raffaele Hospital (IRCCS), Italy

### \*Correspondence:

Franziska Knolle  
franziska.knolle@tum.de

### Specialty section:

This article was submitted to  
Personality and Social Psychology,  
a section of the journal  
Frontiers in Psychology

**Received:** 14 February 2021

**Accepted:** 21 June 2021

**Published:** 28 July 2021

### Citation:

Daimer S, Mihatsch L, Ronan L,  
Murray GK and Knolle F (2021)  
Subjective Impact of the COVID-19  
Pandemic on Schizotypy and General  
Mental Health in Germany  
and the United Kingdom,  
for Independent Samples in May  
and in October 2020.  
Front. Psychol. 12:667848.  
doi: 10.3389/fpsyg.2021.667848

Studies reported a strong impact on mental health during the first wave of the COVID-19 pandemic in March–June, 2020. In this study, we assessed the impact of the pandemic on mental health in general and on schizotypal traits in two independent general population samples of the United Kingdom (May sample N: 239, October sample N: 126; participation at both timepoints: 21) and in two independent general population samples of Germany (May sample N: 543, October sample N: 401; participation at both timepoints: 100) using online surveys. Whereas general psychological symptoms (global symptom index, GSI) and percentage of responders above clinical cut-off for further psychological investigation were higher in the May sample compared to the October sample, schizotypy scores (Schizotypal Personality Questionnaire) were higher in the October sample. We investigated potential associations, using general linear regression models (GLM). For schizotypy scores, we found that loneliness, use of drugs, and financial burden were more strongly corrected with schizotypy in the October compared to the May sample. We identified similar associations for GSI, as for schizotypy scores, in the May and October samples. We furthermore found that living in the United Kingdom was related to higher schizotypal scores or GSI. However, individual estimates of the GLM are highly comparable between the two countries. In conclusion, this study shows that while the general psychological impact is lower in the October than the May sample, potentially showing a normative response to an exceptional situation; schizotypy scores are higher at the second timepoint, which may be due to a stronger impact of estimates of loneliness, drug use, and financial burden. The ongoing, exceptional circumstances within this pandemic might increase the risk for developing psychosis in some individuals. The development of general psychological symptoms and schizotypy scores over time requires further attention and investigation.

**Keywords:** COVID – 19, schizotypy, SPQ, depression, anxiety, SCL-27, mental health, psychosis

## INTRODUCTION

The highly infectious severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2) had developed into an ongoing worldwide pandemic by March 2020 precipitating a global health crisis with nearly 150 million cases and over 3 million deaths by the end of April 2021 (Daly and Robinson, 2021; JHU, 2021). Due to the high risk of infection and the rapid spread of the virus, governments across the world were compelled to implement restrictions and social distancing measures to keep the number of cases and hospitalizations as low as possible. The main aim of this strategy was to prevent the health care system from being overburdened and gain time to develop treatments and vaccines (Han et al., 2020; Kissler et al., 2020). This led to unprecedented changes to everyday life for the people all around the world. In many countries, people were forced to withdraw from usual face-to-face social activities on a large scale, and schools, nurseries, and retailers as well as workplaces were closed at least for some weeks, with workers being required to work at home. The number of permitted social contacts was limited (Kissler et al., 2020). In many countries, restrictions and lock-down measures are still in place in April 2021. The exact measures taken by countries differed vastly, and even countries within Europe with similar developments throughout the pandemic used different strategies in order to deal with the hitherto unknown situation. According to Plümper and Neumayer (2020) government-strategies can be differentiated based on two dimensions: the time to response and the level of stringency of the lockdown policy. Germany, for example, went into lockdown rapidly in Spring 2020 and managed to control the increase of infections efficiently, whereas the United Kingdom delayed lockdown and faced a much higher plateau (Balmford et al., 2020). At the beginning of the pandemic, in March 2020, the government in the United Kingdom pondered with the idea of implementing what has since become known as the Swedish strategy, which avoids a lockdown and allows a relatively high number of infections, in order to reach herd immunity (Plümper and Neumayer, 2020). These strategies might have substantially contributed to the variation in numbers of cases and deaths in each country.

From the start of the pandemic, the World Health Organization, many researchers and clinicians communicated warnings about the consequences of mitigation and suppression measures on mental well-being and mental health (Pfefferbaum and North, 2020; WHO, 2020; Xiong et al., 2020). As expected, the severe restriction of social contacts as well as the fear of the virus or the impact on living conditions had a measurable impact on the mental health of general populations all over the world (Bu et al., 2020; Smith et al., 2020; Xiong et al., 2020). During the first lockdown increased levels of perceived stress and mental distress, COVID-19 related fear, general anxiety and depression and a general decline in mental wellbeing were measured in many countries including, Germany and the United Kingdom (Bäuerle et al., 2020; Smith et al., 2020; Proto and Quintana-Domeque, 2021). Female gender, younger age, being part of an ethnic minority and a low socioeconomic status were associated with a high risk for experiencing mental distress (Bäuerle et al., 2020; Fancourt et al., 2020; Simha et al., 2021). However, also living in

a specific country was associated with lower stress: for example Adamson et al. (2020) found higher perceived stress levels in the United Kingdom than in Germany.

The results of our own study from April–May 2020, confirmed these findings revealing a higher psychological and socioeconomic impact of the pandemic on people resident in the United Kingdom vs. Germany (Knolle et al., 2021). However, both countries reported similarly strong subjective ratings of symptom worsening, with 25% of all responders reported increased levels of anxiety and depressive symptoms [Symptom Check List, SCL-27 (Hardt and Gerbershagen, 2001)], and nearly 10% reported worsening of schizotypy traits measured with Schizotypal Personality Questionnaire [SPQ (Raine, 1991)]. Especially, the findings on the subjective worsening of schizotypy measured with the SPQ are interesting, as to our knowledge no other study has investigated schizotypy traits in the general population within the scope of the current pandemic.

Schizotypy describes a latent personality trait, thought to reflect an underlying vulnerability of developing psychosis or schizophrenia-spectrum disorders (Chapman et al., 1994; Debbané and Barrantes-Vidal, 2015; Schultze-Lutter et al., 2019), though we note that different scholars have conceptualized it in differing ways (Grant et al., 2018). Although the SPQ is commonly used as a measure of schizotypy in the general population, its design was based on diagnostic symptoms of schizotypal personality disorder (Oezgen and Grant, 2018). It can, therefore, be considered as measuring related but not identical traits to other schizotypal questionnaires. Based on a recent review by Preti et al. (2020) the current pandemic poses an especially large risk for people suffering from paranoid or high schizotypal traits, as the measures taken to prevent the spread of the virus might ultimately lead to increased anxiety and depressive symptoms, increased avoidance behaviors, stronger disruption of social contacts, and delayed return to normality in these individuals. Furthermore, studies show links between recent adverse life events (Beards et al., 2013; Betz et al., 2020) or isolation and loneliness (Chau et al., 2019; Le et al., 2019) and schizotypy or psychosis-like experiences. Both these aspects, loneliness and adverse life events, are present in the current pandemic, which might have a worsening effect on schizotypal trait expression in people with pre-existing high schizotypy scores, perhaps leading to increased distress or disability. Additionally, the ongoing uncertainty and the impact on social routines might lengthen the time it takes schizotypal high-scores to return to baseline levels. Preliminary evidence and case reports show an increase in the development of first-episode psychosis linked to the pandemic (Huarcaya-Victoria et al., 2020; Valdés-Flórido et al., 2020) and reactive psychotic disorders in previously healthy individuals (Valdés-Flórido et al., 2020) following the months after start of the pandemic.

When incidence levels of infections decreased during the summer and, as a result, relaxations of the restrictions were initiated (Han et al., 2020; Hetkamp et al., 2020), this also positively impacted the reported mental health status in the general population across both countries. Some studies found a reduction of these scores to a level comparable before the pandemic (Hetkamp et al., 2020; Wang et al., 2020) while others

measured elevated, but no longer worsening, levels (Fancourt et al., 2020; O'Connor et al., 2020; Daly and Robinson, 2021). *Post hoc* comparisons between countries can be challenging due to the use of different questionnaires by different research groups, or different overall developments of the progression of the pandemic and governmental responses. In this study we therefore, investigated the association of the COVID-19 pandemic and the accompanying lockdown with mental health comparing independent samples of the general population of United Kingdom and Germany at two timepoints – the first one during the first lockdown (April/May 2020) and the second after the summer (September/October 2020) when a majority of restrictions had been lifted. Specifically, we examined whether reported levels of depressive symptoms and anxiety, and, in particular, schizotypal scores would change over the summer following the reduction of the restrictions, using the same questionnaire as in the first timepoint. Consistent with other studies, and due to the easing of the restrictions over the summer in Germany and the United Kingdom, we hypothesized to detect lower levels of anxiety and depression in the October compared to the May sample. In contrast, we predicted that SPQ-scores would be similar across the first and the second timepoints, as we expected that the return to baseline levels would take longer for schizotypal traits. In addition, all mental health scores were compared between the Germany and the United Kingdom to provide insight into the impact of political action on the well-being of the population. For clarity, we wish to emphasize that the design of this study is not longitudinal, rather we assess the impact of the pandemic at two different timepoints in highly comparable but different, only partially overlapping samples.

## MATERIALS AND METHODS

### Study Design and Procedure

The questionnaire used in this study assessing mental and psychological health and COVID-19 exposure was designed as an online survey using EvaSys<sup>1</sup>(Electric Paper Evaluationssysteme GmbH, Luneburg, Germany). The questionnaire was available in German and English. For participant recruitment we used a snowball sampling strategy to reach the general public. For the first timepoint, data collection took place from 27/04/2020 to 31/05/2020 and completion of the survey took approximately 35 min; for the second timepoint, data collection took place from 10/09/2020–18/10/2020, and the completion of the survey took approximately 15 min. For each psychological item, the first timepoint survey included a “before the pandemic” or an evaluation of whether or not item strength had changed, which approximately double the time it took to complete the survey. Participation was voluntary. Participants did not receive any compensation.

Ethical approval was obtained from the Ethical Commission Board of the Technical University Munich (250/20 S). All participants provided informed consent.

<sup>1</sup><https://www.evasys.de>

### Outcome Variables

As described in Knolle et al. (2021) in detail, the survey consisted of three parts. The first part, partially comprised of the Coronavirus Health Impact Survey (CRISIS<sup>2</sup>, which assessed demographics (age, gender (not biological sex), education and parental education, living conditions), COVID-19 exposure [infection status, symptoms, contact), mental and physical health questions. In the second part, we assessed the general mental health status (global severity of symptom index (GSI)) using the Symptom Check List (SCL) with 27 items (Hardt and Gerbershagen, 2001; Hardt et al., 2011). The GSI score describes the total expression of symptom strength over all SCL-27 items, combining measures of anxiety, depression, mistrust and vegetative symptoms. Furthermore, we assessed the specific sub-scores verified by Hardt et al. (2011); the sub-scores were dysthymic symptoms, depressive symptoms, symptoms of social phobia, symptoms of mistrust, agoraphobic symptoms, and vegetative symptoms. In the third part, using the dichotomous version of the Schizotypy Personality Questionnaire [SPQ, (Raine, 1991)] we evaluated total schizotypy trait (SPQ-total). We also assessed the subdimensions, using a six-factor model abnormal experiences and beliefs, social anhedonia, paranoid ideation, social anxiety, eccentricity, and disorganized speech (Davies, 2017), as well as the original nine-factor model (Raine, 1991), the three-factor model (Raine et al., 1994) and the four-factor model (Stefanis et al., 2004). During the first timepoint of data collection, we also assessed subjective measures of change for questions on life circumstances, mental health and psychological traits, asking participants to either report on that particular question before the pandemic or report whether the evaluation of the item had increased, decrease or stayed the same.

### Statistical Analysis

Statistical analysis and visualizations were computed using R and R Studio (R Core Team, 2016; R Studio Team, 2020). We first describe demographics and COVID-19 exposure variables, using frequency analysis. For the country comparison we used Wilcoxon rank sum tests or Chi-square test of independence to explore differences between the countries (United Kingdom, Germany) and timepoints (April/May and September/October 2020) on the demographics and the COVID-19 exposure variables.

To further explore the differences between the countries and timepoints in CRISIS variables we conducted robust ANOVAS (Mair and Wilcox, 2020) with country (United Kingdom, Germany) and timepoint [before pandemic (i.e., subjective rating acquired during the first timepoint), April/May, September/October] as between-subjects factor.

In order to identify possible negative associations for experiencing increased general strain and mental distress, we applied Gaussian regression models to assess the connection between the outcome and predictor variables. Our outcome variable was GSI, the total expression of psychological symptoms. In the first basic model, we explored the relationship of

<sup>2</sup><http://www.crisissurvey.org/>

demographic variables (age, gender, education, country of residence, living area, and parenthood) and prior physical and mental health problems with GSI scores. The second model – the harmful and healthy behavior model – was used to investigate the link between healthy and harmful behaviors and the outcome variable. For this purpose, we added sleeping hours per night, days with physical exercise per week, drug, alcohol, media consumption and the degree of perceived loneliness to the basic model conducted previously as predictor variables. In the third model – the COVID-19 context model – we examined the coherence of COVID 19 pandemic and associated restrictions using the following variables: perception of the burden of restrictions, stressful relationship changes, financial impact of the Pandemic, hope for a soon end and suspicion of COVID-19 disease. In addition, we also included the degree of trust in government in the October survey.

We investigated possible associations for expression of schizotypy using Poisson regression models. The dependent variable was the total SPQ-score. We used the same three models as for the investigation of the SPQ as for the GSI. We correct for multiple comparisons in all six models, using an adjusted *p*-value of 0.0083 as the threshold for significance. Models were run for samples separately, using different only partially overlapping samples.

In order to investigate stressful changes in social and family relationships, we used the sum of the degrees of stress experienced in the deterioration in social and family relationships. Excessive media consumption received a positive expression if at least one of the categories of media consumption (television, social media or videogames) was used for more than 4 h per day. The drug score was calculated on the basis of at least one use of marijuana, tranquilizers or other drugs like heroin or other opiates.

## RESULTS

### Whole Sample Description

The first survey (May 2020) was completed by 860 participants. Two participants did not provide consent and were excluded. 6 participants did not consent to sharing the data publicly, and are included in the analysis but will be removed from the open-access data set. In this paper we focused on the comparison of responders living in the United Kingdom (*N* = 239) and in Germany (*N* = 543). In the first survey the majority of respondents were female (72%), 25% were male, 3% diverse or did not provide the information. The age ranged from 18 to 92 years, with a mean 43 (SD 15.5) years and a median of 41 years. The majority of participants were well educated, 60% had a master's degree or higher, and 25% had completed a professional college or a bachelor degree. 48% reported to live in large cities, 12% in suburbs of large cities, 19% in small cities, and 21% in rural areas (see **Table 1**).

The second survey (October 2020) was completed by 550 participants. 22 were excluded from the analysis as they gave not consent to the participation and one he or she did not provide information about the current residency. 69% of the participants

were female, 25% male and 6% did not provide the information. The age ranged from 18 to 93 years (*M* = 42, *SD* = 16.1). Most of the participants had their current residency in Germany (76%, *N* = 401) and 24% in the United Kingdom (*N* = 126). The majority of the sample were well educated, with 45% reporting to have a master's degree or higher and 39% to have completed a professional college or bachelor's degree. 39% lived in a city, 13% in suburbs, 18% in towns and 28% in villages or rural areas (see **Table 1**).

Since participation at the first survey was not required for taking part in the second survey, the two samples are partially overlapping. 121 responders participated in both surveys. The samples did not differ significantly in terms of age ( $X^2 = 100.8$ ,  $p = 0.989$ ) and gender ( $W = 192786$ ,  $p = 0.635$ ) between the timepoints, but in the second survey significantly more participants came from Germany ( $X^2 = 8.55$ ,  $p = 0.014$ ), their living area was more rural ( $W = 182331$ ,  $p = 0.002$ ) and the sample was less educated ( $W = 218690$ ,  $p = 0.018$ ).

### COVID 19 Infection

At the first timepoint, 0.2% of the German sample reported a positive COVID-19 test result, 0.7% reported a diagnosis made by a health care professional without using a test, but based on symptoms and contact to COVID-19 positive individuals, and 14.4% possible symptoms of a COVID-19 infection. 83.8% stated that they had not suspected COVID-19. In the United Kingdom sample, 2.5% reported a medical diagnosis of COVID-19 made by a health care professional without using a test, but based on symptoms and contact to COVID-19 positive individuals, and 18.8% of possible symptoms. None of the respondents reported having received a positive test result. 78.7% reported that they had not previously suspected COVID-19.

At the second timepoint, 0.8% of the German sample reported being positively tested for COVID-19, 0.5% reported receiving a positive diagnosis, without a test, and 16% reported symptoms that could indicate a COVID-19 infection. 82.0% reported not having suspected COVID-19. In the United Kingdom sample, 0.8% reported having received a positive diagnosis, without a test, of COVID-19 and 20.6% had recently experienced symptoms of COVID-19 infection. None of the United Kingdom respondents reported having received a positive test result. 77.0% reported no signs of possible COVID-19 infection (see **Table 1**).

### Results of Robust ANOVAs

Global symptom index scores differed significantly between the two countries ( $p < 0.001$ ) and samples ( $p = 0.04$ ) with higher GSI scores in the May sample and higher scores in the United Kingdom sample compared to the German sample (see **Figure 1**). There was no interaction effect. The results of the robust ANOVAs are shown in **Table 2**. Furthermore, we investigated the development in the small sample of those responders who took part in the survey at both timepoints. The results are similar and are presented in the **Supplementary Materials**.

**TABLE 1** | Demographics and suspected infection among the samples divides by country.

		April – May sample		September – October sample		Difference between GER/United Kingdom		Difference between samples	
		United Kingdom	GER	United Kingdom	GER	W/X <sup>2</sup>	p	W/X <sup>2</sup>	p
N		239	543	126	401			192346	0.008**
Prozent		30.6%	69.4%	23.91%	76.09%				
Age	Mean	39	45.4	40.90	42.67	237.2	<0.001***	100.9	0.989
	SD	16	14.9	16.17	1.14				
Gender	Female	73.6%	71.3%	65.87%	70.32%	159158	0.865	190794	0.786
	Male	24.3%	25.8%	28.57%	24.19%				
	other/NA	2.1%	3%	5.56%	5.49%				
Education	School leavers	0.4%	-	-	-	162346	0.207	218690	0.017*
	8-years	19.3%	13.1%	15.87%	14.46%				
	Prof. college/Bachelor	31.8%	21.6%	38.89%	30.42%				
	Master or higher	47.3%	65%	44.44%	54.61%				
	Missing	1.3%	0.4%	0.79%	0.50%				
Children	Yes	21.3%	30.8%	28.57%	30.17%	155440	0.013*	196510	0.540
	Missing	1.3%	2.6%	0.79%	1.00%				
Living Area	City	20.5%	60.2%	26.19%	42.89%	227803	<0.001***	182331	0.002**
	Suburb	8%	13.1%	11.90%	13.97%				
	Town	36.4%	10.9%	27.78%	14.71%				
	Village or rural Area	34.7%	15.7%	33.33%	25.94%				
	Missing	0.4%	18%	79.00%	2.49%				
Suspected infection	Positive Test	-	0.2%	-	0.75%	160840	0.024*	205006	0.614
	Diagnosis	2.5%	0.7%	0.79%	0.50%				
	Symptoms	18.8%	14.4%	20.63%	15.96%				
	No infection	78.7%	83.8%	76.98%	82.04%				
	Missing	-	0.9%	1.59%	0.75%				

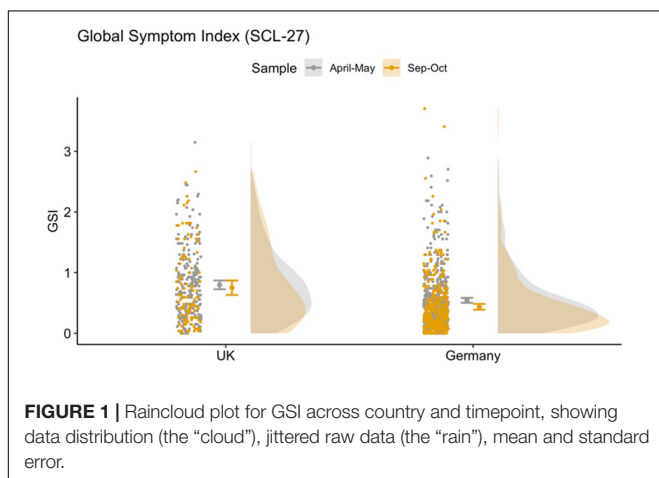
$p < 0.100$ , \* $p < 0.050$ , \*\* $p < 0.010$ , \*\*\* $p < 0.001$ .

In a norm sample (Hardt et al., 2004), 10–15% of the screened population reach the clinical cut-off on the different sub-dimensions, and require additional psychological investigation. As shown in **Figure 2**, there are significant differences between the countries [ $F(1, 7464 = 237.96, p < 0.001)$ ] and samples [ $F(1, 7464 = 12.58, p < 0.001)$ ]. For the sub-dimension of dysthymic symptoms (DYS), the rate fell from 68.38 to 58.82% in the

United Kingdom responders and from 37 to 30% in the German responders who lay above the clinical cut-off; for depressive symptoms (DEP) from 51 to 50% in the United Kingdom and from 39 to 27% in the German responders; for symptoms of social phobia (SOP) from 37 to 34% in the United Kingdom and from 24 to 19% in German responders; for symptoms of mistrust (MIS) from 29 to 26% in the United Kingdom and from 26 to 22% in the German sample; and for agoraphobic symptoms (AGO) from 52 to 32% in the United Kingdom and from 23 to 12% in the German responders. Interestingly, the vegetative symptoms (VEG) increased from 26 to 35% in the United Kingdom and from 14 to 16% in the German responders. The reduced rates were measured in a comparable but different only partially overlapping sample.

SPQ scores (**Figure 3**) also differed significantly between countries with higher scores in the United Kingdom sample ( $p = 0.01$ ). We found a trend toward higher SPQ-scores in the October compared to the May sample ( $p = 0.06$ ). The results of the robust ANOVAS are shown in **Table 2**. Furthermore, we investigated the development of the total SPQ score in the small sample of those responders who took part in the survey at both timepoints. The results are similar and are presented in the **Supplementary Materials**.

Additionally, we investigated the different SPQ subdimensions in an explorative analysis. For completion, we present four



**TABLE 2 |** Overview of means and robust ANOVAS of GSI and SPQ scores, all CRISIS variables and questions concerning life changes due to COVID.

	Mean						Robust ANOVA/M-estimator			Post hoc						
	Before		April - May		September - October		significance			Country (sample comparison)	Before (country comparison)	April - May (country comparison)	Sept. - Oct. (country comparison)	Country x Before	Country x April - May	Country x - Oct. - Sept. -
	United Kingdom	GER	United Kingdom	GER	United Kingdom	GER	ME C	ME TP	Country x sample	Psi	Psi	Psi	Psi	Psi	Psi	Psi
GSI			0.8	0.54	0.75	0.44	0.000 ***	0.043 *	0.973	0.6***			0.2*			0.0
SPQ			13.6	12.25	17.36	13.20	0.011**	0.061 *	0.240	5.7**			-3.8*			-2.5.
<b>Behavior</b>																
Sleep week	2.06	2.06	2.1	2.07	1.86	2.02	i									
Sleep weekend	2.41	2.36	2.3	2.33	2.15	2.33	0.118	0.197	0.232	0.3	-0.3.	0.0.	0.3	-0.3	0.0.	0.3
Exercise	2.84	2.46	3.0	2.55	2.48	2.29	0.000 ***	0.000 ***	0.312	1.0***	-0.2.	0.5***	0.8***	0.0	0.2.	0.2.
Outside	3.89	3.88	3.6	3.00	3.64	3.90	0.121	0.258	0.534	-0.5	0.4**	0.2.	-0.2	0.2*	0.3 *	0.0
<b>Cognition</b>																
Happy content	3.48	3.48	2.8	2.75	2.90	3.11	0.183	0.000 ***	0.250	-0.2.	1.5***	1.0***	-0.5**	0.0	0.2.	0.2.
Concerned	2.23	2.28	2.9	3.10	2.59	2.59	0.098	0.000 ***	0.161	-0.3 *	-1.5***	-0.7***	0.8**	0.2	-0.0	-0.2.
Enjoy activities	3.65	3.86	2.6	3.02	2.79	3.45	0.000 ***	0.000 ***	0.028	-1.3**	1.9**	1.3***	-0.6**	0.2.	0.5***	0.3*
Relaxed	2.91	2.43	3.4	2.89	3.29	2.67	0.001 **	0.000 ***	0.794	1.6**	-0.9**	-0.6***	0.3	0.0	-0.2	-0.2
Restless	1.95	1.89	2.3	2.20	2.25	2.02	0.015 **	0.001 **	0.490	0.4**	-0.7**	-0.4**	0.3*	-0.1	-0.2	-0.1
Tired	2.57	2.62	2.8	2.78	2.74	2.83	0.400	0.006 *	0.858	-0.2	-0.4**	-0.4**	0.0	0.0	0.0	0.1
Focused	2.43	2.29	3.3	2.72	2.97	2.55	0.000 **	0.000 **	0.016	1.1**	-1.3**	-0.8**	0.5*	-0.4**	-0.3*	0.1
Irritated	2.02	2.37	2.5	2.72	2.37	2.49	0.000 **	0.000 **	0.140	-0.66	-0.9**	-0.5**	0.4**	-0.2.	-0.2.	-0.1
Lonely	1.64	1.73	2.2	2.30	2.43	2.04	0.607	0.000 **	0.559	-0.1	-2.5**	-2.1**	0.4*	0.1	0.0	-0.1
Negative Thoughts	2.65	2.55	2.9	2.94	2.82	2.70	0.177	0.000 **	0.607	0.23	-0.7**	-0.3**	0.4*	0.1	0.0	-0.1
<b>Media consumption</b>																
TV digital Media	2.8	2.8	3.26	3.10	3.18	2.94	0.384	0.384	0.384	0.0	0.0**	0.0*	0.0.	0.0**	0.0*	0.0
Social Media	2.5	2.1	2.81	2.42	2.73	2.46	0.000 **	0.026 *	0.554	1.1**	-0.7**	-0.6.	0.0	0.0	0.1.	0.1
Videogame	1.3	1.2	1.55	1.28	1.53	1.26	i									
Print media	2.0	2.4	2.13	2.57	2.20	2.42	0.000 **	0.010 *	0.041 *	-1.5*	-0.8*	-0.3*	0.5	0.4	-0.2*	-0.7*
<b>Substance use</b>																
Alcohol	4.3	3.9	4.36	4.16	4.24	3.84	0.000 **	0.517	0.878	3.4**	-0.6	0.1	0.7.	-0.2	-0.1	0.07
Tobacco	1.5	1.8	1.39	1.93	1.93	1.86	i									
Marihuana	1.2	1.2	1.21	1.19	1.47	1.13	i					0.1				
Opiate Heroin	1.0	1.0	1.01	1.02	1.09	1.05	i									
<b>Life changes</b>																
Time Outside			3.47	3.83	3.71	4.11	0.111	0.151	0.811	-0.8**	-0.5*					0.1
Restrictions stressful			2.92	2.86	2.78	2.59	0.071.	0.008*	0.340	0.3*	0.4**					-0.1
Quality of family relationships			3.23	3.04	3.14	2.96	0.043 *	0.415	0.739	0.3*						0.1

(Continued)

TABLE 2 | Continued

	Mean				Robust ANOVA/M-estimator				Post hoc					
	Before		September - October		significance		Country (sample comparison)		Before (country comparison)		April - May (country comparison)		Sept. - Oct. (country comparison)	
	United Kingdom	GER	United Kingdom	GER	ME	TP	ME	C	TP	ME	C	TP	ME	C
	Country x sample	Country x sample	Country x sample	Country x sample	Country x sample	Country x sample	Country x sample	Country x sample	Country x sample	Country x sample	Country x sample	Country x sample	Country x sample	Country x sample
Change of family (stressful)	2.84	2.47	2.16	2.11	0.038*	0.020*	0.080	0.4*	1.0**	0.3.				
Quality social of Relationships	2.79	2.71	2.43	2.52	0.280	0.000**	0.045	0.2	0.8**	0.4.				
Change of social relations (stressful)	3.05	2.57	2.28	2.23	0.008*	0.000**	0.025	0.5**	1.1**	0.4**				
Financial impact	2.05	1.86	1.82	1.60	0.000**	0.457	0.457	1.9.	0.2	0.2				
Concerned life stability	2.04	2.44	1.86	2.08	0.000**	0.011*	0.265	-1.5**	1.4*	0.7				
No money for food	0.03	0.03	0.04	0.03										
Hopeful for a soon end	2.91	2.51	1.78	2.03	0.013*	0.000**	0.000	-0.6*	2.4**	1.4**				

*psi* = *psihat*, C = Country,  $p < 0.10$ , \* $p < 0.05$ , \*\* $p < 0.01$ , \*\*\* $p < 0.001$ , I = interaction, GER = Germany, UK = United Kingdom, *i* = could not be calculated due to insufficient dispersion or change.

different categorizations based on the six-factor model (Davies, 2017; **Figure 4A**), the original nine factor model (Raine, 1991; **Figure 4B**), the three-factor model (Raine et al., 1994; **Figure 4C**) and the four-factor model (Stefanis et al., 2004; **Figure 4D**). Using a robust ANOVA, we analyzed differences across country of residence and samples. All results are presented in the **Supplementary Table 1**. Factors of social anhedonia and social anxiety (Raine, 1991; Davies, 2017), as well as the interpersonal factor (Raine et al., 1994; Stefanis et al., 2004) show the strongest endorsement in the two countries. The overall trend is similar across all four approaches, revealing significant differences between the two countries and showing higher scores for most subdimension scores in the United Kingdom.

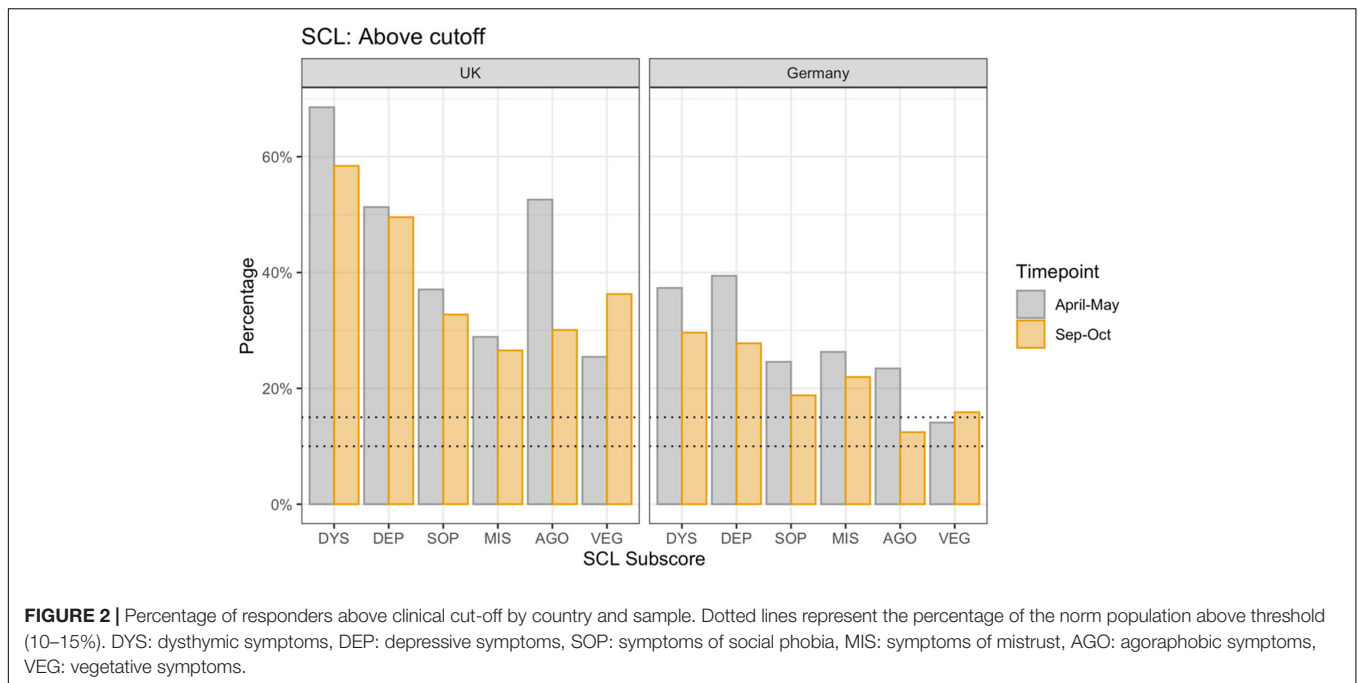
### General Linear Model: General Psychological Symptom Index Effects of Demographic Variables and Prior Physical and Mental Health Problems on GSI Scores (Basic Model)

In both surveys in May and October, age (TP1:  $p < 0.001$ , TP2:  $p = 0.001$ ), country of residence (TP1  $p = 0.005$ , TP2  $p < 0.001$ ) and pre-existing physical (TP1:  $p = 0.002$ , TP2:  $p = 0.005$ ) and mental health problems (TP1:  $p < 0.001$ , TP2:  $p < 0.001$ ) were significantly associated with GSI. Older age and country of residence in Germany are related to lower scores, while the opposite was shown for pre-existing physical and mental health problems. Female gender was associated with lower GSI scores in the May sample (TP1:  $p = 0.031$ , TP2:  $p = 0.779$ ) and higher levels of education was related to lower scores in the October sample (TP1:  $p = 0.962$ ,  $p < 0.001$ ). In the first survey we found that living in a town was significantly connected to higher GSI compared to living in a large city ( $p = 0.045$ ). There was no significant association between having children at home and the outcome in both samples in May and October (TP1:  $p = 0.256$ , TP2:  $p = 0.439$ ) see **Table 3**.

Based on these results, we adjusted for the predictors age, gender, education level, country of residence, having children at home, as well as physical and mental health problems in the following two models. Additionally, all models have been run separately on the small overlapping sample, see **Supplementary Materials**. Results are comparable to the full sample.

### Effects of Harmful and Healthy Behaviors GSI Scores (Harmful and Healthy Behavior Model)

After adjusting for possibly confounding demographic variables from the previous basic model, in the May and October samples drug consumption was associated with higher GSI scores (TP1:  $p = 0.005$ , TP2:  $p < 0.001$ ), as well as excessive media use (TP1:  $B = 0.12$ ,  $t = 3.34$ ,  $p < 0.001$ , TP2:  $B = 0.13$ ,  $t = 2.38$ ,  $p < 0.001$ ). No correlation was found for alcohol consumption (TP1:  $p = 0.602$ , TP2:  $p = 0.740$ ). Sleeping between 6 and 8 h (TP1:  $p < 0.001$ , TP2:  $p < 0.001$ ) and more than 8 h (TP1:  $p < 0.001$ , TP2:  $p < 0.001$ ) compared to sleeping less than 6 h was connected to lower GSI scores. There was no effect of physical exercise on GSI. Feeling lonely had a negative association with GSI scores both on

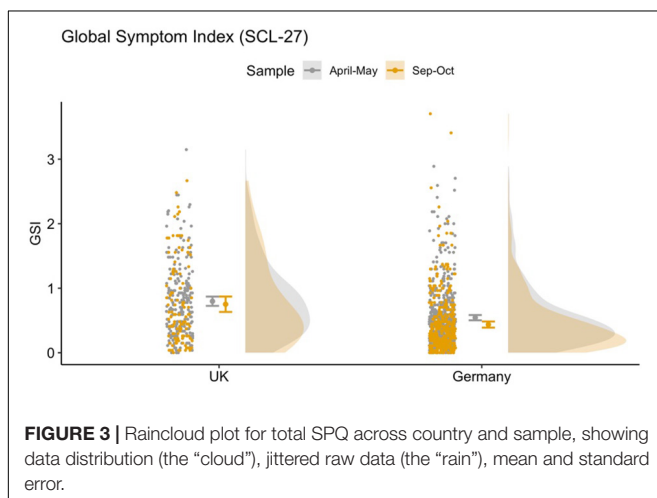


medium levels (TP1:  $p = 0.046$ , TP2:  $p < 0.001$ ) and high levels (TP1:  $p = 0.003$ , TP2:  $p < 0.001$ ) in the May and October samples. See **Table 3** and **Supplementary Materials** for results of the fully overlapping sample.

### Effects of the COVID-19 Pandemic on GSI Scores (COVID-19 Impact Model)

In this third model, we investigated the relationship of factors related to the COVID-19 pandemic with GSI scores. The perception of the restrictions as being stressful was connected with increased GSI scores in the May and October samples (TP1:  $p < 0.001$ , TP2:  $p = 0.013$ ). Financial problems due to the crisis significantly correlated with GSI scores, being higher in the May compared to the October survey. In the May samples, only major financial impact was associated with increased GSI

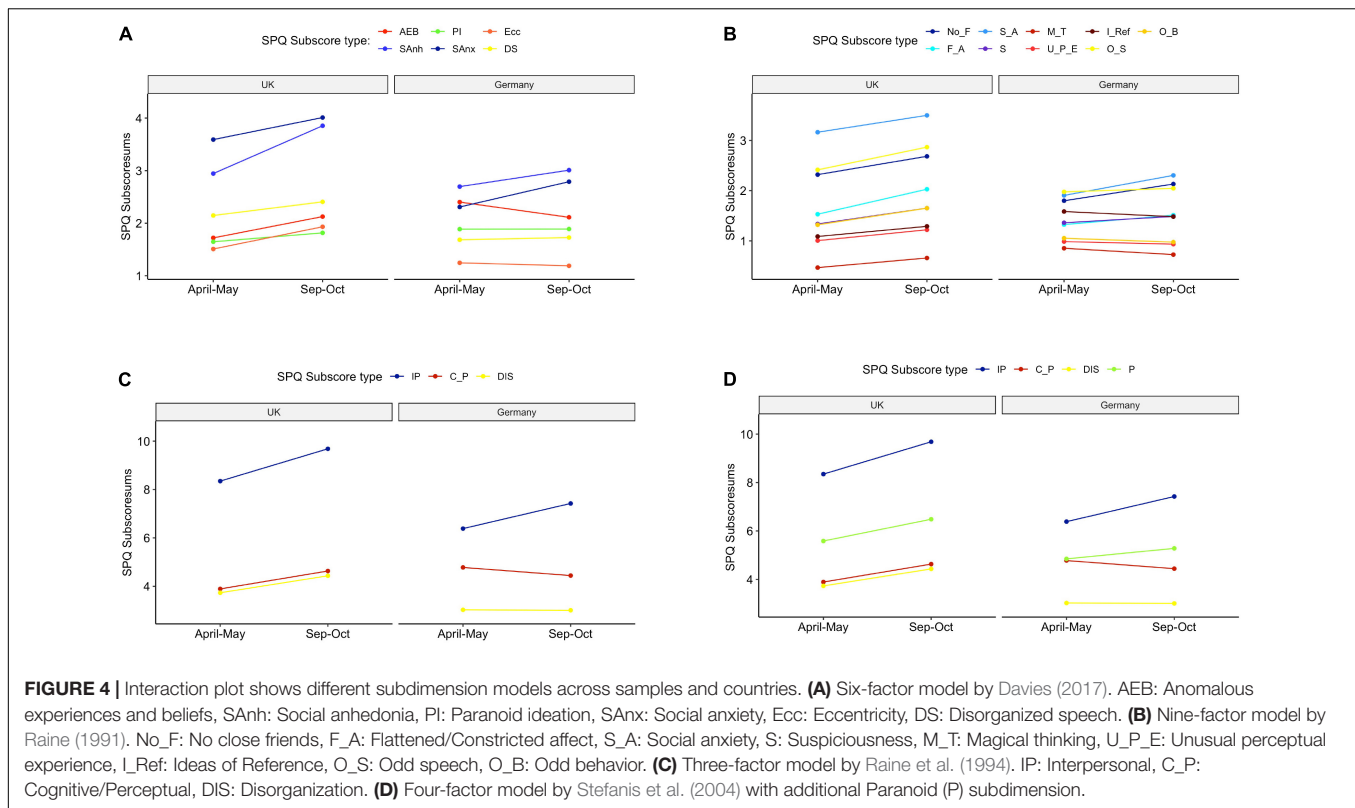
scores ( $p = 0.042$ ). In the October samples, both major financial impact ( $p < 0.001$ ) and medium impact ( $p < 0.001$ ) were negatively related to GSI. Deteriorations in relationships that were experienced as stressful had a negative connection with the outcome in both, the May and October samples, with very stressful changes having a greater association (TP1:  $p < 0.001$ , TP2:  $p < 0.001$ ) on GSI than stressful changes only (TP1:  $p < 0.001$ , TP2:  $p < 0.001$ ). The suspicion of COVID-19 disease or the diagnosis had a negative relationship with GSI scores in both, the May and October samples (TP1:  $p = 0.009$ , TP2:  $p < 0.001$ ). Being hopeful for a soon end of the pandemic did not have a significant association with GSI scores. In the October sample, we also included the degree of trust in the government to lead the country well out of the crisis in our model. However, this predictor had no significant effect on GSI scores. See **Table 3** and **Supplementary Materials** for results of the fully overlapping sample.



### General Linear Model: Schizotypy (SPQ) Effects of Demographic Variables and Prior Physical and Mental Health Problems on SPQ Scores (Basic Model)

In both, the May and October samples, increasing age (TP1:  $p < 0.001$ , TP2:  $p < 0.001$ ), higher education levels (TP1:  $p < 0.001$ , TP2:  $p < 0.001$ ) and female gender (TP1:  $p = 0.088$ , TP2:  $p < 0.001$ ) were connected with lower SPQ scores. The current country of residence in Germany significantly associated with lower scores only in the October sample (TP1:  $p = 0.677$ , TP2:  $p < 0.001$ ). In contrast having children was related to lower outcomes only in the May sample (TP1:  $p = 0.001$ , TP2:  $p = 0.573$ ). In the May samples, living in a suburban ( $p = 0.035$ ) or rural area ( $p = 0.025$ ) compared to a big city were connected with increased





SPQ scores, while in the October sample, living in a small town (TPI:  $p = 0.001$ ) or rural area (TPI:  $p < 0.001$ ) compared to a big city had a decreasing effect on SPQ scores. In addition, there was a trend toward an increased SPQ score when living in a suburb compared to a large city ( $p = 0.074$ ), see **Table 4**.

In the two subsequent models, we adjusted for age, gender, country of residence, place of residence, having pre-existing physical and mental health problems as possible confounding variables. All three models have been calculated for the fully overlapping sample, see **Supplementary Materials**, which are comparable to the full sample.

### Effects of Harmful and Healthy Behavior SPQ Scores (Harmful and Healthy Behavior Model)

After adjusting for possibly confounding variables drug consumption (May samples:  $p < 0.001$ , October samples:  $p < 0.001$ ) as well as excessive media use (May samples:  $p < 0.001$ , October samples:  $p < 0.001$ ) were associated with higher SPQ scores in both samples, whereas alcohol consumption ( $p = 0.004$ ) and medium physical exercise ( $p = 0.036$ ) were connected with lower scores only in the first sample. Physical Exercise at least five times per week had a negative relationship in both, the May and October samples (May samples:  $p = 0.002$ , October samples:  $p < 0.001$ ). Sleeping between 6 and 8 h (May samples:  $p = 0.001$ , October samples:  $p < 0.001$ ) and more than 8 h (May samples:  $p < 0.001$ , October samples:  $p < 0.001$ ) both were associated with lower SPQ-scores compared to sleeping less than 6 h. See **Table 4** and **Supplementary Materials** for results of the fully overlapping sample.

### Effects of the COVID-19 Pandemic on SPQ Scores (COVID-19 Impact Model)

After adjusting for confounds, there was a positive relationship of mistrust in Government of leading the country successfully out of the crisis in the October sample ( $p < 0.001$ ). Perceiving the restrictions as stressful ( $p = 0.043$ ) and being hopeful for a soon end ( $p = 0.022$ ) only had negative effects on SPQ scores in the October sample. Medium financial impact of the CRISIS only had a significant association in the second survey ( $p < 0.001$ ), whereas major financial were positively correlated with SPQ scores in both, the May and October samples (May:  $p < 0.001$ , October:  $p < 0.001$ ). Very stressful relationship changes had only a significant link with the outcome in the second survey ( $p < 0.000$ ). The suspicion or diagnosis of being infected with COVID 19 was associated with an increase of SPQ scores in both, the May and October samples (May samples:  $p < 0.001$ , October samples:  $p = 0.001$ ). See **Table 4** and **Supplementary Materials** for results of the fully overlapping sample.

## DISCUSSION

The current study investigated the association of the COVID-19 pandemic with mental health generally and schizotypy specifically in different, partially overlapping general population samples from the United Kingdom and Germany assessed at two time points – the first during widespread societal restrictions aimed at curbing the spread of the virus (April/May 2020), and the second at a time when the majority of

**TABLE 3 |** Overview over all three conducted models (Basic Model, Harmful and Healthy Behavior, COVID-19 Impact) for both samples for GSI scores.

BIC		April – May sample						September – October sample					
		Basic Model		Harmful and healthy behavior		COVID-19 Impact		Basic Model		Harmful and healthy behavior		COVID-19 Impact	
		874.4		804.2		788.1		641.2		567.4		468.4	
		B	t	B	t	B	t	B	t	B	t	B	t
Intercept		0.67	5.6**	0.68	5.2**	0.45	5.0**	1.15	7.0**	0.97	5.6**	0.59	3.8**
Age		0.00	-3.8**	0.00	-2.9*	-0.00	-3.3*	0.00	-3.3*	-0.004	-2.7*	-0.004	-3.9**
Gender		Reference											
Male		Reference											
Female		0.08	2.2 ~	0.09	2.5 ~	0.07	1.8	0.01	0.3	0.05	1.1	0.04	0.9
Highest Education		0.00	0.1	0.00	0.0	0.00	0.3	-0.06	-3.3*	-0.03	-1.7	-0.03	-1.6
Country		Reference											
United Kingdom		Reference											
GER		-0.12	-2.8 *	-0.14	-3.4*	-0.11	-2.8*	-0.23	-4.3**	-0.20	-4.2**	-0.22	-3.8**
Pre-existing physical health problems		0.14	3.1 *	0.10	2.4 ~	0.11	2.5 ~	0.18	2.8*	0.14	2.4 ~	0.18	3.3*
Pre-existing mental health problems		0.50	11.3**	0.38	8.9**	0.45	10.5**	0.46	7.9**	0.32	5.9**	0.37	7.1**
Children at home		Reference											
No		Reference											
Yes		-0.04	-1.1	-0.02	-0.5	-0.07	-2.0 ~	0.04	0.8	0.05	1.2	-0.08	-2.0 ~
Living area		Reference											
City		Reference											
Suburb		0.06	1.1	0.09	1.8.	0.12	2.3 ~	0.06	0.9	0.04	0.7	0.05	0.8
Town		0.10	2.0 ~	0.12	2.6 ~	0.14	2.8*	-0.02	-0.4	-0.01	-0.1	0.01	0.2
Rural Area		0.03	0.8	0.06	1.3	0.10	2.4 ~	-0.04	-0.7	-0.05	-0.0	0.06	1.2
Alcohol				0.00	0.5					0.01	0.7		
Drug consumption		Reference											
Never		Reference											
At least once				0.1	2.1 ~					0.17	2.8*		
Excessive Media use		Reference											
No		Reference											
Yes				0.1	2.8*					0.11	2.0 ~		
Exercise		Reference											
0 days		Reference											
1–4 days				0.00	0.0					-0.04	-0.8		
>4 days				-0.03	-0.6					-0.12	-1.6		
Sleep		Reference											
<6 h		Reference											
6–8 h				-0.20	-4.3**					-0.24	-4.0**		
>8 h				-0.20	-3.8**					-0.24	-3.0*		
Trust in Government		Reference											
Yes		Reference											
No												0.00	-0.1
Loneliness		Reference											
Not		Reference											
Medium				0.12	2.9*							0.13	2.21 ~
Yes				0.36	8.3**							0.50	8.40**
Restrictions perceived as stressful		Reference											
No		Reference											
Yes						0.02	5.4**					0.05	2.5 ~
Financial impact		Reference											
No		Reference											
Medium						0.01	0.1					0.24	3.9**
Major						0.09	2.0 ~					0.59	7.0**
Hope for a soon end		Reference											
No		Reference											
yes						-0.01	-0.4					-0.03	-0.4
Stressful relationship changes		Reference											
No		Reference											
Few						0.15	4.1**					0.17	3.7**
Many						0.32	5.6**					0.49	7.1**
Suspected of COVID 19		Reference											
No		Reference											
Yes						0.11	2.6 ~					0.19	3.9**

B = Estimate, \* < 0.008, \*\* < 0.001, ~ = only significant when uncorrected.

**TABLE 4 |** Overview over all three conducted models (Basic Model, Harmful and Healthy Behavior, COVID-19 Impact) for both samples for SPQ scores.

		April – May sample						September – October sample					
		Basic Model		Harmful and Healthy Behavior		COVID-19 Impact		Basic Model		Harmful and Healthy Behavior		COVID-19 Impact	
<b>BIC</b>		<b>7056.5</b>		<b>6908.2</b>		<b>6833.2</b>		<b>5268.1</b>		<b>4949.3</b>		<b>4495.3</b>	
		<b>B</b>	<b>z</b>	<b>B</b>	<b>Z</b>	<b>B</b>	<b>z</b>	<b>B</b>	<b>z</b>	<b>B</b>	<b>z</b>	<b>B</b>	<b>z</b>
Intercept		2.97	40.5**	2.92	33.9**	2.81	43.0**	3.78	42.9**	3.48	32.3**	3.45	32.6**
Age		-0.01	-9.2**	-0.01	-6.6**	-0.01	-9.0**	-0.01	-7.7**	-0.01	-5.9**	-0.01	-6.7**
Gender													
Male													
Female		-0.04	-1.7	-0.06	-2.3 ~	-0.04	-1.7	-0.17	-6.1**	-0.16	-5.4**	-0.15	-5.2**
Highest Education		-0.05	-5.2**	-0.04	-4.1**	-0.07	-4.0**	-0.12	-12.4**	-0.07	-6.3**	-0.09	-8.3**
Country													
United Kingdom		Reference											
GER		0.01	0.4	0.02	0.6	0.03	1.2	-0.15	-5.2**	-0.06	-2.1 ~	-0.12	-3.2*
Pre-existing physical health problems		0.22	8.2**	0.17	6.3**	0.23	8.2**	0.20	6.0**	0.19	5.5**	0.24	6.8**
Pre-existing mental health problems		0.44	17.0**	0.37	13.6**	0.43	16.2**	0.56	19.1**	0.43	13.5**	0.42	12.9**
Children at home													
No		Reference											
Yes		-0.08	-3.4*	-0.06	-2.3 ~	-0.09	-3.7**	0.02	0.6	0.02	0.72	-0.13	-4.1**
Living area													
City		Reference											
Suburb		0.07	2.1 ~	0.10	2.9 *	0.10	2.3*	0.07	1.8	0.04	0.89	-0.07	-1.6
Town		0.00	0.3	0.02	0.6	-0.02	-0.5	-0.13	-3.4**	-0.09	-2.3 ~	-0.11	-2.9*
Rural Area		0.07	2.2 ~	0.08	2.6 ~	0.08	2.6 ~	-0.13	-3.9**	-0.11	-3.5**	-0.09	-2.6 ~
Alcohol													
Drug consumption		Reference											
No													
Yes				0.08	2.5*					0.16	4.5**		
Excessive Media consumption		Reference											
No													
Yes				0.22	9.0**					0.21	6.4**		
Exercise		Reference											
No													
1-4 days				-0.06	-2.1 ~					-0.05	-1.7		
Min. 5 days				-0.09	-2.8*					-0.21	-4.1**		
Sleep		Reference											
<6 h													
6-8 h				-0.06	-2.0 ~					-0.27	-7.6**		
>8 h				-0.01	-3.5**					-0.29	-6.2**		
Loneliness		Reference											
Not at all													
Medium				0.02	0.7					0.07	1.8		
Yery				0.14	4.8**					0.26	7.6**		
Trust in Government		Reference											
Yes													
Not at all													
Restrictions perceived as stressful		Reference											
No													
Yes						-0.03	-1.3					0.03	2.1 ~
Financial impact		Reference											
No													
Medium						-0.04	-0.3					0.31	7.9**
Major						0.15	5.3**					0.45	9.2**
End hopeful		Reference											
No													
Yes						-0.04	-1.4					-0.13	-2.5 ~
Stressful relationship chanes		Reference											
No													
Few						0.03	1.1					0.06	1.9
Many						0.02	0.4					0.37	8.3**
Suspected of COVID 19		Reference											
No													
Yes						0.1	3.5**					0.16	4.9**

B = estimate, \*<0.008, \*\*<0.001, ~ = only significant when uncorrected.

these restrictions had been lifted (September/October 2020). Although we are measuring two timepoints, it is not a longitudinal approach, as the samples at both timepoint are only partially overlapping. We are therefore assessing the impact of the pandemic on independent samples from the German and British general population collected in May 2020 and October 2020. The sample from May and October are independent, but highly comparable in terms of distribution of age, and gender. The subjective impact on mental health was quantified using an online survey including questions on the impact on life circumstances, as well as two psychological questionnaires, the Symptom Check List (SCL-27) assessing general psychological symptoms, including depressive symptoms, and symptoms of anxiety, and the Schizotypal Personality Questionnaire (SPQ), assessing schizotypy traits. Furthermore, we assessed the social and economic impact of the pandemic.

We found that the general psychological symptoms (mainly depressive and anxiety symptoms) measured using the Global Symptom Index (GSI, main measure of SCL-27) was significantly lower in the May compared to the October sample in both countries. We were able to confirm these results when running the same analyses in a small sample comparing only those individuals who had taken part at the survey at both timepoints. While during the first timepoint 25–68% of responders were laying above the clinical cut-off for further psychological investigation based on the sub-scores of the SCL-27, at the second timepoint only 12–40% of responders were above clinical threshold. In a normative sample the 10–15% of a cohort reach or exceed the clinical cut-off (Hardt et al., 2004). Schizotypy, however, was higher in the October compared to the May sample, by 4 points in United Kingdom responders, 13.6 to 17.4, and by 1 point in German responders, from 12.3 to 13.2. Furthermore, we investigated the subdimensions of the SPQ. The current literature on the SPQ does not provide a consensus on an optimal dimensional structure of the SPQ. In addition to the original nine-factor structure (Raine, 1991), studies have identified a three-factor structure (Raine et al., 1994; Badoud et al., 2011), four-factor structure (Stefanis et al., 2004; Oezgen and Grant, 2018), a bifactor structure (Preti et al., 2015) as well as a six-factor structure (Davies, 2017). This inconsistency is problematic and may arise because the items of the SPQ introduced measurement error. For completion we investigated the six-, nine-, three- and four-factor models. All models show distinct differences between the two countries, especially in overlapping domains such as social anxiety and the interpersonal scores, where the United Kingdom scores significantly higher than Germany. Interestingly, this analysis shows that while all scores for the subdimensions for the United Kingdom are stable or higher in October compared to May, there is more variation in the score patterns in the German sample.

The results displaying the differences in psychological symptom and schizotypal trait scores confirm our hypotheses. On the one hand, we found that general psychological symptoms (depressive symptoms and symptoms of anxiety) are significantly lower or stay the same comparing samples from October and May, as reported in other studies (Fancourt et al., 2020;

Wang et al., 2020). Fancourt et al. (2020) report for an only United Kingdom cohort that symptoms depression and anxiety stabilized with the introduction of lockdown easing measures from July 2020, whereas we detect a clear decline in symptoms strength. This might be explained by timepoint of data collection, which was conducted in 2 months after the Fancourt study, in September/October 2020. The ability to have social contacts, to resume one's profession, to send children to child care, etc. might have a direct alleviating effect. This shows the possibility that the measured increase in symptoms of anxiety and depression at the onset of the pandemic also resembles a normative psychological response to an exceptional situation. Investigating the sub-scores of the symptom check list (SCL-27) in our study, we found the strongest decrease in agoraphobic symptoms; in the United Kingdom sample, these symptoms decrease by 20% and in the German sample by 10%. This sub-score of the SCL-27, specifically assesses phobic fears of being among others or suppressing actions that could create risks for one's health, like going outside. These behaviors are expected responses during a pandemic, and are therefore likely to reduce when the risk of infection goes down. The only sub-score of the SCL-27 which increased where vegetative symptoms, like dizziness, heart racing, stomach ache, sickness, etc. These symptoms strongly relate to psychosomatic symptoms, which have been reported to have increased significantly in front-line workers (Marinaci et al., 2020; Yi et al., 2020).

On the other hand, and as predicted, we found that schizotypy scores remained the same or were higher at the later timepoint. This is highly interesting, considering that already during the first timepoint nearly 10% of the responders indicated a subjective increase of schizotypy. Recent work shows the impact of adverse life events or loneliness on developing psychotic-like experiences (Beards et al., 2013; Chau et al., 2019; Le et al., 2019; Betz et al., 2020). The social and life-changing consequences of this pandemic (i.e., general reduction of social interaction, job insecurity, experiencing health problems) might therefore provide a long-term risk of schizotypal trait exacerbation in those individuals with high schizotypy scores. Our regression models indicate that indeed loneliness, financial hardship, and drug consumption are predictors for SPQ. The estimates of those three predictors were increased in the October compared to the May sample. In keeping with prior suggestions (Preti et al., 2020), individuals with increased schizotypal traits might show stronger avoidance behaviors, stronger disruption of social contacts, and with that a delayed return to normality, and therefore take longer to reverse the habits established during the first lockdown showing a worsening of schizotypy scores and a delay to return to baseline. However, this hypothesis requires rigorous longitudinal investigations.

In order to identify associations of the impact of the pandemic on psychological symptoms and schizotypy, we ran three sets of regression models separately for the two timepoints. For GSI, we first setup a basic model: During the first survey we identified positive relations between age, being female, living in the United Kingdom, reporting lower mental and physical health prior to the pandemic and living in a town compared to a big city as risk factors, showing an in strong positive association. In the

October samples, we identified an additional positive association with lower education. These results, except for living in the United Kingdom, confirm previous findings (Adamson et al., 2020; Bu et al., 2020; Bäuerle et al., 2020; Fancourt et al., 2020; Smith et al., 2020; Proto and Quintana-Domeque, 2021; Simha et al., 2021). In the harmful and healthy behavior model adjusting for the significant factors of the basic model, we investigated harmful and healthy behavior. We identified the same lowering and increasing association with the outcome in both, the May and October samples. Excessive media consumption and drug consumption contributed to an increased GSI, while longer sleep (>6 h) was negatively associated. Interestingly the effect of drug consumption is twice as high in the second than the May samples. A recent study showed a strong association between newly initiated substance use and increased measures of COVID related fear and worry (Rogers et al., 2020). Those individuals with highest use and fear and worry scores used substances as necessary coping strategies, which might provide an explanation for the increased association between drug use with GSI in the October sample in our study. Regular sleep of more than 6 h and healthy sleep routines are usually predictive of better mental health (Milojevich and Lukowski, 2016), it is therefore not surprising that this is the same during a pandemic. Furthermore, we found that excessive media consumption predicts GSI, which confirms previous findings (Bendau et al., 2020). In the COVID-19 impact model adjusting for the significant factors of the basic model we investigated social and economic impact of the COVID-19 pandemic. We found that while the restrictions themselves and the change in social contacts posed a strong stressor during the May survey it was mainly the financial impact, the change in social contacts and the increased risk of infections which posed the greatest influence during the second timepoint. Already during the first peak in April, Witteveen and Velthorst (2020) linked economic hardship to increased levels of depression and anxiety. During the first peak the economic burden might still be compensated for, however, with the continuing pandemic this burden increases and significantly contributes to mental health decline.

We ran similar regression models to detect potential predictors for schizotypy. In the basic model, we identified similar predictors in both, the May and October samples. While higher age, higher education, and being female were associated with lower outcomes in both, the May and October samples, mental and physical health status before the pandemic were positively correlated. The connection of having children with SPQ scores changed with the continuing of the pandemic from being a negative to a positive predictor; being a United Kingdom resident also correlated with higher outcome scores in the October sample. Living in towns or rural areas was associated with lower scores compared to big cities. The link of annual income with SPQ scores was only recorded during the second survey, with increasing effects on the outcome. Gender differences and younger age have been associated with schizotypy previously (Bora and Baysan Arabaci, 2009), and urbanicity (Fett et al., 2019) as well as lower socioeconomic status (Loch et al., 2017) are often linked to psychotic-like experiences.

In the harmful and healthy behavior model we examined whether harmful and healthy behaviors predicted schizotypy. Adjusting for the significant factors of the basic model, we found the same predictor for both timepoints. While excessive media consumption and drug consumption were linked with higher schizotypy, more exercise and sleep above 6 h showed the opposite relation. Interestingly, however, the association of drug consumption doubled in the October samples and the connection of more exercise tripled in the October samples. The effect of drug use on schizotypy confirms earlier findings reporting that regular cannabis users score higher on schizotypy and psychosis ratings (Nunn et al., 2001). However, it is a critical finding as drug use is also associated with higher conversion from schizotypy to psychosis (Hjorthøj et al., 2018). Therefore, these results are clinically relevant and requires attention in the course of the pandemic. Regular exercise has been identified as an alleviating intervention for early psychosis (Firth et al., 2018), and should be promoted rigorously during a crisis like the current one.

The COVID-19 impact model investigates the relationship of COVID-19 related measures. Here, we see significant worsening comparing the first and the October samples. The association of financial hardship triples, which is independent from annual income. This might show that financial hardship creates a stressor which imposes a risk not only in people with lower socioeconomic status (Loch et al., 2017), but across a wider range of socioeconomic statuses. Furthermore, stress related the change in social contact more than doubled in the October vs. the May samples. This might have been expected that with the continuing course of the pandemic, social isolation might increase, and with that, potentially loneliness too. Loneliness significantly interacts with schizotypy, and has been clinically linked to risk-for psychosis (Chau et al., 2019; Le et al., 2019).

In all models we included country of residence as a predictor, which was significant in most of the models for the May samples and in all models of the October samples. In order to fully understand this relationship, we ran the same regression analysis without country of residence as a predictor (see **Supplementary Tables 6, 7** for results). All main findings remain the same when excluding the country of residence from the models, suggesting that the overall associations, and especially the directionality, is comparable across both countries, but slightly increased in the United Kingdom as indicated by the robust analyses of variance. The reason why United Kingdom residents might suffer a stronger mental health burden is multifold. The delayed start of implementing restrictions and due to that the higher numbers of infections and death (Balmford et al., 2020), followed by a higher unemployment and greater loss in gross income (Bauer and Weber, 2020; Mayhew and Anand, 2020), but also general differences in the health care system might contribute (Kuhlmann et al., 2009). Independently, however, the effects are highly similar, which might be due to the comparability of the samples, and the higher proportion of well-trained and socioeconomically secure responders in both samples.

This study has potential limitations. First, we used online data collection methods, therefore, people without or with

limited access to computers, or less well-versed using these methods would be excluded from the sample. However, in order to maximally ease the accessibility of the questionnaire we provided an online version with smart-phone compatible formatting. Second, we used a snowball sampling method for both timepoint with partially overlapping responders, therefore, the sample is not fully representable of the general population. Although we contacted all participants who had completed the first timepoint and agreed to be re-contacted (71.3%), only 121 participants (14.7%) took part in both timepoints. The results of the study should therefore be interpreted considering the sample's demographics. Furthermore, the reader should be aware that this study is not using a longitudinal approach, it is not showing changes within the same sample. It is, however, comparing two very similar samples at two timepoints within the ongoing pandemic. Third, comparing two countries is problematic as the countries vary on a large number of factors that are not and cannot be accounted for in detail. Therefore, any differences between the countries presented in this study might be linked to baseline differences. However, by specifically asking for a subjective change considering a pre- versus during-pandemic time-point, we minimized this confound. Fourth, we used a self-reporting survey without clinical in-person verifications. Social distancing measures complicate such verification. However, by using a completely voluntary and anonymous format, as well as standardized questionnaires we are minimizing potential effects. Fifth, we are presenting simple regression models without testing for interactions. This approach may not present conclusive results, however, it does allow for comparison with other studies following the same approach, and to generate hypotheses for future research rather than definitive inference. Finally, the usual caveat to observational studies applies, that we are noting associations but cannot infer causality.

In conclusion, we were able to show that whereas general psychological symptoms and percentage of responders above clinical cut-off for further psychological investigation were lower in the sample measured at the second timepoint, following the first peak of the pandemic, schizotypy scores were higher in the October survey. We furthermore found that United Kingdom responders were suffering from a stronger mental health burden than responders from Germany. The financial burden, drug use, the impact of loneliness, and previous mental and physical health problems predicted schizotypy, and general psychological symptoms most strongly, but were stronger in the October samples for schizotypy compared to general psychological symptoms. The differences in the scores over time requires further attention and investigations, to understand whether the

impact on schizotypy increases further, potentially creating a higher risk for developing psychosis.

## DATA AVAILABILITY STATEMENT

The raw data supporting the conclusions of this article will be made available by the authors, without undue reservation.

## ETHICS STATEMENT

The studies involving human participants were reviewed and approved by Ethical Commission Board of the Technical University Munich (reference nr: 250/20 S). The patients/participants provided their written informed consent to participate in this study.

## AUTHOR CONTRIBUTIONS

FK: conceptualization, methodology, data collection, formal analysis and writing (original draft, review and editing), and administration of the project. SD: data collection, formal analysis, and writing (original draft and editing). LM: statistical advice and assistance, discussion, and editing of manuscript. LR: review and editing of manuscript. GM: methodology, discussion, and review and editing of manuscript. All authors contributed to the article and approved the submitted version.

## FUNDING

FK received funding from the European Union's Horizon 2020 (Grant number 754462).

## ACKNOWLEDGMENTS

We thank the participants of the study who contributed their time and further circulated the survey.

## SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: <https://www.frontiersin.org/articles/10.3389/fpsyg.2021.667848/full#supplementary-material>

## REFERENCES

- Adamson, M. M., Phillips, A., Seenivasan, S., Martinez, J., Grewal, H., Kang, X., et al. (2020). International prevalence and correlates of psychological stress during the global COVID-19 pandemic. *Int. J. Environ. Res. Public Health* 17, 1–16. doi: 10.3390/ijerph17249248
- Badoud, D., Chanal, J., Van der Linden, M., Eliez, S., and Debbané, M. (2011). [Validation study of the French schizotypal personality questionnaire in a sample of adolescents: a confirmatory factor analysis]. *Encephale* 37, 299–307. doi: 10.1016/j.encep.2010.08.017
- Balmford, B., Annan, J. D., Hargreaves, J. C., Altoè, M., and Bateman, I. J. (2020). Cross-country comparisons of covid-19: policy, politics and the price of life. *Environ. Resour. Econ.* 76, 525–551. doi: 10.1007/s10640-020-00466-5
- Bauer, A., and Weber, E. (2020). COVID-19: how much unemployment was caused by the shutdown in Germany? *Appl. Econ. Lett.* 1–6.
- Bäuerle, A., Teufel, M., Musche, V., Weismüller, B., Kohler, H., Hetkamp, M., et al. (2020). Increased generalized anxiety, depression and distress during the

- COVID-19 pandemic: a cross-sectional study in Germany. *J. Public Health* 42, 672–678. doi: 10.1093/pubmed/fdaa106
- Beards, S., Gayer-Anderson, C., Borges, S., Dewey, M. E., Fisher, H. L., and Morgan, C. (2013). Life events and psychosis: a review and meta-analysis. *Schizophr. Bull.* 39, 740–747. doi: 10.1093/schbul/sbt065
- Bendau, A., Petzold, M. B., Pyrkosch, L., Maricic, L. M., Betzler, F., Rogoll, J., et al. (2020). Associations between COVID-19 related media consumption and symptoms of anxiety, depression and COVID-19 related fear in the general population in Germany. *Eur. Arch. Psychiatry Clin. Neurosci.* 271, 283–291. doi: 10.1007/s00406-020-01171-6
- Betz, L. T., Penzel, N., Kambeitz-Ilankovic, L., Rosen, M., Chisholm, K., Stainton, A., et al. (2020). General psychopathology links burden of recent life events and psychotic symptoms in a network approach. *NPJ Schizophr.* 6:40. doi: 10.1038/s41537-020-00129-w
- Bora, E., and Baysan Arabaci, L. (2009). Effect of age and gender on schizotypal personality traits in the normal population. *Psychiatr. Clin. Neurosci.* 63, 663–669. doi: 10.1111/j.1440-1819.2009.02011.x
- Bu, F., Steptoe, A., and Fancourt, D. (2020). Loneliness during a strict lockdown: trajectories and predictors during the COVID-19 pandemic in 38,217 United Kingdom adults. *Soc. Sci. Med.* 265:113521. doi: 10.1016/j.socscimed.2020.113521
- Chapman, L. J., Chapman, J. P., Kwapil, T. R., Eckblad, M., and Zinser, M. C. (1994). Putatively psychosis-prone subjects 10 years later. *J. Abnorm. Psychol.* 103:171. doi: 10.1037/0021-843x.103.2.171
- Chau, A. K. C., Zhu, C., and So, S. H.-W. (2019). Loneliness and the psychosis continuum: a meta-analysis on positive psychotic experiences and a meta-analysis on negative psychotic experiences. *Int. Rev. Psychiatr.* 31, 471–490. doi: 10.1080/09540261.2019.1636005
- Daly, M., and Robinson, E. (2021). High-Risk drinking in midlife before versus during the COVID-19 crisis: longitudinal evidence from the united kingdom. *Am. J. Prev. Med.* 60, 294–297. doi: 10.1016/j.amepre.2020.09.004
- Davies, D. J. (2017). *Psychotic Experiences Beyond Psychotic Disorders: From Measurement to Computational Mechanisms*. Ph.D. dissertation. Cambridge: University of Cambridge.
- Debbané, M., and Barrantes-Vidal, N. (2015). Schizotypy from a developmental perspective. *Schizophr. Bull.* 41, S386–S395. doi: 10.1002/9781119125556.devpsy309
- Fancourt, D., Steptoe, A., and Bu, F. (2020). Trajectories of anxiety and depressive symptoms during enforced isolation due to COVID-19 in England: a longitudinal observational study. *Lancet Psychiatr.* 8, 141–149.
- Fett, A.-K. J., Lemmers-Jansen, I. L. J., and Krabbendam, L. (2019). Psychosis and urbanicity: a review of the recent literature from epidemiology to neurourbanism. *Curr. Opin. Psychiatr.* 32:232. doi: 10.1097/ycp.0000000000000486
- Firth, J., Carney, R., Elliott, R., French, P., Parker, S., McIntyre, R., et al. (2018). Exercise as an intervention for first-episode psychosis: a feasibility study. *Early Interv. Psychiatr.* 12, 307–315. doi: 10.1111/eip.12329
- Grant, P., Green, M. J., and Mason, O. J. (2018). Models of schizotypy: the importance of conceptual clarity. *Schizophr. Bull.* 44, S556–S563. doi: 10.1093/schbul/sby012
- Han, E., Tan, M. M. J., Turk, E., Sridhar, D., Leung, G. M., Shibuya, K., et al. (2020). Lessons learnt from easing COVID-19 restrictions: an analysis of countries and regions in Asia Pacific and Europe. *Lancet* 396, 1525–1534. doi: 10.1016/S0140-6736(20)32007-9
- Hardt, J., Dragan, M., and Kappis, B. (2011). A short screening instrument for mental health problems: The Symptom Checklist-27 (SCL-27) in Poland and Germany. *Int. J. Psychiatry Clin. Pract.* 15, 42–49. doi: 10.3109/13651501.2010.523791
- Hardt, J., Egle, U. T., Kappis, B., Hessel, A., and Brähler, E. (2004). Die symptom-checkliste SCL-27. PMP-Psychotherapie-psychosom. *Medizinische Psychol.* 54, 214–223. doi: 10.1055/s-2003-814786
- Hardt, J., and Gerbershagen, H. U. (2001). Cross-validation of the SCL-27: a short psychometric screening instrument for chronic pain patients. *Eur. J. Pain* 5, 187–197. doi: 10.1053/eujp.2001.0231
- Hetkamp, M., Schweda, A., Bäuerle, A., Weismüller, B., Kohler, H., Musche, V., et al. (2020). Sleep disturbances, fear, and generalized anxiety during the COVID-19 shut down phase in Germany: relation to infection rates, deaths, and German stock index DAX. *Sleep Med.* 75, 350–353. doi: 10.1016/j.sleep.2020.08.033
- Hjorthøj, C., Albert, N., and Nordentoft, M. (2018). Association of substance use disorders with conversion from schizotypal disorder to schizophrenia. *JAMA Psychiatr.* 75, 733–739. doi: 10.1001/jamapsychiatry.2018.0568
- Huarcaya-Victoria, J., Herrera, D., and Castillo, C. (2020). Psychosis in a patient with anxiety related to COVID-19: a case report. *Psychiatr. Res.* 289:113052. doi: 10.1016/j.psychres.2020.113052
- JHU (2021). *COVID-19 Dashboard by the Center for Systems Science and Engineering (CSSE) at Johns Hopkins University (JHU)*. Available online at: <https://coronavirus.jhu.edu/map.html> (accessed February 10, 2021).
- Kissler, S., Tedijanto, C., Lipsitch, M., and Grad, Y. H. (2020). Social distancing strategies for curbing the COVID-19 epidemic. *Jinjo Zoki* 49, 40–44. doi: 10.11392/j.sao.49.40
- Knolle, F., Ronan, L., and Murray, G. K. (2021). The impact of the COVID-19 pandemic on mental health in the general population: a comparison between Germany and the UK. *BMC Psychol.* 9:60.
- Kuhlmann, E., Allsop, J., and Saks, M. (2009). Professional governance and public control: a comparison of healthcare in the United Kingdom and Germany. *Curr. Sociol.* 57, 511–528. doi: 10.1177/0011392109104352
- Le, T. P., Cowan, T., Schwartz, E. K., Elvevåg, B., Holmlund, T. B., Foltz, P. W., et al. (2019). The importance of loneliness in psychotic-like symptoms: Data from three studies. *Psychiatr. Res.* 282:112625. doi: 10.1016/j.psychres.2019.112625
- Loch, A. A., Chianca, C., Alves, T. M., Freitas, E. L., Hortêncio, L., Andrade, J. C., et al. (2017). Poverty, low education, and the expression of psychotic-like experiences in the general population of Sao Paulo. *Brazil. Psychiatry Res.* 253, 182–188. doi: 10.1016/j.psychres.2017.03.052
- Mair, P., and Wilcox, R. (2020). Robust statistical methods in R using the WRS2 package. *Behav. Res. Methods* 52, 464–488. doi: 10.3758/s13428-019-01246-w
- Marinaci, T., Carpinelli, L., Venuleo, C., Savarese, G., and Cavallo, P. (2020). Emotional distress, psychosomatic symptoms and their relationship with institutional responses: a survey of Italian frontline medical staff during the Covid-19 pandemic. *Heliyon* 6:e05766. doi: 10.1016/j.heliyon.2020.e05766
- Mayhew, K., and Anand, P. (2020). COVID-19 and the UK labour market. *Oxford Rev. Econ. Policy* 36, S215–S224.
- Miloevich, H. M., and Lukowski, A. F. (2016). Sleep and mental health in undergraduate students with generally healthy sleep habits. *PLoS One* 11:e0156372. doi: 10.1371/journal.pone.0156372
- Nunn, J. A., Rizza, F., and Peters, E. R. (2001). The incidence of schizotypy among cannabis and alcohol users. *J. Nerv. Ment. Dis.* 189, 741–748. doi: 10.1097/00005053-200111000-00002
- O'Connor, R. C., Wetherall, K., Cleare, S., McClelland, H., Melson, A. J., Niedzwiedz, C. L., et al. (2020). Mental health and well-being during the COVID-19 pandemic: longitudinal analyses of adults in the UK COVID-19 Mental health & wellbeing study. *Br. J. Psychiatry* 218, 326–333. doi: 10.1192/bjp.2020.212
- Oezgen, M., and Grant, P. (2018). Odd and disorganized—comparing the factor structure of the three major schizotypy inventories. *Psychiatry Res* 267, 289–295. doi: 10.1016/j.psychres.2018.06.009
- Pfefferbaum, B., and North, C. S. (2020). Mental health and the Covid-19 pandemic. *N. Engl. J. Med.* 383, 510–512.
- Plümper, T., and Neumayer, E. (2020). Lockdown policies and the dynamics of the first wave of the Sars-CoV-2 pandemic in Europe. *J. Eur. Public. Policy* 1–21. doi: 1080/13501763.2020.1847170
- Preti, A., Siddi, S., Vellante, M., Scanu, R., Muratore, T., Gabrielli, M., et al. (2015). Bifactor structure of the schizotypal personality questionnaire (SPQ). *Psychiatr. Res.* 230, 940–950. doi: 10.1016/j.psychres.2015.11.010
- Preti, E., Di Pierro, R., Fanti, E., Madeddu, F., and Calati, R. (2020). personality disorders in time of pandemic. *Curr. Psychiatr. Rep.* 22, 1–9.
- Proto, E., and Quintana-Domeque, C. (2021). COVID-19 and mental health deterioration by ethnicity and gender in the UK. *PLoS One* 16:e0244419. doi: 10.1371/journal.pone.0244419
- R Core Team. (2016). *R: A Language and Environment for Statistical Computing*. Vienna: R Foundation Statistical Computing.
- Raine, A. (1991). The SPQ: a scale for the assessment of schizotypal personality based on DSM-III-R criteria. *Schizophr. Bull.* 17, 555–564. doi: 10.1093/schbul/17.4.555

- Raine, A., Reynolds, C., Lencz, T., Scerbo, A., Triphon, N., and Kim, D. (1994). Cognitive-perceptual, interpersonal, and disorganized features of schizotypal personality. *Schizophr. Bull.* 20, 191–201. doi: 10.1093/schbul/20.1.191
- Rogers, A. H., Shepherd, J. M., Garey, L., and Zvolensky, M. J. (2020). Psychological factors associated with substance use initiation during the COVID-19 pandemic. *Psychiatr. Res.* 293:113407. doi: 10.1016/j.psychres.2020.113407
- R Studio Team (2020). *RStudio: Integrated Development for Rstudio Team*. Boston, MA: PBC. doi: 10.1145/3132847.3132886
- Schultze-Lutter, F., Nenadic, I., and Grant, P. (2019). Psychosis and schizophrenia-spectrum personality disorders require early detection on different symptom dimensions. *Front. Psychiatry* 10:476. doi: 10.3389/fpsy.2019.00476
- Simha, A., Prasad, R., Ahmed, S., and Rao, N. P. (2021). Effect of gender and clinical-financial vulnerability on mental distress due to COVID-19. *Arch. Womens Ment. Health* 23, 775–777. doi: 10.1007/s00737-020-01097-x
- Smith, L., Jacob, L., Yakkundi, A., McDermott, D., Armstrong, N. C., Barnett, Y., et al. (2020). Correlates of symptoms of anxiety and depression and mental wellbeing associated with COVID-19: a cross-sectional study of UK-based respondents. *Psychiatr. Res.* 291, 113138. doi: 10.1016/j.psychres.2020.113138
- Stefanis, N. C., Smyrnis, N., Avramopoulos, D., Evdokimidis, I., Ntzoufras, I., and Stefanis, C. N. (2004). Factorial composition of self-rated schizotypal traits among young males undergoing military training. *Schizophr. Bull.* 30, 335–350. doi: 10.1093/oxfordjournals.schbul.a007083
- Valdés-Flórida, M. J., López-Díaz, Á., Palermo-Zeballos, F. J., Martínez-Molina, I., Martín-Gil, V. E., Crespo-Facorro, B., et al. (2020). Reactive psychoses in the context of the COVID-19 pandemic: clinical perspectives from a case series. *Rev. Psiquiatr. Salud Ment.* 13, 90–94. doi: 10.1016/j.rpsmen.2020.07.004
- Wang, C., Pan, R., Wan, X., Tan, Y., Xu, L., McIntyre, R. S., et al. (2020). A longitudinal study on the mental health of general population during the COVID-19 epidemic in China. *Brain. Behav. Immun.* 87, 40–48. doi: 10.1016/j.bbi.2020.04.028
- WHO (2020). *Mental Health and Psychosocial Considerations During the COVID-19 Outbreak, 18 March 2020*. Geneva: World Health Organization.
- Witteveen, D., and Velthorst, E. (2020). Economic hardship and mental health complaints during COVID-19. *Proc. Natl. Acad. Sci. U.S.A.* 117, 27277–27284. doi: 10.1073/pnas.2009609117
- Xiong, J., Lipsitz, O., Nasri, F., Lui, L. M. W., Gill, H., Phan, L., et al. (2020). Impact of COVID-19 pandemic on mental health in the general population: a systematic review. *J. Affect. Disord.* 277, 55–64.
- Yi, J., Kang, L., Li, J., and Gu, J. (2020). A key factor for psychosomatic burden of frontline medical staff: occupational pressure during the COVID-19 pandemic in China. *Front. psychiatr.* 11:590101. doi: 10.3389/fpsy.2020.590101
- Conflict of Interest:** The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.
- Publisher's Note:** All claims expressed in this article are solely those of the authors and do not necessarily represent those of their affiliated organizations, or those of the publisher, the editors and the reviewers. Any product that may be evaluated in this article, or claim that may be made by its manufacturer, is not guaranteed or endorsed by the publisher.
- Copyright © 2021 Daimer, Mihatsch, Ronan, Murray and Knolle. This is an open-access article distributed under the terms of the Creative Commons Attribution License (CC BY). The use, distribution or reproduction in other forums is permitted, provided the original author(s) and the copyright owner(s) are credited and that the original publication in this journal is cited, in accordance with accepted academic practice. No use, distribution or reproduction is permitted which does not comply with these terms.