Check for updates

OPEN ACCESS

EDITED BY Divya Bhatia, Weill Cornell Medical Center, NewYork-Presbyterian, United States

REVIEWED BY Richard Pratley, AdventHealth, United States Oscar Lorenzo, Health Research Institute Foundation Jimenez Diaz (IIS-FJD), Spain Keiichiro Matoba, Jikei University School of Medicine, Japan

*CORRESPONDENCE Daiji Kawanami, 🛛 kawanami@fukuoka-u.ac.jp

RECEIVED 20 December 2023 ACCEPTED 11 March 2024 PUBLISHED 27 March 2024

CITATION

Muta Y, Kobayashi K, Toyoda M, Tone A, Suzuki D, Tsuriya D, Machimura H, Shimura H, Takeda H, Yokomizo H, Takeshita K, Chin K, Kanasaki K, Tamura K, Miyauchi M, Saburi M, Morita M, Yomota M, Kimura M, Hatori N, Nakajima S, Ito S, Tsukamoto S, Murata T, Matsushita T, Furuki T, Hashimoto T, Umezono T, Takashi Y and Kawanami D (2024), Influence of the combination of SGLT2 inhibitors and GLP-1 receptor agonists on eGFR decline in type 2 diabetes: post-hoc analysis of RECAP study. *Front. Pharmacol.* 15:1358573. doi: 10.3389/fphar.2024.1358573

COPYRIGHT

© 2024 Muta, Kobayashi, Toyoda, Tone, Suzuki, Tsuriya, Machimura, Shimura, Takeda, Yokomizo, Takeshita, Chin, Kanasaki, Tamura, Miyauchi, Saburi, Morita, Yomota, Kimura, Hatori, Nakajima, Ito, Tsukamoto, Murata, Matsushita, Furuki, Hashimoto, Umezono, Takashi and Kawanami. This is an open-access article distributed under the terms of the Creative Commons Attribution License (CC BY). The use, distribution or reproduction in other forums is permitted, provided the original author(s) and the copyright owner(s) are credited and that the original publication in this journal is cited, in accordance with accepted academic practice. No use, distribution or reproduction is permitted which does not comply with these terms.

Influence of the combination of SGLT2 inhibitors and GLP-1 receptor agonists on eGFR decline in type 2 diabetes: post-hoc analysis of RECAP study

Yoshimi Muta¹, Kazuo Kobayashi², Masao Toyoda³, Atsuhito Tone⁴, Daisuke Suzuki⁵, Daisuke Tsuriya⁶, Hideo Machimura⁷, Hidetoshi Shimura⁸, Hiroshi Takeda⁹, Hisashi Yokomizo¹, Kei Takeshita⁶, Keiichi Chin¹⁰, Keizo Kanasaki¹¹, Kouichi Tamura², Masaaki Miyauchi¹², Masuo Saburi¹³, Miwa Morita¹¹, Miwako Yomota¹¹, Moritsugu Kimura³, Nobuo Hatori¹⁴, Shinichi Nakajima¹⁵, Shun Ito¹⁶, Shunichiro Tsukamoto², Takashi Murata^{17,18}, Takaya Matsushita¹³, Takayuki Furuki¹⁹, Takuya Hashimoto⁶, Tomoya Umezono²⁰, Yuichi Takashi¹ and Daiji Kawanami^{1*}

¹Department of Endocrinology and Diabetes, Fukuoka University School of Medicine, Fukuoka, Japan, ²Department of Medical Science and Cardiorenal Medicine, Yokohama City University Graduate School of Medicine, Kanagawa, Japan, ³Division of Nephrology, Endocrinology and Metabolism, Department of Internal Medicine, Tokai University School of Medicine, Kanagawa, Japan, ⁴Department of Internal Medicine, Diabetes Center, Okayama Saiseikai General Hospital, Okayama, Japan, ⁵Suzuki Diabetes Clinic, Kanagawa, Japan, ⁶Division of Endocrinology and Metabolism, 2nd Department of Internal Medicine, Hamamatsu University School of Medicine, Shizuoka, Japan, ⁷Machimura Internal Medicine Clinic, Kanagawa, Japan, ⁸Shimura Clinic, Kanagawa, Japan, ⁹Takeda Clinic, Kanagawa, Japan, ¹⁰Hakuai Clinic, Kanagawa, Japan, ¹¹Department of Internal Medicine 1, Endocrinology and Metabolism, Shimane University Faculty of Medicine, Shimane, Japan, ¹²Miyauchi Diabetes Clinic, Kanagawa, Japan, ¹³Department of Diabetology, Endocrinology and Metabolism, Tokyo Medical University Hachioji Medical Center, Tokyo, Japan, ¹⁴Kobayashi Hospital, Kanagawa, Japan, ¹⁵Sagami Junkanki Clinic, Kanagawa, Japan, ¹⁶Department of Internal Medicine, Sagamihara Red Cross Hospital, Kanagawa, Japan, ¹⁷Department of Clinical Nutrition, National Hospital Organization Kyoto Medical Center, Kyoto, Japan, ¹⁸Diabetes Center, National Hospital Organization Kyoto Medical Center, Kyoto, Japan, ¹⁹Hadano Station South Gate Clinic, Kanagawa, Japan, ²⁰Umezono Internal Medicine Clinic, Kanagawa, Japan

Accumulating evidence has demonstrated that both SGLT2 inhibitors (SGLT2i) and GLP-1 receptor agonists (GLP1Ra) have protective effects in patients with diabetic kidney disease. Combination therapy with SGLT2i and GLP1Ra is commonly used in patients with type 2 diabetes (T2D). We previously reported that in combination therapy of SGLT2i and GLP1Ra, the effect on the renal composite outcome did not differ according to the preceding drug. However, it remains unclear how the initiation of combination therapy is associated with the renal function depending on the preceding drug. In this *post hoc* analysis, we analyzed a total of 643 T2D patients (GLP1Ra-preceding group, n = 331; SGLT2i-preceding group, n = 312) and investigated the differences in annual eGFR decline. Multiple imputation and propensity score matching were performed to compare the annual eGFR decline. The reduction in annual eGFR decline in the SGLT2i-preceding group (pre: $-3.5 \pm 9.4 \text{ mL/min/1.73 m}^2$ /year, post: $-0.4 \pm 6.3 \text{ mL/min/1.73 m}^2$ /year, p < 0.001), was significantly smaller after the

initiation of GLP1Ra, whereas the GLP1Ra-preceding group tended to slow the eGFR decline but not to a statistically significant extent (pre: $-2.0 \pm 10.9 \text{ mL/min/} 1.73 \text{ m}^2/\text{year}$, post: $-1.8 \pm 5.4 \text{ mL/min/} 1.73 \text{ m}^2/\text{year}$, p = 0.83) after the initiation of SGLT2i. After the addition of GLP1Ra to SGLT2i-treated patients, slower annual eGFR decline was observed. Our data raise the possibility that the renal benefits—especially annual eGFR decline—of combination therapy with SGLT2i and GLP1Ra may be affected by the preceding drug.

KEYWORDS

sodium-glucose cotransporter 2 inhibitors, glucagon-like peptide 1 receptor agonists, renal outcome, combination therapy, preceding drug, diabetic kidney disease

1 Introduction

Diabetic kidney disease (DKD) is the leading cause of end-stage kidney disease (ESKD) worldwide (Kim and Kim, 2022). Preventing the onset and progression of DKD is important for preventing ESKD and cardiovascular disease (CVD) (Jankowski et al., 2021). DKD is characterized by albuminuria and estimated glomerular filtration rate (eGFR) decline (Norris et al., 2018). However, recent studies have shown that the prevalence of impaired kidney function with normo-albuminuria is increasing among patients with type 2 diabetes (T2D) (Kume et al., 2019), and those individuals are at increased risk for ESKD and all-cause mortality (Yamamoto et al., 2022). These data suggest that the pathophysiology of DKD is complex.

SGLT2 inhibitors (SGLT2i) and GLP-1 receptor antagonists (GLP1Ra) are widely used for the treatment of T2D and have been shown to have potent glucose-lowering and weight loss effects. Accumulating evidence has demonstrated that SGLT2i and GLP1Ra both have beneficial effects on DKD.

Approximately 40% of T2D patients develop chronic kidney disease (CKD) (Norris et al., 2018). In Japan, 9.9% of patients with type 1 diabetes (T1D) and 15.3% of patients with T2D have a reduced kidney function (eGFR <60) (Shikata et al., 2020). Therefore, it is important to establish an efficient therapeutic strategy against DKD using currently available drugs such as SGLT2i and GLP1Ra.

The International Diabetes Practice Guidelines for CKD, Kidney Disease Improving Global Outcomes (KDIGO) 2022, SGLT2i and metformin are listed as first-line drug therapies for diabetesassociated CKD, and GLP1Ra is preferred as second-line drug therapy (Rossing et al., 2022). Although the renoprotective effects of SGLT2i and GLP1Ra have been reported, the efficacy of their combination therapy has not been fully investigated. Because of side effects, excessive hypoglycemic effects, and medical costs, the initiation of both drugs at the same time is not clinically practiced, and one drug is added after the initiation of treatment with the initial drug. It has not yet been established which drug should be administered first from the viewpoint of renal protection. To clarify this point, we performed the RECAP study (the renoprotective effects of combination treatment with SGLT2i and GLP1Ra in patients with T2D according to their preceding medication). However, in the SGLT2i and GLP1Ra combination therapy, the preceding drug did not affect the renal composite outcome (Kobayashi et al., 2023).

Recently, eGFR decline was implicated as a surrogate endpoint for ESKD (Grams et al., 2019). Therefore, we performed a *post hoc*

analysis of the RECAP study to examine 1) whether the combination therapy of SGLT2i and GLP1Ra is favorable for eGFR decline and 2) how the initiation of combination therapy influences eGFR decline depending on the preceding drug. Evaluations were performed before and after combination therapy with SGLT2i and GLP1Ra in T2D patients.

2 Materials and methods

2.1 Study design

The design of the RECAP study is shown in Supplementary Figure S1. T2D patients who received both SGLT2i and GLP1Ra from April 2010 to December 2021 who had received their preceding medication for at least 6 months, who had received concomitant medication for at least 12 months, and for whom clinical data were available from baseline, the time of drug addition, and the final observation were eligible for inclusion. The following data were collected: sex, age, height, body weight (BW), systolic blood pressure (SBP), diastolic blood pressure (DBP), eGFR, glycated hemoglobin A1c (HbA1c), urinalysis results (urine albumin-to-creatinine ratio (ACR) [mg/g Cr] or qualitative proteinuria), alanine aminotransferase (ALT) level, aspartate aminotransferase (AST) level, platelet count, and concomitant medications (including hypoglycemic drugs, antihypertensive drugs, and statins). eGFR was determined as follows: eGFR (mL/min/1.73 m²) = 194 \times age^{-0.287} × serum creatinine^{-1.094} × (0.739 for women) (Matsuo et al., 2009). Qualitative proteinuria values were converted to albuminuria values using the following formula: predicted ACR = $\exp (5.2659 + 0.2934 \times \log (\min (PCR/50, 1)) + 1.5643 \times \log$ $(\max(\min(PCR/500, 1), 0.1)) + 1.1109 \times \log(\max(PCR/500, 1))$ -0.0773×(if female) + 0.0797×(if diabetic) + 0.1265×(if hypertensive))) (Sumida et al., 2020). Patients with any of the following conditions were excluded from the study: T1D; chronic dialysis; severe liver dysfunction (e.g., liver cirrhosis or severe infection), terminal-stage malignancy, pregnancy, treatment discontinuation, and those who opted out during the course of the study.

The details of the study participants and the dataset used in this study are shown in Supplementary Figure S2. Based on the inclusion criteria, we collected data of 688 patients from 18 medical facilities. Because of the exclusion criteria, 45 patients were excluded and 643 patients (331 patients were previously treated with SGLT2i and later treated with GLP1Ra [SGLT2i-preceding group], and

312 patients were previously treated with GLP1Ra and later treated with SGLT2i [GLP1Ra-preceding group]) were analyzed in this study. Because some data were missing from 643 patients, we performed a multiple imputation (MI) and this dataset complemented by MI was used for the main statistical analysis in this study as full analysis set (FAS). In contrast, after excluding patients with any missing data, 418 patients (227 in the SGLT2i-preceding group, and 191 in the GLP1Ra group) remained. We used this dataset without missing data, that is called complete case analysis set (CCA), to conduct a sensitivity analysis.

As shown in Supplementary Figure S1, patients were first treated with either SGLT2i or GLP1Ra, after more than 6 months, another drug was added and the combination therapy of SGLT2i and GLP1Ra was administered. Therefore, we collected data from three points; at baseline, at the time of drug addition, and at the final observation as shown in Supplementary Figure S1. The median (range) period from baseline to the time of drug addition (monotherapy period) was 23 months (6–114 months). The median (range) period from drug addition to the final observation (combination therapy time) was 31 months (12–85 months). The median (range) period for baseline to the final observation (total observation period) was 59 months (19–134 months) in the total study participants.

The present study was approved by the Institutional Review Board for Clinical Research of Tokai University, Japan on 6 December 2021.

2.2 Office BP measurement

The BP was measured as previously described (Kobayashi et al., 2019). Office BP measurements were performed at each institution using their validated cuff oscillometric devices. According to the JSH 2014 guidelines (Shimamoto et al., 2014), office BP was measured in a quiet environment after resting for a few minutes in a seated position on a chair with legs uncrossed. When two consecutive measurements were taken 1–2 min apart, the average value was defined as the office BP.

2.3 Outcomes

The annual eGFR decline (annual Δ eGFR) during monotherapy (SGLT2i or GLP1Ra) and after combination therapy was evaluated in the SGLT2i-preceding and GLP1Ra-preceding groups. We also evaluated the change in the logarithmic value of ACR (LnACR) in this *post hoc* analysis.

2.4 Statistical method

2.4.1 Missing value analysis

MI was performed to account for missing values (Rubin, 1987). MI is used to replace missing values with other plausible values by creating multiple filling-in patterns to avoid bias caused by missing data. This is recognized as an alternative approach to the analysis of incomplete data (Rubin and Schenker, 1991). In the RECAP study, each missing value was replaced with a set of substituted plausible values by creating 100 complete datasets using MI with the chained equations method (Hershberger and Fisher, 2003; Enders, 2010; Aloisio et al., 2014). For imputation, 100 complete datasets were created using the following covariates: age, sex, height, BW, SBP, DBP, HbA_{1c}, eGFR, LnACR, types of SGLT2i and GLP1Ra, use of concomitant medications (hypoglycemic drugs, antihypertensive drugs, and statins, and period of treatment with either or both SGLT2i or GLP1Ra). The clinical data at baseline, at the time of drug addition, and at the final observation that were associated with the outcome were used for MI (Supplementary Figure S2) (Furukawa et al., 2017).

2.4.2 Statistical analysis for the description of data and comparison

Normally distributed data are presented as the mean \pm standard deviation (SD). Data that showed a skewed distribution were reported as the median [25th percentile, 75th percentile]. For parametric variables, comparisons of clinical characteristics between the two groups were performed using an unpaired *t*-test, The chi-square test was used for nonparametric variables and for categorical data, while a general linear mixed model (GLMM) with Bonferroni correction was used to compare the clinical findings between the three points (at baseline, at the time of drug addition, and at the final observation) as described previously.

Adjusted eGFR decline was calculated by a multiple linear regression analysis using eGFR at baseline, sex, age, BW, BMI, SBP, DBP, HbA1c, and treatment periods in the CCA set.

Multiple regression analysis was performed to identify the independent factors related to change in eGFR decline before and after the combination treatment of SGLT2i and GLP1Ra. A stepwise method was performed using covariates as follows; sex, the duration of T2D, preceding drug, periods of the combination treatment, age, BMI, HbA1c, MAP, eGFR, LnACR at the time of drug addition, changes in BMI, HbA1c, MAP, and LnACR from at the time of drug addition to at the final observation. All statistical analyses were performed using IBM SPSS Statistics 28.0 (IBM Inc., Armonk, NY, USA). Statistical significance was set at p < 0.05.

2.4.3 Propensity score matching

Differences were observed in the baseline characteristics of the SGLT2i-preceding and GLP1Ra-preceding groups. Adjustments were needed for comparison between the groups. Propensity score (PS) matching was conducted to compare the annual eGFR decline and change in albuminuria.

In each dataset built by MI, the PS for the SGLT2i-preceding group was calculated by a logistic analysis that included the following covariates: sex, age, height, BW, body mass index (BMI), SBP, DBP, HbA_{1c} , eGFR, LnACR at baseline, history of T2D, concomitant medications at baseline (hypoglycemic drugs, antihypertensive drugs, and statins), and treatment periods for monotherapy and combination therapy. Because individual PSs were calculated using datasets built by MI, the average PS was used as the representative value. PS matching was performed using representative PS values with the following algorithm: 1:1 nearest neighbor match (caliper value = 0.047, calculated as 0.2 × the SD of PS (Austin, 2011)) with no replacement. In the PS-matching model, the paired *t*-test was used for parametric variables, and the McNemar test was used for categorical data.

	Una	djusted model		PS-matching model					
	GLP1Ra- preceding group, <i>n</i> = 331	SGLT2i- preceding group, <i>n</i> = 312	<i>p</i> -value	GLP1Ra- preceding group, n = 203	SGLT2i- preceding group, <i>n</i> = 203	Standardized difference			
Age (years)	55.7 ± 13.5	56.5 ± 12.7	0.10	57.1 ± 13.6	57.0 ± 13.2	0.007			
Sex (female [%])	152 (46%)	130 (42%)	0.27*	89 (44%)	87 (43%)	0.02			
History of type 2 diabetes >10 years (%)	281 (85%)	237 (76%)	0.006*	165 (81%)	159 (78%)	0.07			
BW (kg)	79.5 ± 20.1	79.4 ± 18.1	0.95	78.7 ± 18.5	78.8 ± 17.0	0.006			
BMI	29.8 ± 6.3	29.5 ± 5.6	0.51	29.4 ± 5.5	29.2 ± 5.3	0.04			
SBP (mmHg)	132.0 ± 18.4	135.4 ± 18.9	0.02	133.1 ± 19.1	134.7 ± 19.4	0.08			
DBP (mmHg)	76.6 ± 12.3	78.7 ± 13.6	0.04	76.7 ± 12.4	78.2 ± 13.5	0.12			
HbA _{1c} (mmol/mol [%])	73.6 ± 18.6 (8.9 ± 1.7)	71.0 ± 17.3 (8.6 ± 1.6)	0.07	72.8 ± 17.8 (8.7 ± 11.6)	71.9 ± 18.2 (8.7 ± 1.7)	0.05			
eGFR (mL/min/1.73 m ²)	78.8 ± 28.7	78.2 ± 26.0	0.79	76.6 ± 26.7	77.7 ± 26.9	0.04			
LnACR	3.75 ± 1.91	3.76 ± 1.97	0.91	3.72 ± 1.90	3.77 ± 1.95	0.003			
Periods of the monotherapy (month)	31.8 ± 23.1	23.9 ± 14.0	<0.001	25.1 ± 18.3	24.7 ± 14.5	0.03			
Periods of the combination therapy (month)	38.8 ± 18.6	28.5 ± 13.5	<0.001	31.6 ± 15.0	31.9 ± 14.0	0.02			
Total periods of the study (month)	70.6 ± 27.0	52.4 ± 15.7	<0.001	56.7 ± 19.4	56.6 ± 14.7	0.006			
Concomitant medicatio	ons								
Sulphonylurea	108 (33%)	91 (29%)	0.34*	58 (29%)	64 (32%)	0.06			
Metformin	169 (51%)	190 (61%)	0.01*	115 (57%)	114 (56%)	0.01			
Insulin	141 (43%)	140 (45%)	0.56*	95 (47%)	90 (44%)	0.05			
Pioglitazone	35 (11%)	51 (16%)	0.03*	29 (14%)	29 (14%)	0			
αGI	40 (12%)	48 (15%)	0.22*	30 (15%)	29 (14%)	0.01			
Glinide	14 (4.2%)	14 (4.5%)	0.87*	11 (5%)	11 (5%)	0			
RAS inhibitor	166 (50%)	160 (51%)	0.77*	108 (53%)	96 (47%)	0.12			
ССВ	128 (39%)	110 (35%)	0.37*	83 (41%)	83 (41%)	0			
b blocker	53 (16%)	49 (16%)	0.92*	33 (16%)	33 (16%)	0			
MRB	14 (4%)	12 (%)	0.81*	10 (5%)	9 (4%)	0.02			
Diuretics	53 (16%)	30 (10%)	0.10*	23 (11%)	25 (12%)	0.03			
Statin	160 (48%)	160 (51%)	0.46*	109 (54%)	98 (45%)	0.11			

TABLE 1 Clinical characteristics at baseline.

The left panel shows the unadjusted model of the full analysis set (FAS) with multiple imputation (MI). The right panel shows the PS-matching model. Values represent the mean \pm SD or n/total n (%). *p*-values were determined by an unpaired t test or *chi-square test. Abbreviations: α GI, alpha glucosidase inhibitor; BMI, body mass index; BW, body weight; DBP, diastolic blood pressure; CCB, calcium channel blocker; eGFR, estimated glomerular filtration; FAS, full analysis set; GLP1Ra, glucagon-like peptide 1 receptor agonists; HbA_{1 α} glycated hemoglobin A_{1 α}; LnACR, logarithmic value of urine albumin-to-creatinine ratio; MAP, mean arterial pressure; MI, multiple imputation; MRB, mineral corticoid receptor blocker; PS, propensity score; RAS, renin-angiotensin system inhibitor; SBP, systolic blood pressure; SGLT2i, sodium-glucose cotransporter inhibitors.

3 Results

3.1 Baseline characteristics

The clinical characteristics at baseline of FAS with MI (n = 643) before and after PS matching in the RECAP study are shown in

Table 1. Before PS matching, the SGLT2i-preceding and GLP1Rapreceding groups showed significant differences in the history of T2D >10 years (76% vs. 85%, p = 0.006), SBP (mmHg) (135.4 ± 18.9 vs. 132.0 ± 18.4, p = 0.02), DBP (mmHg) (78.7 ± 13.6 vs. 76.6 ± 12.3, p = 0.04), period of monotherapy (month) (23.9 ± 14.0 vs. 31.8 ± 23.1, p < 0.001), period of combination therapy (month)

Muta et al.

10.3389/fphar.2024.1358573

TABLE 2 The renal function and eGFR decline during the study periods on unadjusted model (FAS with MI).

A) Chang	es in eGFR	(mL/min/1	.73m²)										
		FAS					CCA						
	At At baseline tim dı add	At the	At the fir	al	<i>p</i> -value*		At A	At the	At the final	<i>p</i> -value*			
		drug addition	Observatio	Baseline vs. addition	Addition vs. final observation	Baseline vs. final observation	basetine	drug addition	Observation	Baseline vs. addition	Addition vs. final observation	Baseline vs. final observation	
Total (<i>n</i> = 643)	78.5 ± 27.4	74.0 ± 26.5	70.8 ± 26.8	<0.001	<0.001	<0.001	78.8 ± 28.6	74.0 ± 27.6	70.1 ± 27.8	<0.001	<0.001	<0.001	
SGLT2i- preceding group ($n =$ 331 in FAS n = 277 in CCA)	78.2 ± 26.0	73.2 ± 26.0	71.4 ± 26.	<0.001	0.06	<0.001	79.1 ± 26.0	74.0 ± 26.4	72.2 ± 26.6	<0.001	0.22	<0.001	
GLP1Ra- preceding group ($n =$ 312 in FAS n = 191 in CCA)	78.8 ± 28.7	74.7 ± 27.0	70.1 ± 27.:	<0.001	0.06	<0.001	78.5 ± 30.6	73.9 ± 28.7	68.3 ± 28.9	<0.001	<0.001	<0.001	
B) Annua	l eGFR dec	line (mL/m	in/1.73 m²/y	ear)									
	FAS								CCA				
From at baseline to at From the time the time of drug addition to addition observa					the time of drug ion to the final observation	g p-value [*] From at baseline to at From the time of drug addition to the fina addition to the fina			he time of drug ion to the final bservation	g p-value*			
Total (1	Total $(n = 643)$		Treatment periods 28 (months)		33.8 ± 17.1				28.3 ± 20.3		34.3 ± 17.3		
		eGFR decline (mL/min/ 1.73 m²/year)		-2.6 ± 9.9		-1.2 ± 6.0			-2.5 ± 10.1		-1.4 ± 6.2		
A		Adjusted eGFR decline (mL/min/1.73 m²/ year)*							-2.1 ± 2.7		-1.6 ± 2.4	0.002	

(Continued on following page)

TABLE 2 (Continued) The renal function and eGFR decline during the study periods on unadjusted model (FAS with MI).

B) Annual eGFR decline (mL/min/1.73 m²/year)											
			FAS	CCA							
		From at baseline to at the time of drug addition	From the time of drug addition to the final observation	<i>p</i> -value*	From at baseline to at the time of drug addition	From the time of drug addition to the final observation	p-value*				
SGLT2i-preceding group (<i>n</i> = 331 in FAS <i>n</i> = 277 in CCA)	Treatment periods (months)	23.9 ± 14.0	28.5 ± 13.5		24.0 ± 14.6	28.7 ± 13.8					
	eGFR decline (mL/min/ 1.73 m²/year)	-3.4 ± 9.5	-0.9 ± 6.5	<0.001	-3.3 ± 9.0	-0.8 ± 6.8	0.004				
	Adjusted eGFR decline (mL/min/1.73 m²/ year)*				-2.3 ± 2.5	-1.7 ± 2.3	0.01				
GLP1Ra-preceding group (n = 312 in FAS n = 191 in CCA)	Treatment periods (months)	31.8 ± 23.1	38.9 ± 23.1		31.9 ± 23.5	39.1 ± 18.6					
	eGFR decline (mL/min/ 1.73 m²/year)	-1.9 ± 10.3	-1.6 ± 5.3	0.69	-1.8 ± 11.0	-2.0 ± 5.7	0.78				
	Adjusted eGFR decline (mL/min/1.73 m²/ year)*				-2.0 ± 2.9	-1.5 ± 2.4	0.06				

Changes in eGFR (A) and Annual eGFR decline (B) are shown respectively. In complete case analysis (CCA) set, values of adjusted eGFR decline were shown. Adjusted eGFR decline = 14.48 - eGFR at baseline $\times 0.087 + 1.504 \times (if female) - 0.069 \times Age + 0.052 \times BW - 0.097 \times BMI - 0.018 \times SBP - 0.032 \times DBP - 0.048 \times HbA1c (mol/l) + Treatment periods(months) \times 0.022$. Values are mean±SD **p*-values were analyzed by paired t test among two groups or by the general linear mixed model with Bonferroni correction among three groups.

Muta et al.

 $(28.5 \pm 13.5 \text{ vs. } 38.8 \pm 18.6, p < 0.001)$, total study period (month) $(52.4 \pm 15.7 \text{ vs. } 70.6 \pm 27.0, p < 0.001)$, use of metformin (61% vs. 51%, p = 0.01) and pioglitazone (16% vs. 11%, p = 0.03) at baseline. In the PS-matching model, the range of standardized differences in covariates was 0.0–0.12. Therefore, the PS-matching model was considered to be well-balanced between the groups.

3.2 Annual eGFR decline on FAS

Table 2A shows that, in the analysis of 643 patients, the eGFR was significantly lower by passage of time, with values (mL/min/ 1.73 m^2) of 78.5 ± 27.4 at baseline, 74.0 ± 26.5 at the time of drug addition, and 70.8 ± 26.8 at the final observation (p < 0.001 between each period). Similar results of eGFR decline were observed in CCA set analysis. Table 2B shows the annual eGFR decline (mL/min/ 1.73 m^2 /year) was -2.6 ± 9.9 in monotherapy with a period of 28.0 ± 19.6 months and -1.2 ± 6.0 in combination therapy with a period of 33.8 ± 17.1 months. The difference was statistically significant (p = 0.005). Table 2 also shows the changes in eGFR and annual eGFR decline depending on the preceding drug in the unadjusted model. A significant difference in annual eGFR decline was observed in patients in the SGLT2i-preceding group before and after the initiation of combination therapy (p < 0.001).

3.3 Adjusted eGFR decline on CCA

The calculated formula is as follows; adjusted eGFR decline = 14.48 - eGFR at baseline $\times 0.087 + 1.504 \times (if female) - 0.069 \times Age + 0.052 \times BW - 0.097 \times BMI - 0.018 \times SBP - 0.032 \times DBP - 0.048 \times HbA1c (mol/L) + Treatment periods(months) <math>\times 0.022$.

Although this adjustment could not eliminate all the effects of confounding factors, annual eGFR decline (mL/min/1.73 m²/year) significantly changed from -2.1 ± 2.7 to -1.6 ± 2.4 (p = 0.002) among all patients. In the analysis depending on the preceding drugs, annual eGFR decline (mL/min/1.73 m²/year) significantly changed from -2.3 ± 2.5 to -1.7 ± 2.3 in SGLT2i-preceding group (p = 0.001), whereas that changed from -2.0 ± 2.9 to -1.5 ± 2.4 in GLP1Ra-preceding group (p = 0.06) (Table 2B).

Multiple regression analysis identified that 1) eGFR at the time of drug addition, 2) change in HbA1c from at the time of drug addition to the final observation, and 3) the preceding of SGLT2i were independent and significant determinants of eGFR decline. These regression coefficient values were -0.10 (95% CI, -0.14, -0.05, p < 0.001), 0.08 (95% CI, 0.01, 0.16, p = 0.02), and 2.68 (95% CI, 0.27, 5.01, p = 0.03), respectively.

3.4 Results of PS-matching model

Using the PS matching method, a matching model of 203 patients in each group was constructed (Supplementary Figure S2).

3.4.1 eGFR decline on PS-matching model

For the evaluation of the eGFR decline before and after combination therapy, we analyzed the eGFR decline separately

for the SGLT2i-preceding and GLP1Ra-preceding groups in the PS-matching model. In the SGLT2i-preceding group, the annual eGFR decline (mL/min/1.73 m²/year) was -3.5 ± 9.4 and -0.4 ± 6.3 in the monotherapy and combination therapy periods, respectively, (p < 0.001). On the other hand, in the GLP1Ra-preceding group, the annual eGFR decline (ml/min/1.73 m²/year) was -2.0 ± 10.9 and -1.8 ± 5.4 in the monotherapy and combination therapy periods, respectively. They did not differ to a statistically significant extent (p = 0.83) (Figure 1).

3.4.2 LnACR on PS-matching model

The LnACR values at baseline, at the time drug addition, and at the final observation were 3.77 ± 1.95 , 3.79 ± 1.99 , and 3.94 ± 2.00 , respectively, in the SGLT2i-preceding group. Those values were 3.72 ± 1.90 , 3.85 ± 1.95 , and 3.78 ± 1.78 , respectively, in the GLP1Ra-preceding group. None of the groups showed a significant difference in LnACR at any of the three time points.

3.4.3 Changes in other clinical findings on PSmatching model

In comparison to baseline, lower HbA_{1c}, BW, SBP and DBP were observed after combination therapy with SGLT2i and GLP1Ra. Table 3 shows the changes in clinical characteristics before and after combination therapy in the PS-matching model. In the monotherapy period, the HbA_{1c} value did not differ between the groups; however, lower HbA1c levels were observed after combination therapy in both the SGLT2i-preceding and GLP1Ra-preceding groups (both p < 0.001). A lighter BW was observed in the monotherapy period (p = 0.03 in the SGLT2i-preceding group and p = 0.007 in the GLP1Ra-preceding group), and further lighter BW was observed in both groups in the combination therapy period (both p < 0.001) (Table 3). There was no difference in SBP or DBP. However, in the GLP1Ra-preceding group, a further decrease in SBP was observed (p = 0.03).

4 Discussion

We performed this *post hoc* study to analyze the relationship between the use of SGLT2i and GLP1Ra combination therapy and renal function in T2D patients. We observed that the annual eGFR decline in the monotherapy period of SGLT2i or GLP1Ra ($-2.6 \pm$ 9.9 mL/min/1.73 m²/year) was significantly lower after combination therapy (-1.2 ± 6.0 mL/min/1.73 m²/year) (p = 0.005). In the PS matching model that adjusted the baseline characteristics of the two groups, in the SGLT2i-preceding group, the eGFR decline was significantly lower after the addition of GLP1Ras in comparison to SGLT2i alone (-3.5 ± 9.4 mL/min/1.73 m²/year to -0.4 ± 6.3 mL/ min/1.73 m²/year, p < 0.001). On the other hand, the annual eGFR decline was not significantly changed by the addition of SGLT2i to the GLP1Ra-preceding group (-2.0 ± 10.9 mL/min/1.73 m²/year to -1.8 ± 5.4 mL/min/1.73 m²/year, p = 0.83).

In our previous study, the decrease in eGFR was slower in patients treated with SGLT2i than in those treated with GLP1Ra during monotherapy (Kobayashi et al., 2022). We hypothesize that this occurred because SGLT2i was administered for a shorter period (24.7 ± 14.5 months) in the present study, and the effect of the initial dip was greater. The mechanisms underlying this discrepancy



drug addition vs. the final observation in the SGLT2i-preceding group. p < 0.001 between each point in the GLP1Ra-preceding group. p-values were analyzed by general linear mixed model). The administration of SGLT2i prior to GLP1Ra significantly reduced the annual eGFR decline. p-values were analyzed by a paired t-test.

remain unclear, but the lack of a significant difference in the annual eGFR decline in the GLP1Ra-preceding group may be due to the initial dip that occurs in the early phase of SGLT2i treatment. However, with an observation period of approximately 2 years, this effect was considered to have disappeared.

Another possible mechanism may be the difference in the mean BP at the time of drug treatment. It is well known that BP control influences the renal function decline (Soejima et al., 2022). We previously reported a lower BP in SGLT2i than in GLP1Ra (Kobayashi et al., 2019). Consistent with this observation, both SBP and DBP were significantly lower during the SGLT2i monotherapy period in the SGLT2i-preceding group. In contrast, GLP1Ra monotherapy showed no change in BP in the GLP1Ra-preceding group. The lower BP during SGLT2i monotherapy may have contributed to the subsequent reduction in annual eGFR decline with combination therapy. These findings suggest that reducing BP before the initiation of combination therapy may be related to the efficient induction of renoprotection.

In this study, no significant changes were observed in the ACR between the two groups. Considering these results, the degree of ACR was relatively low; therefore, it was difficult to determine the effect of combination therapy.

A series of large-scale randomized controlled trials (RCTs) demonstrated that SGLT2i use reduced the risk of kidney disease progression in patients with and without diabetes (Nuffield Department of Population Health Renal Studies and Consortium, 2022). The beneficial effects of SGLT2i on albuminuria and eGFR decline in DKD have been demonstrated by EMPA-REG OUTCOME (Wanner et al., 2016), CANVAS (Perkovic et al., 2018), DECLARE-TIMI58 (Mosenzon et al., 2019), and CREDENCE (Perkovic et al., 2019). These effects were consistent in DAPA-CKD (Heerspink et al., 2020; Heerspink et al., 2021) and EMPA-KIDNEY (The EMPA-KIDNEY Collaborative Group et al., 2023), which included CKD

patients with and without diabetes. Recently, a *post hoc* analysis of SUSTAIN6 and PIONEER6 demonstrated that semaglutide slowed the eGFR decline in T2D patients with a high risk of CVD (Tuttle et al., 2023), suggesting that GLP1Ra has potentially favorable effects on the renal function.

With combination therapy of SGLT2i and GLP1Ra, the preceding drug did not affect the renal outcome in the main analysis of RECAP study (Kobayashi et al., 2023). In this post hoc analysis, the SGLT2ipreceding group had a slower rate of eGFR decline in comparison to the GLP1Ra-preceding group, and multiple regression analysis identified the preceding of SGLT2i as an independent and significant factor of change in eGFR decline. Interestingly, we also demonstrated a combination of both drugs had additive renoprotection. Furthermore, the addition of GLP1Ra to SGLT2i showed an advantage in the decrease in BW. These findings suggest that there is no problem in starting SGLT2i first with the expectation of renoprotection, however, if a stronger protective effect is desired, a combination of GLP1Ra is preferable. In addition, GLP1Ra may be used first in patients with obesity who should prioritize BW reduction. Recent studies demonstrated that SGLT2i had cardioprotective effects in HFpEF (heart failure with preserved ejection fraction) and HFrEF (heart failure with reduced ejection fraction) with or without diabetes (Solomon et al., 2022; Cortes et al., 2023). On the other hand, GLP1Ra can be prescribed only for patients with T2D. Therefore, it remains unclear whether similar results can be observed in patients without T2D.

5 Limitations

The present study was associated with several limitations. First, the data were analyzed as a *post hoc* analysis from a retrospective study and

TABLE 3 Clinical characteristics before and after combination therapy with SGLT2i and GLP1Ra depending on the preceding drug in the PS-matching model.

	The SGLT2i-preceding group ($n = 203$)						The GLP1R-preceding group ($n = 203$)					
A base	At	At the	At the final	<i>p</i> -value*		At	At the	At the final	<i>p</i> -value*			
	Dasetine	drug addition	Observation	Baseline vs. addition	Addition vs. final observation	Baseline vs. final observation	baseune	drug addition	Observation	Baseline vs. addition	Addition vs. final observation	Baseline vs. final observation
HbA _{1c} (mol/ mol [%])	71.9 ± 18.2 (8.7 ± 1.7)	69.3 ± 17.0 (8.5 ± 1.6)	62.4 ± 15.0 (7.9 ± 1.4)	0.21	<0.001	<0.001	72.8 ± 17.8 (8.8 ± 1.6)	70.9 ± 16.2 (8.6 ± 1.5)	92.9 ± 15.2 (7.9 ± 1.4)	0.46	<0.001	<0.001
BW (kg)	78.8 ± 17.0	77.8 ± 16.9	75.5 ± 17.2	0.03	< 0.001	<0.001	78.7 ± 18.5	77.3 ± 18.4	73.6 ± 18.3	0.007	<0.001	< 0.001
SBP (mmHg)	134.7 ± 19.4	130.4 ± 18.6	128.9 ± 17.4	0.003	1.0	<0.001	133.1 ± 19.1	132.5 ± 17.7	129.3 ± 16.1	1.0	0.03	0.01
DBP (mmHg)	78. ± 13.5	75.8 ± 12.4	74.6 ± 12.5	0.01	0.50	<0.001	76.7 ± 12.4	75.8 ± 11.8	74.6 ± 12.3	0.61	0.63	0.03

Treatment periods (months); The SGLT2i-preceding group; from baseline to the time of addition of drug (24.7±14.5), from the time of drug addition to the final observation (31.9±14.0). The GLP1Ra-preceding group; from baseline to at the time of drug addition (25.1±18.3), from the time of drug addition to the final observation (31.6±15.0). Values present the mean±SD. **p*-values were analyzed between two groups by a paired t test, among three groups by a general linear mixed model with Bonferroni correction. Abbreviations: BW, body weight; SBP, systolic blood pressure; DBP, diastolic blood pressure.

may have been insufficient to evaluate eGFR decline. Second, the relatively low numbers of participants may have resulted in insufficient power to detect differences in the effect of treatment on eGFR decline. Third, it remains unclear whether our results can be generalized to T2D patients. The analysis target was patients who required combination therapy with SGLT2i and GLP1Ra for the treatment of T2D. Combination therapy has been used for glycemic management and/or BW reduction but not for renoprotection. Indeed, obese patients were enrolled (mean BMI: $29.2 \pm 5.3 \text{ kg/m}^2$ in the SGLT2i-preceding group and 29.4 \pm 5.5 kg/m² in the GLP1Rapreceding group) in this study, and the influence of combination therapy on obesity-related glomerulopathy was not evaluated sufficiently. Fourth, we could not exclude the possibility that clinicians selected SGLT2i first because of renoprotective properties of SGLT2i. Supplementary Figure S3 shows the distribution of the years at the initiation of treatment depending on the drug. Approximately one-third of patients in GLP1Ra preceding group started GLP1Ra before 2014 when SGLT2i were not commercially available in Japan. In contrast, some patients in the SGLT2ipreceding group may have started SGLT2i based on expectation of renoprotective effects. It is possible that patients at high risk for renal outcomes started SGLT2i first, which may have influenced the eGFR decline in this study. Fifth, the population included in this study was only Japanese. Therefore, the generalizability of these results to other races requires further validation. Finally, the PS-matching could balance known confounding variables as much as possible, including gender, duration of diabetes, HbA1c, BMI and BP to eliminate baseline differences between groups, but it could not ensure that all measured baseline characteristics were matched and consider the influence of unknown variables.

6 Conclusion

In conclusion, this retrospective study demonstrated that SGLT2i and GLP1Ra combination therapy could be beneficial in reducing the annual eGFR decline in patients with T2D. Among the combination therapies, the SGLT2i-preceding group showed a significant reduction in annual eGFR decline, while the GLP1Ra-preceding group did not.

Further studies are required to clarify the additive effect of SGLT2i and GLP1Ra combination therapy on renal protection. In particular, the effects of preceding drugs on the rate of annual eGFR decline should be investigated in a prospective RCT.

Data availability statement

The original contributions presented in the study are included in the article/Supplementary Material, further inquiries can be directed to the corresponding author.

Ethics statement

The studies involving humans were approved by Institutional Review Board for Clinical Research of Tokai University. The studies were conducted in accordance with the local legislation and institutional requirements. Written informed consent for participation was not required from the participants or the participants' legal guardians/next of kin because Since the research targets patients who have received treatment in the past, it is difficult to obtain consent from individual subjects of the research. Therefore, we publicized information on "the fact that the research is being conducted, its contents, methods, the opportunity to refuse participation in the research, and contact information" by publicly announcing the information disclosure document. The information disclosure document was posted on the website of the Institutional Review Board for Clinical Research of Tokai University and in the outpatient of the Division of Nephrology, Endocrinology and Metabolism, Department of Internal Medicine, Tokai University School of Medicine.

Author contributions

YM: Writing-original draft, Writing-review and editing. KaK: Writing-original draft, Writing-review and editing. MT: Writing-original draft, Writing-review editing. and Writing-original draft, Writing-review editing. AT: and DS: Writing-original draft, Writing-review and editing. DT: Writing-original draft, Writing-review and editing. HM: Writing-original draft, Writing-review editing. and Writing-original Writing-review HS: draft, and editing. HT: Writing-original draft, Writing-review editing. and HY: Writing-original draft, Writing-review and editing. KeT: Writing-original draft, Writing-review editing. and KC: Writing-original draft, Writing-review and editing. KeK: Writing-original draft, Writing-review editing. and KoT: Writing-original draft, Writing-review editing. and MaM: Writing-original draft, Writing-review editing. and MS: Writing-original draft, Writing-review and editing. MiM: Writing-original draft, Writing-review and editing. MY: Writing-original draft, Writing-review editing. and MK: Writing-original draft, Writing-review editing. and NH: Writing-original draft, Writing-review and editing. SN: Writing-original draft, Writing-review editing. and SI: Writing-original draft, Writing-review editing. and ST: Writing-original draft, Writing-review editing. and TMu: Writing-original draft, Writing-review and editing. TMa: Writing-original draft, Writing-review and editing. TF: Writing-original draft, Writing-review and editing. TH: Writing-original draft, Writing-review and editing. TU: Writing-original draft, Writing-review and editing. YT: Writing-original draft, Writing-review and editing. DK: Writing-original draft, Writing-review and editing.

Funding

The author(s) declare that no financial support was received for the research, authorship, and/or publication of this article.

Acknowledgments

We are grateful to all participants and acknowledge the support of the members of the RECAP study who contributed considerably to data collection.

Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

Publisher's note

All claims expressed in this article are solely those of the authors and do not necessarily represent those of their affiliated organizations, or those of the publisher, the editors and the reviewers. Any product that may be evaluated in this article, or claim that may be made by its manufacturer, is not guaranteed or endorsed by the publisher.

References

Aloisio, K. M., Swanson, S. A., Micali, N., Field, A., and Horton, N. J. (2014). Analysis of partially observed clustered data using generalized estimating equations and multiple imputation. *Stata J.* 14, 863–883. doi:10.1177/1536867x1401400410

Austin, P. C. (2011). Optimal caliper widths for propensity-score matching when estimating differences in means and differences in proportions in observational studies. *Pharm. Stat.* 10, 150–161. doi:10.1002/pst.433

Cortes, M., Lorenzo, O., Lumpuy-Castillo, J., Martinez-Albaladejo, S., Taibo-Urquia, M., Pello, A. M., et al. (2023). Dapagliflozin improved cardiac function and structure in diabetic patients with preserved ejection fraction: results of a single centre, observational prospective study. *J. Clin. Med.* 12, 6698. doi:10.3390/jcm12206698

Enders, C. K. (2010). Applied missing data analysis. New York: Applied missing data.

Furukawa, K., Preston, D. L., Misumi, M., and Cullings, H. M. (2017). Handling incomplete smoking history data in survival analysis. *Stat. Methods Med. Res.* 26, 707–723. doi:10.1177/0962280214556794

Grams, M. E., Sang, Y., Ballew, S. H., Matsushita, K., Astor, B. C., Carrero, J. J., et al. (2019). Evaluating glomerular filtration rate slope as a surrogate end point for ESKD in clinical trials: an individual participant meta-analysis of observational data. *J. Am. Soc. Nephrol.* 30, 1746–1755. doi:10.1681/ASN.2019010008

Heerspink, H. J. L., Jongs, N., Chertow, G. M., Langkilde, A. M., Mcmurray, J. J. V., Correa-Rotter, R., et al. (2021). Effect of dapagliflozin on the rate of decline in kidney function in patients with chronic kidney disease with and without type 2 diabetes: a prespecified analysis from the DAPA-CKD trial. *Lancet Diabetes Endocrinol.* 9, 743–754. doi:10.1016/S2213-8587(21)00242-4

Heerspink, H. J. L., Stefansson, B. V., Correa-Rotter, R., Chertow, G. M., Greene, T., Hou, F. F., et al. (2020). Dapagliflozin in patients with chronic kidney disease. *N. Engl. J. Med.* 383, 1436–1446. doi:10.1056/NEJMoa2024816

Hershberger, S. L., and Fisher, D. G. (2003). A note on determining the number of imputations for missing data. *Struct. Equ. Model. A Multidiscip. J.* 10, 648–650. doi:10. 1207/s15328007sem1004_9

Jankowski, J., Floege, J., Fliser, D., Bohm, M., and Marx, N. (2021). Cardiovascular disease in chronic kidney disease: pathophysiological insights and therapeutic options. *Circulation* 143, 1157–1172. doi:10.1161/CIRCULATIONAHA.120.050686

Kim, M. K., and Kim, D. M. (2022). Current status of diabetic kidney disease and latest trends in management. J. Diabetes Investig. 13, 1961–1962. doi:10.1111/jdi.13895

Kobayashi, K., Toyoda, M., Hatori, N., Sakai, H., Furuki, T., Sato, K., et al. (2022). Comparison of the blood pressure management between sodium-glucose cotransporter 2 inhibitors and glucagon-like peptide 1 receptor agonists. *Sci. Rep.* 12, 16106. doi:10. 1038/s41598-022-20313-5

Kobayashi, K., Toyoda, M., Kimura, M., Hatori, N., Furuki, T., Sakai, H., et al. (2019). Renal effects of sodium glucose co-transporter 2 inhibitors in Japanese type 2 diabetes mellitus patients with home blood pressure monitoring. *Clin. Exp. Hypertens.* 41, 637–644. doi:10.1080/10641963.2018.1529778

Supplementary material

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fphar.2024.1358573/ full#supplementary-material

SUPPLEMENTARY FIGURE S1

Patients with T2D who were treated with both SGLT2i and GLP1Ra from April 2010 to December 2021. Patients who had been on monotherapy for at least 6 months and on combination therapy for at least 12 months were registered in the study.

SUPPLEMENTARY FIGURE S2

A total of 688 patients were registered, and 45 patients were excluded. The data of 643 patients (SGLT2i-preceding group, n = 312; GLP1Ra-preceding group, n = 331) were analyzed as the FAS. The multiple imputation method was applied to FAS data. Propensity score matching was conducted to compare the annual Δ eGFR and LnACR.

SUPPLEMENTARY FIGURE S3

Distribution of years at the initiation of treatment for each type of drug. Approximately one-third of patients in GLP1Ra preceding group started GLP1Ra before 2014 when SGLT2i were not commercially available in Japan. The vertical axis shows years and the horizontal axis shows numbers of patients.

Kobayashi, K., Toyoda, M., Tone, A., Kawanami, D., Suzuki, D., Tsuriya, D., et al. (2023). Renoprotective effects of combination treatment with sodium-glucose cotransporter inhibitors and GLP-1 receptor agonists in patients with type 2 diabetes mellitus according to preceding medication. *Diab Vasc. Dis. Res.* 20, 14791641231222837

Kume, S., Araki, S. I., Ugi, S., Morino, K., Koya, D., Nishio, Y., et al. (2019). Secular changes in clinical manifestations of kidney disease among Japanese adults with type 2 diabetes from 1996 to 2014. *J. Diabetes Investig.* 10, 1032–1040. doi:10.1111/jdi.12977

Matsuo, S., Imai, E., Horio, M., Yasuda, Y., Tomita, K., Nitta, K., et al. (2009). Revised equations for estimated GFR from serum creatinine in Japan. *Am. J. Kidney Dis.* 53, 982–992. doi:10.1053/j.ajkd.2008.12.034

Mosenzon, O., Wiviott, S. D., Cahn, A., Rozenberg, A., Yanuv, I., Goodrich, E. L., et al. (2019). Effects of dapagliflozin on development and progression of kidney disease in patients with type 2 diabetes: an analysis from the DECLARE-TIMI 58 randomised trial. *Lancet Diabetes Endocrinol.* 7, 606–617. doi:10.1016/S2213-8587(19)30180-9

Norris, K. C., Smoyer, K. E., Rolland, C., Van Der Vaart, J., and Grubb, E. B. (2018). Albuminuria, serum creatinine, and estimated glomerular filtration rate as predictors of cardio-renal outcomes in patients with type 2 diabetes mellitus and kidney disease: a systematic literature review. *BMC Nephrol.* 19, 36. doi:10.1186/s12882-018-0821-9

Nuffield Department of Population Health Renal Studies Group, SGLT2 inhibitor Meta-Analysis Cardio-Renal Trialists' ConsortiumSGLT2 inhibitor Meta-Analysis Cardio-Renal Trialists' Consortium (2022). Impact of diabetes on the effects of sodium glucose co-transporter-2 inhibitors on kidney outcomes: collaborative metaanalysis of large placebo-controlled trials. *Lancet* 400, 1788–1801. doi:10.1016/S0140-6736(22)02074-8

Perkovic, V., De Zeeuw, D., Mahaffey, K. W., Fulcher, G., Erondu, N., Shaw, W., et al. (2018). Canagliflozin and renal outcomes in type 2 diabetes: results from the CANVAS Program randomised clinical trials. *Lancet Diabetes Endocrinol.* 6, 691–704. doi:10. 1016/S2213-8587(18)30141-4

Perkovic, V., Jardine, M. J., Neal, B., Bompoint, S., Heerspink, H. J. L., Charytan, D. M., et al. (2019). Canagliflozin and renal outcomes in type 2 diabetes and nephropathy. *N. Engl. J. Med.* 380, 2295–2306. doi:10.1056/NEJMoa1811744

Rossing, P., Caramori, M. L., Chan, J. C. N., Heerspink, H. J. L., Hurst, C., Khunti, K., et al. (2022). Executive summary of the KDIGO 2022 clinical Practice guideline for diabetes management in chronic kidney disease: an update based on rapidly emerging new evidence. *Kidney Int.* 102, 990–999. doi:10.1016/j.kint.2022.06.013

Rubin, D. B. (1987). Multiple imputation for nonresponse in surveys. John Wiley and Sons.

Rubin, D. B., and Schenker, N. (1991). Multiple imputation in health-care databases: an overview and some applications. *Stat. Med.* 10, 585–598. doi:10.1002/sim. 4780100410

Shikata, K., Kodera, R., Utsunomiya, K., Koya, D., Nishimura, R., Miyamoto, S., et al. (2020). Prevalence of albuminuria and renal dysfunction, and related clinical factors in

Japanese patients with diabetes: the Japan Diabetes Complication and its Prevention prospective study 5. J. Diabetes Investig. 11, 325–332. doi:10.1111/jdi.13116

Shimamoto, K., Ando, K., Fujita, T., Hasebe, N., Higaki, J., Horiuchi, M., et al. (2014). The Japanese society of hypertension guidelines for the management of hypertension (JSH 2014). *Hypertens. Res.* 37, 253–390. doi:10.1038/hr.2014.20

Soejima, H., Ogawa, H., Morimoto, T., Okada, S., Matsumoto, C., Nakayama, M., et al. (2022). Kidney function deterioration is dependent on blood pressure levels: 11.2 year followup in diabetic patients. *Heart Vessels* 37, 1873–1881. doi:10.1007/s00380-022-02085-0

Solomon, S. D., Mcmurray, J. J. V., Claggett, B., De Boer, R. A., Demets, D., Hernandez, A. F., et al. (2022). Dapagliflozin in heart failure with mildly reduced or preserved ejection fraction. *N. Engl. J. Med.* 387, 1089–1098. doi:10.1056/NEJMoa2206286

Sumida, K., Nadkarni, G. N., Grams, M. E., Sang, Y., Ballew, S. H., Coresh, J., et al. (2020). Conversion of urine protein-creatinine ratio or urine dipstick protein to urine albumincreatinine ratio for use in chronic kidney disease screening and prognosis: an individual participant-based meta-analysis. *Ann. Intern Med.* 173, 426–435. doi:10.7326/M20-0529 The EMPA-KIDNEY Collaborative Group, Herrington, W. G., Staplin, N., Wanner, C., Green, J. B., Hauske, S. J., Emberson, J. R., et al. (2023). Empagliflozin in patients with chronic kidney disease. *N. Engl. J. Med.* 388, 117–127. doi:10.1056/ NEJMoa2204233

Tuttle, K. R., Bosch-Traberg, H., Cherney, D. Z. I., Hadjadj, S., Lawson, J., Mosenzon, O., et al. (2023). *Post hoc* analysis of SUSTAIN 6 and PIONEER 6 trials suggests that people with type 2 diabetes at high cardiovascular risk treated with semaglutide experience more stable kidney function compared with placebo. *Kidney Int.* 103, 772-781. doi:10.1016/j.kint.2022.12.028

Wanner, C., Inzucchi, S. E., Lachin, J. M., Fitchett, D., Von Eynatten, M., Mattheus, M., et al. (2016). Empagliflozin and progression of kidney disease in type 2 diabetes. *N. Engl. J. Med.* 375, 1801–1802. doi:10.1056/NEJMc1611290

Yamamoto, Y., Hanai, K., Mori, T., Yokoyama, Y., Yoshida, N., Murata, H., et al. (2022). Kidney outcomes and all-cause mortality in people with type 2 diabetes exhibiting non-albuminuric kidney insufficiency. *Diabetologia* 65, 234–245. doi:10. 1007/s00125-021-05590-5