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RECEIVED 01 March 2024 ACCEPTED 07 January 2025 PUBLISHED 03 February 2025

#### CITATION

Mei P, Huang L, Lin L, Chen Y and Guo X (2025) The prognostic utility of the ratio of lymphocyte to monocyte in patients with metastatic colorectal cancer: a systematic review and meta-analysis. *Front. Oncol.* 15:1394154. doi: 10.3389/fonc.2025.1394154

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# The prognostic utility of the ratio of lymphocyte to monocyte in patients with metastatic colorectal cancer: a systematic review and meta-analysis

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**Background:** Although advancements in treatment have resulted in improved overall outcomes for patients diagnosed with colorectal cancer, the prognosis remains unfavorable for individuals with metastatic colorectal cancer (mCRC). The association between lymphocyte/monocyte ratio (LMR) and clinical outcomes in mCRC patients is a subject of controversy. To systematically evaluate the correlation between LMR and prognostic factors in individuals with mCRC, we conducted this meta-analysis.

**Methods:** The databases PubMed, Embase, Web of Science, and the Cochrane Library were systematically searched for all relevant studies from their inception until October 26, 2024. Study selection was conducted based on predetermined inclusion and exclusion criteria. The primary outcomes of interest included prognosis measures such as overall survival (OS), progression-free survival (PFS), disease-free survival (DFS), and cancer-specific survival (CSS) in patients with metastatic colorectal cancer. Random-effects models or fixed-effects models were used to determine the pooled risk ratio (HR) and corresponding 95% confidence interval (CI) for each outcome indicator. Additionally, the pooled odds ratio (OR) and its corresponding 95% CI were calculated for LMR and clinicopathological characteristics.

**Results:** Fourteen studies involving 3,089 patients were included in the analysis. The pooled analysis found that high LMR was correlated with better OS (HR: 0.55, 95% CI: 0.49-0.62, p<0.00001), PFS (HR: 0.68, 95% CI: 0.57-0.81, p<0.0001) and CSS(HR: 0.55, 95% CI: 0.32-0.95, p=0.03),The prognostic value of high LMR values for DFS(HR: 0.93, 95% CI: 0.78-1.12, p=0.46) in patients with metastatic rectal cancer was not found to be significant. We performed subgroup analyses based on study characteristics to confirm the robustness of our findings. Further clinicopathological analysis showed no significant difference between patients with elevated LMR and those without elevated LMR.

**Conclusions:** In conclusion, the results demonstrate a robust correlation between elevated LMR levels and a favorable prognosis in terms of overall survival (OS), progression-free survival (PFS), and cancer-specific survival (CSS) among patients diagnosed with metastatic colorectal cancer. However, further

high-quality prospective studies are warranted to validate our findings since the majority of current investigations have relied on retrospective study designs.

**Systematic review registration:** https://www.crd.york.ac.uk/prospero/display\_record.php?ID=CRD42024496467, identifier CRD42024496467.

KEYWORDS

lymphocyte-to-monocyte ratio, metastatic colorectal cancer, survival, prognosis, meta-analysis

## **1** Introduction

Colorectal cancer (CRC) ranks as the fourth most common cause of cancer-related deaths worldwide, resulting in nearly 900,000 fatalities annually (1). Metastasis is initially observed in approximately 20% to 30% of patients upon colorectal cancer diagnosis, and about 10% to 25% of patients develop metachronous metastasis following treatment-oriented surgery (2). Metastatic colorectal cancer (mCRC) carries a grim prognosis, with a 5-year survival rate of less than 20% (3). Many factors can predict the prognosis of colorectal cancer, such as tumor stage, cell differentiation, vascular invasion, nerve invasion, etc. However, some patients with good prognostic factors still have poor prognosis. Therefore, finding other novel biomarkers to predict the prognosis of colorectal cancer and help select the best treatment strategy remains a challenge in current clinical practice.

It is increasingly acknowledged that tumor occurrence is not solely determined by tumor-related factors but also significantly influenced by the immune status of the host (4-6). Recent studies have demonstrated that systemic inflammation facilitates tumor metastasis through various mechanisms (7). Therefore, the inflammatory response within the host system plays a pivotal role in tumor initiation, progression, malignant transformation, invasion, and metastasis. Alterations in peripheral blood cell composition including lymphocytes, monocytes, neutrophils, and platelets can be used to assess systemic inflammation. Parameters such as neutrophil-to-lymphocyte ratio (NLR), platelet-to-lymphocyte ratio (PLR), and systemic immune inflammation index (SII) are employed to characterize these changes (8). A large number of studies have shown that these biomarkers play a prognostic role in different tumors such as lung cancer, colorectal cancer, kidney cancer, and melanoma; higher NLR and PLR indicate worse prognosis (9-11). In recent years, there have been proposals to use the combination of lymphocytes and monocytes for measuring the host systemic inflammatory response, specifically the lymphocyte-to-monocyte ratio (LMR). The LMR, which can be easily calculated from a complete blood count of peripheral blood, has been reported as a novel prognostic indicator based on inflammation in recent years. Studies have shown that an increase in LMR is closely associated with

a favorable prognosis in patients with pancreatic cancer, lung cancer, liver cancer, ovarian cancer, and breast cancer (12), although some authors disagree due to differences in study design and sample size (13–16). A previous meta-analysis conducted by Hamid et al. (17) reported that LMR was superior to PLR as a predictor of long-term prognosis in colorectal patients.

Furthermore, previous studies have demonstrated that a lowpreconditioning lymphocyte-to-monocyte ratio (LMR) is an important prognostic biomarker for poor survival in rectal cancer patients who undergo radical resection or chemotherapy (18, 19). To the best of our knowledge, there are currently no systematic reviews and meta-analyses establishing an association between LMR and survival in metastatic colorectal cancer (mCRC) patients. Therefore, we conducted this meta-analysis to investigate the prognostic significance of LMR in patients diagnosed with mCRC, aiming to provide robust evidence-based medical support and objective data for clinical decision-makers.

## 2 Materials and methods

#### 2.1 Literature search

This systematic review study process followed the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA 2020) statement (20), and the protocol was registered with the International Prospective Registry of Systematic Reviews (PROSPERO: CRD42024496467). Two investigators, the MPP and LL, were responsible for developing the search strategy. We conducted a systematic literature search at PubMed, Cochrane Library, Embase, and Web of Science for eligible studies from database construction until October 26, 2024. The search strategy is based on the following keywords: colorectal tumor, "intestine tumor", "cecum tumor", "colon tumor", "Colorectal tumor" Neoplasms, prognosis, survival, Prognostic Factors, outcome, Lymphocytes, Monocytes, lymphocyte ratio(MLR) and lymphocyte monocyte ratio(LMR), detailed search strategies are shown in Supplementary Table S1. Additionally, the two researchers independently screened titles and abstracts, obtained full-text articles, and evaluated them to identify qualified studies. Any disagreements in the literature search were resolved through consensus.

### 2.2 Study selection

The inclusion criteria are as follows:(1) The diagnosis of mCRC was made based on pathological examination (or) the current clinical practice guidelines; (2) association of pretreatment LMR with overall survival (OS), disease free survival (DFS), cancer special survival (CSS), or other survival rates were reported; 3)cox regression; (4) Risk ratio (HR) with 95% confidence interval (CI) can be extracted or calculated from the literature; (5) Patients were divided into high and low LMR groups according to the truncation value; (6) The full text of the article is published; (7) The data collected primarily from distant (M) metastatic colorectal cancer. The exclusion criteria are as follows: (1) Repeatedly published studies or secondary analyses, in vitro experiments, animal experiments, reviews, letters, guidelines, case reports, pathological mechanisms, conference abstracts, or systematic reviews; (2) Colorectal cancer patients without metastasis; (3) There is no cutoff value; (4) Overlapping or duplicate data.

#### 2.3 Data extraction

Two researchers (PP.M, LL) independently screened all articles identified by the database. First, EndNote X 9.0 software was used to remove duplicate literature, excluding case reports, conference abstracts, letters, and review articles. Then, the titles and abstracts of the remaining literature were screened to exclude studies that were not relevant to the topic, and the full texts and supplementary information of the remaining studies were reviewed to identify eligible studies. Finally, data were extracted according to a unified extraction table, which included information such as first author name, year of publication, country, Study type, sample size, patient age, study period, TNM stage, follow-up time, cut-off value, and HR (95% CI) of Study center, OS, PFS, DFS, and CSS. The two researchers (PP.M,LL) independently cross-checked all included papers and extracted data. Any disputed articles were referred to a third researcher (XT,G) for consensus resolution.

#### 2.4 Quality assessment

The Newcastle-Ottawa Quality Assessment Scale (NOS) was used for independent evaluation in terms of selection, comparability, and outcome (21). A total of 9 items were extracted, and studies achieving scores of 1 as well as 7-9 in each item were considered to be of high quality (22).

#### 2.5 Statistical analysis

The pooled HR and 95% CI were calculated, and the MLR correlation results were converted into LMR format as necessary to evaluate the prognostic value of LMR in mCRC patients. It should be noted that for studies reporting MLR data (23, 24), we took the reciprocal of the corresponding HR values and confidence intervals,

exchanging the upper and lower confidence limits to convert MLR into LMR values for ease of statistical analysis. Heterogeneity was assessed using Cochran's Q test and Higgins I<sup>2</sup> statistic. The I<sup>2</sup> value was used to evaluate the degree of heterogeneity, with I<sup>2</sup>  $\leq$  25%, 25%  $<I^2<75\%$ , and I<sup>2</sup> $\geq$ 75% indicating low, moderate, and high levels of heterogeneity respectively. Subsequently, a random effects model was employed for data analysis when I<sup>2</sup> exceeded 50%; otherwise, a fixed effects model was applied. Subgroup analysis and sensitivity analysis were conducted to validate the robustness of OS and PFS results. Funnel plots as well as Egger's and Begg's tests were performed to assess potential publication bias. A significance threshold of P<0.05 was adopted for evaluating STATA 15.0 software and Review Manager 5.

## **3** Results

#### 3.1 Study characteristics

Employing the search strategy mentioned above, an initial literature search yielded 1016 records. However, after removing duplicates, only 823 studies remained eligible for inclusion. Further screening based on title and abstract resulted in the exclusion of 791 records, leaving a final selection of 32 studies for comprehensive evaluation. Among these, seventeen additional studies were excluded primarily due to insufficient availability of relevant data required for survival information computation. Consequently, this meta-analysis includes fourteen selected publications spanning from years 2015 to 2023 and involving a collective cohort comprising a total sample size of n=3199 patients (13, 23–36) (Figure 1).

Eligible studies had the following baseline characteristics. Among the included articles, 14 examined the prognostic effect of LMR on OS, 3 examined DFS, 3 examined CSS, and 3 examined PFS. Eight studies were conducted in Asia and six in Europe. All studies were cohort studies, with 13 articles being retrospective and 2 being prospective. All cohort studies were published in English. All included studies consisted of two groups: high and low LMR. The study characteristics, patient baseline data, and study results of the included studies are listed in Table 1.

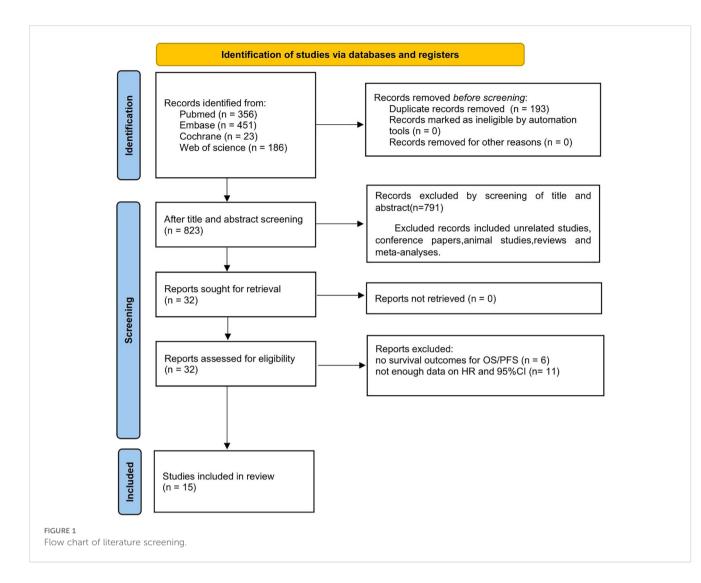
# 3.2 The quality assessment of the included studies

The methodological quality of the 15 included studies was assessed using the Newcastle-Ottawa scale (NOS), a well-established tool for evaluating research rigor. Scores on this scale, which range from 5 to 8 stars, are presented in Supplementary Table S2.

#### 3.3 Meta-analysis results

#### 3.3.1 The effect of LMR on OS

A total of 14 studies have reported the relationship between LMR levels and OS in patients with metastatic colorectal cancer (13, 23–31, 33–36), involving a total of 3082 patients. Heterogeneity



test found no significant heterogeneity among the studies (I<sup>2</sup> = 37%, p=0.008). Therefore, we performed a meta-analysis using a fixed effects model. Pooled analysis showed that patients with metastatic colorectal cancer with high LMR values were significantly associated with better OS (HR: 0.55, 95%CI: 0.49-0.62, p < 0.00001, Figure 2A).

#### 3.3.2 The effect of LMR on PFS

A total of 3 studies have reported the relationship between LMR level and PFS in patients with metastatic colorectal cancer (24, 26, 27), involving a total of 716 patients. Heterogeneity test found no significant heterogeneity among the studies ( $I^2 = 0\%$ , p=0.84). Therefore, we performed a meta-analysis using a fixed effects model. Pooled analysis showed that patients with metastatic colorectal cancer with high LMR values were significantly associated with better PFS (HR: 0.68, 95%CI: 0.57-0.81, p < 0.0001, Figure 2B).

#### 3.3.3 The effect of LMR on DFS

A total of 3 studies have reported the relationship between LMR levels and DFS in patients with metastatic colorectal cancer (30, 32, 35), involving a total of 709 patients. Heterogeneity test found

significant heterogeneity among studies ( $I^2 = 6\%$ , p=0.35). Therefore, we performed a meta-analysis using a fixed effects model. Pooled analysis showed no significant association between patients with metastatic colorectal cancer with high LMR values and DFS (HR: 0.93, 95%CI: 0.78-1.12, p=0.46, Figure 2C).

#### 3.3.4 The effect of LMR on CSS

A total of 3 studies have reported the relationship between LMR levels and CSS in patients with metastatic colorectal cancer (29, 30, 32), involving a total of 569 patients. Heterogeneity test found no significant heterogeneity among the studies ( $I^2 = 72\%$ , p=0.03). Therefore, we performed a meta-analysis using a random effects model. Pooled analysis showed that patients with metastatic colorectal cancer with high LMR values were significantly associated with better CSS (HR: 0.55, 95%CI: 0.32-0.95, p=0.03, Figure 2D).

### 3.4 Subgroup analysis

To detect potential heterogeneity, we performed subgroup analyses based on sample size, cutoff values, study centers, year of

Author	Year	Sources of patients	Study design	sample size	Age	Duration	TNM stage	Study center	Treatment	Follow- up (median)	LMR cut-off	Outcome	NOS score
Shibutani et al. (13)	2015	Japan	retrospective	104	64	2005-2010	IV	Single center	С	NA	3.38	OS	6
Basile et al. (23)	2020	Italy	retrospective	528	NA	2009-2018	IV	Multicenter	С	55	2.04	OS	8
Li et al. (24)	2022	China	retrospective	196	NA	2010-2015	IV	Single center	S	NA	3.27	OS,PFS	6
Facciorusso et al. (25)	2016	Italy	retrospective	127	66	2003-2012	IV	Single center	RFA	63	3.96	OS	7
Kuramochi et al. (26)	2021	Japan	prospective	32	67	2016-2017	IV	Multicenter	С	NA	3.18	OS,PFS	8
Lin et al. (27)	2016	China	retrospective	488	54	2005-2013	IV	Single center	С	23.5	3.11	OS,PFS	7
Lisanti et al. (28)	2020	Italy	retrospective	168	NA	2005-2018	IV	Multicenter	С	51.5	2.22	OS	7
Neal et al. (29)	2015	UK	prospective	312	66	2006-2010	IV	Single center	S	29.7	2.35	OS,CSS	8
Neofytou et al. (30)	2015	UK	retrospective	140	NA	2005-2012	IV	Single center	Neoadjuvant + S	33	3	OS,CSS,DFS	7
Ouyang et al. (31)	2023	China	retrospective	110	59	2019-2022	IV	Single center	I\I+C\I+T\I +T+C	NA	1.49	OS	7
Ozawa et al. (32)	2015	Japan	retrospective	117	63	1997-2012	IV	Multicenter	C+S	39	3	CSS,DFS	8
Peng et al. (33)	2017	China	retrospective	150	NA	2000 -2012	IV	Single center	C/S/RFA	36	2.82	OS	8
Song et al. (34)	2015	Korea	retrospective	177	52	2006-2013	IV	Single center	S+C+R	3.1	3.4	OS	7
Wang et al. (35)	2019	China	retrospective	452	57	2002-2016	IV	Single center	S+/R/C/RFA	28	3.4	OS,DFS	8
Zager et al. (36)	2020	Israel	retrospective	98	59.2	2009-2018	IV	Single center	S+C	NA	4.4	OS	6

NOS, Newcastle–Ottawa scale; OS, overall survival; PFS, progression-free survival; DFS, disease free survival; CSS, cancer special survival; LMR, lymphocyte/monocyte ratio; S, surgery; C, chemotherapy; R, radiotherapy; RFA, radiofrequency ablation; I, immunotherapy; T, targeted therapy; NA, not available.

TABLE 2 Pooled HRs for	or OS in	subgroup	analyses.
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Subgroup	Number of Studies	Patient	HR [95%CI]	P value	l <sup>2</sup>
Total	14	3082	0.55 [0.49-0.62]	<0.00001	37%
Sample size					
≤150	7	761	0.44 [0.34-0.55]	<0.00001	0%
>150	7	2321	0.60 [0.52-0.69]	<0.00001	49%
LMR cut-off					
LMR≤3	6	1408	0.49 [0.40-0.59]	<0.00001	28%
LMR>3	8	1674	0.60 [0.52-0.69]	<0.00001	37%
Study center					
Multicenter	3	728	0.43 [0.32-0.57]	<0.00001	0%
Single center	11	2354	0.57 [0.50-0.64]	<0.00001	37%
Publish year					
2015-2017	7	1498	0.56 [0.48-0.65]	<0.00001	0%
2018-2023	7	1584	0.54 [0.45-0.65]	<0.00001	62%
Region					
Asia	8	1709	0.59 [0.51-0.69]	<0.00001	44%
Europe	6	1373	0.50 [0.41-0.60]	<0.00001	22%

R, hazard ratio; CI, confidence Interval; OS, overall survival; LMR, lymphocyte/ monocyte ratio.

publication, and region. Table 2 shows the results of these analyses. Our findings suggest that patients with high LMR had better OS regardless of study center (single or multi-center), region of publication (Asia or Europe), year of publication (2015-2017 or 2018-2023), sample size ( $\leq 150$  or >150), and cutoff ( $\leq 3$  or >3). However, in the articles published from 2018 to 2023, the heterogeneity test found significant heterogeneity across studies ( $I^2 = 62\%$ ).

# 3.5 Correlation between LMR and clinicopathological parameters

Information on LMR and tumor size in mCRC patients was retrieved from four cohorts, which were summarized (OR=1.75, 95%CI: 1.17-2.61, P=0.007, Supplementary Figure S1A), indicated that patients with elevated LMR have smaller tumor size. Four studies reported an association between high LMR and carcinoembryonic antigen, however, the combined data did not show such an association (OR=1.52, 95%CI: 0.83-2.78, P=0.18, Supplementary Figure S1B). Three studies reported the time to metastasis of LMR in mCRC patients, the combined OR of 1.36 (95% CI: 0.89-2.08, P =0.16, Supplementary Figure S1C) suggested that elevated LMR was not associated with time to metastasis for these patients. In four studies evaluating LMR and the level of tumor differentiation in CRC patients, the results showed no correlation (OR=0.50, 95% CI: 0.24-1.04, P =0.06, Supplementary Figure S1D) (see Table 3).

### 3.6 Sensitivity analysis

Sensitivity analysis assesses the impact of each study on the overall outcome by removing the included studies one at a time. We conducted a sensitivity analysis on LMR and OS, and found that the effect size of each study after the removal of each study in turn had little change and was still within the original range, which proved that no single study affected the results of OS (Figure 3), indicating that the results were relatively credible.

## 3.7 Publication bias

Among the above results, there were enough articles (>10 studies) included in the analysis of the relationships between LMR at baseline and OS. To assess publication bias, HR for OS and its associated 95% CI were aggregated and evaluated using funnel plots as well as Begg and Egger tests. The shape of the funnel plot indicates that there is publication bias in the included LMR and OS studies (Egger's P=0.003 and Begg's P=0.006; Figure 4A). Next, we used trim-and-fill methods to evaluate the symmetry of the funnel chart, by supplementing unpublished research, and the final result showed that the difference in effect sizes before and after the Trim-and-fill method was NS, indicating that publication bias had little effect on the results of the Meta-analysis (Figure 4B). However, due to a limited number of meaningfully evaluated studies (<10 studies), no publication bias analysis was performed for others.

## 4 Discussion

Tumor-related inflammation is still an important research area, and a large number of studies have shown its influence on the occurrence and progression of cancer (37). Blood-derived parameters provide an easily accessible and repeatable method for assessing systemic inflammation as an objective biomarker for predicting patient outcomes (38, 39). However, there is evidence that a high lymphocyte-to-monocyte ratio is associated with improved survival

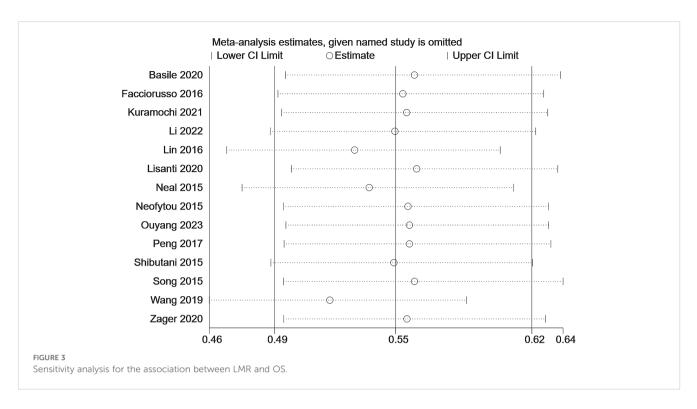
						Hazard Ratio	Hazard Ratio	
		Study or Subgroup	log[Hazard Ratio]	SE	Weight	IV, Fixed, 95% CI	IV, Fixed, 95% Cl	
	_	Basile 2020		0.1961	9.4%			-
,		Facciorusso 2016	-0.7133			0.49 [0.29, 0.83]	<b>_</b>	
,		Kuramochi 2021	-0.7985			0.45 [0.26, 0.78]	_ <b>_</b>	
		Li 2022	-0.5798			0.56 [0.35, 0.90]		
		Lin 2016		0.1417		0.66 [0.50, 0.87]		
		Lisanti 2020	-1.1712			0.31 [0.16, 0.60]	(	
		Neal 2015	-0.4463			0.64 [0.47, 0.87]		
		Neofytou 2015	-0.8916			0.41 [0.22, 0.76]	<b>_</b>	
		Ouyang 2023	-1.3093			0.27 [0.11, 0.66]		
		Peng 2017		0.2767		0.43 [0.25, 0.74]	_ <b>_</b>	
		Shibutani 2015	-0.5447			0.58 [0.31, 1.09]		
		Song 2015		0.1717				
		Wang 2019	-0.1165		11.6%			
		Zager 2020	-1.3093			0.27 [0.10, 0.73]		
		20301 2020	1.0000	0.0000	1.170	0.21 [0.10, 0.10]		
		Total (95% CI)			100.0%	0.55 [0.49, 0.62]	•	
		Heterogeneity: Chi <sup>2</sup> =	20.76, df = 13 (P = 0	.08); I <sup>2</sup> =	37%			
		Test for overall effect:	Z = 9.82 (P < 0.0000	)1)			Favours [High LMR] Favours [Low LMR]	
						Hazard Ratio	Hazard Ratio	
		Study or Subgroup	log[Hazard Ratio]	SE	Weight	IV, Fixed, 95% CI	IV, Fixed, 95% Cl	_
E		Kuramochi 2021	-0.4005		16.9%	0.67 [0.44, 1.02]		
Ľ		Li 2022		0.1605	30.2%	0.63 [0.46, 0.86]		
		Lin 2016	-0.3425	0.1211	53.0%	0.71 [0.56, 0.90]		
					400.0%	0 60 10 57 0 041		
		Total (95% CI)	0.00 +/- 0.00 - 0.0	0.17-00		0.68 [0.57, 0.81]	•	
		Heterogeneity: Chi <sup>2</sup> =			0		0.01 0.1 1 10 100	
		Test for overall effect:	Z = 4.41 (F < 0.0001	,			Favours [High LMR] Favours [Low LMR]	
		Study or Subgroup	log[Hazard Ratio]	SE	Woight	Hazard Ratio IV, Fixed, 95% Cl	Hazard Ratio IV, Fixed, 95% Cl	
		Neofytou 2015	-0.1863		19.3%	0.83 [0.55, 1.25]		-
0				0.2358				
		OTOWO 2015						
		Ozawa 2015 Wong 2019				1.27 [0.80, 2.02]		
		Ozawa 2015 Wang 2019			65.5%		+	
		Wang 2019			65.5%	0.90 [0.72, 1.13]	•	
		Wang 2019 Total (95% CI)	-0.1054	0.1139	65.5% <b>100.0</b> %		•	
		Wang 2019 <b>Total (95% CI)</b> Heterogeneity: Chi <sup>2</sup> =	-0.1054 2.12, df = 2 (P = 0.35	0.1139	65.5% <b>100.0</b> %	0.90 [0.72, 1.13]		
		Wang 2019 Total (95% CI)	-0.1054 2.12, df = 2 (P = 0.35	0.1139	65.5% <b>100.0</b> %	0.90 [0.72, 1.13]	0.01 0.1 1 10 100 Favours [High LMR] Favours [Low LMR]	
		Wang 2019 <b>Total (95% CI)</b> Heterogeneity: Chi <sup>2</sup> =	-0.1054 2.12, df = 2 (P = 0.35	0.1139	65.5% <b>100.0</b> %	0.90 (0.72, 1.13) 0.93 (0.78, 1.12)	Favours [High LMR] Favours [Low LMR]	
		Wang 2019 Total (95% Cl) Heterogeneity: Chi <sup>a</sup> = Test for overall effect:	-0.1054 2.12, df = 2 (P = 0.38 Z = 0.74 (P = 0.46)	0.1139 5); I <sup>z</sup> = 6%	65.5% <b>100.0</b> %	0.90 (0.72, 1.13) 0.93 (0.78, 1.12) Hazard Ratio	Favours [High LMR] Favours [Low LMR] Hazard Ratio	
		Wang 2019 Total (95% CI) Heterogeneity: Chi <sup>2</sup> = Test for overall effect: Study or Subgroup	-0.1054 2.12, df = 2 (P = 0.38 Z = 0.74 (P = 0.46) log[Hazard Ratio]	0.1139 5); I <sup>2</sup> = 6% <u>SE</u>	65.5% 100.0% 6	0.90 [0.72, 1.13] 0.93 [0.78, 1.12] Hazard Ratio IV, Random, 95%	Favours [High LMR] Favours [Low LMR] Hazard Ratio .CI IV, Randon, 95% CI	
Г		Wang 2019 Total (95% CI) Heterogeneity: Chi <sup>a</sup> = Test for overall effect: <u>Study or Subgroup</u> Neal 2015	-0.1054 2.12, df = 2 (P = 0.3; Z = 0.74 (P = 0.46) <u>log[Hazard Ratio]</u> -0.2231	0.1139 5); I <sup>2</sup> = 6% <u>SE</u> 0.1059	65.5% 100.0% 6 <u>Weight</u> 44.3%	0.90 [0.72, 1.13] 0.93 [0.78, 1.12] Hazard Ratio <u>IV, Random, 95%</u> 0.80 [0.65, 0.9	Favours [High LMR] Favours [Low LMR] Hazard Ratio CI IV, Random, 95% CI 98]	
C		Wang 2019 Total (95% CI) Heterogeneifly: Chi <sup>2</sup> = Test for overall effect: Study or Subgroup Neal 2015 Neofytou 2015	-0.1054 2.12, df = 2 (P = 0.38 Z = 0.74 (P = 0.46) log[Hazard Ratio] -0.2231 -0.755	0.1139 5); I <sup>2</sup> = 6% <u>SE</u> 0.1059 0.3429	65.5% 100.0% 6 <u>Weight</u> 44.3% 27.7%	0.90 [0.72, 1.13] 0.93 [0.78, 1.12] Hazard Ratio <u>IV, Random, 95%</u> 0.80 [0.65, 0.9 0.47 [0.24, 0.9	Favours [High LMR]           Hazard Ratio           CI         N, Random, 95% CI           38]	
Ľ		Wang 2019 Total (95% CI) Heterogeneity: Chi <sup>a</sup> = Test for overall effect: <u>Study or Subgroup</u> Neal 2015	-0.1054 2.12, df = 2 (P = 0.3; Z = 0.74 (P = 0.46) <u>log[Hazard Ratio]</u> -0.2231	0.1139 5); I <sup>2</sup> = 6% <u>SE</u> 0.1059 0.3429	65.5% 100.0% 6 <u>Weight</u> 44.3% 27.7%	0.90 [0.72, 1.13] 0.93 [0.78, 1.12] Hazard Ratio <u>IV, Random, 95%</u> 0.80 [0.65, 0.9	Favours [High LMR]           Hazard Ratio           CI         N, Random, 95% CI           38]	
Ľ		Wang 2019 Total (95% CI) Heterogeneifly: Chi <sup>2</sup> = Test for overall effect: Study or Subgroup Neal 2015 Neofytou 2015	-0.1054 2.12, df = 2 (P = 0.38 Z = 0.74 (P = 0.46) log[Hazard Ratio] -0.2231 -0.755	0.1139 5); I <sup>2</sup> = 6% <u>SE</u> 0.1059 0.3429	65.5% 100.0% 6 <u>Weight</u> 44.3% 27.7%	0.90 [0.72, 1.13] 0.93 [0.78, 1.12] Hazard Ratio <u>IV, Random, 95%</u> 0.80 [0.65, 0.9 0.47 [0.24, 0.9	Favours [High LMR] Favours [Low LMR] Hazard Ratio I. N. Random, 95% Cl	
Ľ	)	Wang 2019 Total (95% CI) Heterogeneity: Chi <sup>2</sup> = Test for overall effect: Study or Subgroup Neal 2015 Neofytou 2015 Ozawa 2015 Total (95% CI) Heterogeneity: Tau <sup>2</sup> =	-0.1054 2.12, df = 2 (P = 0.38 Z = 0.74 (P = 0.46) -0.2231 -0.755 -1.0498 : 0.17; Chi <sup>2</sup> = 7.05, df	0.1139 5); I <sup>2</sup> = 6% <u>SE</u> 0.1059 0.3429 0.3393	65.5% 100.0% 6 Weight 44.3% 27.7% 28.0% 100.0%	0.90 [0.72, 1.13] 0.93 [0.78, 1.12] Hazard Ratio <u>IV, Random, 95%</u> 0.80 [0.65, 0.9 0.47 [0.24, 0.9 0.35 [0.18, 0.6 0.55 [0.32, 0.9	Favours [High LMR] Favours [Low LMR] Hazard Ratio IV, Random, 95% Cl Ratio IV, Rativ I	 
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-	)	Wang 2019 Total (95% CI) Heterogeneity: Chi <sup>2</sup> = Test for overall effect: Study or Subgroup Neal 2015 Neofytou 2015 Ozawa 2015 Total (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect:	-0.1054 2.12, df = 2 (P = 0.3 Z = 0.74 (P = 0.46) -0.2231 -0.755 -1.0498 : 0.17; Chi <sup>2</sup> = 7.05, df Z = 2.14 (P = 0.03)	0.1139 5); I <sup>2</sup> = 6% 0.1059 0.3429 0.3393 f= 2 (P =	65.5% 100.0% 44.3% 27.7% 28.0% 100.0% 0.03);   <sup>2</sup> =	0.90 [0.72, 1.13] 0.93 [0.78, 1.12] Hazard Ratio <u>IV, Random, 95%</u> 0.80 [0.65, 0.9 0.47 [0.24, 0.9 0.35 [0.18, 0.6 0.55 [0.32, 0.9 72%	Favours [High LMR] Favours [Low LMR] Hazard Ratio CI V. Random, 95% CI 98] 92] 95] 0.01 0.1 1 10 10 Favours [High LMR] Favours [Low LMR]	
-	)	Wang 2019 Total (95% CI) Heterogeneity: Chi <sup>2</sup> = Test for overall effect: Study or Subgroup Neal 2015 Neofytou 2015 Ozawa 2015 Total (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect:	-0.1054 2.12, df = 2 (P = 0.3 Z = 0.74 (P = 0.46) -0.2231 -0.755 -1.0498 : 0.17; Chi <sup>2</sup> = 7.05, df Z = 2.14 (P = 0.03)	0.1139 5); I <sup>2</sup> = 6% 0.1059 0.3429 0.3393 f= 2 (P =	65.5% 100.0% 44.3% 27.7% 28.0% 100.0% 0.03);   <sup>2</sup> =	0.90 [0.72, 1.13] 0.93 [0.78, 1.12] Hazard Ratio <u>IV, Random, 95%</u> 0.80 [0.65, 0.9 0.47 [0.24, 0.9 0.35 [0.18, 0.6 0.55 [0.32, 0.9 72%	Favours [High LMR] Favours [Low LMR] Hazard Ratio CI V, Random, 95% CI 98] 92] 95] 0.01 0.1 1 10 10	

outcomes in these patients, and LMR remains a valuable predictor of long-term prognosis (40). In this study, reliable evidence from 15 studies in 3199 cases was comprehensively analyzed to explore the prognostic significance of LMR in mCRC patients. Combined hazard ratios consistently showed that increased LMR was significantly

associated with improved OS (HR: 0.55, 95%CI: 0.49-0.62, p < 0.00001). Subgroup analysis was performed to further evaluate the relationship between LMR and OS. Subgroup analysis stratified by truncation value, sample size, study center, region, and publication year was consistent with the combined results. In addition, higher

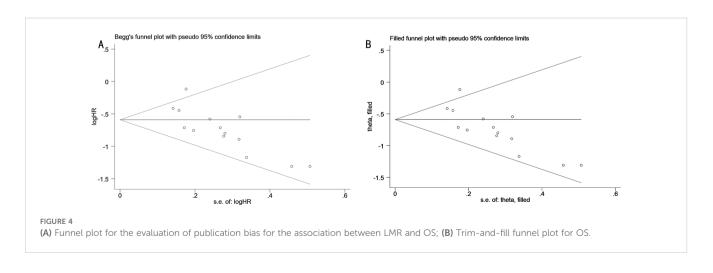
	Number of Studies	Meta-anal	ysis results	Effect model	Meta-analysis results		
		OR	Р		l <sup>2</sup> (%)	Р	
Primary tumor size	4	1.75 [1.17-2.61]	0.007	fix-effects model	0%	0.58	
CEA	4	1.52 [0.83-2.78]	0.18	fix-effects model	0%	0.44	
Time to metastasis	3	1.36 [0.89-2.08]	0.16	fix-effects model	0%	0.42	
Tumor differentiation	4	0.50 [0.24-1.04]	0.06	Random-effects model	71%	0.02	

primary tumor size, > 5 cm or < 5 cm; CEA, > CEA cut-off; time to metastasis, Synchronous transfer vs Metachronous transfer; Tumor differentiation, well or moderate vs poor. LMR, lymphocyte/monocyte ratio; mCRC, metastatic colorectal cancer; OR, odd ratio; CEA, Carcinoembryonic Antigen.



baseline LMR was positively associated with better PFS and CSS, suggesting that higher LMR may signal better outcomes in patients with mCRC. Exceptionally, high LMR values had no significant prognostic value for DFS in patients with metastatic colorectal cancer. Clinicopathological factors in different groups of patients affect their long-term outcomes. We compared four major clinicopathological factors in patients with elevated and normal LMR, and the results were comparable between the two groups. Finally, sensitivity analysis and publication bias assessment were performed, and the results were not statistically significant. Together, our findings suggest that LMR holds promise as a potential peripheral blood biomarker and an effective tool for stratifying patients who may benefit from treatment for mCRC.

In the past few decades, LMR as an indicator of inflammation, which can be measured in plasma or serum, has been more popular, easier to estimate, and more readily available than other indicators of systemic inflammation. Pre-treatment LMR has been shown in many studies to reflect systemic inflammation and is positively associated with the prognosis of various solid tumors, such as melanoma, classical Hodgkin lymphoma, and gastric cancer (41, 42). The lymphocyte and monocyte levels, which serve as indicators of antitumor immunity and tumor burden (43), are crucial for the determination of LMR. A decreased count of lymphocytes can impair the immune response against tumor cells, where T lymphocytes play a vital role in recognizing and eliminating these cells, thereby inhibiting tumor cell proliferation and metastasis (44). Tumor-infiltrating lymphocytes (TILs) are pivotal for cell-mediated antitumor immune responses, with increased TILs being associated with improved outcomes in cancer patients. CD4 T cell infiltration triggers the activation of CD8 T cells, promoting apoptosis and



cytotoxic activity against cancer cells (45, 46). Studies have shown that reduced lymphocyte counts can lead to reduced survival in various cancers and also diminish the effectiveness of ICIs (27, 47). In addition to lymphocytes, circulating monocytes possess the capacity to differentiate into tumor-associated macrophages (TAMs), which play a pivotal role in the context of tumor-associated inflammation. TAMs exhibit dual functionality by attracting tumor-associated chemokines to recruit peripheral blood monocytes and by modulating innate immune responses and regulatory T cells, thereby exerting suppressive effects on anti-tumor immunity while also promoting angiogenesis and extracellular matrix degradation. Consequently, these effects promote the occurrence and progression of tumors (25). Therefore, monocytes have a crucial function in the tumor immune microenvironment, where TAMs can be identified as high cancer burden markers (48). In this case, the lymphocyte-tomonocyte ratio (LMR) serves as a comprehensive indicator reflecting both the host's immune status and the extent of tumor development. A decreased LMR, characterized by reduced lymphocyte count and elevated monocyte count, signifies compromised antitumor immunity and increased tumor burden (34). Consequently, LMR holds promise as a more effective prognostic marker for predicting outcomes in patients with mCRC.

To our knowledge, this meta-analysis is the first to comprehensively examine the prognostic significance of LMR in patients with mCRC. Previous meta-analyses of colorectal cancer survival and systemic inflammatory measures have mostly mixed metastatic and non-metastatic colorectal cancer together (16, 18, 19), which may obscure the true relationship between outcomes and prognostic measures, especially given differences in treatment strategies and survival. To the best of our knowledge, only one prior study has conducted a meta-analysis on the prognostic significance of platelet-to-lymphocyte ratio (PLR) in patients with metastatic colorectal cancer (49). However, this meta-analysis represents the first comprehensive evaluation of the prognostic value of LMR in patients with metastatic colorectal cancer. Our study included a large number of qualified, high-quality studies, further confirming that elevated LMR is associated with good OS. These findings are consistent with previous studies of LMR and prognosis in other tumor types. In addition, the subgroup analysis was consistent with the combined results. However, this conclusion may have been influenced by the small number of studies included. Therefore, the prognostic value of LMR needs to be further validated through large-scale and rigorously designed studies.

Although this study provides more detailed outcome measures that enhance the reliability of the conclusions and provide more robust evidence-based medical insights, our meta-analysis has some limitations. First, all eligible studies were from Asia and Europe, mostly from China and Japan, which limits generalizability in regions such as Africa or the Americas. Second, we do not impose restrictions on nationality or region in selecting eligible studies. Only Englishlanguage publications were considered. Therefore, it remains necessary to validate the prognostic effect of LMR in mCRC patients treated worldwide. Third, the majority of the included studies were retrospective in nature and had limited sample sizes. Only three studies assessed the significance of LMR in predicting PFS, which may not provide sufficient evidence to establish robust findings. Additionally, there was insufficient data available to support the consistency of outcomes between LMR and DFS as well as CSS. Fourth, the studies included in the analysis had varying LMR cutoff values, ranging from 1.49 to 4.4. To avoid the inherent heterogeneity caused by data inconsistencies in this meta-analysis, a standard cutoff value for LMR is required for future studies. Fifth, because the treatment protocols used in each trial are not uniform, such as radiotherapy and chemotherapy alone, radiotherapy and chemotherapy addet the reliability of the conclusion. In the future, the effect of the same treatment regimen on the relationship between LMR and prognosis needs to be further discussed.

## 5 Conclusion

In conclusion, this study shows that high levels of LMR are strongly associated with survival outcomes in patients with metastatic colorectal cancer, as well as favorable outcomes for OS, CSS, and PFS. The present findings underscore the capacity of LMR to serve as an independent prognostic biomarker in patients diagnosed with mCRC, thereby offering valuable insights for treatment decision-making. In the future, the clinical value of these biomarkers still needs to be further evaluated in prospective, large-scale, and multicenter studies.

## Data availability statement

The original contributions presented in the study are included in the article/Supplementary Material. Further inquiries can be directed to the corresponding authors.

## Author contributions

PM: Conceptualization, Data curation, Formal analysis, Methodology, Project administration, Writing – original draft, Writing – review & editing. LH: Conceptualization, Investigation, Supervision, Writing – review & editing. LL: Data curation, Formal analysis, Investigation, Methodology, Software, Writing – review & editing. YC: Funding acquisition, Project administration, Resources, Supervision, Writing – review & editing. XG: Funding acquisition, Investigation, Methodology, Resources, Supervision, Validation, Writing – review & editing.

## Funding

The author(s) declare financial support was received for the research, authorship, and/or publication of this article. This study

was financially supported by the National Natural Science Foundation of China (81873308 and 82305230); Shanghai University of Traditional Chinese Medicine Industrial Development Center Combined Medical and Nursing Science and Technology Innovation Project (YYKC-2021-01-14); Future Plan for Traditional Chinese Medicine development of Science and Technology of Shanghai Municipal Hospital of Traditional Chinese Medicine (WL-HBQN-2122013K); Future Plan for Traditional Chinese Medicine Inheritance and Development of Shanghai Municipal Hospital of Traditional Chinese Medicine (WLJH2021ZY-MZY024 and WLJH2021ZY-GZS003); Shanghai Municipal Science Commission Project (23Y11920800).

## Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

## References

1. Ferlay J, Colombet M, Soerjomataram I, Parkin DM, Piñeros M, Znaor A, et al. Cancer statistics for the year 2020: An overview. *Int J Cancer*. (2021). doi: 10.1002/ ijc.v149.4

2. Biller LH, Schrag D. Diagnosis and treatment of metastatic colorectal cancer: A review. JAMA. (2021) 325:669–85. doi: 10.1001/jama.2021.0106

3. Malakorn S, Ouchi A, Hu CY, Sandhu L, Dasari A, You YN, et al. Tumor sidedness, recurrence, and survival after curative resection of localized colon cancer. *Clin Colorectal Cancer*. (2021) 20:e53–60. doi: 10.1016/j.clcc.2020.08.007

4. Planchard D, Popat S, Kerr K, Novello S, Smit EF, Faivre-Finn C, et al. Metastatic non-small cell lung cancer: ESMO Clinical Practice Guidelines for diagnosis, treatment and follow-up. *Ann Oncol.* (2018) 29:iv192–237. doi: 10.1093/annonc/mdy275

5. Bhardwaj PV, Abdou YG. The evolving landscape of immune checkpoint inhibitors and antibody drug conjugates in the treatment of early-stage breast cancer. *Oncologist.* (2023) 28:832–44. doi: 10.1093/oncolo/oyad233

6. Chen DS, Mellman I. Elements of cancer immunity and the cancer-immune set point. *Nature*. (2017) 541:321-30. doi: 10.1038/nature21349

7. Pikuła A, Skórzewska M, Pelc Z, Mlak R, Gęca K, Sędłak K, et al. Prognostic value of systemic inflammatory response markers in patients undergoing neoadjuvant chemotherapy and gastrectomy for advanced gastric cancer in the eastern european population. *Cancers (Basel).* (2022) 14:1997. doi: 10.3390/cancers14081997

8. Leone P, Buonavoglia A, Fasano R, Solimando AG, De Re V, Cicco S, et al. Insights into the regulation of tumor angiogenesis by micro-RNAs. *J Clin Med.* (2019) 8:1–20. doi: 10.3390/jcm8122030

9. Guthrie GJ, Charles KA, Roxburgh CS, Horgan PG, McMillan DC, Clarke SJ. The systemic inflammation-based neutrophil-lymphocyte ratio: experience in patients with cancer. *Crit Rev Oncol Hematol.* (2013) 88:218–30. doi: 10.1016/j.critrevonc.2013.03.010

10. Suh KJ, Kim SH, Kim YJ, Kim M, Keam B, Kim TM, et al. Post-treatment neutrophil-to-lymphocyte ratio at week 6 is prognostic in patients with advanced non-small cell lung cancers treated with anti-PD-1 antibody. *Cancer Immunol Immunother*. (2018) 67:459–70. doi: 10.1007/s00262-017-2092-x

11. Urakawa S, Yamasaki M, Goto K, Haruna M, Hirata M, Morimoto-Okazawa A, et al. Peri-operative monocyte count is a marker of poor prognosis in gastric cancer: increased monocytes are a characteristic of myeloid-derived suppressor cells. *Cancer Immunol Immunother*. (2019) 68:1341–50. doi: 10.1007/s00262-019-02366-0

12. Stotz M, Pichler M, Absenger G, Szkandera J, Arminger F, Schaberl-Moser R, et al. The preoperative lymphocyte to monocyte ratio predicts clinical outcome in patients with stage III colon cancer. *Br J Cancer*. (2014) 110:435–40. doi: 10.1038/bjc.2013.785

13. Shibutani M, Maeda K, Nagahara H, Ohtani H, Sakurai K, Yamazoe S, et al. Prognostic significance of the lymphocyte-to-monocyte ratio in patients with metastatic colorectal cancer. *World J Gastroenterol.* (2015) 21:9966–73. doi: 10.3748/wjg.v21.i34.9966

14. Wu QB, Wang M, Hu T, He WB, Wang ZQ. Prognostic role of the lymphocyteto-monocyte ratio in patients undergoing resection for nonmetastatic rectal cancer. *Med (Baltimore)*. (2016) 95:e4945. doi: 10.1097/MD.000000000004945

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## Supplementary material

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fonc.2025.1394154/ full#supplementary-material

#### SUPPLEMENTARY FIGURE 1

(A) Forest plots for the association between LMR and primary tumor size;
 (B) Forest plots for the association between LMR and CEA; (C) Forest plots for the association between LMR and time to metastasis; (D) Forest plots for the association between LMR and Tumor differentiation.

15. Oh SY, Kim YB, Suh KW. Prognostic significance of systemic inflammatory response in stage II colorectal cancer. *J Surg Res.* (2017) 208:158–65. doi: 10.1016/ j.jss.2016.08.100

16. Tan D, Fu Y, Tong W, Li F. Prognostic significance of lymphocyte to monocyte ratio in colorectal cancer: A meta-analysis. *Int J Surg.* (2018) 55:128–38. doi: 10.1016/j.ijsu.2018.05.030

17. Hamid HKS, Davis GN, Trejo-Avila M, Igwe PO, Garcia-Marín A. Prognostic and predictive value of neutrophil-to-lymphocyte ratio after curative rectal cancer resection: A systematic review and meta-analysis. *Surg Oncol.* (2021) 37:101556. doi: 10.1016/j.suronc.2021.101556

18. Sanger CB. Research perspective on prognostic significance of lymphocyte-tomonocyte and platelet-to-lymphocyte ratio in rectal cancer: A systematic review, metaanalysis, and meta-regression. *Dis Colon Rectum*. (2022) 65:188. doi: 10.1097/ DCR.0000000000002350

19. Portale G, Bartolotta P, Azzolina D, Gregori D, Fiscon V. Prognostic role of platelet-to-lymphocyte ratio, neutrophil-to-lymphocyte, and lymphocyte-to-monocyte ratio in operated rectal cancer patients: systematic review and meta-analysis. *Langenbecks Arch Surg.* (2023) 408:85. doi: 10.1007/s00423-023-02786-8

20. Page MJ, McKenzie JE, Bossuyt PM, Boutron I, Hoffmann TC, Mulrow CD, et al. The PRISMA 2020 statement:an updated guideline for reporting systematic reviews. *BMJ*. (2021) 372:n71. doi: 10.1136/bmj.n71

21. Wells G, Shea B, O'connell D, Peterson J, Welch V, Losos M, et al. *The Newcastle-Ottawa Scale (NOS) for assessing the quality of nonrandomised studies in meta-analyses* (2014). Ottawa (ON: Ottawa Health Research Institute (Accessed December 11, 2013).

22. Kim SR, Kim K, Lee SA, Kwon SO, Lee JK, Keum N, et al. Effect of red, processed, and white meat consumption on the risk of gastric cancer: an overall and dose? Response meta-analysis. *Nutrients*. (2019) 11:826. doi: 10.3390/nu11040826

23. Basile D, Garattini SK, Corvaja C, Montico M, Cortiula F, Pelizzari G, et al. The MIMIC study: prognostic role and cutoff definition of monocyte-to-lymphocyte ratio and lactate dehydrogenase levels in metastatic colorectal cancer. *Oncologist.* (2020) 25:661–8. doi: 10.1634/theoncologist.2019-0780

24. Li Q, Chen L, Jin H, Zhao Y, Hao Z, Ma X. Pretreatment inflammatory markers predict outcomes and prognosis in colorectal cancer patients with synchronous liver metastasis. *Clin Med Insights Oncol.* (2022) 16:1-11. doi: 10.1177/11795549221084851

25. Facciorusso A, Del Prete V, Crucinio N, Serviddio G, Vendemiale G, Muscatiello N. Lymphocyte-to-monocyte ratio predicts survival after radiofrequency ablation for colorectal liver metastases. *World J Gastroenterol.* (2016) 22:4211–8. doi: 10.3748/wjg.v22.i16.4211

26. Kuramochi H, Yamada T, Yoshida Y, Matsuda A, Kamiyama H, Kosugi C, et al. The pre-treatment lymphocyte-to-monocyte ratio predicts efficacy in metastatic colorectal cancer treated with TAS-102 and bevacizumab. *Anticancer Res.* (2021) 41:3131–7. doi: 10.21873/anticanres.15098

27. Lin GN, Liu PP, Liu DY, Peng JW, Xiao JJ, Xia ZJ. Prognostic significance of the pre-chemotherapy lymphocyte-to-monocyte ratio in patients with previously untreated metastatic colorectal cancer receiving FOLFOX chemotherapy. *Chin J Cancer*. (2016) 35:5. doi: 10.1186/s40880-015-0063-1

28. Lisanti C, Basile D, Parnofiello A, Bertoli E, Andreotti VJ, Garattini SK, et al. The SENECA study: Prognostic role of serum biomarkers in older patients with metastatic colorectal cancer. *J Geriatr Oncol.* (2020) 11:1268–73. doi: 10.1016/j.jgo.2020.06.010

29. Neal CP, Cairns V, Jones MJ, Masood MM, Nana GR, Mann CD, et al. Prognostic performance of inflammation-based prognostic indices in patients with resectable colorectal liver metastases. *Med Oncol.* (2015) 32:144. doi: 10.1007/s12032-015-0590-2

30. Neofytou K, Smyth EC, Giakoustidis A, Khan AZ, Williams R, Cunningham D, et al. The preoperative lymphocyte-to-monocyte ratio is prognostic of clinical outcomes for patients with liver-only colorectal metastases in the neoadjuvant setting. *Ann Surg Oncol.* (2015) 22:4353–62. doi: 10.1245/s10434-015-4481-8

31. Ouyang H, Xiao B, Huang Y, Wang Z. Baseline and early changes in the neutrophil-lymphocyte ratio (NLR) predict survival outcomes in advanced colorectal cancer patients treated with immunotherapy. *Int Immunopharmacol.* (2023) 123:110703. doi: 10.1016/j.intimp.2023.110703

32. Ozawa T, Ishihara S, Kawai K, Kazama S, Yamaguchi H, Sunami E, et al. Impact of a lymphocyte to monocyte ratio in stage IV colorectal cancer. *J Surg Res.* (2015) 199:386–92. doi: 10.1016/j.jss.2015.06.014

33. Peng J, Li H, Ou Q, Lin J, Wu X, Lu Z, et al. Preoperative lymphocyte-tomonocyte ratio represents a superior predictor compared with neutrophil-tolymphocyte and platelet-to-lymphocyte ratios for colorectal liver-only metastases survival. *Onco Targets Ther.* (2017) 10:3789–99. doi: 10.2147/ OTT.S140872

34. Song A, Eo W, Lee S. Comparison of selected inflammation-based prognostic markers in relapsed or refractory metastatic colorectal cancer patients. *World J Gastroenterol.* (2015) 21:12410–20. doi: 10.3748/wjg.v21.i43.12410

35. Wang YY, Liu ZZ, Xu D, Liu M, Wang K, Xing BC. Fibrinogen-albumin ratio index (FARI): A more promising inflammation-based prognostic marker for patients undergoing hepatectomy for colorectal liver metastases. *Ann Surg Oncol.* (2019) 26:3682–92. doi: 10.1245/s10434-019-07586-3

36. Zager Y, Hoffman A, Dreznik Y, Jacoby H, Cordoba M, Horesh N, et al. Cytoreductive surgery plus hyperthermic intraperitoneal chemotherapy in patients with peritoneal carcinomatosis from colorectal cancer: The prognostic impact of baseline neutrophil-lymphocyte, plateletel-lymphocyte and lymphocyte-monocyte ratios. *Surg Oncol.* (2020) 35:321–7. doi: 10.1016/j.suronc.2020.09.002

37. Motono N, Mizoguchi T, Ishikawa M, Iwai S, Iijima Y, Uramoto H. Prognostic impact of cancer inflammation prognostic index for non-small cell lung cancer. *Lung.* (2023) 201:603–10. doi: 10.1007/s00408-023-00649-z

38. Huang R, Zheng Y, Zou W, Liu C, Liu J, Yue J. Blood biomarkers predict survival outcomes in patients with hepatitis B virus-induced hepatocellular carcinoma treated with PD-1 inhibitors. *J Immunol Res.* (2022) 2022:3781109. doi: 10.1155/2022/3781109

39. Trinh H, Dzul SP, Hyder J, Jang H, Kim S, Flowers J, et al. Prognostic value of changes in neutrophil-to-lymphocyte ratio (NLR), platelet-to-lymphocyte ratio (PLR) and lymphocyte-to-monocyte ratio (LMR) for patients with cervical cancer undergoing definitive chemoradiotherapy (dCRT). *Clin Chim Acta*. (2020) 510:711–6. doi: 10.1016/j.cca.2020.09.008

40. Aoyama T, Kato A, Maezawa Y, Hashimoto I, Hara K, Komori K, et al. Lymphocyte to monocyte ratio is an independent prognostic factor in patients with esophageal cancer who receive curative treatment. *Anticancer Res.* (2024) 44:339–46. doi: 10.21873/anticanres.16817

41. Li ZM, Huang JJ, Xia Y, Sun J, Huang Y, Wang Y, et al. Blood lymphocyte-tomonocyte ratio identifies high-risk patients in diffuse large B-cell lymphoma treated with R-CHOP. *PloS One.* (2012) 7:e41658. doi: 10.1371/journal.pone.0041658

42. Porrata LF, Ristow K, Colgan JP, Habermann TM, Witzig TE, Inwards DJ, et al. Peripheral blood lymphocyte/monocyte ratio at diagnosis and survival in classical Hodgkin's lymphoma. *Haematologica*. (2012) 97:262–9. doi: 10.3324/haematol.2011.050138

43. Goto W, Kashiwagi S, Asano Y, Takada K, Takahashi K, Hatano T, et al. Predictive value of lymphocyte-to-monocyte ratio in the preoperative setting for progression of patients with breast cancer. *BMC Cancer.* (2018) 18:1137. doi: 10.1186/s12885-018-5051-9

44. Yang J, Liao D, Chen C, Liu Y, Chuang TH, Xiang R, et al. Tumor-associated macrophages regulate murine breast cancer stem cells through a novel paracrine EGFR/ Stat3/Sox-2 signaling pathway. *Stem Cells.* (2013) 31:248–58. doi: 10.1002/stem.1281

45. Grimm M, Feyen O, Hofmann H, Teriete P, Biegner T, Munz A, et al. Immunophenotyping of patients with oral squamous cell carcinoma in peripheral blood and associated tumor tissue. *Tumour Biol.* (2016) 37:3807–16. doi: 10.1007/s13277-015-4224-2

46. Chan JC, Chan DL, Diakos CI, Engel A, Pavlakis N, Gill A, et al. The lymphocyte-to-monocyte ratio is a superior predictor of overall survival in comparison to established biomarkers of resectable colorectal cancer. *Ann Surg.* (2017) 265:539-46. doi: 10.1097/SLA.00000000001743

47. Chen F, Zhuang X, Lin L, Yu P, Wang Y, Shi Y, et al. New horizons in tumor microenvironment biology: challenges and opportunities. *BMC Med.* (2015) 13:45. doi: 10.1186/s12916-015-0278-7

48. Koh YW, Kang HJ, Park C, Yoon DH, Kim S, Suh C, et al. The ratio of the absolute lymphocyte count to the absolute monocyte count is associated with prognosis in Hodgkin's lymphoma: correlation with tumor-associated macrophages. *Oncologist.* (2012) 17:871–80. doi: 10.1634/theoncologist.2012-0034

49. Wang J, Li J, Wei S, Xu J, Jiang X, Yang L. The ratio of platelets to lymphocytes predicts the prognosis of metastatic colorectal cancer: A review and meta-analysis. *Gastroenterol Res Pract.* (2021) 2021:9699499. doi: 10.1155/2021/9699499