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# Impact of obesity on all-cause and cause-specific mortality among critically ill men and women: a cohort study on the eICU database

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**Background:** The effect of obesity on intensive care unit outcomes among critically ill patients and whether there are sex differences have not been well investigated. We sought to determine the association between obesity and 30-day all-cause and cause-specific mortality among critically ill men and women.

**Methods:** Adult participants who had body mass index (BMI) measurements were included from the eICU database. Participants were divided into six groups according to BMI (kg/m<sup>2</sup>) categories (underweight, <18.5; normal weight, 18.5–24.9; overweight, 25–29.9; class I obesity, 30–34.9; class II obesity, 35–39.9; class III obesity, ≥40). A multivariable adjusted logistic model was conducted with odds ratios (ORs) and 95% confidence intervals (CIs). A cubic spline curve based on the generalized additive model was used to represent the nonlinear association. Stratified analysis and sensitivity analysis were also performed.

**Results:** A total of 160,940 individuals were included in the analysis. Compared with the class I obesity category, the underweight and normal weight categories had higher all-cause mortality, and the multivariable adjusted ORs were 1.62 (95% CI: 1.48-1.77) and 1.20 (95% CI: 1.13-1.27) for the general population, 1.76 (95% CI: 1.54-2.01) and 1.22 (95% CI: 1.13-1.32) for men, and 1.51 (95% CI: 1.33-1.71) and 1.16 (95% CI: 1.06-1.27) for women, respectively. Accordingly, multivariable adjusted ORs for the class III obesity category were 1.14 (95% CI: 1.05-1.24) for the general population, 1.18 (95% CI: 1.05-1.33) for men, and 1.10 (95% CI: 0.98-1.23) for women. With cubic spline curves, the association between BMI and all-cause mortality was U-shaped or reverse J-shaped. Similar findings were observed for cause-specific mortality, with the underweight category associated with a higher risk of mortality. Class III obesity increased the risk of cardiovascular death among men (OR 1.51; 95% CI: 1.23-1.84) and increased the risk of other-cause death among women (OR 1.33; 95% CI: 1.10-1.61).

**Conclusion:** The obesity paradox appears to be suitable for all-cause and cause-specific mortality among critically ill men and women. However, the protective effect of obesity cannot be extended to severely obese individuals. The association between BMI and cardiovascular mortality was sex-specific and was more pronounced among men than among women.

#### KEYWORDS

obesity, sex difference, all-cause mortality, cause-specific mortality, critically ill patients



GRAPHICAL ABSTRACT

## Introduction

The prevalence and disease burden of obesity is increasing worldwide, posing a substantial public health challenge and clinical concern. By 2025, the global prevalence of obesity is expected to reach 18% for men and 21% for women, and severe obesity will surpass 9% in women and 6% in men (1). Epidemiological studies have revealed that obesity is strongly correlated with a series of pathophysiological disturbances, including atherosclerosis, diabetes, hypertension, chronic obstructive pulmonary disease, renal insufficiency and cancer (2-4). In recent years, the evidence that obesity promotes valvular heart disease, cardiomyopathy and multimorbidity has been further consolidated (5-7). These problems are direct consequences of excessive fat mass or indirect consequences of obesity-related metabolic dysfunction. Due to the high metabolic activity of adipose tissue, abnormal and detrimental adipocyte secretion patterns promote chronic proinflammatory, prothrombotic and procoagulant states. Although obesity carries a range of disease risks, elevated BMI is paradoxically associated with better survival in various clinical settings, including heart failure, atrial fibrillation, nephropathy, sepsis, acute respiratory distress syndrome and critical illness (8-13). This so-called 'obesity paradox' phenomenon appears to be more pronounced among men according to several studies (14, 15). However, recent studies on coronavirus disease 2019 (COVID-19) have consistently shown that obesity is independently correlated with severe outcomes and mortality from COVID-19 infection (16, 17). Its pathophysiological mechanisms involve obesity-induced weakened immune response, hypercoagulation and metabolic disorder (18). Therefore, the existence of the obesity paradox in different populations remains controversial.

Critically ill patients admitted to the intensive care unit (ICU) have a variety of systemic diseases, which are more dangerous and have a higher risk of death. Obesity brings greater diagnostic challenges (CT or ultrasound image quality reduction), increased operation difficulty (such as tracheal intubation), and pharmacokinetic and pharmacodynamic changes, which may complicate acute diseases and weaken the effectiveness of evidence-based interventions. Therefore, it is imperative to understand the impact of obesity on the clinical prognosis of these patients. However, reliable data on the relationship between obesity and mortality in critical settings are scarce and discrepant, showing positive, zero, or negative correlations (19, 20). Some studies reported positive results but involved only all-cause mortality and no cause-specific mortality. Moreover, it is now believed that fat mass and distribution vary by sex, and whether there is a sex difference in the association between obesity and mortality is also a matter of concern that has not been well assessed. To address this evidence gap, we analysed data from a large contemporary multicentre ICU cohort to explore whether there is an obesity paradox in all-cause and cause-specific mortality among critically ill patients, and if the obesity paradox exists, the existence of a sex-related difference therein.

# Methods

## Study participants

Data were extracted from the publicly available eICU Collaborative Research Database (eICU-CRD). The eICU-CRD is a telemedicine system developed by Philips Healthcare in cooperation with the Laboratory for Computational Physiology (LCP) of the

Massachusetts Institute of Technology to optimize the management of critically ill patients (21). The LCP has previously successfully shared the Medical Information Mart for Intensive Care (MIMIC) database to support academic research and ICU quality improvement (22). The eICU-CRD is a complete and expanded dataset independent of MIMIC, which collects comprehensive clinical data of more than 200,000 ICU encounters from 208 U.S. hospitals. This high-quality data integration system contains a large amount of information on demographic profiles, vital signs, disease severity scores, laboratory parameters, fluid balance, medications, diagnostic codes, treatments, survival status, and hospital-level data, including regional location, teaching status, bed capacity, etc. All data were deidentified, and the requirement for informed consent from patients was waived. Data are free to access after completing the required training course and signing a usage agreement. This study was conducted in accordance with the Declaration of Helsinki. The study is exempt from institutional review board approval due to the retrospective design, lack of direct patient intervention, and the security schema, for which the re-identification risk was certified as meeting safe harbor standards by an independent privacy expert (Privacert, Cambridge, MA) (Health Insurance Portability and Accountability Act Certification no. 1031219-2). One author (Shan Li) obtained database access and was responsible for data extraction (certification number: 46622370). We included individuals admitted to the ICU from 2014 to 2015. The exclusion criteria were (1) age under 18 years, (2) no weight or height data available, and (3) BMI < 10 kg/m<sup>2</sup> or > 60 kg/m<sup>2</sup>. Finally, 160,940 individuals were included in the analysis.

## Exposure

The primary exposure of interest was BMI, calculated by the formula BMI (kg/m<sup>2</sup>) = weight/height<sup>2</sup>. For this calculation, the weight and height documented at ICU admission were used. According to the international classification criteria, individuals were divided into six categories: underweight, <18.5 kg/m<sup>2</sup>; normal weight, 18.5–24.9 kg/m<sup>2</sup>; overweight, 25–29.9 kg/m<sup>2</sup>; class I or mild obesity, 30–34.9 kg/m<sup>2</sup>; class II or moderate obesity, 35–39.9 kg/m<sup>2</sup>; and class III or severe obesity,  $\geq$ 40 kg/m<sup>2</sup>.

## Outcomes

The primary outcome was all-cause mortality within 30 days of ICU admission. The secondary outcomes were cardiovascular mortality, infectious mortality, and other-cause mortality. ICU death statistics were determined according to the International Classification of Diseases codes (9th revision). All-cause death was defined as death caused by any cause from the date of admission until the time of death. Cardiovascular death was defined as death from diseases with ICD-9 codes 390–459. Infectious disease death was defined as death from diseases with ICD-9 codes 320–326, 460–488, 566–567, 590, 595, 597, 614–616, 680–686, and 995. Noncardiovascular and noninfectious causes death were defined as other-cause death. Death from acute myocardial infarction was defined as death from disease with ICD-9 code 410, death from heart failure was defined as death from disease with ICD-9 code 428, death from sepsis was defined as death from a condition with ICD-9 code 995, death from ischaemic stroke was

defined as death from conditions with ICD-9 codes 430–432, and death from intracranial haemorrhage was defined as death from conditions with ICD-9 codes 433–434.

## Covariates

The following factors were considered for covariate selection: (1) individual-level factors, including age, sex, ethnicity, heart rate, mean blood pressure, disease severity score (Acute Physiology, Age and Chronic Health Evaluation [APACHE] score and Glasgow Coma Scale [GCS] score); (2) clinical risk factors, including primary disease at admission (cardiovascular disease, respiratory disease, digestive disease, genitourinary disease, neurological disease, endocrine disease, trauma, other infectious disease [nonrespiratory, nonurinary, nondigestive tract infections disease]) and prehospital comorbidities (coronary artery disease, stroke/transient ischaemic attack, diabetes mellitus, hypertension, congestive heart failure, peripheral artery disease, chronic obstructive pulmonary disease, renal dysfunction); (3) important treatments, including mechanical ventilation, dialysis and vasoactive drugs; and (4) hospital-level factors, including admission source, geographic location and discharge year.

## Statistical analysis

Statistical analyses were performed using R software (version 3.6.1)<sup>1</sup> and EmpowerStats (X&Y Solutions, Inc., Boston, MA).<sup>2</sup> Statistical significance was defined as a 2-sided *p* value <0.05. Continuous variables are presented as the means with standard deviations (SDs) or medians with interquartile ranges (IQRs) and analysed using unpaired t tests or Mann–Whitney U tests depending on their distribution. Categorical variables are presented as numbers with percentages and were compared using the chi-square or Fisher's exact test. The R package multiple imputation by chained equation ( $N_{imputation} = 5$ ) was used to account for missing data (13.3% for APACHE score, 2.5% for GCS score).

A multivariable logistic regression model was used to examine adjusted ORs for the association between BMI on a categorical scale and all-cause and cause-specific mortality, with the BMI category related to the lowest mortality as a reference. Three models were constructed: Model I unadjusted, Model II adjusted by age, sex and ethnicity, and Model III adjusted by all covariates without selection. The cubic spline curves based on the generalized additive model that adjusted for all covariates were used to visually display the nonlinear relationship between BMI on a continuous scale and all-cause and cause-specific mortality. Stratified analysis was conducted to examine the interaction between BMI and stratified covariates on all-cause mortality by including two or multiple interaction terms with adjustment for predefined covariates. Several sensitivity analyses were conducted to evaluate the robustness of the primary analysis. First, we excluded deaths that occurred within the first 48h of ICU entry to determine whether the association between BMI and mortality could be explained by a reverse

<sup>1</sup> http://www.r-project.org

<sup>2</sup> http://www.empowerstats.com

causality of severe disease. Second, we performed a complete case analysis using only the complete data of all covariates to test whether the missing data distorted the current findings. Third, we plotted Kaplan– Meier survival curves by taking the length of ICU stay time as an underlying time scale and censoring at discharge or death to assess whether different statistical methods might change the results.

# Results

## **Baseline characteristics**

A total of 160,940 individuals were included in this study (mean [SD] age, 63.2 [17.1] years; 87,226 men [54.2%] and 123,959 Caucasians [77.0%]), 14,568 (9.1%) all-cause deaths, 5,565 (3.5%) cardiovascular deaths, 4,308 (2.7%) infectious disease deaths and 4,695 (2.9%) other-cause deaths were recorded within 30 days of ICU admission. The average BMI was 28.7 (SD 7.6) kg/m<sup>2</sup>, and 58,518 (36.4%) individuals had class I to class III obesity. Figure 1 shows the distribution of BMI categories in the overall population and among men and women. Individuals with a higher BMI were younger and had a higher prevalence of diabetes mellitus, hypertension, chronic heart failure and renal dysfunction. Individuals with a lower BMI were older, had higher APACHE scores, and had a higher prevalence of respiratory disease and digestive disease. More dependence on mechanical ventilation was observed among individuals with class III obesity. Both underweight and class III obese patients had longer ICU stays. Underweight individuals accounted for 4.4% of the total population, resulting in 14.4% of all-cause deaths, which was approximately twice that of class I obese individuals. Baseline characteristics classified by BMI category are shown in Table 1.

## Obesity and all-cause mortality

With a multivariable logistic regression model, the association between BMI on a categorical scale and all-cause mortality was U-shaped in the general population and in men, with both a low and high BMI correlated with a greater risk of mortality. However, a reverse J-shaped association was noted in women, with only a low BMI increasing the risk of mortality. Underweight individuals had the highest mortality in the overall population (OR 1.62; 95% CI: 1.48-1.77), followed by those with normal weight (OR 1.20; 1.13-1.27), and the lowest mortality was observed among class I obese individuals. Underweight and normal weight men had corresponding odds ratios of 1.76 (95% CI: 1.54-2.01) and 1.22 (95% CI: 1.13-1.32) compared with class I obese men, respectively Similarly, underweight and normal weight among women were also independently associated with all-cause mortality, with odds ratios of 1.51 (95% CI: 1.33-1.71) and 1.16 (95% CI: 1.06-1.27), respectively, after adjustment for potential confounders. In the class III obese category, the multivariable adjusted odds ratios for all-cause mortality were 1.14 (95% CI: 1.05-1.24) for the overall population, 1.18 (95% CI: 1.05-1.33) for men and 1.10 (95% CI: 0.98–1.23) for women (Figure 2; Supplementary Table 1).

Based on the cubic spline curve, the association between BMI on a continuous scale and all-cause mortality was also U-shaped in the general population and in men, whereas it was reverse J-shaped in women. The risk inflection point correlated with the lowest all-cause mortality was  $28.3 \text{ kg/m}^2$  in the general population,  $28.2 \text{ kg/m}^2$  in men and  $28.3 \text{ kg/m}^2$  in women (Figure 3). Among the overall population, men and women with a BMI below the risk inflection point accounted for 54.6% (87906), 54.9% (47922) and 53.5% (39401), respectively. Before the corresponding risk inflection points, for every  $5 \text{ kg/m}^2$ decrease in BMI, the risk of all-cause mortality increased by 18% in the whole population, 21% in men and 15% in women. After the inflection points, the risk of death plateaued. For every  $5 \text{ kg/m}^2$ increase in BMI, the risk of all-cause mortality in the whole population, men and women increased by only 1%(Supplementary Table 2).

In stratified analysis, no clear evidence of a statistical interaction between BMI category and stratified variables on all-cause mortality was found. Compared with the class I obesity category, both the underweight category and normal weight category were independently associated with higher all-cause mortality, although the CI risk estimates were slightly wider in certain groups due to the relatively small number of individuals and events (Supplementary Table 3).

## Obesity and cause-specific mortality

A U-shaped association between BMI and cardiovascular mortality was observed in the general population and in men, while a reverse J-shaped association was noted in women. Low BMI was consistently associated with increased cardiovascular mortality in the overall population and both sexes. However, the association between class III obesity and cardiovascular mortality was more pronounced among men (OR 1.51; 95% CI: 1.23-1.84) than among women (OR 1.03; 95% CI: 0.85-1.24) (P for interaction 0.0046) (Supplementary Figure 1). Regarding infectious disease mortality, there was a consistent monotonic decreased risk with increasing BMI in the general population and both sexes. Low BMI was strongly associated with an increased risk of infectious disease mortality, while high BMI was not related to it. The association between BMI and other-cause mortality exhibited U-shaped in the general population and in women but reverse J-shaped in men. Contrary to cardiovascular death, the relationship between class III obesity and other-cause mortality was significant among women (OR 1.33; 95% CI: 1.10-1.61) but not significant among men (OR 1.16; 95% CI: 0.95-1.42). Moreover, class II obese women also had an increased risk of other-cause mortality (OR 1.25; 95% CI:



#### TABLE 1 Baseline characteristics of individuals by BMI categories.

BMI, kg/m²	<18.5	18.5–24.9	25.0–29.9	30.0-34.9	35.0-39.9	≥40	p value		
N (%)	7,067 (4.4)	47,768 (29.7)	47,587 (29.6)	29,818 (18.5)	15,042 (9.3)	13,658 (8.5)			
Age, years	64.2±19.1	64.1±19.0	64.3±16.7	62.8±15.5	61.1±14.9	58.6±14.2	<0.001		
Male, <i>n</i> (%)	4,056 (57.4)	22,112 (46.3)	19,170 (40.3)	12,874 (43.2)	7,440 (49.5)	8,062 (59.0)	<0.001		
Ethnicity									
Caucasian, n (%)	5,327 (75.4)	36,638 (76.7)	36,710 (77.1)	23,243 (77.9)	11,622 (77.3)	10,419 (76.3)			
African American, <i>n</i> (%)	952 (13.5)	4,962 (10.4)	4,724 (9.9)	3,207 (10.8)	1,876 (12.5)	2,058 (15.1)			
Hispanic, n (%)	206 (2.9)	1,846 (3.9)	2,001 (4.2)	1,065 (3.6)	492 (3.3)	393 (2.9)			
Asian, <i>n</i> (%)	212 (3.0)	1,219 (2.6)	802 (1.7)	299 (1.0)	101 (0.7)	53 (0.4)			
Other/unknown, <i>n</i> (%)	370 (5.2)	3,103 (6.5)	3,350 (7.0)	2,004 (6.7)	951 (6.3)	735 (5.4)			
BMI, kg/m²	$16.7 \pm 1.5$	$22.2 \pm 1.8$	$27.3 \pm 1.4$	$32.2 \pm 1.4$	$37.1 \pm 1.4$	$45.9\pm5.0$	<0.001		
Heart rate, bpm	106±32	101±33	98±33	98±32	99±32	$100\pm32$	<0.001		
Mean blood pressure, mmHg	81±41	83±41	86±42	88±43	$89 \pm 44$	89±45	<0.001		
Severity score									
APACHE score	52 (34–71)	49 (31-67)	47(30-64)	45(30-64)	45(29-64)	46(29–64)	<0.001		
GCS	14 (11–15)	15 (12–15)	15 (12–15)	15 (12–15)	15 (12–15)	15 (12–15)	<0.001		
Primary reason of ICU admission									
Cardiovascular disease, $n$ (%)	1,744 (24.7)	15,694 (32.9)	19,504 (41.0)	12,877 (43.2)	6,173 (41.0)	4,850 (35.5)	<0.001		
Respiratory disease, <i>n</i> (%)	2,053 (29.1)	8,772 (18.4)	7,125 (15.0)	4,524 (15.2)	2,581 (17.2)	2,942 (21.5)	<0.001		
Digestive disease, <i>n</i> (%)	834 (11.8)	5,553 (11.6)	5,096 (10.7)	2,841 (9.5)	1,435 (9.5)	1,157 (8.5)	<0.001		
Genitourinary disease, <i>n</i> (%)	365 (5.2)	2,438 (5.1)	2,239 (4.7)	1,547 (5.2)	868 (5.8)	965 (7.1)	<0.001		
Neurological disease, n (%)	474 (6.7)	3,181 (6.7)	2,964 (6.2)	1,784 (6.0)	878 (5.8)	764 (5.6)	<0.001		
Endocrine disease, $n$ (%)	363 (5.1)	2,266 (4.7)	1,519 (3.2)	768 (2.6)	365 (2.4)	341 (2.5)	<0.001		
Trauma, <i>n</i> (%)	221 (3.1)	2,323 (4.9)	1,984 (4.2)	1,085 (3.6)	379 (2.5)	277 (2.0)	<0.001		
Other infectious disease, $n$ (%)	296 (4.2)	1,793 (3.8)	1,667 (3.5)	1,155 (3.9)	662 (4.4)	778 (5.7)	<0.001		
Other disease, <i>n</i> (%)	717 (11.1)	5,748 (12.0)	5,489 (11.5)	3,237 (10.8)	1,701 (11.3)	1,584 (11.6)	<0.001		
Pre-admission comorbidities									
Coronary artery disease, <i>n</i> (%)	966 (13.7)	8,272 (17.3)	9,524 (20.0)	6,078 (20.4)	3,009 (20.0)	2,316 (17.0)	<0.001		
Stroke/TIA, n (%)	687 (9.7)	4,709 (9.9)	4,710 (9.9)	2,704 (9.1)	1,235 (8.2)	1,031 (7.5)	<0.001		
Diabetes mellitus, <i>n</i> (%)	572 (8.1)	4,674 (9.8)	5,291 (11.1)	4,221 (14.2)	2,711 (18.0)	3,007 (22.0)	<0.001		
Hypertension, <i>n</i> (%)	2,522 (35.7)	19,372 (40.6)	21,825 (45.9)	14,641 (49.1)	7,685 (51.1)	7,117 (52.1)	<0.001		
Congestive heart failure, <i>n</i> (%)	724 (10.2)	5,769 (12.1)	6,058 (12.7)	4,267 (14.3)	2,468 (16.4)	2,806 (20.5)	<0.001		

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(Continued)

#### TABLE 1 (Continued)

BMI, kg/m²	<18.5	18.5–24.9	25.0–29.9	30.0-34.9	35.0-39.9	≥40	<i>p</i> value
Peripheral arterial disease, $n$ (%)	306 (4.3)	2087 (4.4)	2,150 (4.5)	1,266 (4.2)	632 (4.2)	526 (3.9)	0.024
Chronic obstructive pulmonary disease, $n$ (%)	1,527 (21.6)	6,474 (13.6)	5,533 (11.6)	3,646 (12.2)	2,132 (14.2)	2,399 (17.6)	<0.001
Renal dysfunction, <i>n</i> (%)	662 (9.4)	5,177 (10.8)	5,308 (11.2)	3,467 (11.6)	1866 (12.4)	1860 (13.6)	<0.001
Therapeutics							
Mechanical ventilation, <i>n</i> (%)	1,800 (25.5)	10,748 (22.5)	10,844 (22.8)	7,301 (24.5)	4,050 (26.9)	4,395 (32.2)	<0.001
Dialysis, n (%)	280 (4.0)	1927 (4.0)	1715 (3.6)	1,056 (3.5)	541 (3.6)	467 (3.4)	<0.001
Vasoactive drugs, n (%)	259 (3.7)	1,686 (3.5)	1760 (3.7)	1,042 (3.5)	526 (3.5)	481 (3.5)	0.624
Admission source							<0.001
Emergency department, <i>n</i> (%)	3,926 (55.6)	25,341 (53.1)	23,582 (49.6)	13,979 (46.9)	7,028 (46.7)	6,592 (48.3)	
Acute care/floor, <i>n</i> (%)	1,381 (19.5)	8,013 (16.8)	7,598 (16.0)	4,838 (16.2)	2,609 (17.3)	2,584 (18.9)	
Other, <i>n</i> (%)	1760 (24.9)	14,414 (30.2)	16,407 (34.5)	11,001 (36.9)	5,405 (35.9)	4,482 (32.8)	
Geographic location							<0.001
Midwest, <i>n</i> (%)	2,110 (29.9)	15,092 (31.6)	15,825 (33.3)	10,475 (35.1)	5,551 (36.9)	5,237 (38.3)	
South, <i>n</i> (%)	2,280 (32.3)	14,262 (29.9)	14,127 (29.7)	8,822 (29.6)	4,345 (28.9)	4,076 (29.8)	
West, <i>n</i> (%)	1,400 (19.8)	9,802 (20.5)	9,694 (20.4)	5,700 (19.1)	2,706 (18.0)	2,250 (16.5)	
Northeast, n (%)	451 (6.4)	3,122 (6.5)	3,156 (6.6)	2,130 (7.1)	1,132 (7.5)	1,025 (7.5)	
Other, <i>n</i> (%)	826 (11.7)	5,490 (11.5)	4,785 (10.1)	2,691 (9.0)	1,308 (8.7)	1,070 (7.8)	
Hospital discharge year							0.703
2014, n (%)	3,348 (47.4)	22,416 (46.9)	22,364 (47.0)	14,136 (47.4)	7,023 (46.7)	6,417 (47.0)	
2015, <i>n</i> (%)	3,719 (52.6)	25,352 (53.1)	25,223 (53.0)	15,682 (52.6)	8,019 (53.3)	7,241 (53.0)	
Length of stay, days	5.1 (2.8-8.9)	4.8 (2.5-8.3)	4.7 (2.5-8.1)	4.8 (2.6-8.2)	4.9 (2.6-8.7)	5.1 (2.8-9.2)	<0.001
All-cause death, <i>n</i> (%)	1,016 (14.4)	4,830 (10.1)	4,007 (8.4)	2,329 (7.8)	1,248 (8.3)	1,138 (8.3)	<0.001
Cardiovascular death, <i>n</i> (%)	289 (4.1)	1,641 (3.4)	1,654 (3.5)	1,008 (3.4)	523 (3.5)	450 (3.3)	0.040
Infectious-cause death, n (%)	385 (5.4)	1,560 (3.3)	1,112 (2.3)	627 (2.1)	311 (2.1)	313 (2.3)	<0.001
Other-cause death, <i>n</i> (%)	342 (4.8)	1,629 (3.4)	1,241 (2.6)	694 (2.3)	414 (2.8)	375 (2.7)	< 0.001

Values are mean (standard deviation), median (inter-quartile range) or number (percentage). APACHE, acute physiology, age and chronic health evaluation. GCS, glasgow coma score.

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1.03–1.52) (Figure 2; Supplementary Table 4). These findings were reconfirmed by a cubic spline model with BMI as a continuous variable (Figure 4).

# Obesity and specific disease-related mortality

Regarding fatal myocardial infarction and fatal ischaemic stroke, no significant association was found between BMI and disease-related death. A J-shaped association between BMI and fatal heart failurerelated death was observed, with a plateau at approximately a BMI of 25 to 30 kg/m<sup>2</sup>. When BMI exceeded this plateau, the risk of heart failure-related death increased significantly. There was a reverse J-shaped relationship for sepsis-related death, with a low BMI associated with high mortality, whereas a high BMI was not. A strongly monotonic decreased risk for intracranial haemorrhage-related death with increasing BMI was detected, and this significant negative correlation dominated the association between BMI and all stroke deaths, including ischaemic and haemorrhagic stroke (Figure 5).





#### FIGURE 3

Multivariable adjusted odds ratios for all-cause mortality according to BMI on a continuous scale among (A) Overall population, (B) Men and (C) Women. Odds ratios (solid line) and 95% confidence intervals (dashed lines) were from cubic spline curves based on the generalized additive model. Arrows indicate BMI associated with the lowest mortality.



(C) wom model.

## Sensitivity analysis

To evaluate the possible impact of reverse causality from severe illness, we examined the association between BMI and risk of mortality by excluding deaths that occurred within the first 48 h of ICU entry. The overall odds ratio was similar, only slightly attenuated (Supplementary Figure 2; Supplementary Table 5). The results from complete case analyses that included only individuals with complete data on all covariates were consistent with those of the main analysis, and the findings were greatly similar for men and women separately (Supplementary Figure 3; Supplementary Table 6). Finally, the result from Kaplan–Meier survival analysis considering



mortality as a time-to-event variable with length of ICU stay as the timescale was also consistent with that of the primary analysis (Supplementary Figure 4).

# Discussion

In a large multicentre ICU cohort, we found a striking U-shaped or reverse J-shaped association between BMI and all-cause and causespecific mortality among critically ill men and women, independent of obesity-related comorbidities and other potential confounding factors. Both underweight and normal weight individuals had a greater risk of death than their obese counterparts. These findings suggest that obesity exerts a protective effect on all-cause and causespecific mortality among men and women, consistent with the obesity paradox. However, this protective effect appears not to extend to individuals with severe obesity (class III obesity). The relationship between severe obesity and cardiovascular mortality diverged between men and women. The current results confirm that the obesity paradox remains apparent among critically ill patients, but it is not applicable to severely obese patients. There is a sex difference in the impact of severe obesity on cause-specific mortality. These findings provide more information for predicting disease prognosis and improving the quality of ICU management.

In the past 30 years, the prevalence of obesity and the burdens of obesity-related diseases have gradually increased globally. It is predicted that as the prevalence of obesity in the general population increases, the incidence of obesity among critically ill patients will also increase. A meta-analysis reported that approximately one-third of ICU patients were obese, and nearly 7% were morbidly obese (20). The high incidence of obesity in this study was consistent with previous results, with 36.3% of obese patients and 8.5% of severely obese

patients. The association pattern of obesity and adverse outcomes has been investigated in some relatively small studies in the critical care field, with the obesity paradox existing in short-term and long-term all-cause mortality (23, 24). Akinnusi et al. reported a U-shaped correlation between BMI and mortality, with worse survival among underweight (BMI <18.5 kg/m<sup>2</sup>) and morbidly obese (>40 kg/m<sup>2</sup>) patients (20). Oliveros et al. found a lower mortality among obese patients (BMI 30.0-39.9  $kg/m^2)$  but not among morbidly obese patients (BMI>40 kg/m<sup>2</sup>) when using normal weight patients as a reference (25). In alignment with previous studies, this analysis showed that obese individuals had a better 30-day survival rate, although they had a higher incidence of clinical comorbidities, including hypertension, diabetes mellitus, heart failure and renal insufficiency, and the differences in these comorbidity patterns may be the major confounding factors affecting the clinical prognosis. Underweight individuals had an approximately 1.7-fold increased risk, and normal weight individuals had a 1.2-fold increased risk of all-cause mortality compared to their class I obese counterparts, which was noted among the overall population, among men and women. Among class III obese individuals, increased all-cause mortality was observed among men but not among women, resulting in a U-shaped association among the overall population and among men and a reverse J-shaped association among women. These findings support the existence of the obesity paradox, but the survival benefit does not extend to class III obese individuals, especially men. Furthermore, extremely close BMI inflection points for all-cause mortality were generated by the cubic spline curves, with 28.3 kg/m<sup>2</sup> for the whole population,  $28.2 \text{ kg/m}^2$  for males and  $28.3 \text{ kg/m}^2$  for females. There are several potential explanations for the obesity paradox. First, adipocytes positively regulate worsening inflammatory processes by secreting immunomodulatory substances such as leptin and interleukin-10, thereby improving survival during severe illness (26). Second, high cholesterol and lipoprotein levels in obese individuals may provide the precursors for adrenal steroid hormone synthesis to combat lethal stress (27). Third, adipose tissue also affords important nutritional reserves for critically ill patients with highly catabolic status and negative energy balance (28). Fourth, underweight individuals are usually more vulnerable and have less positive responses to supportive therapy (29). Finally, disparities in medical care may also lead to survival differences. Due to subconsciously entrenched concerns about obesity, obese patients often receive earlier and more aggressive management and are assigned closer monitoring, higher care standards and a lower threshold for transfer to the ICU. Indeed, this analysis showed that obese patients had higher rates of mechanical ventilation usage, partly reflecting more aggressive interventions.

The association pattern between BMI and cardiovascular death was largely consistent with that of all-cause death, supporting the obesity paradox. Notably, an obviously increased cardiovascular mortality was found among class III obese men but not women. Sex hormones may play an important role in determining fat mass and distribution. Oestrogen increases fat deposition, while testosterone inhibits fat deposition, so men tend to have less fat mass than women (30). In addition, because oestrogen blocks the androgen effect by downregulating the androgen receptor, women tend to accumulate more subcutaneous fat but less visceral fat than men (31). Visceral fat appears to be the major pathogenic fat depot associated with cardiovascular and metabolic alterations. Its proinflammatory, prothrombotic and low-fibrinogen milieus have a negative impact on cardiovascular protection and metabolic regulation, while subcutaneous fat acts more as a metabolic reserve, helping other tissues defend against lipotoxicity (32). These mechanisms could partly explain our findings that extremely obese men still face an increased risk of cardiovascular death, while women may be exempt due to the heterogeneity in adipose distribution. A recent study of a large cohort of women with coronary artery disease treated with drug-eluting stents also showed that the adjusted risk estimates for cardiovascular mortality among severely obese women were not statistically significant (33), which was in line with our findings. In addition, a cardiovascular magnetic resonance study explained this issue from an imaging perspective, that is, there was a sex-specific difference in left ventricular remodelling among obese subjects (34). Men predominantly exhibited concentric hypertrophy, while women presented a combination of eccentric and concentric hypertrophy. Concentric hypertrophy is proven to be more closely associated with cardiovascular mortality than eccentric hypertrophy.

The obesity paradox among patients with pneumonia and sepsis has been observed, despite evidence supporting that obesity impairs the immune response and increases susceptibility to infection (35). In this study, there was a consistent reverse J-shaped association between BMI and infectious disease mortality across the whole population, men and women. Only underweight and normal weight individuals had an increased risk, and when BMI exceeded 25 kg/m<sup>2</sup>, the risk of infectious disease death no longer increased but tended to decrease. The potential link between obesity and lower infectious disease mortality may be related to adipocytes positively regulating worsening inflammatory processes, high lipid levels neutralizing circulating endotoxin, and adipocytes providing adrenal steroid synthesis precursors and energy storage (26-28).

Other-cause deaths in this study included trauma, cancer, and uncommon disease-related deaths. Due to the relatively small number of events, separate analysis of a single disease could not be performed. Obesity had a protective effect on risk-adjusted mortality among individuals who died of noncardiovascular and noninfectious causes. However, this protective effect did not extend to severely obese individuals. Severely obese women remained at significantly increased risk of death compared with their mildly obese counterparts.

Previous studies have shown that obesity has a contradictory protective effect on heart failure (14, 15). However, it has also been suggested that the obesity paradox disappears after adjusting for B-type natriuretic peptide levels (36). We found that obesity was positively correlated with 30-day mortality among patients with acute heart failure. One possible explanation is that high BMI in the acute phase may be due to fluid retention rather than fat accumulation, affecting short-term prognosis, while cardiac cachexia and tissue hypoperfusion may contribute to worse long-term prognosis. These findings indicate that obesity may have different impacts on short-term and long-term prognoses among patients with heart failure. A prior heart failure study also showed that high BMI had a protective effect on 1-year mortality but not on 30-day mortality (37). The monotonous negative correlation between BMI and intracranial haemorrhage-related death was in line with expectations. A prospective study among 1.3 million British women revealed a robust relationship between low BMI and haemorrhagic stroke-related death (38). The trend in sepsis-related death was consistent with that of infectious disease death, with sepsis-related deaths accounting for 79.3% (3,414 of 4,308) of infectious disease deaths.

The unequal presentation of the obesity paradox between sexes has been reported. Studies among patients with heart failure and cardiogenic shock showed that the obesity paradox occurred only among men and not among women (14, 39). Clark et al. found that both women and men with systolic heart failure were affected by the obesity paradox (15). Our study showed that the obesity paradox apparently existed among both men and women, which was identified by BMI as a categorical variable and a continuous variable. However, the impact of the obesity paradox did not extend to severely obese individuals, and there was a sex difference between extremely high BMI and cause-specific mortality. Severe obesity increased cardiovascular deaths among men and increased othercause deaths among women, leading to increased cardiovascular deaths and other-cause deaths among severely obese individuals in the overall population. According to these findings, we have several considerations. First, BMI may not be a perfect anthropometric indicator for characterizing obesity due to its inherent limitations in assessing body composition and fat distribution. Obese individuals may have increased lean mass or more favourable subcutaneous fat distribution than visceral fat distribution, and these clinical phenotypes may confuse the findings of the obesity paradox. However, there is no corresponding suspicion when BMI is considered a risk predictor for pathophysiological disorders. Therefore, the defects of evaluation indicators cannot completely deny the obesity paradox. The obesity paradox may indicate a lack of comprehension of the complex pathophysiological link between

obesity and clinical outcomes, requiring further study. Second, reports on the obesity paradox have brought a confusing message to clinicians and policy-makers, leading to misguided healthy lifestyle management. However, given that obesity is a significant contributor to various pathophysiological dysfunctions and causes a substantial multimorbidity burden, the debate of the obesity paradox should not reduce efforts to control obesity while awaiting further evidence. Moreover, our study also showed that severe obesity led to worse survival. Third, the current results are generally consistent with and further extend previous reports on the obesity paradox in various clinical milieus. Although this paradox exists among both men and women, it cannot be extended to severely obese individuals. The increased cardiovascular death among severely obese men drove the increased all-cause mortality risk, while the increased other-cause death among severely obese women led to an upward trend of all-cause mortality risk. However, infectious deaths did not appear to be involved. Therefore, in addition to focusing on the greater risk among underweight and normal weight patients, clinicians should pay special attention to the risk of cardiovascular death among severely obese men and other-cause death among severely obese women and manage potential complications and risk factors that may compromise survival. Finally, the association between BMI and disease-specific mortality underscores that the impact of obesity on mortality may be subdivided and cannot be simply summarized in terms of the obesity paradox. Developing more accurate and targeted predictors to provide precise and personalized assessments is needed for future research.

## Strengths and limitations

This study included more than 160,000 ICU patients from a contemporary multicentre database. It was heterogeneous in terms of disease composition, ICU type and admission source, yielding a certain extrapolation validity for the study results. The model was extensively adjusted for confounding factors and had significant statistical power. Moreover, we extended the existing view of the obesity paradox and posited that it cannot be extended to severely obese individuals and that there was a sex difference in the impact of obesity on cause-specific mortality. Several limitations need to be considered. First, given that the retrospective design is inherently limited, we could not prove a causal relationship between obesity and mortality. Second, we broadly adjusted for confounding factors in multivariate analysis, including not only disease type and clinical comorbidities but also mechanical ventilation, dialysis, and vasoactive drug usage. Obesity leads to increased use and duration of mechanical ventilation, requires more frequent dialysis to achieve sufficient clearance, and affects the titration of vasoactive drugs, which may have an impact on mortality. However, obesity may be a net result of complex interactions between genetic, behavioural and environmental factors. Residual confounding factors, including dietary habits, smoking history, alcohol consumption, physical activity, income and socioeconomic status, may be involved, which were not extracted from the database. Third, we did not have any information about abdominal obesity or adipose distribution, such as waist circumference and waist-hip ratio, which may have an additional impact on the outcomes. Fourth, the subset of severely obese individuals was relatively small in number, which may have limited the statistical power of this group. Fifth, this study has a large heterogeneity in ethnic composition, and the majority of the cohort is Caucasians, accounting for 77% of the total population, which limits the extrapolation of the current results to other ethnic populations. Finally, given the regional differences in the definition of obesity based on BMI, a large proportion of individuals from the United States and European countries may restrict the extrapolation of these findings.

# Conclusion

With the rapid development of the global economy and the general improvement of living standards, obesity is likely to become an increasingly prominent concern in the ICU. Our study provides new evidence on the obesity paradox, which is a well-known phenomenon in a variety of disease entities and is still evident among critically ill patients. Although the protective effect of obesity on all-cause and cause-specific mortality is largely consistent among men and women, this effect cannot be extended to severely obese individuals. Special attention needs to be paid to cardiovascular death risk among severely obese men and other-cause death risk among severely obese women.

# Data availability statement

Publicly available datasets were analyzed in this study. This data can be found at: https://www.physionet.org/content/eicu-crd-demo/2.0.1/.

## Ethics statement

The studies involving human participants were reviewed and approved by the institutional review board of the Massachusetts Institute of Technology. The ethics committee waived the requirement of written informed consent for participation.

# Author contributions

SL has fully obtained all the data and takes responsibility for the integrity of the data and the accuracy of the analysis. SL and HL contributed to the concept and study design. SL and WZ contributed to the acquisition, statistical analysis, and interpretation of the data. SL and ZF contributed to the drafting of the manuscript. All authors contributed to the article and approved the submitted version.

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# **Conflict of interest**

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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## References

1. NCD Risk Factor Collaboration (NCD-RisC). Trends in adult body-mass index in 200 countries from 1975 to 2014: a pooled analysis of 1698 population-based measurement studies with 19.2 million participants. *Lancet.* (2016) 387:1377–96. doi: 10.1016/S0140-6736(16)30054-X

2. Afshin A, Forouzanfar MH, Reitsma MB, Sur P, Estep K, Lee A, et al. Health effects of overweight and obesity in 195 countries over 25 years. *N Engl J Med.* (2017) 377:13–27. doi: 10.1056/NEJMoa1614362

3. Wormser D, Kaptoge S, Di Angelantonio E, Wood AM, Pennells L, Thompson A, et al. Separate and combined associations of body-mass index and abdominal adiposity with cardiovascular disease: collaborative analysis of 58 prospective studies. *Lancet*. (2011) 377:1085–95. doi: 10.1016/S0140-6736(11)60105-0

4. Lauby-Secretan B, Scoccianti C, Loomis D, Grosse Y, Bianchini F, Straif K. Body fatness and cancer-viewpoint of the IARC working group. *N Engl J Med.* (2016) 375:794–8. doi: 10.1056/NEJMsr1606602

5. Kaltoft M, Langsted A, Nordestgaard BG. Obesity as a causal risk factor for aortic valve stenosis. J Am Coll Cardiol. (2020) 75:163–76. doi: 10.1016/j.jacc.2019.10.050

 Robertson J, Lindgren M, Schaufelberger M, Adiels M, Björck L, Lundberg CE, et al. Body mass index in young women and risk of cardiomyopathy: a long-term follow-up study in Sweden. *Circulation*. (2020) 141:520–9. doi: 10.1161/CIRCULATIONAHA.119.044056

7. Kivimäki M, Strandberg T, Pentti J, Nyberg ST, Frank P, Jokela M, et al. Body-mass index and risk of obesity-related complex multimorbidity: an observational multicohort study. *Lancet Diabet Endocrinol.* (2022) 10:253–63. doi: 10.1016/S2213-8587(22)00033-X

8. Sharma A, Lavie CJ, Borer JS, Vallakati A, Goel S, Lopez-Jimenez F, et al. Metaanalysis of the relation of body mass index to all-cause and cardiovascular mortality and hospitalization in patients with chronic heart failure. *Am J Cardiol.* (2015) 115:1428–34. doi: 10.1016/j.amjcard.2015.02.024

9. Lavie CJ, Pandey A, Lau DH, Alpert MA, Sanders P. Obesity and atrial fibrillation prevalence, pathogenesis, and prognosis: effects of weight loss and exercise. *J Am Coll Cardiol.* (2017) 70:2022–35. doi: 10.1016/j.jacc.2017.09.002

10. Lu JL, Molnar MZ, Naseer A, Mikkelsen MK, Kalantar-Zadeh K, Kovesdy CP. Association of age and BMI with kidney function and mortality: a cohort study. *Lancet Diabet Endocrinol*. (2015) 3:704–14. doi: 10.1016/S2213-8587(15)00128-X

11. Pepper DJ, Sun J, Welsh J, Cui X, Suffredini AF, Eichacker PQ. Increased body mass index and adjusted mortality in ICU patients with sepsis or septic shock: a systematic review and meta-analysis. *Crit Care.* (2016) 20:181. doi: 10.1186/s13054-016-1360-z

12. Ni YN, Luo J, Yu H, Wang YW, Hu YH, Liu D, et al. Can body mass index predict clinical outcomes for patients with acute lung injury/acute respiratory distress syndrome? A meta-analysis. *Crit Care.* (2017) 21:36. doi: 10.1186/s13054-017-1615-3

13. Sakr Y, Alhussami I, Nanchal R, Wunderink RG, Pellis T, Wittebole X, et al. Being overweight is associated with greater survival in ICU patients: results from the intensive care over nations audit. *Crit Care Med.* (2015) 43:2623–32. doi: 10.1097/CCM.000000000001310

14. Hong S, Lee JH, Kim KM, Lee JW, Youn YJ, Ahn MS, et al. Is there a sex-related difference in the obesity paradox in systolic heart failure? Sex-related difference in the obesity paradox. *Yonsei Med J.* (2018) 59:57–62. doi: 10.3349/ymj.2018.59.1.57

15. Clark AL, Chyu J, Horwich TB. The obesity paradox in men versus women with systolic heart failure. *Am J Cardiol.* (2012) 110:77–82. doi: 10.1016/j.amjcard.2012.02.050

16. Tartof SY, Qian L, Hong V, Wei R, Nadjafi RF, Fischer H, et al. Obesity and mortality among patients diagnosed with COVID-19: results from an integrated health care organization. *Ann Intern Med.* (2020) 173:773–81. doi: 10.7326/M20-3742

17. Suleyman G, Fadel RA, Malette KM, Hammond C, Abdulla H, Entz A, et al. Clinical characteristics and morbidity associated with coronavirus disease 2019 in a series of patients in metropolitan Detroit. *JAMA Netw Open*. (2020) 3:e2012270. doi: 10.1001/jamanetworkopen.2020.12270

18. Steenblock C, Schwarz PEH, Ludwig B, Linkermann A, Zimmet P, Kulebyakin K, et al. COVID-19 and metabolic disease: mechanisms and clinical management. *Lancet Diab Endocrinol.* (2021) 9:786–98. doi: 10.1016/S2213-8587(21)00244-8

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## Supplementary material

The Supplementary material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fnut.2023.1143404/ full#supplementary-material

19. Hogue CW Jr, Stearns JD, Colantuoni E, Robinson KA, Stierer T, Mitter N, et al. The impact of obesity on outcomes after critical illness: a meta-analysis. *Intensive Care Med.* (2009) 35:1152–70. doi: 10.1007/s00134-009-1424-5

20. Akinnusi ME, Pineda LA, El Solh AA. Effect of obesity on intensive care morbidity and mortality: a meta-analysis. *Crit Care Med.* (2008) 36:151–8. doi: 10.1097/01. CCM.0000297885.60037.6E

21. Pollard TJ, Johnson AEW, Raffa JD, Celi LA, Mark RG, Badawi O. The eICU collaborative research database, a freely available multi-center database for critical care research. *Sci Data*. (2018) 5:180178. doi: 10.1038/sdata.2018.178

22. Johnson AE, Pollard TJ, Shen L, Lehman LWH, Feng M, Ghassemi M, et al. MIMIC-III, a freely accessible critical care database. *Sci Data*. (2016) 3:160035. doi: 10.1038/sdata.2016.35

23. Hutagalung R, Marques J, Kobylka K, Zeidan M, Kabisch B, Brunkhorst F, et al. The obesity paradox in surgical intensive care unit patients. *Intensive Care Med.* (2011) 37:1793–9. doi: 10.1007/s00134-011-2321-2

24. Zhou D, Wang C, Lin Q, Li T. The obesity paradox for survivors of critically ill patients. Crit Care. (2022) 26:198. doi: 10.1186/s13054-022-04074-1

25. Oliveros H, Villamor E. Obesity and mortality in critically ill adults: a systematic review and meta-analysis. *Obesity*. (2008) 16:515-21. doi: 10.1038/oby.2007.102

26. Tilg H, Moschen AR. Adipocytokines: mediators linking adipose tissue, inflammation and immunity. *Nat Rev Immunol.* (2006) 6:772-83. doi: 10.1038/ nri1937

27. Murch O, Collin M, Hinds CJ, Thiemermann C. Lipoproteins in inflammation and sepsis. I. Basic science. *Intensive Care Med.* (2007) 33:13–24. doi: 10.1007/s00134-006-0432-y

28. Marques MB, Langouche L. Endocrine, metabolic, and morphologic alterations of adipose tissue during critical illness. *Crit Care Med.* (2013) 41:317–25. doi: 10.1097/CCM.0b013e318265f21c

29. Reeves GK, Balkwill A, Cairns BJ, Green J, Beral V. Hospital admissions in relation to body mass index in UK women: a prospective cohort study. *BMC Med.* (2014) 12:45. doi: 10.1186/1741-7015-12-45

30. Mayes JS, Watson GH. Direct effects of sex steroid hormones on adipose tissues and obesity. *Obes Rev.* (2004) 5:197–216. doi: 10.1111/j.1467-789X.2004.00152.x

31. Wajchenberg BL. Subcutaneous and visceral adipose tissue: their relation to the metabolic syndrome. *Endocr Rev.* (2000) 21:697–738. doi: 10.1210/edrv.21.6.0415

32. Lee SJ, Shin SW. Mechanisms, pathophysiology, and Management of Obesity. N Engl J Med. (2017) 376:1491–2. doi: 10.1056/NEJMc1701944

33. Faggioni M, Baber U, Afshar AE, Giustino G, Sartori S, Sorrentino S, et al. Effects of body mass index on clinical outcomes in female patients undergoing percutaneous coronary intervention with drug-eluting stents: results from a patient-level pooled analysis of randomized controlled trials. *JACC Cardiovasc Interv.* (2018) 11:68–76. doi: 10.1016/j.jcin.2017.06.060

34. Rider OJ, Lewandowski A, Nethononda R, Petersen SE, Francis JM, Pitcher A, et al. Gender-specific differences in left ventricular remodelling in obesity: insights from cardiovascular magnetic resonance imaging. *Eur Heart J.* (2013) 34:292–9. doi: 10.1093/ eurhearti/ehs341

35. Nie W, Zhang Y, Jee SH, Jung KJ, Li B, Xiu Q. Obesity survival paradox in pneumonia: a meta-analysis. BMC Med. (2014) 12:61. doi: 10.1186/1741-7015-12-61

36. Pozzo J, Fournier P, Lairez O, Vervueren PL, Delmas C, Elbaz M, et al. Obesity paradox: origin and best way to assess severity in patients with systolic HF. *Obesity*. (2015) 23:2002–8. doi: 10.1002/oby.21216

37. Shah R, Gayat E, Januzzi JL Jr, Sato N, Cohen-Solal A, DiSomma S, et al. Body mass index and mortality in acutely decompensated heart failure across the world: a global obesity paradox. *J Am Coll Cardiol.* (2014) 63:778–85. doi: 10.1016/j. jacc.2013.09.072

38. Kroll ME, Green J, Beral V, Sudlow CLM, Brown A, Kirichek O, et al. Adiposity and ischemic and hemorrhagic stroke: prospective study in women and meta-analysis. *Neurology.* (2016) 87:1473–81. doi: 10.1212/WNL.000000000003171

39. Kwon W, Lee SH, Yang JH, Choi KH, Park TK, Lee JM, et al. Impact of the obesity paradox between sexes on in-hospital mortality in cardiogenic shock: a retrospective cohort study. *J Am Heart Assoc.* (2022) 11:e024143. doi: 10.1161/JAHA.121.024143