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RECEIVED 07 September 2023

ACCEPTED 31 October 2023

PUBLISHED 14 November 2023

CITATION

Zhao J and Du Y (2023), The impact and spatial effect of rural revitalization on agricultural carbon dioxide emissions: a case study of Henan Province. *Front. Environ. Sci.* 11:1290126. doi: 10.3389/fenvs.2023.1290126

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The impact and spatial effect of rural revitalization on agricultural carbon dioxide emissions: a case study of Henan Province

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Clarifying the influencing mechanism of rural revitalization on agricultural carbon emissions is crucial for attaining carbon peaking and carbon neutrality. This study utilized spatial econometric model, mediating effect model and dual fixed effect model to explore the influence and spatial impact of rural revitalization on carbon emissions from agricultural land by using the panel data of prefectural cities in Henan Province. Results indicate that rural revitalization exerts a notable beneficial influence on carbon emissions, as its improvement results in a rise in such emissions. Furthermore, rural revitalization demonstrates a favorable spatial spillover effect on agricultural carbon emissions in neighboring cities. Agricultural GDP and mechanical technological progress act as intermediate factors, as rural revitalization promotes carbon emissions from agriculture by fostering economic development and technological advancements. Heterogeneity analysis indicates that the correlation between rural revitalization and greenhouse gas emissions from agriculture is nonlinear, as moderate and low levels of rural revitalization promote agricultural carbon emissions, while higher levels exhibit a negative effect. Thus, rural revitalization exhibits an inflection point effect on agricultural carbon emissions.

KEYWORDS

agricultural carbon emission, rural revitalization, spatial durbin model, heterogeneity analysis, Henan Province

1 Introduction

Excessive emissions of carbon dioxide, covering carbon dioxide (CO₂), methane (CH₄), and nitrous oxide (N₂O) emissions, have had detrimental effects such as global temperature increase, climate crisis, and elevated sea levels. As a result, governments worldwide have placed significant emphasis on addressing this issue (Douglas, 1991). In September 2020, at the UNGA (United Nations General Assembly), China declared its commitment to peaking CO₂ discharges before 2030 and pursuing carbon neutrality by 2060. However, achieving these goals requires in-depth research on carbon footprint across different industries. Agriculture, in particular, represents a significant origin of carbon emissions, making up around 14% of global greenhouse gas emissions (Nduagu and Gates, 2015). Global agricultural greenhouse gas emissions have nearly doubled from 2,752 Mt CO₂ equivalent in 1961 to 5,294 Mt CO₂ equivalent in 2016 (Mondal et al., 2018). Over the period from 1980 to 2020, China's The overall greenhouse gas emissions from agriculture have risen from 665 million metric tons of CO₂ equivalent to 970 million tons, exhibiting a fluctuating growth trend and an overall increase of

nearly 46% (Guttikunda et al., 2014). While China's agricultural sector contributes approximately 17% to the world's whole agricultural carbon emissions (Liu and Xiao, 2020), it is crucial to note that the proportion of farming emissions in China is relatively lower compared to industrial emissions. Nonetheless, the total emissions remain substantial. Furthermore, given that China's level of agricultural development is still lower in comparison to developed countries, there is significant potential for reducing carbon emissions agricultural field. Therefore, it is crucial to recognize the key factors contributing to carbon footprint in agriculture and propose effective measures to reduce releases in agricultural production. These efforts will contribute to the fulfillment of the "dual carbon" objective and the high-quality development of China's modernization.

Early research related to agricultural carbon emissions primarily concentrated on measuring emissions and analyzing their temporal and spatial characteristics (Tristram and West, 2002; Tian et al., 2014; Xiong et al., 2016a; Xiong et al., 2016b; Huang et al., 2019; Chen et al., 2020; Liu et al., 2023). As research progressed, scholars started to explore the factors influencing agricultural carbon emissions from various perspectives. External factors include carbon taxes, economics, renewable energy sources, agricultural technology, urbanization, and policy (Liu et al., 2017; Ismael et al., 2018; Ridzuan et al., 2020; Dumortier and Elobeid, 2021; Liu et al., 2021; Alam et al., 2023; Li et al., 2023; Wojewodzki et al., 2023; Xia et al., 2023). Internal factors mainly revolve around land use (Liu et al., 2023). For instance, Pugh (Pugh et al., 2015) used the LP-GUESS model to find that people often underestimate the greenhouse gas emissions during land cover change and the potential carbon absorption from future reforestation. Guo (Guo et al., 2021) concluded that increasing the planting area of certain crops, such as wheat, soybeans, vegetables and sorghum, can reduce carbon emissions built on statistics from three provinces in Northeast China. Hu (Hu et al., 2016) observed that adopting intercropping using conservation agriculture principles for maize and wheat in China's Hexi Corridor can reduce water usage in arid areas and reduce carbon emissions. While existing research has extensively discussed agricultural greenhouse gas emissions from multiple facets, limited research has examined the effect of rural revitalization, a major innovation in modern rural development theory and practice (Liu et al., 2020), on the carbon footprint of agriculture. As rural revitalization progresses, agricultural production will be optimized, agricultural carbon emissions have the potential to be reduced. Rural revitalization plays a crucial role in promoting the adoption of sustainable agricultural production methods. Traditional farming practices often contribute significantly to greenhouse gas emissions, but rural revitalization policies encourage farmers to embrace environmentally friendly approaches such as organic farming and precision agriculture. These sustainable practices have the potential to make a substantial impact in reducing greenhouse gas emissions. Most existing literature primarily concentrates on the connotation and realization path of rural revitalization strategy (Gao et al., 2023), with only a few exploring

agricultural and emission reduction strategies from the perspective of rural revitalization (Zhou et al., 2022). Few studies have utilized econometric models to quantify the influence and spatial consequences of rural revitalization on agricultural greenhouse gas emissions. Given the variations in economic development and natural endowments across different regions, the process of rural revitalization and development may exhibit spatial dependence and spillover effects. As a significant farming province, Henan has 7.51 million hectares arable land, ranking third in China, with low urbanization rate less than 60%. Therefore, this study takes the province of Henan as an example, and employs panel data from 18 cities at the prefecture level within Henan province from 2001 to 2020. The study uses Spatial Moran's global index to examine the correlation and agglomeration characteristics of agricultural carbon emissions. Additionally, a spatial panel data econometric model is used to empirically examine the influence and spatial impact of rural revitalization on agricultural greenhouse gas emissions. The outcomes of this study aim to offer insights for agricultural carbon reduction policies in Henan province.

2 Theoretical analysis and research hypothesis

Rural revitalization is tightly linked to agricultural carbon emissions. The 19th Communist Party of China National Congress introduced the rural revitalization strategy, which encompasses the general policy of "flourishing industry, livable ecology, Rural style civilization, efficient governance, and thriving livelihood." The strategy similarly outlines specific demands for rural revitalization across five dimensions.

"Flourishing industry" is a driving force behind rural revitalization (Wang et al., 2021), aiming to promote economic growth in rural regions. Nevertheless, the development related to the agricultural sector heavily relies on machinery, leading to a long-term equilibrium connection between the growth in the farming industrial economic system and mechanization (Juhász, 2018). This heavy mechanization results in increased usage of diesel and electricity, thus contributing to higher carbon dioxide emissions in agriculture.

"Ecological livability" is a fundamental aspect of rural revitalization (Yang et al., 2023). The use of pesticides and fertilizers in rural areas exhibits a direct influence on the natural environment. Excessive usage relating to these chemicals can deteriorate the ecological balance and diminish air quality for local residents. Moreover, it can contribute to increased carbon dioxide emissions from agriculture. Furthermore, the expansion of rural postal routes brings convenience to rural residents, but it also promotes the use of mechanical transportation, leading to a rise in carbon dioxide emissions.

"Rural style civilization" serves as the essence of the rural revitalization strategy (Ru, 2023). Rural residents' investments

in education, culture, and entertainment reflect their emphasis on cultural and educational values. The improvement of education levels is helpful for the widespread adoption of advanced agricultural cultivation methods and the enhancement of productivity. Ultimately, this has the ability to help curb agricultural carbon emissions.

“Efficient governance” serves as the organizational backbone for rural revitalization (Leck and Simon, 2013). An increased budget in the urban and rural affairs sector often focuses on rural infrastructure development, such as the construction of farmland and canals. These endeavors invariably require additional manpower, resources, and machinery, thereby contributing to higher agricultural carbon emissions.

“Thriving livelihood” represents the ultimate objective of the rural revitalization tactic (Zhuo et al., 2021). It could be measured through the disposable income of rural residents. If disposable income increases, it ensures a higher standard of material wellbeing for rural residents. Consequently, people will have more time and energy to engage in various cultural activities. This, in turn, facilitates the widespread adoption of advanced green and low-carbon agricultural technology as well as professional agricultural knowledge. Ultimately, this helps to curb agricultural carbon emissions.

A comprehensive analysis reveals that the effect of rural revitalization concerning agricultural carbon emissions can be uncertain, leading to both positive and negative outcomes. Therefore, we propose Hypothesis 1.

H1. The effect of rural revitalization on farming greenhouse gases is uncertain.

Countryside revitalization will facilitate the enhancement of agricultural infrastructure and production factors, with agricultural production technology being a crucial resource that spreads rapidly through inter-regional exchanges. Technological upgrades and spillover effects will expedite the decrease of farming carbon emissions in nearby areas. Furthermore, the development of countryside revitalization may enhance the competitive advantage of relevant enterprises, serving as benchmarks for other companies to learn from. This will further generate an “imitation result” and “demonstration result,” reinforcing the spatial effect of countryside revitalization regarding farming carbon emissions. Therefore, we propose Hypothesis 2.

H2. Rural revitalization exhibits a notable spatial spillover impact on carbon emissions in agriculture.

Rural revitalization aims to achieve a prosperous life (Stokes and Seto, 2019). During the course of rural revitalization and progress, promoting the expansion of the agricultural economy is crucial. Zang et al. conducted research on Xinjiang Province in China using relevant data from 2002 to 2020 to examine the association between agricultural economy and agricultural greenhouse gas emissions. The consequences revealed that the farming economy had a notable influence on the intensity of agricultural greenhouse gas emissions (Zang et al., 2022). Zhang et al. also discovered a reciprocal cause-

and-effect relationship between agricultural carbon emissions and agricultural economic growth in both the short and long term (Zhang et al., 2019).

Furthermore, the level of mechanization plays a crucial role in agricultural production technology. The advancement of rural revitalization is closely linked to agricultural mechanization, which impacts farming scale, agricultural labor productivity, and industrial structure (Wang et al., 2022). Consequently, it also affects agricultural carbon emissions. Drawing from these findings, we propose Hypothesis 3.

H3. Agricultural GDP and the level of mechanical technology progress serve as intermediate factors in the correlation between rural revitalization and the release of carbon emissions from agricultural activities.

3 Data and methods

3.1 Description of variables

3.1.1 Dependent variables

The dependent factor in this study refers to carbon emissions from agriculture, which encompass CH₄ emissions from rice farming and N₂O emissions generated by fertilizers as well as soils. Additionally, it includes CO₂ emissions produced by fertilizers, insecticides, agricultural plastics, ploughing, agricultural equipment, and agricultural water management (Xing and Yan, 2000; Hu et al., 2010; Min and Hu, 2012). The calculation method for agricultural carbon emissions is provided in Formula (1) as proposed by Tian et al. (Tian et al., 2012).

Regarding N₂O emissions from different crops, this study synthesizes the research findings of Wang (Wang, 1997), Yu et al. (Yu et al., 1995), Su et al. (Su et al., 1992), Huang et al. (Huang et al., 1995), and Qiu et al. (Qiu et al., 2010). The N₂O emission parameters for corn, rice, winter wheat, soybean, vegetables, in addition to other dry crops were set as 2.532, 0.24, 1.75, 2.29, 4.944, and 0.95 (kg·hm⁻²) respectively. The emission coefficients for other emission sources are presented in Table 1. N₂O and CH₄ emissions were converted using the results of the sixth IPCC survey: 1 metric ton of CH₄ is equal to 27.2 tons of CO₂, and 1 metric ton of N₂O is equal to 273 metric tons of CO₂. (CO₂ emission coefficient = carbon emission coefficient * 44/12).

$$E = \sum E_i = \sum T_i * \& \quad (1)$$

Where T_i is in reference to the actual cultivated area of rice, the real utilization of nitrogen fertilizer, the actual use of chemical fertilizer, the actual use of pesticides, the actual utilization of farming film, the real usage of diesel fuel, the actual irrigation zone, and the actual cultivated area of crops. $\&$ indicates coefficients of agricultural GHG emissions.

3.1.2 Key explanatory variable

The Key explanatory variable in the investigation refers to the holistic measure of rural rejuvenation. It is grounded in the new era of rural revival strategy, which focuses on the aim of rural

TABLE 1 Emission coefficients of agricultural GHG emissions.

Emission source	Emission coefficient	Data source
Plowing	CO ₂ : 1,146.2 kg hm ⁻²	China Agricultural University
Fertilizer	CO ₂ : 3.2840 kg kg ⁻¹	ORNL, United States
Pesticides	CO ₂ : 18.0917 kg kg ⁻¹	ORNL, United States
Agriplastic	CO ₂ : 18.9933 kg kg ⁻¹	Nanjing Agricultural University
Farm diesel	CO ₂ : 3.1863 kg kg ⁻¹	IPCC
Field irrigation	CO ₂ : 91.667 kg hm ⁻²	Dubey (Dubey and Lal, 2009)
Nitrogen fertilizer application	N ₂ O: 0.0125 kg kg ⁻¹	IPCC
Rice farming	CH ₄ : 236.7 kg hm ⁻²	Guidelines for compiling provincial greenhouse gas inventories

TABLE 2 Indicator system of rural revitalization.

Dimension	Meaning	Weight	Description	Unit
Prosperous industry	Economic level of the primary industry	0.122	Primary output value/rural population	100 million yuan/10,000 people
	Rural industrialization level	0.142	Rural electricity consumption/rural population	Millions of kilowatt-hours per million people
	Rural mechanization level	0.028	Total power/sown area related to farm equipment	Ten thousand kW hours per hectare
Livable ecology	Chemical input strength	0.021	quantity of fertilizer used in farming production/ planted area	Ton per hectare
	Convenience of living infrastructure	0.097	Rural mail line	km
	Rural population employment	0.104	Rural employment	Ten thousand of people
Civilized village style	Cultural level of consumption among rural residents	0.006	Education and entertainment expenditure/ <i>per capita</i> total expenditure consumption	%
	Health and wellness consumption degree of rural residents	0.017	Health expenditure/total expenditure average individual consumption	%
	Strength level of rural teachers	0.013	Number of countryside faculty/agricultural population	%
Effective governance	Budget amount for urban and rural affairs of each municipality	0.092	Spending on urban as well as rural affairs	100 million yuan
	Villagers' autonomy level	0.044	Number of village committees/Number of resident rural population	Per 10,000 people
	Rural governance level	0.026	Rural minimum guaranteed number/rural population	%
Prosperous life	Abundance of farmers	0.221	Average disposable income per agricultural worker	Yuan
	Food expenditure as a proportion of total <i>per capita</i> expenditure level	0.033	Food expenditure Per individual/Total expenditure	%
	The level of urban-rural income gap	0.034	Per person in cities disposable income/rural disposable income	%

revival: flourishing industry, livable ecology, rural style civilization, efficient governance, and thriving livelihood. The indicator system in favor of rural revitalization was constructed and is presented in Table 2. To avoid subjectivity and uncertainty, the entropy method was employed to assess weights for the relevant indexes. Firstly, these indicators were standardized, and then the information entropy of each indicator was calculated. Finally, the weights for each indicator could be determined according to the information entropy. The detailed steps can be found in references (Zhao et al., 2018).

3.1.3 Control factors

The selected Controlled factors in this investigation include urbanization, industrial structure, scale of cultivated land, investment intensity of farmers, and investment in the primary industry. The concept of urbanization is multidimensional and encompasses economic, social, land, population, and ecological aspects (Chen et al., 2018). As scholars have deepened their understanding of urbanization, it is widely believed that a single indicator is insufficient to capture its complexity. Therefore, this study adopts a multidimensional urbanization index as a control

variable, which includes the commonly used dimensions of economic urbanization, Population growth in urban areas, and Urban land development.

Urban economic growth is defined as the ratio between non-agricultural economy to the combined economy of agriculture, industry, and services. Population urbanization represents the proportion of residents to the total urban population. Land urbanization is assessed by the proportion of urban built-up area to the administrative area. The structure of the industrial sector is captured by the proportion of the GDP of the agriculture sector to the total industrial GDP. The scale of cultivated land is represented through the comparison of the actual sown cultivated area to the quantity of employees in the primary industry. The investment intensity of rural households is indicated by the ratio of fixed investment by farmers to the population working in the agricultural sector. Investment in fixed assets in the agricultural sector is measured by the volume of fixed investment in that sector.

3.2 Data sources

The information utilized in this research includes agricultural carbon emissions data, indicators of rural revitalization at all levels, and controlled variables. The aforementioned information was obtained from the Henan Provincial Statistical Compilation and Rural Statistical Compilation spanning from 2001 to 2020. In instances where there were a few absent data points, they were supplemented using trend analysis and linear fitting methods.

3.3 Methods

3.3.1 Spatial correlation test

The Global Moran's coefficient model was utilized to examine the geographic autocorrelation of total agricultural carbon footprint. The equation utilized is as follows:

$$I = \frac{n \sum_{i=1}^n \sum_{j=1}^n W_{ij} (x_i - \bar{x})(x_j - \bar{x})}{\sum_{i=1}^n \sum_{j=1}^n W_{ij} \sum_{i=1}^n (x - \bar{x})^2} \quad (2)$$

where I stands for the Global Moran's coefficient I , and n denotes the quantity of measured cities.

The variables x_i and x_j represent the carbon emissions from agricultural activities in prefecture-level cities, while \bar{x} denotes the average carbon emissions from agricultural activities. The term W_{ij} denotes the spatial weight matrix. The range of possible values for Global Moran's I is between -1 and 1.0 indicates that agricultural carbon emissions are spatially unrelated and randomly distributed. When the value is negative, it suggests negative spatial correlation, indicating dispersion. Conversely, when the value is positive, it suggests positive spatial correlation, indicating agglomeration.

3.3.2 Spatial Durbin model

Empirical analysis was conducted to examine the influence of rural revitalization regarding carbon emissions from agriculture in the province of Henan, with the objective of testing the hypothesis. The spatial metrology model utilized is described as:

$$Y_{it} = c\mathcal{E} + cP \sum_{j=1, j \neq i}^n w_{ij} y_{it} + c\&x + c\beta x_{it} + cy \sum_{j=1, j \neq i}^n w_{ij} (x + x_{it}) + \sigma_i + \omega_{it} \quad (3)$$

where $c\mathcal{E}$ represents a constant value; cP refers to the spatial autocorrelation coefficient; w_{ij} refers to the nested weight matrix of economic geography; $c\&$ and $c\beta$ refer to regression coefficients; cy refers to the lag term coefficient; σ_i refers to regional fixed effects; ω_{it} refers to the random error term.

3.3.3 Intermediary effect model

This study was inspired by the findings of Deng et al. (Deng and Zhang, 2021). It employed a panel data intermediate effect model to analyze the intermediary effect of rural revitalization regarding carbon emissions in agriculture for the purpose of examining the hypothesis.

The initial part of the intermediate effect model, represented by Eq. 6, focused on the main effect of rural revitalization. The second and third stage models, represented by Eqs 4, 5 respectively, were used to examine the mediating effects.

In this study, a parallel intermediate effect analysis was conducted to assess the mediating effects of variables.

$$M_{it} = a + a'x + \sum SG_{it} + \sigma_i + \varepsilon_{it} \quad (4)$$

Where M_{it} denotes the intermediate variable; a relates to the intercept value; a' refers to the coefficient of rural revitalization.

$$Y_{it} = b + b'x + cM_{it} + \sum SG_{it} + \sigma_i + \varepsilon_{it} \quad (5)$$

Where b refers to the intercept term; b' refers to the coefficient of rural revitalization; c refers to the coefficient of the intermediate variable.

3.3.4 Dual fixed effect model

To identify the impacts of various levels of rural revitalization on the greenhouse gas emissions from agriculture in Henan Province, a grouping regression approach was employed. The double fixed regression model used is as follows:

$$Y_{it} = c + \delta x + \sum SG_{it} + \sigma_i + v_t + \varepsilon_{it} \quad (6)$$

where i represents to place; t represents to time; Y_{it} represents to agricultural carbon emissions; c represents to the intercept term; δ represents to the rural revitalization coefficient; x refers to rural revitalization; S refers to the coefficient of control variable; G_{it} refers to the control variable; σ_i refers to regional fixed effects; v_t refers to time-fixed effect. ε_{it} refers to the random error term.

4 Results and discussion

4.1 Spatio-temporal characteristics

From 2001 to 2020, carbon footprint of agriculture in Henan Province exhibited an overall pattern of initially increasing afterward decreasing, as depicted in Figure 1. These values appear to be higher compared to the findings of Wei et al. (Wei et al., 2023). This disparity can be attributed to two aspects: firstly,

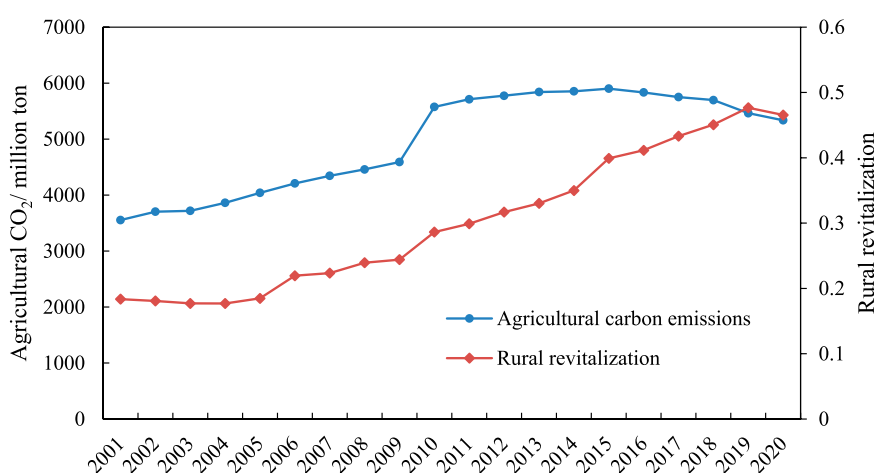


FIGURE 1
Change trend of agricultural CO₂ and rural revitalization from 2001 to 2020.

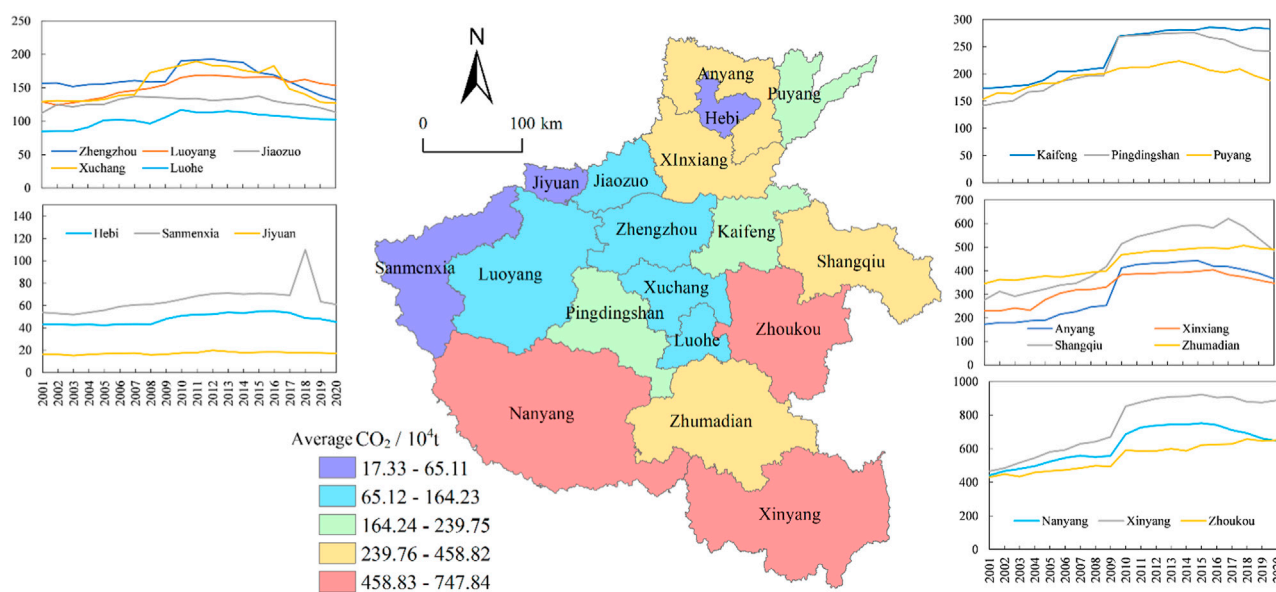


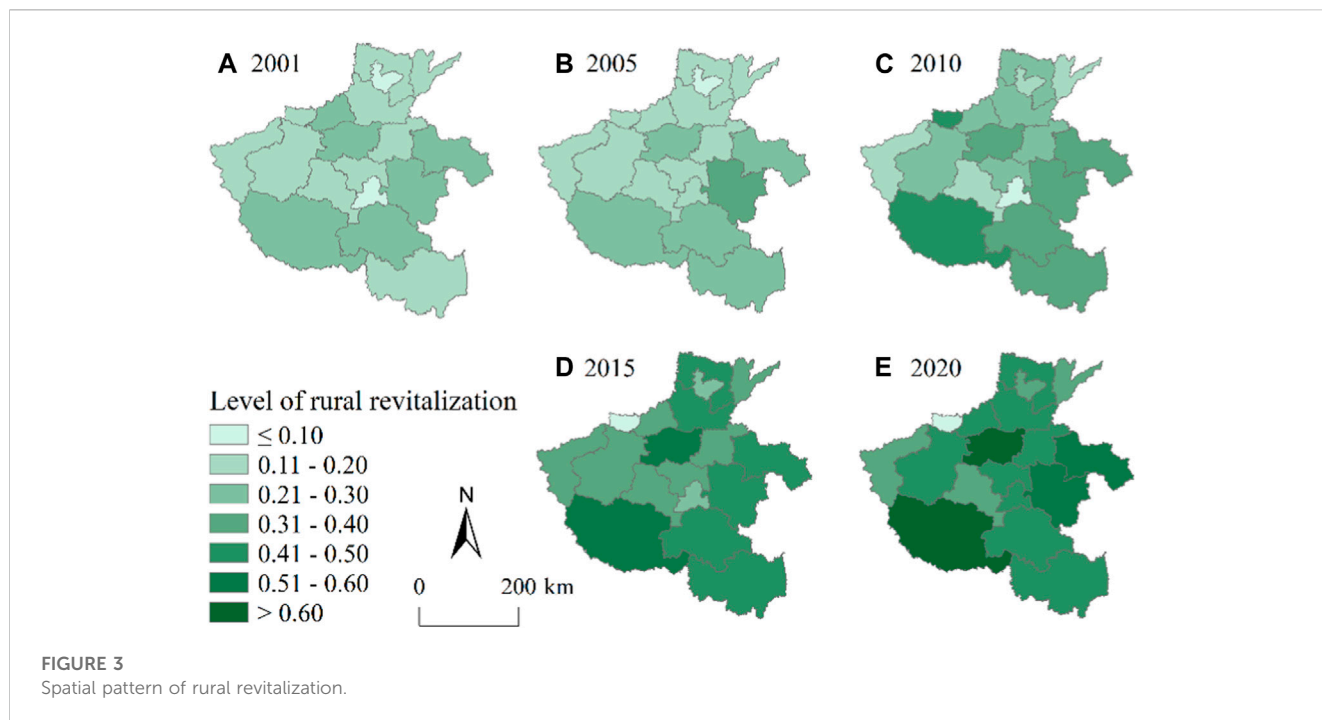
FIGURE 2
Spatial and temporal trend of agricultural carbon emissions.

the emission factors were converted to carbon dioxide emission factors; secondly, the greenhouse gases (N₂O and CH₄) were also calculated and converted to an equivalent of carbon dioxide. Specifically, there was an upward trend in emissions from 2001 to 2015, followed by a decline in total carbon emissions between 2015 and 2020. This change in trend can be attributed to the implementation of a series of green development strategies by China, particularly the introduction of green and low-carbon agricultural policies in 2015. As a result, the effective control of pesticides and fertilizers (Lal, 2004) led to a suppression of carbon footprint of agriculture.

As shown in Figure 1, the rural revitalization curve in Henan Province exhibited an upward trend. From 2001 to 2004, the curve

remained relatively flat. However, from 2004 to 2019, it showed a consistent and rapid growth trend, characterized by stability. In the period of 2019–2020, there was a slight decrease in the rural revitalization curve. This pattern can be attributed to the inclusion of the “Three Agricultural Questions” since 2004. As a result, local cities diligently implemented national policies, which acted as a driving force for the continuous growth of rural revitalization.

Based on Figure 2, Xinyang City, Nanyang City, and Zhoukou City ranked as the top three cities in Henan Province with high carbon footprint of agriculture. Jiyuan City, Hebi City, and Sanmenxia City ranked as the three cities with the lowest agricultural carbon emissions. This disparity can be attributed to



the fact that Xinyang City, Nanyang City, and Zhoukou City have larger areas of cultivated land and a greater emphasis on agricultural development. Consequently, they tend to consume more fertilizers and pesticides, resulting in higher agricultural carbon emissions. Conversely, Jiyuan City, Hebi City, and Sanmenxia City may have a different agricultural composition, leading to lower carbon emissions in this sector. Furthermore, by referring to other figures, we observe that some cities exhibited a relatively stable trend, such as Hebi City, Sanmenxia City, and Jiyuan City. Conversely, cities like Zhengzhou City, Luoyang City, and Jiaozuo City exhibited a tendency of initially growing followed by declining. Finally, cities like Nanyang City, Xinyang City, and Zhoukou City experienced a significant initial increase followed by a subsequent decrease in agricultural carbon emissions.

The Equal Interval method was employed to classify the level of rural revitalization, as illustrated in Figure 3. The rural revitalization level in Henan Province was relatively low in 2001. However, over time, the level of rural revitalization improved. Specifically, by 2005, the eastern part of Henan Province showed an elevated level of rural revitalization. By 2010, the southwestern region of Henan Province exhibited an elevated level of rural revitalization. During the year 2015, the southern and central areas of Henan Province reached a notable extent of rural revitalization. Finally, in the year 2020, the southwestern, eastern, and central cities of Henan Province demonstrated a notable extent of rural revitalization. This shift at the level of rural revitalization can be attributed to the varying levels of importance placed on rural development by local governments. Additionally, different regions possess distinct natural endowments, contributing to the disparities in rural revitalization and development.

According to Figure 4, the worldwide Moran's I index for total greenhouse gas emissions from agriculture in the region of Henan Province showed a consistently positive trend from 2001 to 2020. The index showed a variable growth pattern, starting at 0.264 in

2001 and reaching 0.275 in 2020. This finding indicates that the spatial dispersion of agricultural carbon footprint located in Henan Province was not random. Instead, it displayed significant spatial agglomeration characteristics. Furthermore, the degree of spatial agglomeration has progressively increased, suggesting a content of carbon footprint of agriculture in specific areas.

In accordance with Rangel et al. (Rangel et al., 2010), the Moran diagram is commonly accustomed to explore the spatial association pattern of a specific variable among neighboring cities in a sample region. In this study, agricultural carbon footprint data from cities in Henan Province were selected for the years 2001–2020. A local Moran scatter plot was then created to analyze the spatial agglomeration variations in carbon emissions from agriculture within this group of cities. Then, according to the Moran scatter plot of agricultural carbon emissions, cities were categorized into four quadrants: High-High level, Low-High level, Low-Low level, and High-Low level, displayed in Table 3.

Table 3 indicates that, in general, there was a dominant presence of high and low concentration in 2001, accounting for 94.44% of the total. In 2020, the dominance shifted to high agglomeration, low agglomeration, and low agglomeration, accounting for 88.89% of the total. This observation suggests spatial polarization characteristics in urban carbon footprint of farming within the province of Henan.

In the year of 2001, the cities in the first quadrant (H-H) including Zhumadian, Xinyang, Zhoukou, Shangqiu, and Nanyang exhibited heightened carbon footprint from agriculture. Furthermore, he nearby cities had enhanced agricultural greenhouse gas emissions. This indicates that these urban areas contribute to the agricultural carbon footprint of adjacent cities. Luohe, Xuchang, and Pingdingshan were situated in second quadrant (L-H), suggesting that these urban areas had lower agricultural carbon emissions themselves, but the surrounding cities had higher emissions. Additionally, the surrounding cities had spillover effects on the agricultural carbon emissions of Luohe, Xuchang, and

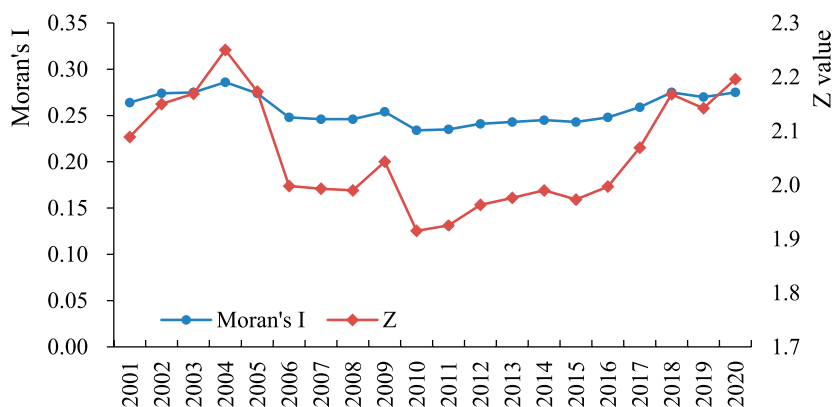


FIGURE 4
Global Moran' I of agricultural carbon emissions 2001–2020.

TABLE 3 Spatial agglomeration of agricultural carbon emissions in cities of Henan Province.

Year	H-H	L-H	L-L	H-L
2001	Zhumadian, Xinyang, Zhoukou, Shangqiu, Nanyang	Luohe, Xuchang, Pingdingshan	Sanmenxia, Hebi, Jiyuan, Jiaozuo, Luoyang, Zhengzhou, Puyang, Kaifeng, Anyang	Xinxiang
2020	Zhumadian, Xinyang, Zhoukou, Shangqiu, Nanyang	Luohe, Xuchang, Pingdingshan, Zhengzhou, Puyang, Hebi	Jiyuan, Jiaozuo, Luoyang, Sanmenxia, Kaifeng	Anyang, Xinxiang

Pingdingshan. The urban areas within the third quadrant (L-L), including Sanmenxia, Hebi, Jiyuan, Jiaozuo, Luoyang, Zhengzhou, Puyang, Kaifeng, and Anyang, had lower agricultural carbon emissions both within the cities themselves and in the surrounding cities, with no significant spillover effects. Xinxiang was located in the fourth quadrant (H-L), indicating that Xinxiang had high Carbon footprint of agriculture. Conversely, surrounding cities had lower emissions. Xinxiang Played a crucial role in driving agricultural greenhouse gas emissions in the surrounding cities. Compared to 2001, the high-concentration pattern remained unchanged in 2020. However, in 2020, Zhengzhou, Puyang, and Hebi shifted from the low-low clustering pattern to the low-high agglomeration pattern, indicating a rise in greenhouse gas emissions in the cities encircling them. Anyang shifted from a low-low agglomeration form to a high-low agglomeration form. This indicates an increase in Anyang's carbon footprint of agriculture, efficiently promoting farming-related carbon emissions locally.

4.2 Overall regression analysis

The analysis of carbon footprint of agriculture revealed significant spatial dependency, leading to the construction of a spatial econometric approach to examine the consequence and spatial effects regarding rural revitalization on carbon footprints in agriculture. Several tests, including the Hausmann test, LM test, LR test, Wald test, were conducted to evaluate the methodology. The conclusions indicated that the regional fixed effect model was the most suitable for analysis, and it was not feasible to simplify the spatial Dubin model into a spatial lag model or spatial error model.

Therefore, a region-fixed Spatial Durbin model was adopted. The outcomes are shown in Table 4.

Regarding the main explanatory variable, the coefficient of rural revitalization was found to be significantly favorable at the 1% level. This suggests that countryside revitalization has a significant contribution to carbon emissions. This finding is consistent with the results of Zhou et al. (Zhou et al., 2022), which revealed a rise in emissions *per capita* in rural areas during the rural revitalization process in China. This may be attributed to the promotion of agricultural mechanization and increased mechanization levels through rural revitalization. Mechanization shows a favorable influence on agricultural carbon footprint (Guo et al., 2022). Furthermore, rural revitalization may lead to increased fertilizer usage, thereby boosting agricultural carbon footprint.

In relation to control variables, urbanization was found to possess a significant effect on reducing carbon footprint from agriculture. This finding contradicts the conclusion of Magazzino et al. (Magazzino et al., 2023), which may be due to differences in the research scope and regional heterogeneity, leading to variations in the results. Economic urban expansion, through the advancement of non-agricultural industrial sectors, has facilitated the advancement of green and low-carbon technologies, thus minimizing carbon footprint from agriculture. Population urbanization has led to rural labor migration to urban areas, resulting in more efficient management of rural land and a reduction in agricultural carbon emissions. Land urbanization, characterized by urban expansion and transforming farmland into urban construction land, has also made a contribution to the overall decrease in carbon footprint of agriculture by reducing the total agricultural land area.

TABLE 4 Results of Durbin model.

Variables	Main	Wx
Rural revitalization	682.401*** (8.11)	222.135 (1.32)
Economic urbanization	-5.399*** (-4.77)	6.814*** (3.03)
Urbanization of population	-2.827* (-1.87)	8.990*** (3.60)
Land urbanization	-4.855*** (-6.54)	-8.169*** (-3.42)
Industrial structure	-0.011 (-0.25)	-0.401** (-2.10)
Cultivated land scale	2.807* (1.74)	-0.869 (-0.34)
Intensity of fixed asset investment of rural households	-0.011*** (-2.68)	0.022*** (4.61)
Investment in fixed assets in primary industry	0.073* (1.93)	-1.152*** (-5.41)
rho	0.248** (2.02)	
Observations	360	
R-squared	0.6603	

(*** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$).

The scale of cultivated land was found to have a substantial positive effect on carbon footprint of agriculture in Henan Province at a significance level of 5%, which aligns with the finding of Wang et al. (Wang et al., 2022). Larger-scale farming operations tend to utilize more agricultural equipment, which positively influences carbon footprint of agriculture. The intensity of fixed asset investment of rural households was substantially negative at 1% level. This result is inconsistent with the finding of Lin (Lin and Xu, 2018), possibly attributed to variations in research scope, resulting in disparities in the outcomes. This suggests that rural households' fixed asset investment, which often prioritize the use of green production machines, effectively curbs carbon emissions from agriculture. Additionally, fixed investment in the primary industry had a significant positive effect on agricultural carbon footprint at a 1% level. This can be attributed to the increase in fixed investment in the primary industry, which promotes agricultural mechanization, leading to higher agricultural electricity consumption and diesel usage, consequently boosting agricultural carbon emissions.

4.3 Comparative analysis about spatial effects

The direct, indirect, total effects of the core explanatory variable (rural revitalization) and control variables on agricultural carbon footprint are presented within Table 5. Regarding the core explanatory variable, an increase of one unit in the degree of rural revitalization was found to lead to a significant increase of 690.730 units in agricultural carbon footprint within the locality. Additionally, agricultural carbon footprint in nearby regions also increased by 550.450 units, passing the significance analysis at 1% level. The evidence suggests that rural revitalization has a significant impact on spatial spillover effects and hypothesis 2 is confirmed. These findings indicate that rural revitalization has both a direct effect on agricultural carbon footprint in the locality and an indirect effect through spillover effects on neighboring areas. The increase in

agricultural carbon footprint can be attributed to the expanded use of fertilizers and agricultural machine in rural revitalization efforts. Fertilizers contribute to the release of N_2O emissions into the air, while mechanized operation leads to a greater utilization of diesel and gasoline in agriculture, resulting in increased agricultural greenhouse gas emissions. Furthermore, when the extent of rural revitalization improves in a particular region, it serves as a demonstration effect for neighboring areas, prompting them to also increase their level of rural revitalization and subsequently boosting carbon emissions from agriculture. In summary, rural revitalization has a notable favorable influence on both the region itself and its neighboring areas in terms of agricultural carbon footprint.

The control variables in this research have shown interesting effects on agricultural carbon footprint.

4.3.1 Economic urbanization

Increase of one unit in economic urbanization led to a reduction of 5.299 units in local agricultural carbon footprint. However, it also resulted in an increase of 7.375 units in agricultural carbon footprint in adjacent areas. It suggests that when the non-agricultural economy in a region flourishes, the local agricultural sector may weaken, leading to a shift in demand for agricultural products to neighboring areas. This, in turn, promotes agricultural development in the neighboring areas and increases their carbon emissions.

4.3.2 Population urbanization

An increase of one unit in population urbanization resulted in a reduction of 2.685 units in local agricultural carbon emissions. However, it also led to an increase of 11.192 units in agricultural carbon emissions in neighboring areas. This indicates that urbanization can help curb local agricultural carbon footprint, but it holds a promoting influence on agricultural carbon footprint in neighboring areas. The higher level of urbanization in a region attracts labor from neighboring areas, where the rural labor force may still rely on outdated and extensive farming methods, contributing to increased carbon emissions in agriculture.

TABLE 5 Decomposition of the influence of explanatory variables on agricultural carbon footprint.

Variables	Direct effect	Indirect effect	Total effect
Rural revitalization	690.730*** (8.68)	550.450*** (2.61)	1,241.180*** (6.00)
Economic urbanization	-5.229*** (-4.55)	7.375*** (2.65)	2.146 (0.84)
Urbanization of population	-2.685* (-1.85)	11.192*** (3.41)	8.507*** (2.63)
Land urbanization	-5.018*** (-6.63)	-12.766*** (-3.33)	-17.784*** (-4.13)
Industrial structure	-0.020 (-0.47)	-0.559** (-2.10)	-0.579** (-2.06)
Cultivated land scale	2.924* (1.84)	0.201 (0.06)	3.125 (0.87)
Intensity of fixed asset investment of rural households	-0.011** (-2.57)	0.024*** (4.87)	0.013** (2.33)
Investment in fixed assets in primary industry	0.045 (1.19)	-1.546*** (-4.68)	-1.501*** (-4.39)
rho	0.248** (2.02)		
Observations	360		
R-squared	0.6603		

(*** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$).

4.3.3 Land urbanization

The total consequence, direct consequence, and indirect consequence of land urbanization were all significantly negative. This may be influenced by the competition and cooperation among neighboring governments. An increase in land urbanization in the local region also leads to a growth in neighboring areas, thereby further inhibiting carbon emissions in the neighborhood. This exhibits a notable detrimental overflow effect.

4.3.4 Industrial structure

The total consequence, direct consequence, and indirect consequence of industrial structure were all unfavorable, but indirect consequence was significantly higher rather than direct consequence. The development of the agricultural economy stimulates technological progress, which flows into neighboring areas and reduces carbon emissions in agriculture. This suggests that advancements in farming technology can hold a favorable effect on reducing agricultural carbon footprint.

4.3.5 Cultivated land scale

The total consequence, direct consequence, and indirect consequence of cultivated land scale were all positive, but only the direct consequence was statistically significant. This implies that enlarging the magnitude cultivated land in the local vicinity hold a positive effect on reducing agricultural carbon footprint, but the indirect consequence and overall consequence were not significant.

4.3.6 Rural household fixed assets investment intensity

The direct effect of rural household fixed asset investment intensity was significantly negative, indicating that investment behavior in advanced farming concepts among local farmers can help reduce local agricultural carbon footprint. However, the indirect consequence was positive significantly, suggesting that the spread of farming ideas across administrative lines is limited, and neighboring areas are less influenced by changes in farming practices in the studied region.

4.3.7 Fixed asset investment in the first industry

The direct effect of fixed asset investment in the first industry was positive, while the indirect effect was significantly negative. This suggests that increasing fixed asset investment in the first industry promotes energy use in agriculture, potentially contributing to higher carbon emissions. However, neighbors may complement each other in terms of learning and adopting green technologies, leading to a net reduction in agricultural carbon emissions.

Overall, these findings highlight the complex interplay between different factors and their effects on agricultural carbon emissions. It emphasizes the importance of considering various socioeconomic and environmental factors when formulating strategies to mitigate carbon emissions in agriculture.

4.4 Analysis of intermediate effect

According to Wen et al. (Wen et al., 2004), they employed an intermediate effect model to investigate the process and mechanism of rural revitalization on agricultural carbon footprint. Farming GDP and mechanical technological development were chosen as intermediate variables. The findings are summarized in Tables 6, 7.

Table 6 shows the outcomes of the model, including the coefficient estimates and statistical significance of the variables. It shows the direct relationships between rural revitalization, agricultural GDP, mechanical technological progress, and agricultural carbon footprint. The coefficients indicate the magnitude and the path of the relationships.

Model (1) shown in Table 6, rural revitalization has a substantial favorable impact on mechanical technological progress with a parameter of 2,293.729, indicating that rural revitalization promotes advancements in mechanical technology. Model (2) reveals that rural revitalization hold a significant positive influence on agricultural carbon footprint with a coefficient of 495.314. This suggests that rural revitalization contributes significantly to agricultural carbon emissions. Additionally, mechanical technological progress is found to hold a significant

TABLE 6 Intermediate effect model (the level of mechanical technological progress as the intermediate variable).

Variables	Model (1)	Model (2)
	Mechanical technological progress	Agricultural carbon emission
Rural revitalization	2,293.729*** (8.34)	495.314*** (5.02)
Mechanical technological progress	--	0.107*** (5.87)
Control variable	Control	control
R-squared	0.9006	0.9643
Regional fixed effect	Control	control
Time-fixed effect	Control	control

(*** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$).

TABLE 7 Intermediate effect model (agricultural GDP as the intermediate variable).

Variables	Model (1)	Model (2)
	Agricultural GDP	Agricultural carbon emission
Rural revitalization	792.419*** (5.34)	378.566*** (5.58)
Agricultural GDP	--	0.458*** (18.58)
Control variable	control	control
R-squared	0.8607	0.9811
Regional fixed effect	control	control
Time-fixed effect	control	control

(*** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$).

favorable influence on agricultural carbon footprint with a parameter of 0.107. This implies that advancements in mechanical technology promote agricultural carbon emissions. Therefore, the action pathway for mechanical technological progress is as follows: Rural revitalization $\uparrow \rightarrow$ Mechanical technological progress $\uparrow \rightarrow$ Agricultural carbon emissions \uparrow .

As for model (1) of Table 7, rural revitalization holds a favorable influence on agricultural GDP with a coefficient of 792.419, indicating that rural revitalization promotes an increase in agricultural GDP. Model (2) shows that rural revitalization significantly contributes to agricultural carbon footprint with a parameter of 378.566. Furthermore, agricultural GDP hold a positive influence on agricultural carbon footprint with a parameter of 0.458. This indicates that farming GDP significantly promotes agricultural carbon emissions. Hence, the action pathway for agricultural GDP is as follows: Rural revitalization $\uparrow \rightarrow$ Agricultural GDP $\uparrow \rightarrow$ Agricultural carbon emissions \uparrow .

The study also utilized bootstrap sampling and Sobel analysis to investigate the intermediate effects of mechanical technological progress and agricultural GDP. The results of the Sobel analysis confirmed the findings from Tables 6, 7, showing similar significance levels for the intermediate model. Additionally, the bootstrap sampling results supported the presence of a significant intermediate effect. Therefore, we conclude that the assessment of the intermediate effect model is robust.

4.5 Robustness test

This study followed the approach of He et al. (He et al., 2022) to address the issue of endogeneity in the analysis of rural revitalization. The backward treatment method was applied to test the endogeneity. The Spatial Durbin model was employed, and the residual term was extracted. This residual term was then included as a new independent variable in the Durbin model. The p -values of the residual coefficient were found to be 0.377, 0.382, 0.386, 0.389, and 0.373, indicating that the model did not suffer from endogeneity.

To ensure the reliability of the spatial model findings, the research utilized the replacement space weight matrix method for robustness analysis. The economic geography nested matrix was replaced with the inverse distance square matrix. The results, as shown in Table 8, revealed that the signs of the core explanatory variables remained the same as in Table 4. These findings indicate that the spatial effect analysis conducted in this study exhibited a certain level of robustness.

4.6 Heterogeneity analysis

The accelerated evolution of the agricultural economy has generated an increasing imbalance in economic development among regions. This imbalance is reflected in various aspects

TABLE 8 Robustness test of replacing weight matrix.

Variables	Main	Wx	Direct effect	Indirect effect	Total effect
Rural revitalization	669.640*** (8.70)	173.292 (1.35)	682.213*** (9.44)	467.449*** (2.72)	1,149.662*** (6.68)
Economic urbanization	-5.004*** (-4.59)	4.128** (2.29)	-4.852*** (-4.45)	3.837* (1.80)	-1.015 (-0.53)
Urbanization of population	-3.998*** (-2.70)	4.240** (2.06)	-3.930*** (-2.82)	4.302* (1.83)	0.372 (0.19)
Land urbanization	-4.489*** (-6.46)	-6.524*** (-2.67)	-4.740*** (-6.58)	-10.334*** (-2.89)	-15.074*** (-3.81)
Industrial structure	-0.011 (-0.26)	-0.168** (-1.97)	-0.017 (-0.42)	-0.233** (-2.00)	-0.250* (-1.87)
Cultivated land scale	2.627* (1.69)	-5.777*** (-2.58)	2.553* (1.69)	-6.532** (-2.51)	-3.978* (-1.68)
Intensity of fixed asset investment of rural households	-0.013*** (-3.28)	0.041*** (7.28)	-0.012*** (-2.95)	0.049*** (8.29)	0.037*** (6.44)
Investment in fixed assets in primary industry	0.071* (1.84)	-0.791*** (-5.62)	0.039 (1.03)	-1.031*** (-4.67)	-0.992*** (-4.26)
rho	0.247*** (2.59)				
Observations	360				
R-squared	0.7020				

(***p < 0.01, **p < 0.05, *p < 0.1).

TABLE 9 Influence of diverse levels of rural revitalization on agricultural carbon footprint.

Variables	Low-level rural revitalization	Medium-level rural revitalization	High-level rural revitalization
Rural revitalization	126.546 (1.24)	737.995*** (4.04)	-183.220** (-2.33)
Economic urbanization	0.176 (0.17)	-3.322 (-1.62)	2.895*** (3.11)
Urbanization of population	2.673 (1.64)	-8.355*** (-3.39)	-3.726 (-1.07)
Land urbanization	1.079* (1.71)	-6.086*** (-2.71)	0.156 (0.18)
Cultivated land scale	3.650 (0.69)	6.391 (1.07)	3.087*** (2.68)
Intensity of fixed asset investment of rural households	-0.007 (-1.01)	-0.019** (-2.02)	-0.002 (-0.48)
Investment in fixed assets in primary industry	0.051*** (3.56)	-0.017*** (-2.95)	-0.001 (-0.63)
Individual fixation effect	Yes	Yes	Yes
Time-fixed effect	Yes	Yes	Yes
Observations	119	120	118
R-squared	0.9864	0.9856	0.9956

(***p < 0.01, **p < 0.05, *p < 0.1).

of rural life, including agricultural economy, rural culture, and resources allocation in agricultural production. These factors, in turn, impact agricultural carbon emissions. This study focused on different levels of rural revitalization and divided the sample into low, medium, and high levels of rural revitalization. By regressing the data based on these groups, the study aimed to determine the consequences of rural revitalization regarding agricultural carbon footprint and examine the heterogeneity.

The results, as presented in Table 9, showed that for areas with low-level rural revitalization, there was a positive but insignificant impact on agricultural carbon emissions. This suggests that the impact of rural revitalization on agricultural carbon footprint is minimal when the degree of rural revitalization is insufficient. However, when countryside revitalization reaches a medium level, it notably stimulates the

growth of agricultural carbon footprint, which matches with the regression results of the Durbin model mentioned earlier. Interestingly, when the level of rural revitalization further improves, the effect on agricultural carbon emissions shifts to a negative state at the 1% level. It indicates that the association between rural revitalization and agricultural carbon footprint is not straightforward and hold an inflection point effect. Moreover, the study found that with further improvement in rural revitalization, it can effectively curb agricultural carbon emissions.

Overall, these findings demonstrate the complex and nonlinear correlation between rural revitalization and agricultural carbon footprint, emphasizing the criticality of reaching an optimal level of rural revitalization to efficiently manage carbon footprint within the farming sector.

5 Conclusions and policy suggestions

5.1 Conclusion

On the basis of study on panel data for 18 prefecture-level cities in Henan Province from 2001 to 2020, the study examined the correlation between rural revitalization and agricultural carbon footprint. The analysis used systematic calculations to determine agricultural carbon footprint and employed Moran's index to identify spatial agglomeration patterns in these emissions. Additionally, a Spatial Durbin model was utilized to analyze the process and spatial effects of rural revitalization on agricultural carbon footprint. The analysis demonstrated that the global Moran's index analysis revealed a significant positive spatial autocorrelation in agricultural carbon footprint of Henan Province, indicating clustering patterns. The initial findings of analysis indicated that rural revitalization has a significant positive influence on agricultural carbon footprint. Intermediate effect model showed that rural revitalization promotes a growth in carbon footprint by stimulating farming GDP and mechanical technological progress. The spatial effect decomposition results demonstrated that rural revitalization has a notable spatial spillover effect. Improvements in the local rural revitalization degree effectively encourage the growth of agricultural carbon footprint in adjacent regions. The analysis about heterogeneity revealed an inflection point influence of rural revitalization on agricultural carbon footprint. Specifically, low and medium-level rural revitalization positively influenced carbon emissions, while further improvements in rural revitalization had a detrimental suppressive impact on these releases. These findings highlight the spatial characteristics and effects of rural revitalization on agricultural carbon footprint. It underscores the importance of considering spatial dynamics and optimizing rural revitalization strategies to efficiently address carbon footprint in agricultural areas.

5.2 Policy suggestions

Drawing from the research results, the subsequent recommendations are proposed.

- (1) Improving farming patterns according to local conditions: Consider economic, social, and environmental factors when implementing the rural revitalization strategy. Guide farmers to optimize farming patterns to mitigate the potential increase in agricultural carbon emissions resulting from rural revitalization. For example, promote the use of organic fertilizers to reduce greenhouse gas emissions, encourage the adoption of biotechnology and biological pest control methods, and provide subsidies for low-carbon agricultural machinery and soil testing formula fertilization.
- (2) Accelerating the advancement and application of eco-friendly agricultural core innovations: Recognize the role about agricultural GDP and mechanical technological progress in the correlation between rural revitalization and agricultural carbon footprint. Invest in exploration and advancement of eco-friendly core innovations to reduce emissions. Promote the adoption of technological advancements to mitigate agricultural carbon emissions.
- (3) Promoting inter-regional exchanges and cooperation: Recognize the spatial spillover effect of rural revitalization about neighboring areas' agricultural carbon emissions. Facilitate the exchange and cooperation between different regional agricultural departments to share technology and experiences. Accelerate the spread of agricultural technology and promote the adoption of sustainable farming practices to decrease carbon emissions.
- (4) Enhancing the degree of rural revitalization: Understand the inflection point influence of rural revitalization on agricultural carbon footprint. When the rural revitalization' level reaches a certain threshold, its impact changes from positive promotion to negative inhibition. Therefore, focus on advancing the level of rural revitalization through the Rural Revitalization Strategy, which contributes towards the decrease of agricultural carbon footprint.

By implementing these suggestions, it is possible to effectively address the carbon footprint within the farming industry while promoting rural revitalization and ensuring sustainable development.

Data availability statement

The original contributions presented in the study are included in the article/Supplementary material, further inquiries can be directed to the corresponding author.

Author contributions

JZ: Conceptualization, Funding acquisition, Methodology, Supervision, Writing–review and editing. YD: Data curation, Software, Writing–original draft.

Funding

The author(s) declare that no financial support was received for the research, authorship, and/or publication of this article.

Acknowledgments

The authors are grateful to reviewers and editors for helpful comments and suggestions.

Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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