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Psychometric properties of the critical thinking disposition scale in Peruvian adolescents

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Background: The disposition to critical thinking equips students with the capacity for objective analysis, informed decision-making, and problem-solving. Furthermore, it enhances intellectual autonomy and adaptability within academic contexts.

Objectives: To assess the psychometric properties of the Critical Thinking Disposition Scale in Peruvian adolescents.

Methodology: The psychometric adaptation process was implemented to evaluate the relevance, clarity, and comprehension of the items through expert judgment and pilot testing. Utilizing a sample of 499 students aged 10-17 years (M = 15.8, SD = 1.30), a confirmatory factor analysis (CFA) was conducted.

Results: The scale demonstrated validity and reliability, yielding optimal fit and reliability indices that support its unidimensional structure.

Conclusion: The validation of the Critical Thinking Disposition Scale in the Peruvian context provides a valuable instrument for assessing and fostering the propensity of Peruvian students to employ critical thinking skills, with implications for professional practice and educational policy. Furthermore, the study is significant because it shows that all factor values were greater than 0.40, which represents that the items contribute significantly to the evaluation of adolescents' disposition to critical thinking.

KEYWORDS

disposition to critical thinking, psychometric properties, adolescents, reliability, validity

1 Introduction

Critical thinking touches on a crucial topic in education since it is an essential indicator of quality in the teaching-learning process (Alsaleh, 2020), and its development is a priority in contemporary education (Aderoncele-Acosta et al., 2020).

Various investigations have demonstrated the effectiveness of critical thinking in academic activities. For example, Almulla and Al-Rahmi (2023) found that students who develop critical thinking have better learning capacity, and their self-efficacy increases (Namaziandost et al., 2023), increasing their performance (Martín-Raugh et al., 2023). Critical thinking promotes intellectual autonomy, independence (Fu et al., 2023), self-regulation, and self-assessment (Bernal et al., 2020). In this way, students

can make informed decisions and develop their self-organization skills favorably, becoming a key factor in dealing with adverse situations (Van Tonder et al., 2022).

Critical thinking also becomes relevant in the digital age (Leuwol et al., 2023; Sel, 2022) and its development is of significant interest within the field of education in general (Nunnally and Bernstein, 1995; Campo et al., 2023). It allows adolescents to evaluate and discern the information disseminated by digital media and the use they make of it, which constitutes a safeguard for their physical, mental, and social integrity, given the amount of false information that exists and can be harmful to them (Gouseti et al., 2024).

Recent reports highlight the looming danger of the digital age for teenagers (Rauscher and Badenhorst, 2021). For example, the Global Standards of Digital Inteligence Institute (2023) pointed out, after a study in 30 countries worldwide, the following risks for children and adolescents in the digital age. Of the participants, 45% face cyberbullying, 29% consume violent and sexual content, and 28% receive online threats. Likewise, the United Nations International Children's Emergency Fund (United Nations International Children's Emergency Fund, 2024) expresses great concern about these and other dangers that the digital age brings for adolescents.

Given this problem, timely intervention by national and international educational authorities is necessary to implement appropriate policies for children, adolescents, and young people (Miller, 2023). In this sense, promoting the development of critical thinkers can be a possible and effective alternative solution (Aguilera and Pandya, 2021).

Critical thinking also benefits students on a personal level. It allows them to consolidate their identity by creating a space for reflection on their experiences, beliefs, and values through self-knowledge and critical reflection. This knowledge is used for decision-making and developing the ability to resolve conflicts and maintain interpersonal relationships (Nunnally and Bernstein, 1995; Branje et al., 2021). Furthermore, Haag et al. (2022) assert that the development of critical thinking skills is a priority in the twenty-first century and enables students to construct their identity through active participation and reflective practice.

Critical thinking is also relevant for the development of a sense of social responsibility and civic spirit in adolescents by encouraging analysis, interpretation, and dialogue, allowing them to assume an active role in democratic, civic, and social participation for a committed participation in their environment (Fuentes-Moreno et al., 2020), helping to improve their self-esteem.

In the context described above, critical thinking constitutes an important skill for adolescents on a personal, academic, and socio-emotional level, contributing to their optimal development and preparing them for their transition to youth and adulthood (Bernal et al., 2020). In this way, characteristics of personality, such as honesty, empathy and solidarity, have been shown to be related to critical thinking and that, when combined, enhance the ability to make correct decisions in adolescents (Merma-Molina et al., 2022). However, when talking about critical thinking, it is necessary to mention that this does not only consist of the cognitive part Guerci de Siufi (2008), which has to do with the skills of searching for information, analysis, interpretation, and evaluation for decisionmaking (Richards et al., 2020); but it also includes a dispositional aspect, little addressed in research (Fandiño Parra et al., 2021), especially in the adolescent population. This disposition is the tendency, inclination, attitude, and mental habits that a person has to think critically and reflectively. That is the openness to listening to different perspectives or points of view for analysis based on reason and a solid argument (Ennis, 1996). Facione et al. (1995) assert that the development of critical thinking skills is essential for sound decision-making and effective problem-solving. Halpern (1998) posits that this constitutes a learning experience that would enable the adolescent to navigate new contexts, with enhanced prospects of attaining academic success. However, the development and application of these skills in various contexts would not be possible, but thanks to motivation and dispositional adaptability (Perkins et al., 1993).

This attitudinal aspect is fundamental for developing critical thinking skills (Lun et al., 2023). It is directly related to reflective learning and predicts the ability to make decisions after reflecting, analyzing, and verifying information (Karahan et al., 2023) and influences their learning abilities (Daradoumis and Arguedas, 2020).

Previous studies highlight the importance of a disposition toward critical thinking (Aysu, 2023) points out that the inclination to explore new ideas, reflective questioning, and proactive learning allow the development of the skills of this type of thinking (Rauscher and Badenhorst, 2021).

There are some scales to measure critical thinking skills (such as the California Critical Thinking Skills Test [CCTST, (Facione and Facione, 1994)] and The Watson–Glaser Critical Thinking Appraisal [WGCTA, (Watson, 1980)] used in various investigations around the world. Cui et al. (2021) mention that although there are various tools to measure critical thinking skills, a better understanding is required regarding the inclination to use these skills in everyday life, that is, the disposition.

In this sense, some known scale to measure the disposition to critical thinking are the California Critical Thinking Disposition Inventory (CCTDI), which combines critical thinking skills and dispositions into 75 items grouped into seven factors: inquisitiveness, open-mindedness, systematicity, analyticity, truthseeking, critical thinking self-confidence; maturity (Facione and Facione, 1992; Orhan, 2022), the Critical Thinking Disposition Assessment (CTDA), which measures the disposition to critical thinking through 19 items grouped into three factors: systematicity and analyticity, inquisitiveness and conversance, and maturity and skepticism (Yuan et al., 2014), the Watson-Glaser Critical Thinking Appraisal (WGCTA), which, although primarily focuses on critical thinking skills, can be used to quantify the predisposition to apply these skills in various contexts (El Hassan and Madhum, 2007) through 80 items grouped into five factors: interpretation, analysis, evaluation, inference, and explanation (Arslan and Demirtas, 2016; Kemper et al., 2019).

On the other hand, Quinn et al. (2020) point out a notable need for more consensus on measuring the dispositional dimension of critical thinking. Thus, the need for a valid and reliable instrument to adequately measure this construct is evident (Butler, 2024).

Faced with this need, after a thorough theoretical review, Sosu (2013) developed the Critical Thinking Disposition Scale (CTDS). This scale is the most relevant and appropriate for measuring the disposition to critical thinking since, unlike the previous ones, it focuses exclusively on a person's attitude to using critical thinking skills (Bakhtiari-Dovvombaygi et al., 2024). This relevance

and conceptual specificity are fundamental from a psychometric point of view, as pointed out by Kline (2015). Furthermore, the CTDS is the closest to the educational context, as it addresses the issue of commitment to critical thinking, highlighting the importance of maintaining the motivation to think critically, a constancy that is part of the educational task (Álvarez-Huerta et al., 2023). The CTDS has also demonstrated efficacy in administration due to its brevity compared to other scales, ensuring enhanced efficiency in data collection and mitigation of respondent fatigue (Arslan and Demirtas, 2016; Kemper et al., 2019). This makes it more appropriate for other extensive scales, as indicated by psychometric theory (Nunnally and Bernstein, 1994). Although it has received criticism for its brevity and simplicity, its use in recently published studies supports its reliability (Abanto-Ramírez et al., 2024; Karakuş, 2024; Li et al., 2024).

The CTDS has been validated and adapted to various contexts. In Türkiye, Akin et al. (2015) verified the bifactor structure of the CTDS with good fit indices ($x^2 = 53.24$, df = 40, RMSEA = 0.040, NFI = 0.90, NNFI = 0.96, GFI = 0.96, AGFI = 0.93, CFI = 0.97, IFI = 0.97 and SRMR = 0.046) and reliability ($\alpha = 0.78$).

Likewise, Gerdts-Andresen et al. (2022) evaluated the psychometric properties of CTDS in Norwegian university students by testing the two factors proposed by Sosu (2013) with good fit indices in the analysis performed (GFI = 0.95, TLF = 0.94, CFI = 0.96, RMSEA = 0.040, SRMR = 0.053) and an α = 0.76. Likewise, Orhan (2023) supports the two-factor structure. For its part, Yockey (2016) and Bravo et al. (2020) support the unifactorial structure of the scale by reporting that they present better parsimony and fit indices [CFI = 0.976; RMSEA = 0.089 (0.081,0.097); SRMR = 0.031; ECV = 0.889]. These discrepancies in the factor structure may be due to cultural and specific reasons of the various study populations, which makes it necessary to analyze the psychometric properties of this scale, as suggested by Anastasi and Urbina (1997).

Evaluating the critical thinking disposition scale is essential to understand Peruvian students' critical thinking disposition better. Therefore, verifying whether a scale is appropriate and relevant for Peruvian students is crucial. Few studies demonstrate the psychometric properties of this scale in the adolescent population. Most have only focused on university students, and there is one in the Peruvian context. This situation makes it difficult to measure this construct reliably and accurately, giving rise to possible inaccuracies and biases in the results obtained (American Educational Research Association, 2014). Therefore, this research aims to evaluate the psychometric properties of the Critical Thinking Disposition scale in Peruvian adolescents. This will allow educational actors, managers, and teachers to diagnose their students' willingness to think critically and design effective interventions to develop these skills.

2 Materials and methods

2.1 Design and participants

This study is of an instrumental type (Ato et al., 2013), using a convenience sampling method. The sample was selected using the electronic calculator Soper (2024), which considered several TABLE 1 Detailed description of the study sample.

Characteristi	cs	N	%	
N = 499				
Sex	Male	266	53.3	
	Female	233	46.7	
Year of study	First	205	41.1%	
	Second	27	5.4% 42.7%	
	Third	213		
	Forth	22	4.4%	
	Fifth	32	6.4%	
Age	12	27	5.4%	
	13	21	4.2%	
	14	17	3.4%	
	15	35		
	16	261	52.3%	
	17	138	27.7%	

TABLE 2 Content validity from Aiken's V.

Items	Aiken's V	CI 95%		
		Low	Up	
DPC1	0.93	0.79	0.98	
DPC2	0.99	0.87	1.00	
DPC3	0.97	0.83	0.99	
DPC4	1.00	0.89	1.00	
DPC5	0.98	0.85	1.00	
DPC6	1.00	0.89	1.00	
DPC7	1.00	0.89	1.00	
DPC8	1.00	0.89	1.00	
DPC9	1.00	0.89	1.00	
DPC10	1.00	0.89	1.00	
DPC11	1.00	0.89	1.00	
F1	0.99	0.87	1.00	

factors: the number of observed and latent variables in the model, the expected effect size ($\lambda = 0.10$), the desired statistical significance ($\alpha = 0.05$), and the level of statistical power ($1-\beta = 0.80$). According to these parameters, the minimum sample required for the study was 87 participants. However, 499 students were recruited, with a majority proportion of men (53.3%) aged between 12 and 17 years (M = 15.8, SD = 1.30). Most students were in their third year of secondary school, representing 42.7% of the sample (Table 1).

2.2 Instruments

The Critical Thinking Readiness Scale was developed by Sosu (2013) based on the theoretical proposals of the APA Delphi Report (Facione, 1990), Californian critical thinking inventory

(Facione and Facione, 1992; Perkins et al., 1993; Halonen, 1995; Ennis, 1996; Halpern, 1998). Its objective is to measure two dispositional domains of critical thinking: Critical Openness (the willingness to receive new ideas and critically evaluate them) and Reflective Skepticism (the tendency to reflect on past experiences and question current evidence). The Spanish version of this scale, validated in young university students, was used for this study (Bravo et al., 2020) because it is more suitable for the target population due to its cultural and linguistic adaptation compared to the original version in English, as well as the demonstration of validity and reliability (Muñiz et al., 2013). Its structure is unifactorial and consists of 11 items that can be directly quantified on a Likert scale from 1 to 5 (1 = Strongly disagree, 2 = Disagree, 3 = Undecided, 4 = Agree and 5 = Strongly agree). It has proven reliable ($\alpha = 0.777$) and has adequate psychometric properties (CFI = 0.976; RMSEA = 0.089; SRMR = 0.031). The version of the instrument used retains all the items of Sosu's initial proposal, for example "I usually try to think about the bigger picture during a discussion."

2.3 Procedure

The Peruvian Union University Ethics Committee approved the research protocol, with reference number 2023-CE-EPG-00110. Permission was requested from the management of the educational institutions to which the survey was applied. Permission was obtained from the parents and/or guardians of the students who participated in the study, considering they are minors. In addition, each student gave their consent to participate. Free and voluntary participation was taken into account, as well as the confidentiality and anonymity of the data collected, following the principles of the Declaration of Helsinki (Puri et al., 2009).

2.4 Analysis

To evaluate the content validity of the CTDS, five teachers with experience in teaching and research in the educational and psychometric field participated, who rated the relevance of the items and provided suggestions for improvement (Suárez-Alvarez et al., 2018). Responses were analyzed using Aiken's V coefficient (V = 1, p = 0.031; Aiken, 1985; Penfield and Giacobbi, 2004). A pilot test was carried out with 30 high school students through group discussion to verify the items' comprehensibility and detect possible semantic errors and inconsistencies (Muñiz and Fonseca-Pedrero, 2019).

Then, construct validity was determined. First, the descriptive statistics of the Critical Thinking Disposition Scale items (mean, standard deviation, asymmetry, and kurtosis) were analyzed using the FACTOR Analysis program version 10.1. In the second stage, the validity evidence based on the internal structure was analyzed through a Confirmatory Factor Analysis (CFA) utilizing the freely available R software within its RStudio interface, using the Lavaan package (Rosseel, 2012). The estimation method was Robust Weighted Least Squares (WLSMV), and structural equation modeling (SEM) was considered, in order to establish the existence of relationships between latent variables and observed variables,

TABLE 3	Preliminary	analysis	of the	items	of the	Critical	Thinking
Disposition	on Scale.						

	М	SD	g1	g2
n = 399	,			
Item 1	3.58	0.90	-1.11	1.38
Item 2	3.84	0.87	-1.41	2.66
Item 3	4.08	0.82	-1.48	3.64
Item 4	3.92	0.82	-1.23	2.72
Item 5	3.64	0.79	-0.93	1.54
Item 6	4.16	0.80	-1.44	3.54
Item 7	3.94	0.80	-1.38	3.38
Item 8	3.89	0.83	-1.1	2.48
Item 9	3.82	0.80	-1.08	2.18
Item 10	3.88	0.73	-1.12	2.86
Item 11	4.19	0.79	-1.61	4.69

TABLE 4 DPC reliability analysis.

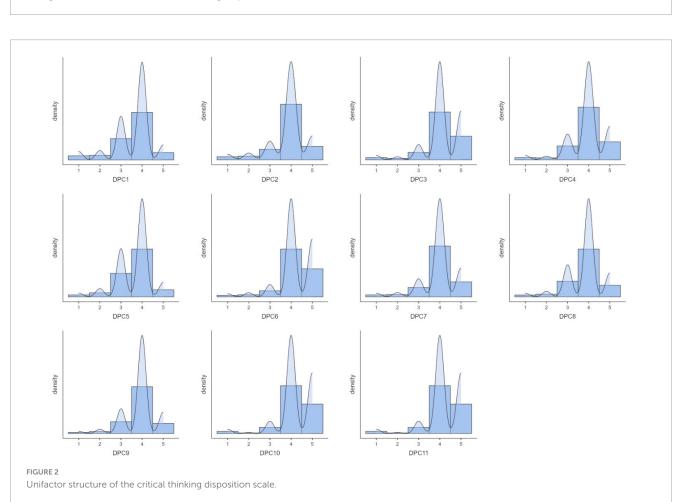
	Coefficient α	Coefficient ω
Item 1	0.853	0.856
Item 2	0.845	0.849
Item 3	0.845	0.849
Item 4	0.843	0.847
Item 5	0.852	0.856
Item 6	0.846	0.851
Item 7	0.847	0.852
Item 8	0.846	0.85
Item 9	0.85	0.854
Item 10	0.834	0.836
Item 11	0.834	0.836

thereby enabling the verification of a proposed structural model (Brown, 2015; Kline, 2016), It is beneficial to examine the factor structure of a scale and its psychometric properties (Byrne, 2010). The comparative fit index (CFI), adjusted goodness of fit index (AGFI), and Tucker-Lewis Index (TLI) were estimated. Also, the parameters for the root mean square error of approximation (RMSEA) and the root mean square error rate (SRMR) were taken into account, following the criteria proposed by Hu and Bentler (1999), who point out that the CFI and TLI should be greater than 0.9 and the RMSEA less than 0.08. Finally, the JAMOVI statistical software was used to evaluate reliability through Cronbach's alpha and Omega coefficients.

3 Results

3.1 Content validity

The content validation criterion was obtained from the evaluation of five specialists in the educational area, who rated



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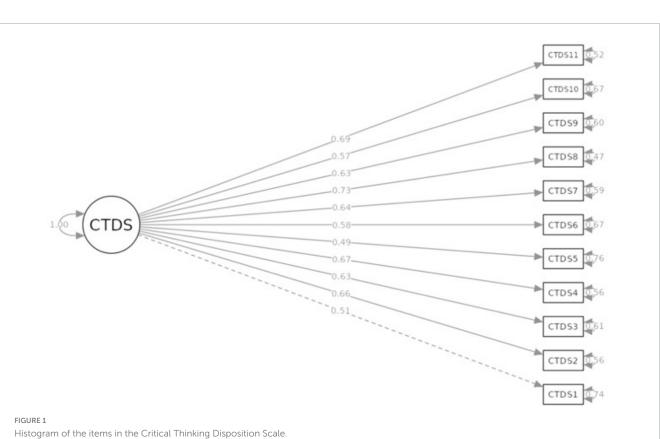


TABLE 5 Fit indices of the models evaluated by confirmatory factor analysis (CFA) of the study instrument.

Model	χ2	df	CFI	TLI	RMSEA		SRMR
	_	_	_	_	Value	(90%) CI	_
11 items	112.42	35	0.963	0.952	0.067	(0.061,0.081)	0.042

df, degree of freedom; CFI, comparative fit index; TLI, Tucker-Lewis's index; RMSEA, root mean square error of approximation; CI, confidence interval.

TABLE 6 Matrix of correlations of Critical Thinking Disposition Scale (CTDS) items.

	DPC1	DPC2	DPC3	DPC4	DPC5	DPC6	DPC7	DPC8	DPC9	DPC10	DPC11
DPC1	-	-	-	-	-	-	-	-	-	-	-
DPC2	0.344***	-	-	-	-	-	-	-	-	-	-
DPC3	0.257***	0.448***	-	-	-	-	-	-	-	-	-
DPC4	0.228***	0.435***	0.418***	-	-	-	-	-	-	-	-
DPC5	0.264***	0.336***	0.297***	0.322***	-	-	-	-	-	-	-
DPC6	0.365***	0.318***	0.372***	0.331***	0.257***	-	-	-	-	-	-
DPC7	0.305***	0.286***	0.292***	0.358***	0.298***	0.414***	-	-	-	_	-
DPC8	0.310***	0.340***	0.400***	0.395***	0.285***	0.278***	0.373***	-	-	-	-
DPC9	0.247***	0.277***	0.297***	0.299***	0.259***	0.254***	0.369***	0.415***	-	-	-
DPC10	0.329***	0.391***	0.388***	0.478***	0.311***	0.421***	0.345***	0.351***	0.379***	-	-
DPC11	0.329***	0.391***	0.388***	0.478***	0.311***	0.421***	0.345***	0.351***	0.379***	1.000***	-

***The correlation is significant at the 0.01 level.

the items based on the criteria of clarity, congruence, context, and belonging to the domain of the construct on a scale of 0 = "The validation criterion is not met" to 6 = "The validation criterion is met." As a result, the minimum value of V = 0.96 (CI 95% = 0.80; 0.99) for item 1, V = 0.93 (CI 95% = 0.79; 0.98), and values equal to greater than V = 0.97 for the rest of the items (CI 95% = 0.83; 0.99) for the rest of the items, considering them suitable for subsequent statistical analysis. These findings indicate that the item wording demonstrates clarity and unambiguity, uniformly measures the construct without contradictions, and aligns with the theoretical framework supporting the construct they aim to assess. No items were eliminated, as illustrated in Table 2.

3.2 Pilot test

Following recommendation of Wilson (2005), a qualitative pilot test was carried out using a discussion group made up of 30 adolescents with characteristics similar to the population under study. Each one evaluated the comprehensibility of the items and possible writing errors. The members of the pilot group easily understood all the items. This suggests that, upon administration of the scale to the participants, the items were readily comprehensible and did not elicit confusion or ambiguities. No difficulties were reported in interpreting the meaning of each item, and the language employed was appropriate for the sample's characteristics in terms of age, cultural context, and educational level.

3.3 Preliminary analysis of items and reliability

Table 3 shows the mean, standard deviation, asymmetry, and kurtosis of the eleven Critical Thinking Disposition Scale items.

The mean of the items ranges from 3.58 to 4.19, and its standard deviation exceeds 0.70 in all cases. The asymmetry indices (g1) of all items were within the range \pm 2, indicating a symmetrical data distribution (Finney and DiStefano, 2006). On the other hand, the kurtosis indices (g2) are less than three in most items, indicating a relatively normal distribution (de Winter et al., 2016; DeCarlo, 1997). Several items exhibit a kurtosis slightly exceeding three (3,6,7), indicating heavier tails and greater concentration in the mean values, as well as a more elongated tail (leptokurtic, in the case of item 11); this observation is common, in practice, within the social sciences (Cain et al., 2017). See Figure 1 and Table 3. In addition, it presents the reliability indices by internal consistency using Cronbach's alpha, ordinal alpha, and McDonald's Omega methods.

The findings have shown that the scale generally presents an expected internal consistency of Cronbach $\alpha = 0.857$, ordinal $\alpha = 0.870$, and McDonald's $\omega = 0.860$.

3.4 Confirmatory factor analysis

Validity evidence based on internal structure was analyzed through a CFA. The goodness-of-fit indices of the original model showed a good fit ($\chi^2 = 112.42$, df = 190, p = 0.000; CFI = 0.949; TLI = 0.937; RMSEA = 0.056 and SRMR < 0.036). As shown in Figure 2 and Table 5. This result represents an appropriate fit for the single-factor model of CTDS since it has a TLI and CFI greater than 0.90 and an RMSEA and SRMR less than 0.08, which indicates a good fit (Bentler, 1990; Hu and Bentler, 1999). Furthermore, the moderate to high correlations (0.22–1.00) presented in Table 6 indicate that the items share variance, which substantiates the univariate structure of CTDS (Nunnally and Bernstein, 1994).

4 Discussion and conclusion

The disposition to critical thinking is essential for education because it allows adaptation to personal, social, and professional demands Tiruneh et al. (2014). A decent education is characterized by the number of opportunities for thinking and practice given to students (Freire, 2021). Despite the importance of critical thinking in the school population, the literature does not provide evidence for the causes or results for populations under 18 (Çelik İskifoğlu et al., 2022). In this sense, the research objective was to validate the Critical Thinking Disposition scale in Peruvian adolescents.

The Critical Thinking Disposition Scale validation analyses in Peruvian adolescents were developed using a psychometric methodology (Ato et al., 2013). The Peruvian version of the Critical Thinking Disposition Scale showed good clarity considering a population of adolescents between 12 and 17 years old with direct approval of the 11 items and instructions. Likewise, the AFC findings confirm the unidimensional structure of the instrument in Spanish, with items that present factor loadings following what is recommended (Dominguez-Lara and Fernández-Arata, 2019). These results are similar to those reported by Bravo et al. (2020), where the single-factor structure was the most consistent considering the construct of disposition to critical thinking. This one-factor model is further based on recent theoretical evidence (Yockey, 2016), which can explain most of the variance of each item on the scale. This result differs from the original model (Sosu, 2013), which maintains two main dimensions.

These findings suggest both theoretical and practical implications. The results of this study confirm the application of the theory on dispositions to critical thinking, which identifies a set of attitudes that define the personal attitude and disposition to value and use critical thinking in personal, professional, and civic matters (Facione et al., 2021). In this sense, a unidimensional scale can assess the disposition to critical thinking (Bravo et al., 2020). Highlight that this scale has 11 items that ensure speed and practicality in its application, ultimately promoting research into this construct. This scale also aims to motivate Peruvian adolescent students to develop critical thinking skills and attitudes. As a result, it is expected that students will become more involved in academic activities, developing strategies that improve their learning capacity and decision-making attitudes, which will translate into greater academic success (Büyüközkan et al., 2020; Rivas et al., 2023). Furthermore, this scale is a solid tool that allows universities to find students with a greater disposition to critical thinking and, in turn, implement strategies that improve students' critical thinking skills. For this reason, this scale constitutes a fundamental axis in university evaluations that seeks to understand the disposition of adolescent students toward critical thinking. Finally, this study can help teachers find the determinants of dispositional critical thinking and improve their levels.

In relation to educational policy and based on the University Law (Ministerio de Educacion, 2014), the legislation establishes the training of professionals with high professional quality and a full sense of social responsibility, therefore, it is important to promote critical thinking skills, necessary in the analysis, evaluation and solution of society's problems. Therefore, relevant evaluation through reliable instruments such as the critical thinking disposition scale will help in achieving this objective. In this sense, the main contribution of the present study is to provide a tool that can help researchers study and evaluate the disposition to critical thinking in adolescent students in a Peruvian context, which is a current need (Pérez-Morán et al., 2021). Therefore, this critical thinking readiness scale can help develop public policies that address students' critical thinking attitudes and abilities. In this sense, this questionnaire allows for collecting information that supports actions to improve students' critical thinking skills during their adolescent years and guiding improvement actions in this field.

This research presents limitations: First, the students were selected conveniently, so probability samples must be developed to avoid representation bias. Although the sampling method has some advantages since in many cases random sampling is not feasible or practical due to the high cost and time to be developed (Etikan, 2016). Therefore, it is generalizable only to the surveyed population (Andrade, 2021). Furthermore, empirical research on the proposed instrument is limited to adolescent students only; additional studies could expand this scope to measure beyond institutional, educational intervention, understanding that a high percentage of students study outside the educational system. Finally, considering the sample's gender, 53.3% of the participants were women, so it is advisable to repeat the study with a more balanced gender distribution. Developing similar research by adding interviews and focus groups is also important.

In conclusion, the results showed that the Critical Thinking Disposition Scale in Peruvian adolescents shows adequate internal validity, based on content analysis and confirmatory factor analysis, and adequate reliability, evaluating the internal consistency of CTDS and whether the items are coherent with each other and reflect the same underlying construct. It is concluded that the Critical Thinking Readiness Scale in Peruvian adolescent students presents adequate psychometric properties and allows a significant evaluation of the student's willingness to develop critical thinking skills and attitudes.

Data availability statement

The raw data supporting the conclusions of this article will be made available by the authors, without undue reservation.

Ethics statement

The studies involving humans were approved by Comité de ética de la Escuela de Posgrado de la Universidad Peruana Unión. The studies were conducted in accordance with the local legislation and institutional requirements. Written informed consent for participation in this study was provided by the participants' legal guardians/next of kin.

Author contributions

HM-H: Conceptualization, Investigation, Writing – original draft. JT-C: Investigation, Methodology, Writing – review

and editing. CA-R: Project administration, Supervision, Validation, Writing – review and editing.

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