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Hybrid time series and ANN-based ELM model on JSE/FTSE closing stock prices

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Given the numerous factors that can influence stock prices such as a company's financial health, economic conditions, and the political climate, predicting stock prices can be quite difficult. However, the advent of the newer learning algorithm such as extreme learning machine (ELM) offers the potential to integrate ARIMA and ANN methods within a hybrid framework. This study aims to examine how hybrid time series models and an artificial neural network (ANN)-based ELM performed when analyzing daily Johannesburg Stock Exchange/Financial Times Stock Exchange (JSE/FTSE) closing stock prices over 5 years, from 15 June 2018 to 15 June 2023, encompassing 1,251 data points. The methods used in the study are autoregressive integrated moving average (ARIMA), ANN-based ELM, and a hybrid of ARIMA-ANN-based ELM. The ARIMA method was used to model linearity, while nonlinearity was modeled using an ANN-based ELM. The study further modeled both linearity and non-linearity using the hybrid ARIMA-ANN-based ELM model. The model was then compared to identify the best model for closing stock prices using error matrices. The error metrics revealed that the hybrid ARIMA-ANN-based ELM model performed better than the ARIMA [1, 6, 6] and ANN-based ELM models. It is evident from the literature that better forecasting leads to better policies in the future. Therefore, this study recommends policymakers and practitioners to use the hybrid model, as it yields better results. Furthermore, researchers may also delve into assessing the effectiveness of models by utilizing additional conventional linear models and hybrid variants such as ARIMA-generalized autoregressive conditional heteroskedasticity (GARCH) and ARIMA-EGARCH. Future studies could also integrate these with non-linear models to better capture both linear and non-linear patterns in the data.

KEYWORDS

artificial neural networks, ARIMA, extreme learning machine, hybrid models, stock price prediction

1 Introduction

The study investigated the hybrid time series and artificial neural network (ANN)-based extreme learning machine (ELM) model on Johannesburg Stock Exchange/Financial Times Stock Exchange (JSE/FTSE) closing stock prices. Traditional linear time series forecasting methods, such as the autoregressive integrated moving average (ARIMA) [1], the autoregressive conditional heteroskedasticity (ARCH) [2], the generalized autoregressive conditional heteroskedasticity (GARCH) [3], and many more have been proven to be effective for linear forecasting while providing poor performance for nonlinear forecasting [4]. The ARIMA model is a highly influential and commonly employed linear time series model [5]. ARIMA's popularity originates from its statistical properties, as well as the well-known Box-Jenkins (BJ) methodology, used in model development [6].

Numerous nonlinear models, such as support vector regression (SVR) [7–9], ANNs [10–12], and deep learning [13–15] have been proposed as alternative approaches for addressing the issue of nonlinearity, with ANNs being one of the most recognized and significant models [16]. Neural networks (NNs) are appealing for predicting tasks because they offer several benefits over traditional forecasting models. First, ANNs have adaptive nonlinear feature mapping properties that may accurately estimate any continuous measurable function with arbitrary accuracy. Second, as ANNs are nonparametric and data-driven models, they place minimal preconditions on the fundamental mechanism by which data are produced. This characteristic makes ANNs less prone to the model misspecification issue than many parametric nonlinear techniques [17]. Third, ANNs are naturally adaptable; the adaptability suggests that the network's generalization skills continue to be reliable and accurate in a non-stationary environment with changing environmental conditions [18]. Finally, while classic trigonometric expansions, polynomials, and splines employ exponentially various parameters to attain the same estimation rate, ANN models only utilize linearly various parameters [19].

Furthermore, some features obtained using the ANNs method become less valuable over time due to changes in the functional relationships between price series [20]. The limitations of the ANN methodology include a rapid rate of convergence and a high likelihood of becoming stuck in local minima, an excessive number of tunable parameters, a slow learning rate, long calculation time, and over-tuning [21]. To tackle these restrictions, the ELM model was created, which has been reported to have a high predictive capacity [22]. Manssouri et al. ([23], p. 7445) defined ELM as “a learning algorithm for feedforward NNs with a single hidden layer.” This method, including the backpropagation (BP) learning technique [24], offers various advantages over conventional learning techniques.

Based on the shortcomings of the existing approaches, the current study proposes a hybrid approach of ARIMA and ANN-based ELM since the proposed ARIMA model is incapable of handling nonlinear interactions and the ANN-based ELM model is unable to handle both nonlinear and linear patterns equally on its own. The hybrid approach was developed to increase the degree of accuracy of time series forecasts. The hybrid approach proposed in this study is an approach integrated with good adaptability to both linear and nonlinear situations, which are commonly encountered in complexly structured periodical time series. The hybrid approach will be generalized to other settings and therefore only limited to ARIMA and ANN-based ELM using the closing stock prices.

Due to the large number of factors that can affect stock prices, including the financial prosperity of the firm, the state of the economy, and the political environment, making predictions about stock prices can be challenging. The introduction of the relatively recent learning algorithm ELM opens up the possibility of direct ARIMA and ANN methods within a hybrid framework. This enables the development of a novel hybrid forecasting model that merges linear ARIMA with the nonlinear capabilities of ELM. Therefore, the objective of this study is to propose a model that can be used to effectively model the South African (JSE/FTSE closing) stock prices. The rest of the study is organized as follows: Section 2 presents the literature review, Section 3 presents the

research methodology, Section 4 discusses the data analysis and interpretation of results, and Section 5 provides the conclusion.

2 Literature review

Khan and Alghulaiakh [25] employed and compared ARIMA models using 5 years of historical Netflix stock data and two customized ARIMA (p, d, q) models to create a precise stock forecasting model. The model's accuracy was determined and compared using autocorrelation functions (ACFs), PACFs, and the mean absolute percentage of error (MAPE). After numerous tests, ARIMA (1, 1, 33) demonstrated correct outcomes in its calculations, demonstrating the capability of the ARIMA model on time series to provide reliable stock forecasts that will assist stock investors in their investment decisions.

Khanderwal and Mohanty [26] provided a thorough explanation of how to construct an ARIMA model for predicting stock prices. The selected ARIMA (0,1,0) model experimental findings demonstrated evidently that ARIMA models can estimate stock costs in a short-run manner with adequate accuracy. This could help stock market investors make effective investment decisions. With the results obtained, ARIMA models were completely effective in the short-term prediction market with emerging prediction techniques.

Milačić et al. [24] carried out a study to develop and utilize an ANN with an ELM to predict the gross domestic product (GDP) growth rates. The GDP-added values from the manufacturing, agricultural, industrial, and service sectors were used to forecast GDP growth rates. The predicted capacities of the ANN models proposed were compared using the root mean square error (RMSE), Pearson coefficient (r), and coefficient of determination (R^2) indicators. The back propagation (BP) and ELM were compared with the predicted values' accuracy level. The results of the simulation showed that, depending on the inputs used, ANN with ELM can forecast GDP positively. When it comes to GDP applications, particularly GDP estimates, the ELM approach may work well.

To enhance the precision of time series forecasting, the study by Pan et al. [27] employed an autoregression (AR) and NN-based ELM hybrid model. Two unknown prediction sets were employed along with the known portion of the set (TR), which is assumed to span the years ranging from 1700 to 1920. From 1921 to 1955, the first (PR1) and, from 1956 to 1979, the second (PR2) were used to test the proposed method's forecasting accuracy. The findings obtained from the hybrid model were contrasted with those produced by the AR and the NN-based ELM. In this study, Pan et al. [27] applied Normalized Mean Square Error (NMSE) and RMSE statistical measures to the experiment. The findings showed that the hybrid model outperformed the NN-based ELM and AR models when evaluated against various types of time series data.

Nonlinear and linear models can be used independently or in combination with a variety of methods for forecasting time series. Research shows that combining nonlinear and linear models can improve forecast accuracy. Büyüksahin and Ertekin [28], in this study, present a novel hybrid ARIMA-ANN algorithm that functions in broader contexts. GbpUsd, Lynx, Sunspot, and Intraday are the four datasets used to estimate the outcomes of

the suggested hybrid strategy and the other methodologies (ANN, ARIMA Zhang's, Khashei-Bijari's, Naïve, and Babu-Reddy's). To compare accuracy and performance, mean absolute error (MAE), mean square error (MSE), and mean absolute scaled error (MASE) are used. The findings of the experiments demonstrate that approaches for deconstructing the primary data and integrating models that are both linear and nonlinear through the process of hybridization are significant indicators of the methods' forecasting capability. These results are used to combine the Empirical Mode Decomposition (EMD) approach with the suggested hybrid method, producing more predictable components. The results of this study demonstrate that integrating a hybrid technique with EMD with any of the other methods applied independently can also be a helpful technique for improving the level of prediction precision achieved with traditional hybrid techniques.

3 Materials and methods

The current study uses daily JSE/FTSE All Index closing price data from 15th June 2018 to 15th June 2023. The data were obtained from the Index Market Data on the JSE's website (jse.co.za). The analysis was carried out using Python and R Studio software packages. The following subsections discuss the models that are used in the study.

3.1 ARIMA

The modern approach to time series analysis is defined by the Box and Jenkins [29] process. The Box and Jenkins (BJ) method seeks to construct an ARIMA model from an observed time series. In particular, the technique focuses on stationary processes, passing over helpful preliminary transformations of data [30]. The ARIMA models have dominated time series forecasting for a very long time in several fields [31]. According to Carvajal et al. [32], the model assumes that a variable's future value is a linear function of multiple recent observations and random errors, where p is the number of AR terms, d is the number of non-seasonal differences, and q is the number of moving average (MA) lags. The "integrated" step, which changes the time series to turn a non-stationary time series into a stationary time series, allows it to handle non-stationary time series data [33]. The general form of the ARIMA model is denoted as ARIMA (p, d, q), and the model is given as follows:

$$Y_t = \varphi_1 Y_{t-1} + \varphi_2 Y_{t-2} + \dots + \varphi_p Y_{t-p} + e_t + \theta_1 e_t - 1 + \theta_2 e_t - 2 + \dots + \theta_q e_t - q, \quad (1)$$

where Y_t is the response variable being predicted at time t , $Y_{t-1}, Y_{t-2}, \dots, Y_{t-p}$ is the response variable at time lags $t-1, t-2, \dots, t-p$, respectively, $\varphi_1, \varphi_2, \dots, \varphi_p$ and $\theta_1, \theta_2, \dots, \theta_q$ are the AR and MA parameters, respectively, and e_t 's are the white noise. This study used three iterative BJ procedures, as explained below. The first step in the Box-Jenkins process is to determine whether or not the time series data is stationary by plotting the graph to obtain a general idea of the data and to understand the trend. If the mean and variance of a series do not change over time, then the series has stationarity. The sample ACF also makes the

data visible, in addition to using the graphic representations of the data across time to assess if they are stationary or non-stationary. If the time series data is not stationary, it will be transformed for stationarity. This study places a greater emphasis on logarithm differencing. Even though there have been a lot of stationarity tests suggested in the literature, this study will use the Augmented Dickey-Fuller (ADF) and the Kwiatkowski-Phillips-Schmidt-Shin (KPSS) tests. Additionally, the study will also use correlograms in support of the formal tests.

3.1.1 ADF test

Dickey and Fuller [34] researched stationarity testing originally and conceptualized the technique as "testing for a unit root." The hypothesis for the ADF unit root test is as follows:

$$H_0 : \beta_0 = 0 \text{ (The data is non-stationary or is a unit root)} \quad (2)$$

$$H_1 : \beta_1 \neq 0 \text{ (The data is stationary or there is no unit root)} \quad (3)$$

The ADF test statistic has the following form:

$$DF_t = \frac{y}{SE(y)} \quad (4)$$

where y is the least squares (LS) coefficient estimate of the y coefficient and $SE(y)$ is the standard error of the LS estimate of the y coefficient from the regression model.

3.1.2 The Kwiatkowski, Phillips, Schmidt and Shin test

Kwiatkowski et al. [35] proposed a Lagrange Multiplier (LM) test (the KPSS test) to evaluate the trend and/or level of stationarity. In other words, the null hypothesis assumes a stationary process. A conservative testing approach would assume the unit root as an alternative to the null hypothesis and consider it as a stationary process. Therefore, when the null hypothesis is not accepted, it is evident that the series has a unit root [36]. The KPSS test hypothesis is as follows:

$$H_0 : \sigma_e^2 = 0 \quad (5)$$

$$H_1 : \sigma_e^2 \neq 0 \quad (6)$$

Under H_0 of $e_t \sim NIID(0, \sigma_e^2)$, the KPSS test statistic has the following form:

$$LM = \frac{\sum_{t=1}^T S_t^2}{\hat{\sigma}_e^2} \quad (7)$$

where,

$$\hat{\sigma}_e^2 = \frac{\sum_{t=1}^T e_t^2}{T} \quad (8)$$

and

$$S_t = \sum_{i=1}^t e_i, t = 1, \dots, T \quad (9)$$

where e_i are the residuals obtained from the regression of Y_t on a constant and a time trend [36].

TABLE 1 Behavior of ACF and PACF.

| | AR(p) | MA(q) | ARMA (p, q) |
|------|----------------------|----------------------|-------------|
| ACF | Dies down | Cuts off after lag p | Dies down |
| PACF | Cuts off after lag q | Dies down | Dies down |

Source: Suwardo et al. [55].

3.1.3 Correlograms

The autocorrelation function (ACF) and partial autocorrelation function (PACF) are used to provide formal test techniques in addition to graphical stationarity checking. To analyse the time series data and attempt to identify the functional form of the data, ACF and PACF are plotted in the correlograms [37]. The AR (p) and MA (q) are determined from the analysis of the ACF and PACF; and d indicates the number of differences applied. The AR coefficients ϕ's and MA coefficients θ's are estimated from the model based on p, d, and q [38]. The time series values are considered to have stationarity if the ACF of the time series swiftly decays or dies off. If the ACF plot decays very slowly, the time series values are non-stationary. Table 1 illustrates how the various models behave on the ACF and PACF.

In the second step, the parameters of the models identified (defined) in the first step are estimated. The maximum likelihood (ML) approach is employed in the model estimation. In a standard Gaussian, the likelihood function is given as follows:

$$\text{Log}L = -\frac{T}{L} \ln(2\pi) - \frac{T}{L} \ln \sigma^2 - \frac{1}{2\sigma^2} \sum_{i=1}^T \varepsilon_i^2 \quad (10)$$

where T is the time at t = 1, 2, . . . , T of time series data and σ and ε are the constant variance and the error term, respectively. The logarithm of the probability of the observed data from the fitted model is shown by the log-likelihood [37]. The model with the maximum log-likelihood is chosen. To select the most appropriate ARIMA model, the Akaike Information Criterion (AIC) is applied to all competing models. The AIC proposed by Akaike [39] is a method that uses in-sample fit to determine the likelihood of a model forecasting future values [40]. The model with the lowest AIC value is the best one among all the other estimated models [41]. The equation used to estimate the AIC is as follows:

$$\text{AIC} = -2 * \ln(L) + 2 * k \quad (11)$$

where L is the value of the likelihood and k is the number of estimated parameters.

In the Box–Jenkins technique, the third step is diagnostic testing, which entails standard testing procedures on the estimations and the error terms' statistical properties (weak white noise assumption and normality assumption) [42]. The adequacy of the model can be evaluated using both formal testing techniques and graphical testing methods. The Ljung–Box [43] test will be used to check the overall acceptability of the overall model. The hypothesis for Ljung–Box is formulated as follows:

$$H_0 = \text{model is adequate} \quad (12)$$

$$H_1 = \text{model is not adequate} \quad (13)$$

The test statistic for Ljung–Box is computed using the following equation:

$$Q^* = n^1 (n^1 + 2) \sum_{l=1}^k \frac{1}{n^1 - l} r_l^2(\hat{a}) \quad (14)$$

where n¹ = n - d, n is the number of observations used in the estimated model and d is the level of non-seasonal differencing employed in transforming the initial time series data into stationary data. The r_l²(â) represents the square of the residual's autocorrelation at lag l. If Q* is larger (significantly larger from zero), it is that the autocorrelation residuals are considerably distinct from 0 as a collection, and the estimated model's random shocks are autocorrelated. Since the model will be rejected, one should consider repeating the model-building cycle [30]. The Jarque–Bera test, which will be used to test the normality of the residuals of the model, is a common statistical test used to test for normality in return series. This assumption refers to the degree to which data follows a normal distribution. The JB test for normality comes from a Chi-square distribution, which is calculated using the skewness and kurtosis with two degrees of freedom. The following is the formulation of the hypothesis that will be tested:

$$H_0 : E(\varepsilon^2) = 0 \text{ (Data is normally distributed)} \quad (15)$$

$$H_1 : E(\varepsilon^2) \neq 0 \text{ (Non - normal distribution of data)} \quad (16)$$

The JB test statistic is computed using the following equation:

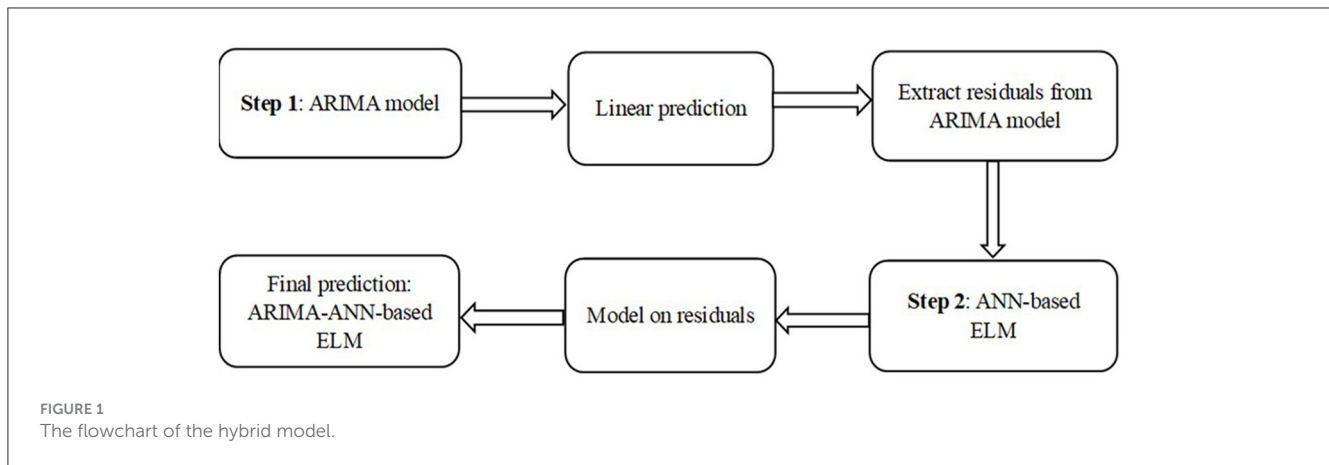
$$JB = N^* \left[\frac{\text{skewness}}{6} + \frac{(\text{kurtosis} - 3)^2}{4} \right] \quad (17)$$

where N is the number of observations. The likelihood that the given series is drawn from a normal distribution decreases with increasing JB value.

3.2 ANN-based ELM

Tokar and Johnson [44] defined ANN-based ELM as a “fast-training artificial intelligence (AI) approach for prediction that employs a Single Layer Feedforward Neural Network (SLFN) to build a relationship between complicated nonlinear dependent and independent variables”. This method has the advantage of not requiring any knowledge of the complexity of the process under study. The adoption of a nonlinear activation function and the capability of randomly learning input weights increased the popularity of this technique among researchers [45]. When training input, the layer's ANN-based ELM hidden node is independent of the hidden layer. This indicates that hidden nodes were independent of the input training set [46]. For N arbitrary distinct inputs samples (u_i, t_i), where u = [u_{i1}, u_{i2} u_{im}]^T ∈ Rⁿ and t_i = [t_{i1}, t_{i2} t_{im}]^T ∈ R^m. The following formula can be used to represent SLFNs with hidden neurons:

$$\sum_{i=1}^{\hat{N}} \beta_i g_i(u_i) = \sum_{i=1}^{\hat{N}} \beta_i g(w_i \bullet u_j + b_i) = 0_j \quad j = 1, \dots, N \quad (18)$$



where $w_i = [w_{i1}, w_{i2}, \dots, w_{in}]^T$ is the weight vector between the input nodes and the i th hidden node, $\beta = [\beta_{i1}, \beta_{i2}, \dots, \beta_{in}]$ is the weight vector between the i th hidden node and the output nodes, b_i denotes the threshold of the i th hidden node, and $w_i \bullet u_i$ is the inner product of w_i and u_i [45, 47]. The SLFN can approximate n vectors means that there exist b_i, w_i such that

$$\sum_{j=1}^{\tilde{N}} \|O_j - T_j\| = 0 \tag{19}$$

The term β_i of Equation 18 is estimated as follows:

$$\sum_{i=1}^{\tilde{N}} \beta_i g(w_i \bullet u_j + b_i) = t_j, j = 1, \dots, N \tag{20}$$

Equation 20 can also be written as follows:

$$H\beta = T, \tag{21}$$

where H is the hidden layer output matrix of the NNs given by

$$H = (w_1 \dots w_n, b_1 \dots b_n, u_1 \dots u_n) = \begin{bmatrix} g(w_1 \bullet u_1 + b_1) & \dots & g(w_{\tilde{N}} \bullet u_1 + b_{\tilde{N}}) \\ \vdots & \dots & \vdots \\ g(w_1 \bullet u_N + b_1) & \dots & g(w_{\tilde{N}} \bullet u_N + b_{\tilde{N}}) \end{bmatrix}_{N \times \tilde{N}} \tag{22}$$

where β is the weights connecting the hidden and output layers computed using the following:

$$\beta = \begin{bmatrix} \beta_1^T \\ \vdots \\ \beta_{\tilde{N}}^T \end{bmatrix}_{\tilde{N} \times M} \tag{24}$$

where T is the target values of N vectors in the training dataset given by

$$T = \begin{bmatrix} t_1^T \\ \vdots \\ t_N^T \end{bmatrix}_{N \times M} \tag{25}$$

TABLE 2 The results of the descriptive statistics of JSE/FTSE closing prices.

| | |
|--------------------|------------|
| Mean | 62,625.840 |
| Median | 59,408.680 |
| Maximum | 80,791.360 |
| Minimum | 37,963.010 |
| Standard deviation | 8,751.866 |
| Skewness | 0.220 |
| Kurtosis | -0.887 |
| Observations | 1,251 |

where H_0 is referred to as the generalized Moore-Penrose inverse of matrix H . When the number of hidden neurons and training samples is equal, SLFNs may approximate the training samples with no error. Numerous techniques, such as singular value decomposition (SVD), iterative approaches, orthogonal projection methods, and orthogonalization methods, can be used to determine H_0 . It was shown that SLFNs with randomly generated hidden nodes and with a pervasive piecewise continuous activation function may universally approximate any continuous target function. The SVD approach is employed to compute H_0 [45].

3.3 Hybrid ARIMA-ANN-based ELM

The goal of hybrid models is to lower the probability of employing an incorrect model by fusing different models to lower the chance of failure and provide more accurate results [19]. The ARIMA and ANN-based ELM models have found success in their respective nonlinear or linear areas. However, none of these is a universal model that can be applied to all situations. Traditional models' approximation to complicated nonlinear issues and ANN-based ELM approximation to linear problems may be completely incorrect, especially in situations with both nonlinear and linear correlation structures. Unlike previously introduced hybrid forecasting models, which normally handle the original forecasting models as distinct linear or nonlinear units, the suggested hybrid model is an integrated model that can respond

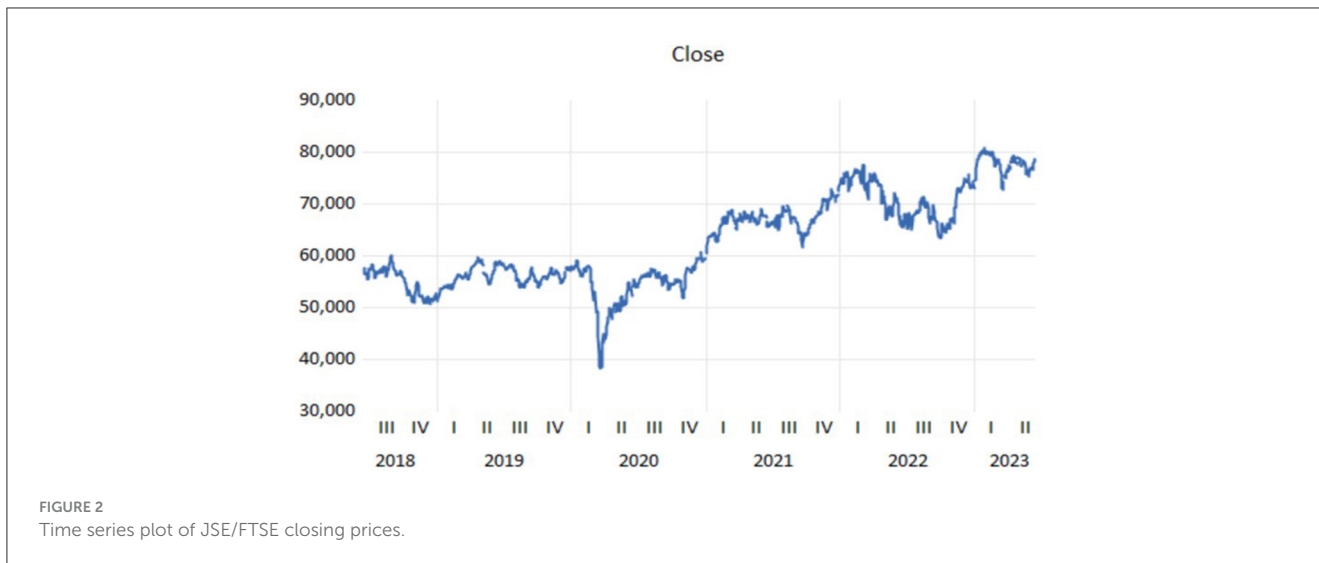


FIGURE 2 Time series plot of JSE/FTSE closing prices.

TABLE 3 ADF and KPSS test results of JSE/FTSE closing prices.

| Returns | ADF test statistics | Probability | KPSS test statistics | Probability |
|----------------------------|---------------------|-------------|----------------------|-------------|
| JSE/FTSE at level | -2.866 | 0.212 | 12.212 | 0.010 |
| JSE/FTSE at log difference | -10.600 | 0.000 | 0.064 | 0.100 |

well to both linear and nonlinear conditions, which are common in complex frequent time series.

According to a few researchers in hybrid linear and nonlinear models, it would be reasonable to assume that a time series is made up of a linear autocorrelation structure and a nonlinear component [48], which is given as follows:

$$y_t = N_t + L_t \tag{26}$$

where L_t represents the linear component and N_t represents the nonlinear component. The data must be used to estimate these two components. The study first allows ARIMA to model the linear component, and only the nonlinear relationship will be present in the residuals from the linear model [48]. Let e_t represent the residual at time t from the linear model, then

$$e_t = y_t - \hat{L}_t \tag{27}$$

where \hat{L}_t is the forecast value for time t from the estimated relationship. By modeling the residuals using ANN-based ELM, the nonlinear relationship can be revealed [48]. With n input nodes, the ANN-based ELM model for the residuals will be as follows:

$$e_t = f(e_{t-1}, \dots, e_{t-n}) + \varepsilon_t \tag{28}$$

where f is the nonlinear function that is established by the NN and ε_t is the random error. It should be noted that the error term may not be random if the model f is inappropriate [48]. Consequently, it

is crucial to identify the right model. The forecast from Equation 28 is represented as \hat{N}_t , then the combined forecast will be as follows:

$$\hat{y}_t = \hat{L}_t + \hat{N}_t \tag{29}$$

The hybrid model uses the unique characteristics and strengths of both the ANN-based ELM and the ARIMA models to identify various patterns. To improve the overall effectiveness of modeling and forecasting, it may be useful to model linear and nonlinear patterns independently using several models before combining the forecasts [48, 49]. The flowchart of the hybrid model is presented in Figure 1.

3.4 Model evaluation

Three evaluation metrics, such as MAPE, MSE, and MAE, are employed to measure the predictability of the methods, which are computed from the following equations:

$$MAPE = \frac{1}{N} \sum_{i=t}^N \frac{|x_i - \hat{x}_i|}{|x_i|} \times 100\% \tag{30}$$

$$MSE = \frac{1}{N} \sum_{i=1}^N (x_i - \hat{x}_i)^2 \tag{31}$$

$$MAE = \frac{1}{N} \sum_{i=t}^N \frac{|x_i - \hat{x}_i|}{|x_i|} \tag{32}$$

where n is the number of values and \hat{x} is the forecast value [50]. The model with the lowest values of MAPE, MSE, and MAE will be selected and proposed as an adequate algorithm for predicting purposes.

4 Results and discussion

This section of the study presents the results of the data analysis and the discussion of the results obtained from the data analysis. Table 2 presents the results of the descriptive statistics. Descriptive

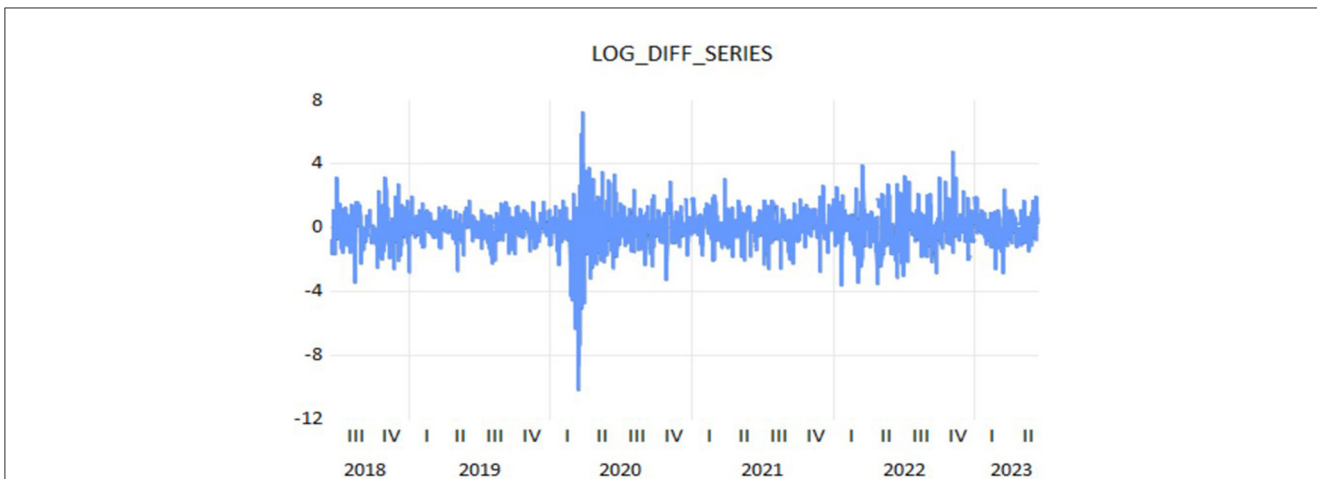


FIGURE 3 Time series plot of the log differenced data.

| Autocorrelation | Partial Correlation | AC | PAC | Q-Stat | Prob |
|-----------------|---------------------|-----------|--------|--------|-------|
| | | 1 -0.013 | -0.013 | 0.2176 | 0.641 |
| | | 2 0.012 | 0.011 | 0.3846 | 0.825 |
| | | 3 0.061 | 0.061 | 5.0170 | 0.171 |
| | | 4 -0.029 | -0.028 | 6.0799 | 0.193 |
| | | 5 0.011 | 0.009 | 6.2327 | 0.284 |
| | | 6 -0.076 | -0.079 | 13.422 | 0.037 |
| | | 7 0.088 | 0.090 | 23.105 | 0.002 |
| | | 8 -0.072 | -0.072 | 29.612 | 0.000 |
| | | 9 -0.021 | -0.013 | 30.158 | 0.000 |
| | | 10 -0.020 | -0.036 | 30.657 | 0.001 |
| | | 11 0.046 | 0.064 | 33.278 | 0.000 |
| | | 12 0.006 | -0.004 | 33.327 | 0.001 |
| | | 13 -0.037 | -0.020 | 35.105 | 0.001 |
| | | 14 0.065 | 0.038 | 40.529 | 0.000 |
| | | 15 0.032 | 0.048 | 41.804 | 0.000 |
| | | 16 -0.044 | -0.050 | 44.239 | 0.000 |
| | | 17 -0.070 | -0.072 | 50.503 | 0.000 |
| | | 18 -0.055 | -0.072 | 54.412 | 0.000 |
| | | 19 0.001 | 0.014 | 54.414 | 0.000 |
| | | 20 0.055 | 0.079 | 58.262 | 0.000 |

FIGURE 4 ACF and PACF plots of the log differenced data.

TABLE 4 Model selection of ARIMA models for log-differenced data.

| Returns | Models | AIC |
|----------|-----------------|------------|
| JSE/FTSE | ARIMA [1, 6, 6] | -7,366.347 |
| | ARIMA [1, 6, 7] | — |
| | ARIMA [1, 6, 7] | -7,365.513 |
| | ARIMA [1, 7, 7] | -7,366.271 |

statistics are used to describe the dataset used in the study. The results in Table 2 revealed that both the mean and median are positive, suggesting that the closing prices increase slightly over time. The skewness coefficient shows that time series data is skewed to the right. A negative kurtosis demonstrates that the distribution is relatively flat. The kurtosis value of the time series data is less than 3, which reveals that the distribution has characteristics of a platykurtic distribution. The standard deviation value is 8,751.866,

and this large value suggests that the data points are more spread apart from the mean. This signifies that the time series data has a higher level of variability or dispersion.

4.1 The findings from the Box–Jenkins procedure

The Box–Jenkins procedure used in the study follows a three-step approach. The three-step approach is outlined in the following subsections.

4.1.1 Model identification

To create a BJ model, first determine whether the series is stationary and observe any patterns. The plot provides a first

TABLE 5 Parameter estimates of fitted ARIMA [1, 6, 6].

| | Estimate | SE | z-value | Pr (> z) |
|------|----------|-------|---------|-----------|
| ar 1 | -1.691 | 0.778 | -2.172 | 0.029 |
| ar 2 | -1.566 | 1.240 | -1.263 | 0.207 |
| ar 3 | -1.543 | 1.095 | -1.409 | 0.159 |
| ar 4 | -0.880 | 1.081 | -0.815 | 0.415 |
| ar 5 | -0.069 | 0.557 | -0.123 | 0.901 |
| ar 6 | -0.012 | 0.045 | -0.278 | 0.781 |
| ma 1 | 0.697 | 0.778 | 0.897 | 0.369 |
| ma 2 | -0.110 | 0.474 | -0.232 | 0.816 |
| ma 3 | 0.063 | 0.136 | 0.465 | 0.642 |
| ma 4 | -0.694 | 0.073 | -9.503 | <2e-16 |
| ma 5 | -0.876 | 0.554 | -1.581 | 0.114 |
| ma 6 | -0.081 | 0.628 | -0.128 | 0.898 |

TABLE 6 JB test results of the fitted ARIMA [1, 6, 6] model residuals.

| Returns | JB test statistics | Probability |
|-----------|--------------------|-------------|
| Residuals | 2,867.500 | <2.2e-16 |

TABLE 7 Diagnostic test results of the fitted ARIMA model residuals.

| Model | Ljung-Box Q* | p-value |
|-----------------|--------------|---------|
| ARIMA [1, 6, 6] | 25.679 | 0.177 |

indication of the expected nature of the time series as presented in Figure 2.

The time series plot of JSE/FTSE closing prices is presented above in Figure 2. The plot reveals irregular fluctuations, which implies that the mean and variance change over time. This is a demonstration that this time series is non-stationary by eye inspection. Through logarithm differencing, trend behavior in non-stationary data may be changed. As stated, this study places a greater emphasis on logarithmic differencing. In any time series study, it is significant to understand the concept of stationarity and its definition regarding keeping statistical properties constant over time. The stationarity test makes the study of model estimation and forecasts easier. The study made use of ADF and KPSS tests.

The stationary test results for ADF and KPSS are presented in Table 3. The null hypothesis of a unit root (non-stationary series) is not rejected, since the p-value associated with the ADF test is greater than the 5% significance level. This indicates that the time series is not stationary and has a trend or other types of dependency over time. Furthermore, the KPSS test findings have a p-value less than the significant level of 5%, the null hypothesis is rejected, and the time series data is non-stationary. Both tests provide evidence that differencing/logarithm is necessary. The results of the ADF test indicate that the time series is stationary after log differencing since its p-value is less than the 5% significance level, thus rejecting the null hypothesis of non-stationary. The KPSS results are also consistent with the ADF test results, which also provide evidence that, at a 5% significance level, the log-differenced time series is

indeed stationary. The time series plot of the log-differenced data is presented in Figure 3.

The figure shows fluctuation around the mean of zero, suggesting that the series remains constant to the mean. In addition to graphical stationarity checking, the ACF and PACF are employed for formal testing and identify the order of ARIMA models.

Considering Figure 4 of the ACF and PACF plots, it is observed that the PACF plot suggests AR [6] and AR [7]. The ACF plot also suggests MA [6] and MA [7]. In the ACF plot, spikes are additionally observed at higher lags. The strikes from lag 8 have been discarded since Tsoku et al. (42, p. 764) specified that “spikes at greater lags are commonly ignored to simplify the initial tentatively recognized model.” The plots indicated several AR and MA lags; therefore, the AIC approach will be employed as the selection criteria to select an ARIMA model.

4.1.2 Model estimation and selection

The findings indicated the existence of ARIMA models with various lag orders. However, not all the ARIMA models found could be used in this study. Hence, choosing an appropriate model is crucial. Table 4 gives a summary of the models that are compared. According to the findings in Table 5, the chosen ARIMA model for log-differenced data is ARIMA [1, 6, 6] based on the lowest AIC value. The findings further revealed that ARIMA [1, 6, 7] did not converge, therefore the model was not considered. The next step is to compute the parameter estimation of the selected model. The results presented in Table 5 show the coefficients of AR and MA components that resulted from modeling the selected model using the ML method.

Based on the coefficients summarized in Table 5, the estimated parameters generate the following equation:

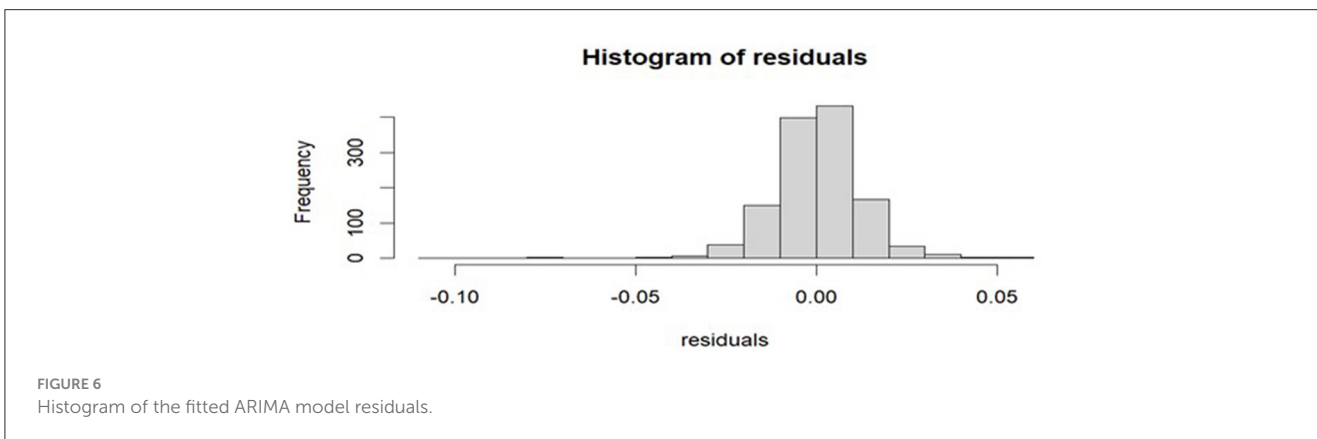
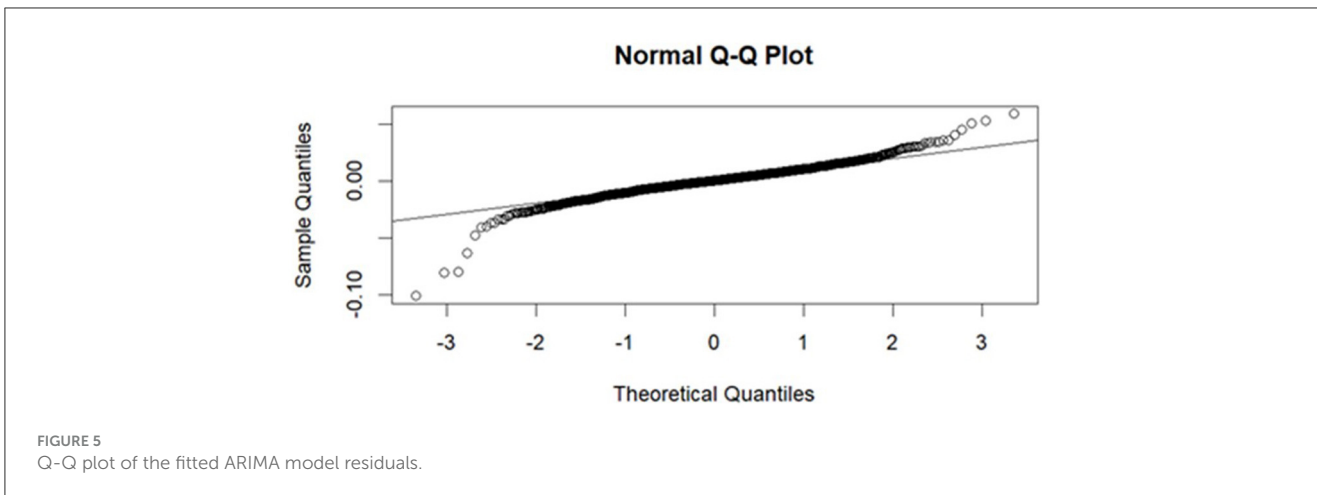
$$\begin{aligned}
 Y_t = & -1.691Y_{t-1} - 1.566Y_{t-2} - 1.543Y_{t-3} - 0.880Y_{t-4} \\
 & - 0.069Y_{t-5} - 0.0123 Y_{t-6} + 0.697e_{t-1} - 0.110e_{t-2} \\
 & + 0.063e_{t-3} - 0.6940e_{t-4} - 0.876e_{t-5} - 0.081e_{t-6} + \varepsilon_t \quad (33)
 \end{aligned}$$

The next step is to run a diagnostic check to ensure that the chosen model is appropriate for further analysis.

4.1.3 Diagnostic checking results

The best-fit model is determined by how well the residual analysis is carried out in time series modeling [51]. The diagnostic test results of the fitted ARIMA model are presented in Tables 6, 7 and Figures 5, 6.

The JB test was employed to determine the normality of the residuals, and the findings are shown in Table 6. The findings suggest that the null hypothesis of normality in the residuals is not accepted, and the conclusion is that the residual distribution is not normal. The graphical representations of residual diagnostics are displayed in Figures 5, 6. It is evident from Figure 5 that almost all the points are either on the 45-degree line or very near to it, but a few of the points are slightly separated from the Quantile-Quantile (QQ) line. Furthermore, the residuals of the fitted model are slightly negatively skewed and normally distributed, as shown by the histogram in Figure 6.



To further test the model’s adequacy, this study performed the Ljung–Box test for residuals. The Q^* statistic (25.679) of the residuals and p -value (0.177) are given in Table 7, where the null hypothesis is not rejected since the p -value exceeds the significant level of 5%. Therefore, it can be concluded that there is sufficient statistical evidence that the selected model is adequate. This suggests that the fitted ARIMA [1, 6, 6] model is sufficient enough to be used for further analysis.

4.2 ANN-based ELM

As the second step of the analysis, ANN-based ELM was applied. The dataset was divided into test data and training data to assess the forecasting abilities of the model. The test data is used to assess the forecasted model, while the training data is used to build the model. The dataset was split into 80% training and 20% testing. The log-differenced data set has test and train sizes of 250 and 1,000 rows, respectively. With an 80%/20% splitting approach, predictive models can obtain better prediction performance, as outlined by Abdulkareem et al. [52]. The model is trained using the training data set, and then the test dataset is used to assess its performance. The study evaluated the model’s performance using three evaluation metrics, and the results are presented in Table 8.

TABLE 8 Evaluation metrics for the differenced data.

| Model | MAPE | MSE | MAE |
|----------------------------|---------------|--------------|--------------|
| ARIMA [1, 6, 6] | 99.999 | 2,499.970 | 49.999 |
| ANN-based ELM | 26.665 | 0.046 | 0.213 |
| Hybrid ARIMA-ANN-based ELM | 11.421 | 0.008 | 0.089 |

4.3 Hybrid ARIMA-ANN-based ELM

Different ARIMA models were obtained for the initial step of building hybrid models in ARIMA-ANN-based ELM modeling, but for better results, the log-differenced data were fitted using the ARIMA [1, 6, 6] model. The hybrid model was performed by modeling the residuals from the ARIMA [1, 6, 6] model using an ANN-based ELM. The accuracy measurements, to show which model was chosen, were used to determine which model can be used for forecasting between ARIMA[1, 6, 6], ANN-based ELM, and hybrid ARIMA-ANN-based ELM, and the results are provided in Table 8. The results in Table 8 indicate that the hybrid ARIMA-ANN-based ELM yields better prediction results for closing prices since it has the lowest values of MAPE, MSE, and MAE. Therefore, the hybrid ARIMA-ANN-based ELM was selected as a better-performing model as compared to the individual models.

5 Conclusion

The study modeled the closing stock prices using three models, namely, ARIMA, ANN-based ELM, and a hybrid of ARIMA-ANN-based ELM to determine the most appropriate model. The findings demonstrated that there were four competing ARIMA models, namely, ARIMA [1, 6, 6], ARIMA [1, 6, 7], ARIMA [1, 6, 7], and ARIMA [1, 7, 7]. The chosen ARIMA model for log-differenced data was found to be ARIMA [1, 6, 6] based on the lowest AIC value.

The diagnostic test was computed and revealed that the chosen model was adequate. The dataset was then divided into 80% for training and 20% for testing for ANN-based ELM. The model was trained using the training data set, and its performance was evaluated using the test dataset. These methods, however, may not be sufficient when modeling time series data with both linear and nonlinear characteristics simultaneously, as stated by Bulut and Hudaverdi [53]. Hence, this study proposed a hybrid of both ARIMA and ANN-based ELM to model both linearity and nonlinearity simultaneously. The study developed a hybrid approach for time series prediction that aims to address the limitations of prior hybrid methods by eliminating strong assumptions. It was evident from the findings that the hybrid model performed better than the individual models. It was also evident that the hybrid model improved the performance of the ARIMA and ANN-based ELM when computed individually.

The study by Khan et al. [54] also integrated the ARIMA and ANN models. The findings from the study by Khan et al. [54] demonstrated that ANN performs more effectively in forecasting than the ARIMA model, but when their forecasts are combined, the hybrid ARIMA-ANN outperformed in terms of forecast accuracy, ANN, and ARIMA. The study recommends policymakers and practitioners to use the hybrid model as it yields better results. In the future, researchers may delve into assessing the effectiveness of models by using additional conventional linear models and hybrid variants such as ARIMA-GARCH and ARIMA-EGARCH. Future studies could also integrate these with nonlinear models to better capture both linear and nonlinear patterns in the data. This approach offers a promising avenue for enhancing model performance and gaining deeper insights into complex phenomena.

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Data availability statement

Publicly available datasets were analyzed in this study. This data can be found at: Johannesburg Stock Exchange (jse.co.za).

Author contributions

OM: Writing – original draft, Writing – review & editing. JT: Writing – original draft, Writing – review & editing. DM: Writing – original draft, Writing – review & editing.

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Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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