



Metastases Growth Patterns *in vivo*—A Unique Test Case of a Metastatic Colorectal Cancer Patient

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Colorectal cancer (CRC) is one of the most common causes of cancer-related mortality worldwide. Most cases of deaths result from metastases, assumed to be shed, in many cases, before disease detection. Providing reliable predictions of the metastases' growth pattern may help planning treatment. Available mathematical tumor growth models rely mainly on primary tumor data, and rarely relate to metastases growth. The aim of this work was to explore CRC lung metastases growth patterns. We used data of a metastatic CRC patient, for whom 10 lung metastases were measured while untreated by seven serial computed tomography (CT) scans, during almost 3 years. Three mathematical growth models—Exponential, logistic, and Gompertzian—were fitted to the actual measurements. Goodness of fit of each of the models to actual growth was estimated using different scores. Factors affecting growth pattern were explored: size, location, and primary tumor resection. Exponential growth model demonstrated good fit to data of all metastases. Logistic and Gompertzian growth models, in most cases, were overfitted and hence unreliable. Metastases inception time, calculated by backwards extrapolation of the fitted growth models, was 8–19 years before primary tumor diagnosis date. Three out of ten metastases demonstrated enhanced growth rate shortly after primary tumor resection. Our unique data provide evidence that exponential growth of CRC lung metastases is a legitimate approximation, and encourage focusing research on short-term effects of surgery on metastases growth rate.

SIGNIFICANCE

Providing reliable predictions of the metastases' growth pattern using mathematical models may help determining the optimal treatment plan that fits a given patient best and maximizes the probability of cure.

Keywords: lung metastases, mathematical growth models, exponential growth, logistic growth, gompertzian growth, primary tumor resection

INTRODUCTION

Colorectal cancer (CRC) is one of the most common causes of cancer-related morbidity and mortality worldwide. Most cases of deaths result from development of metastatic disease [1]. CRC has a slow natural history (i.e., development of disease) that provides a great opportunity for early detection and prevention strategies. Surgery is the main curative treatment, but despite

complete resection of the primary tumor, metastatic disease might develop in a significant number of patients [2].

The exact dynamics of tumor and metastasis formation is not well-established. It is assumed that many (if not most) of metastases are shed before primary tumor is even detectable [1, 3, 4]. Hence, preventing metastases growth by adjuvant or perioperative treatments is indicated in many cases after resection of primary tumor [5, 6]. Providing reliable predictions of the metastases' growth pattern using mathematical models may help determining the optimal treatment plan.

Growth laws of primary tumors are thoroughly investigated [7, 8], however, not many mathematical models are dealing with metastases growth dynamics in humans. Many of the mathematical models for primary tumor growth are based on fitting *in-vivo* data to relatively simple growth models, such as exponential, logistic, Gompertzian, or power law [9, 10], and models for metastases dynamics rely on the same laws. The Gompertzian law is considered most reliable, because it was found that generally, doubling time of tumors usually decreases with time. Nevertheless, the assumption of exponential growth is preferred over logistic or Gompertzian because it includes one less parameter, which reduces the degree of freedom in the model, consequently reducing the difficulty in getting numerical convergence with limited amount of data. Hence, exponential law is often assumed, at least for the first period of growth [10–12]. However, this assumption is hard to prove *in vivo*, since there are very few available data of untreated metastases growth in humans. Moreover, diversity between patients, and between metastases of the same patient, further increases the challenge when trying to find growth patterns that can be used as predictors.

Here, we describe a CRC patient with 10 lung metastases, for which uncommon data of *in vivo* growth over time is available. The metastases were followed and measured—while untreated—for over 2 years. Our aims were:

- To describe metastases growth pattern and decide which of the three models—Gompertzian, exponential, logistic—fits best.
- To determine whether factors such as location and size of metastases have an effect on growth pattern and rate.
- To estimate natural history of the disease (i.e., time of onset of metastases).

MATERIALS AND METHODS

Data

A 65 years old patient was diagnosed with rectal cancer TNM stage [13] T3N0 (and colon polyp containing superficial cancer TNM stage T1N0). A CT scan at the time of first diagnosis showed also 8 mm nodules in the lungs. A PET-CT scan did not show FDG uptake in these nodules, which may have implied that these nodules are not malignant. Fifty-four days after first diagnosis, the primary tumor in rectum (and colonic polyp) were resected. On post-surgery follow up, six additional CT scans were conducted, roughly every 6

months, in which 10 lung metastases were evidently growing. During this time period systemic treatment (chemotherapy, targeted treatment) was offered, but not administered, because of personal preference of the patient. The measured volumes of metastases at these seven timepoints (marked as W1–W7) are reported in **Table 1**. See **Figure S1** for examples of CT tomographic images and **Figure S2** for illustration of the locations of all diagnosed metastases (marked #1–#10) in the lungs.

Modeling

Based on the data available, we wanted to set a growth model (exponential, Gompertzian, or logistic) for each of the 10 metastases, and assess the values of growth rate parameters.

Exponential growth was modeled by the equation:

$$\Psi(t) = N_0^{\text{exp}} e^{\lambda t}, \quad (1)$$

where $\Psi(t)$ is the metastasis volume at time t , counted from the day of primary tumor resection, N_0^{exp} is the size of metastasis at $t = 0$, and λ is the growth rate parameter.

Logistic growth was modeled by the equation:

$$\Theta(t) = \frac{K^{\text{logistic}}}{1 + \left(\frac{K^{\text{logistic}}}{N_0^{\text{logistic}}} - 1 \right) e^{-rt}}, \quad (2)$$

where $\Theta(t)$ is metastasis volume at time t , N_0^{logistic} is the size of metastasis at $t = 0$, K^{logistic} is the limiting tumor size—carrying capacity, and r is a rate parameter.

Gompertzian growth was modeled by:

$$\Phi(t) = K^{\text{gomp}} e^{\ln\left(\frac{N_0^{\text{gomp}}}{K^{\text{gomp}}}\right) e^{-\beta t}}, \quad (3)$$

where $\Phi(t)$ is metastasis volume at time t , N_0^{gomp} is the size of metastasis at $t = 0$, K^{gomp} is the limiting tumor size and β is a rate parameter.

Direct fit of the data, by numerical minimization of the sum of squared errors (SSE) was done for each of the metastases separately, to optimize the parameter values for each of the three equations: N_0^{exp} and λ in Equation (1), N_0^{logistic} , K^{logistic} , and r in Equation (2) and N_0^{gomp} , K^{gomp} and β in Equation (3). Specifically, the minimization was done for errors of the model predictions of the log-volume of tumor size:

$$\text{SSE} = \sum_{i=1}^n (\ln(f(t_i, p)) - \ln(Y_i))^2, \quad (4)$$

where Y_i is the observed metastasis volume at time t_i and $f(t_i, p)$ is predicted metastasis volume at the same time, as calculated by each of the model Equations (1)–(3), depending on the estimated parameters vector p . The minimization procedure was performed using the Matlab functions *lsqnonlin* and *nlinfit*.

TABLE 1 | Metastases sizes measured by CT scans of the patient, at different times marked W1–W7.

	W1	W2	W3	W4	W5	W6	W7
Date	17/10/2012	08/05/2013	06/10/2013	22/04/2014	05/10/2014	27/04/2015	29/09/2015
MET1	0.014	0.016	0.050	0.099	0.177	0.326	0.776
MET2	0.178	0.236	0.309	0.754	1.466	3.613	6.589
MET3	0.004	0.101	0.197	0.356	0.544	0.940	1.371
MET4	0.128	0.330	0.506	0.921	2.384	5.370	9.292
MET5	0.108	0.205	0.349	0.674	1.039	3.933	14.547
MET6	–	0.077	0.347	0.479	0.887	3.031	4.475
MET7	–	0.058	0.197	0.410	0.675	1.565	2.138
MET8	0.292	0.807	4.944	8.548	12.718	32.654	66.693
MET9	0.108	0.209	0.361	0.543	0.954	1.338	1.897
MET10	0.175	0.429	2.719	8.045	19.250	55.708	91.538

Primary tumor was resected on 10/12/2012, 54 days after W1. First row is date of the CT scan, and other rows are metastases volumes in cm^3 .

The fit was done for each metastasis using data of all available measurements in time, including at time W1, conducted 54 days before resection. Indeed, the growth law and rate may change between W1 and W2 due to the resection, however we assumed that the time between W1 and resection time was short enough that it would change the measure only slightly, within the measurement error.

Goodness of Fit Analysis

Different criteria for the goodness of fit were compared, in order to determine the best growth model for each of the metastases, and the reliability of the estimated parameter values [10, 14]. For this purpose, the root of mean square of errors (RMSE) was calculated for each of the three models that were fitted to each of the 10 metastases.

$$RMSE = \sqrt{MSE} = \sqrt{\frac{SSE}{(n - P)}}, \tag{5}$$

where *SSE* is defined by Equation (4). The MSE is normalized to the number of measurements (*n*) available for the specific metastasis, and to the number of model parameters (*P*), to enable fair comparison between exponential model (where *P* = 2) and the other models (where *P* = 3).

Another criterion used for the goodness of fit of predicted curves to the data was the adjusted coefficient of determination:

$$R^2 = 1 - \frac{n - 1}{n - P} \frac{SSE}{SST}, \tag{6}$$

where $SST = \sum_{i=1}^n (\ln(Y_i) - \overline{\ln(Y_i)})^2$, and $\overline{\ln(Y_i)}$ is the time average of the observed log-volumes measured at all *n* time-points. This metric quantifies how much of the variability in the data is described by the model, as the denominator is proportional to the data variance. In this case, *R*² is adjusted

to the number of measurements (*n*) and number of model parameters (*P*).

To quantify the reliability of the estimated parameter values, the variance-covariance matrix of parameters was calculated, in the context of non-linear least squares regression:

$$Cov = MSE \cdot (J^T J)^{-1}, \tag{7}$$

where *J* is the Jacobian of the model as a function of the parameters vector *p*: $J_{ij} = \frac{\partial f(t_i, p)}{\partial p_j}$. *p_j* is the *j*th element of *p*. The variance of an estimated parameter *p_j* is defined by the diagonal element of the covariance matrix, *Cov_{jj}*. This is a measure of the sensitivity of model prediction to the estimated value of the parameter *p_j*.

Evaluation of Metastases' Natural History

After the best fitted models are chosen, and their reliability is established, the fitted models can be used to estimate the time of onset of metastasis. For this purpose, the fitted curve with estimated parameters for each metastasis *k* was extrapolated backwards to determine the time of onset of metastasis (*T_k*), defined as time of appearance of the first malignant cell, adopting the evaluation of $10^{-9} cm^3$ for the volume of a single tumor cell. For example, in case of an exponential model the value for *T_k* was derived from $\Psi(t = T_k) = 10^{-9} cm^3$. This method was also used to assess the time of metastasis' size reaching to the threshold enabling detection by CT scan (*D_k*), approximated as $0.002 cm^3$.

RESULTS

Fitting and Comparing Growth Models

For every one of the metastases, values for the parameters of each of the three growth models examined were fitted to the dataset of all available measurements at times W1–W7. Metastases #6 and #7 were not detectable at W1 timepoint (see **Table 1**). For metastasis #3, the measure at W1 was omitted from the fit since it was very small—close to the limit of detection. The parameters' optimal values, as well as different scores for goodness of fit

(see section Materials and Methods), are presented in **Tables 2, 3**. In general, all growth models provided good fit for most of metastases, as their predicted curves are within or close to the measurement error bounds (**Figure 1**). This accuracy reflects in the adjusted R squared values (**Table 3**) which are almost all >0.94 . Unlike R^2 , the SSE and RMSE values are not normalized to variability of observed values. Hence, comparing SSE or RMSE values of different metastases would reflect the variability in absolute values of metastases volumes: metastases that grow to high volumes would have higher SEE and RMSE values. However, their values can be used to compare between goodness of fit of different growth models for the same metastasis, as detailed below.

For metastases numbered 1, 2, 4, 5, comparing goodness of fit of the three models shows that the exponential model demonstrated the closest prediction to actual growth measurements for these metastases in all three scores (see **Table 3**). Logistic and Gompertzian models converged with extremely high values of carrying capacity parameter K (marked orange in **Table 2**), which means that they essentially degenerate into exponent. The variance of K could not be calculated in those cases, because the Jacobian was singular or close to singular, i.e., curve fit does not depend on the value of K . This may point on redundancy in these models.

For metastases numbered 6, 8, 10 the exponential fit scores were inferior than those of the other two models. However, Gompertzian fit has the same problem of parameter redundancy, where curve fit does not depend on the value of K (marked orange in **Table 2**). The logistic fit is also not reliable in these cases, since the variance of the parameter K is very large, 10–100 times its value (marked pink in **Table 2**). Hence, in these metastases, exponential model is also the preferable one.

For metastases 3, 7, 9 the logistic curve seems reliable, and it has better scores than exponential. However, these metastases are very small in size hence measurement error is relatively large, and exponential curve is also within this error. Gompertzian fit is not reliable from the same reasons mentioned above for other metastases.

For all metastases described above, it seems that using exponential approximation for the growth law is a good enough approximation, at least for a range of 2 years from the time of primary tumor detection and resection (for the first timepoints of measure W1–W5, or W2–W5 for metastases #3, #6, and #7).

Variability of Growth Rates of Metastases

Looking at the fitted exponential model parameters, the value of exponent of the growth rate λ (see Equation 1) is in the same order of magnitude for all metastases, and its value is estimated to be the average of their fitted values: 1.48 years^{-1} , with standard deviation of 0.34 years^{-1} (**Table 2**). Their distribution (assumed to be normal) is presented in **Figure 2A**. Note, that the metastases most distant from this mean value are #9 and #10, which are both located in the left lung, while all other metastases are in the right lung. Other than that, no relation was found between fitted growth rates to the metastasis location in the lungs, nor

TABLE 2 | Values of estimated optimal parameters, for each of the 10 observed metastases, for the three fitted models (see Equations 1–3), along with their variance (see Equation 7), presented in parentheses.

	Exponential			Logistic			Gompertz		
	N_0^{exp} [cm^3] (var)	λ [years^{-1}] (var)	$N_0^{logistic}$ [cm^3] (var)	$K^{logistic}$ [cm^3] (var)	r [years^{-1}] (var)	N_0^{gomp} [cm^3] (var)	K^{gomp} [cm^3] (var)	β [years^{-1}] (var)	
MET #1	0.0136 (4.1E-06)	1.4016 (2.2E-05)	0.0136 (4.1E-06)	2.33E+06 (2.33E+06)	1.4008 (2.2E-05)	0.0132 (4.4E-06)	5.44E+08 (4.4E-08)	0.0621 (4.7E-08)	
MET #2	0.1495 (5.8E-04)	1.2945 (2.6E-05)	0.1495 (5.8E-04)	2.75E+07 (2.75E+07)	1.2943 (2.6E-05)	0.1468 (7.7E-04)	6.14E+08 (8.1E-08)	0.0629 (8.1E-08)	
MET #3	0.0758 (5.2E-05)	1.0604 (7.6E-06)	0.0638 (4.0E-05)	2.6565 (0.8103)	1.3243 (3.8E-05)	0.0562 (3.3E-05)	22.31 (593.1)	0.2692 (1.5E-05)	
MET #4	0.1595 (1.7E-04)	1.4504 (6.4E-06)	0.1594 (1.7E-04)	1.37E+08 (1.37E+08)	1.4496 (6.4E-06)	0.1536 (2.3E-04)	6.46E+08 (6.46E+08)	0.0717 (2.2E-08)	
MET #5	0.1012 (6.4E-04)	1.5730 (6.1E-05)	0.1012 (6.4E-04)	3.13E+06 (3.13E+06)	1.5725 (6.1E-05)	0.0992 (8.2E-04)	1.08E+09 (1.08E+09)	0.0741 (1.8E-07)	
MET #6	0.0565 (2.7E-04)	1.6017 (7.3E-05)	0.0506 (5.3E-04)	15.00 (1.495)	1.7406 (5.1E-04)	0.0411 (1.9E-04)	2.82E+03 (2.82E+03)	0.1954 (1.0E-06)	
MET #7	0.0465 (1.3E-04)	1.4452 (5.0E-05)	0.0318 (7.9E-05)	2.8551 (1.4E+00)	1.9492 (2.3E-04)	0.0207 (5.6E-05)	9.8627 (104)	0.4930 (8.2E-05)	
MET #8	0.5395 (2.3E-02)	1.7831 (7.7E-05)	0.4525 (1.9E-02)	76.82 (2822)	2.1618 (3.3E-04)	0.4187 (9.2E-03)	1245.44 (1245.44)	0.3410 (1.7E-06)	
MET #9	0.1426 (1.3E-04)	0.9631 (6.1E-06)	0.1309 (4.7E-05)	3.3190 (6.5E-01)	1.2171 (1.8E-05)	0.1301 (4.0E-05)	38.1023 (1940)	0.2266 (1.0E-05)	
MET #10	0.2821 (5.1E-03)	2.2069 (6.3E-05)	0.2282 (2.3E-03)	118.51 (2637)	2.6072 (1.3E-04)	0.2249 (1.9E-03)	5.37E+04 (5.37E+04)	0.2430 (4.2E-07)	

Colored in pink are parameters for which calculated variance is very high (> 10 times the parameter value). Colored in orange are parameters for which the variance could not be calculated, because the Jacobian in Equation (7) was singular or close to singular.

TABLE 3 | Different measures of the goodness of fit, for each of the 10 observed metastases, for the three fitted models.

	SSE			RMSE			Adjusted R^2		
	Exp	Logistic	Gomp	Exp	Logistic	Gomp	Exp	Logistic	Gomp
MET #1	0.267	0.267	0.297	0.231	0.258	0.272	0.977	0.976	0.967
MET #2	0.317	0.317	0.415	0.252	0.281	0.322	0.968	0.968	0.947
MET #3	0.046	0.016	0.009	0.107	0.074	0.054	0.988	0.996	0.997
MET #4	0.080	0.080	0.112	0.126	0.141	0.167	0.993	0.993	0.988
MET #5	0.761	0.762	0.973	0.390	0.436	0.493	0.948	0.948	0.917
MET #6	0.443	0.424	0.401	0.333	0.376	0.366	0.950	0.953	0.940
MET #7	0.305	0.137	0.083	0.276	0.214	0.167	0.958	0.981	0.985
MET #8	0.959	0.703	0.525	0.438	0.419	0.362	0.949	0.963	0.965
MET #9	0.076	0.020	0.017	0.123	0.071	0.066	0.986	0.996	0.996
MET #10	0.782	0.340	0.402	0.396	0.292	0.317	0.972	0.988	0.982

See Equations (4)–(6).

to its size at the time of detection. The distribution of initial metastases sizes (at time W_1), and the lack of correlation to growth rate λ can be seen in **Figure 2B**. The similarity of exponential growth rates of metastases can be also seen in **Figure 3A**, where fitted models for all metastases are presented on the same graph (note that the figure is presented in log-scale, and volumes are also normalized to the initially detected volume, at time W_1).

Metastases' Natural History

If we assume each metastasis has followed the same growth law since its formation, then for each metastasis k the onset time (i.e., time of emergence of the first malignant clonogenic cell,) T_K , could be estimated. The earliest possible detection time (i.e., time of metastasis' size reaching to the threshold enabling detection by CT scan,) D_k , could also be evaluated. These evaluations were obtained by extrapolating backwards of the fitted exponential growth model, assuming growth rate was the same through all the time of metastasis' existence. Results, presented in **Figure 3**, show that according to the model, all metastases were formed 8–19 years before primary tumor was detected (**Figure 3B**), however, earliest possible time on which they could be detected, assuming detection limit is 0.002 cm^3 , was years later $-1-5$ years before primary tumor detection (This can be seen in **Figure 3B**, and more clearly in **Figure 3A**).

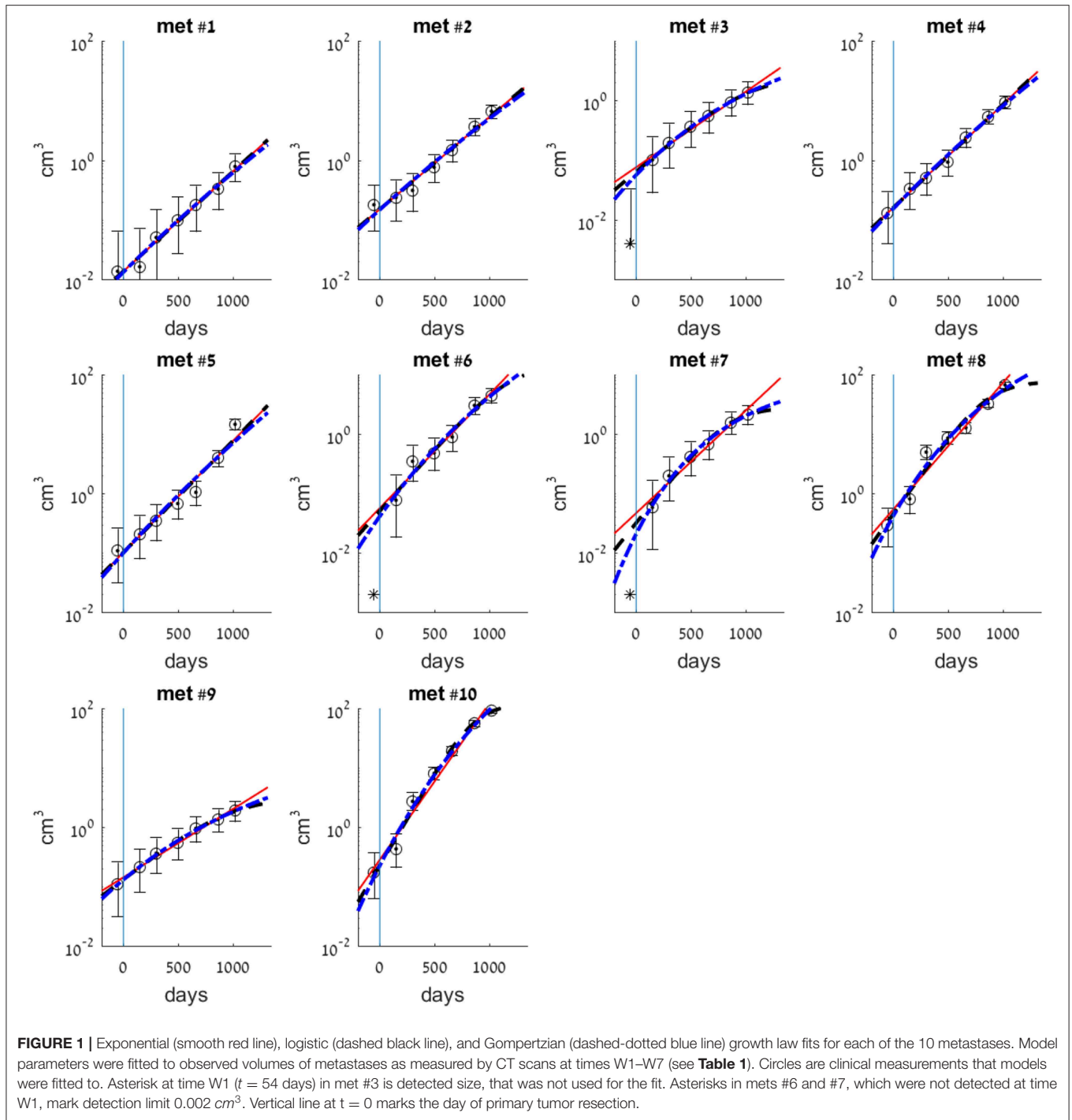
For metastases #3, #6, and #7, the fitted models imply that they have reached detection limit 3–4 years before primary detection time (see **Figure 3A**), and were far larger than this threshold at the day of disease detection (see smooth lines compared to the asterisks in corresponding subplots in **Figure 1**). This result stands in contrast with the fact that they were not observed at the CT done on primary tumor detection (timepoint W_1). We assume that these metastases were either undetectable because their size was below detection limit, or small enough to be missed at the scan, i.e., their size was above detection limit but close to it. Either way, for these metastases it seems that growth law was *not* the same all the time; it was dramatically changed in the 6 months between W_1 and

W_2 , the beginning of the period for which exponential curve was well-fitted. The enhanced growth rate for these metastases between W_1 and W_2 was evaluated by assuming exponential growth and fitting it to these two timepoints (taking maximal possible metastasis size at W_1 , when it was not detectable, as 0.002 cm^3). Results showed its minimal possible value was 5.85, 6.55, and 6.05 years^{-1} for metastases #3, #6, and #7, respectively. This is at least four times higher than the growth rate at the following period of time, between W_2 and W_7 (**Table 2**). For all other metastases, the exponential curve is well-fitted to the measure at W_1 , which means that the exponential growth rate remained the same.

DISCUSSION

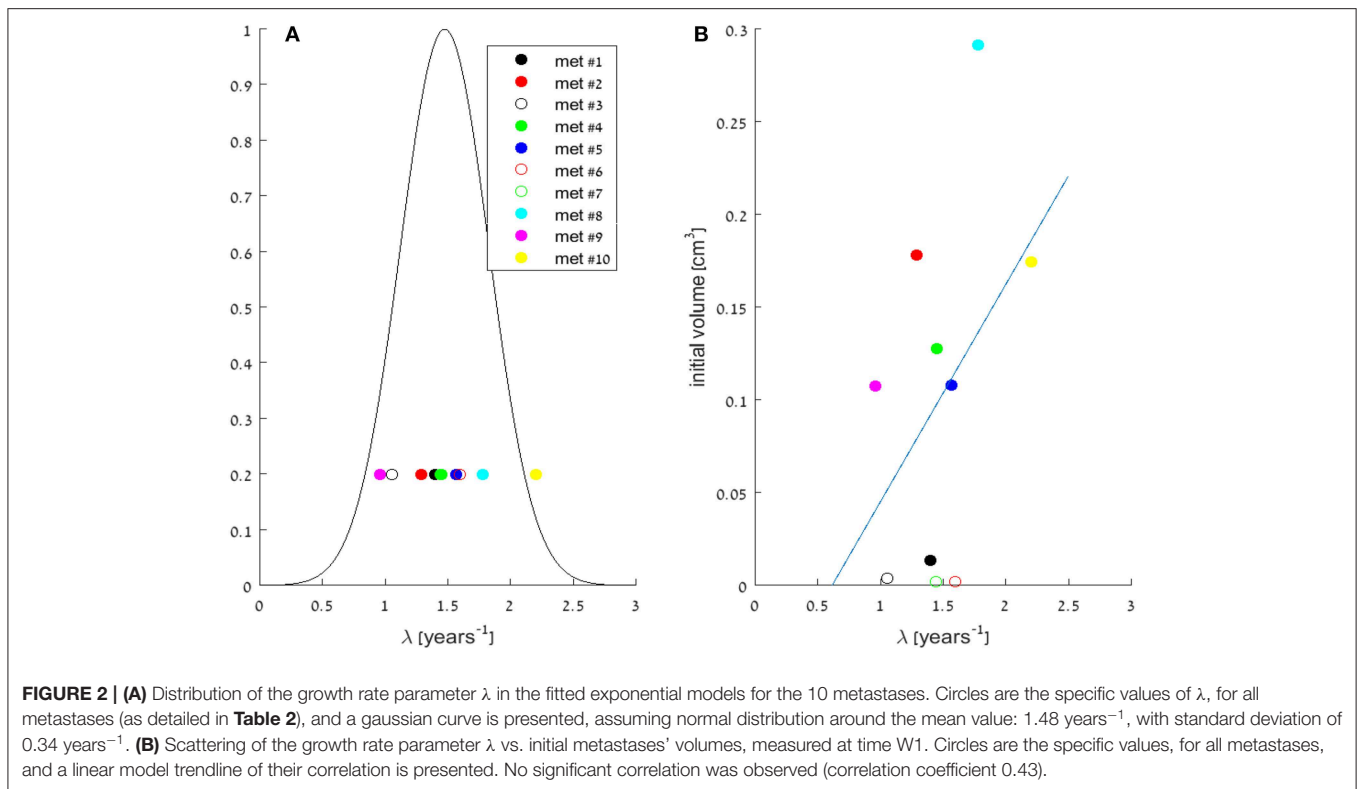
Understanding metastases growth is crucial for treating cancer patients. However, little is known about the dynamics of untreated metastases, because such dynamical data in humans are rare. In this paper, we used rare data of a metastatic CRC patient, for which CT measurements of growth of 10 untreated lung metastases during 3 years are available. We aimed to examine the common hypothesis that metastases growth rate can be approximated as exponential. Our results showed that all the metastases could be regarded as growing exponentially, at least for the first 2 years after disease detection and primary tumor resection. Logistic and Gompertzian growth models were also examined, but in most cases, they are overfitted and could not be used.

In addition, we found evidence that the exponential growth curve does not always demonstrate the closest prediction to actual growth measurements throughout all the follow-up time period. That was true for the first time period after primary tumor resection. Our results imply that some of the metastases (#3, #6, and #7) grew more rapidly between the time of first diagnosis and shortly after primary tumor surgery, while during the period of the next 2 years the growth rate was exponential with a constant, slower rate. This result is supported by literature describing implications of surgery on metastases growth rate,



especially in the short term. There is emerging evidence that the stress response caused by surgery as well as anesthesia and analgesia may promote growth of pre-existing micro-metastasis [5, 15–18]. Such post-surgery metastatic acceleration (PSMA) might be related to surgical stress through several mechanisms, such as suppression of anti-tumor immune response, stimulatory effects on tumor cells, and activation of the coagulation system [5, 16–19]. There are mathematical models that assume PSMA

is caused by removal of the suppression that the primary tumor induces on metastases, through systemic inhibition of angiogenesis [15, 20, 21]. This mechanism may explain changes in growth rate between the time before primary resection and the time after it (which was not examined in this work), but it does not explain the change in the growth rate during the period after resection, i.e., in the first month after resection compared to the next following years. Our results show that the most



significant impact of surgery is in the short term, at least for some of the metastases. It implies that other mechanisms, which decay shortly after surgery, for example—increased angiogenesis factors production due to wound healing [22], are dominant and should be investigated.

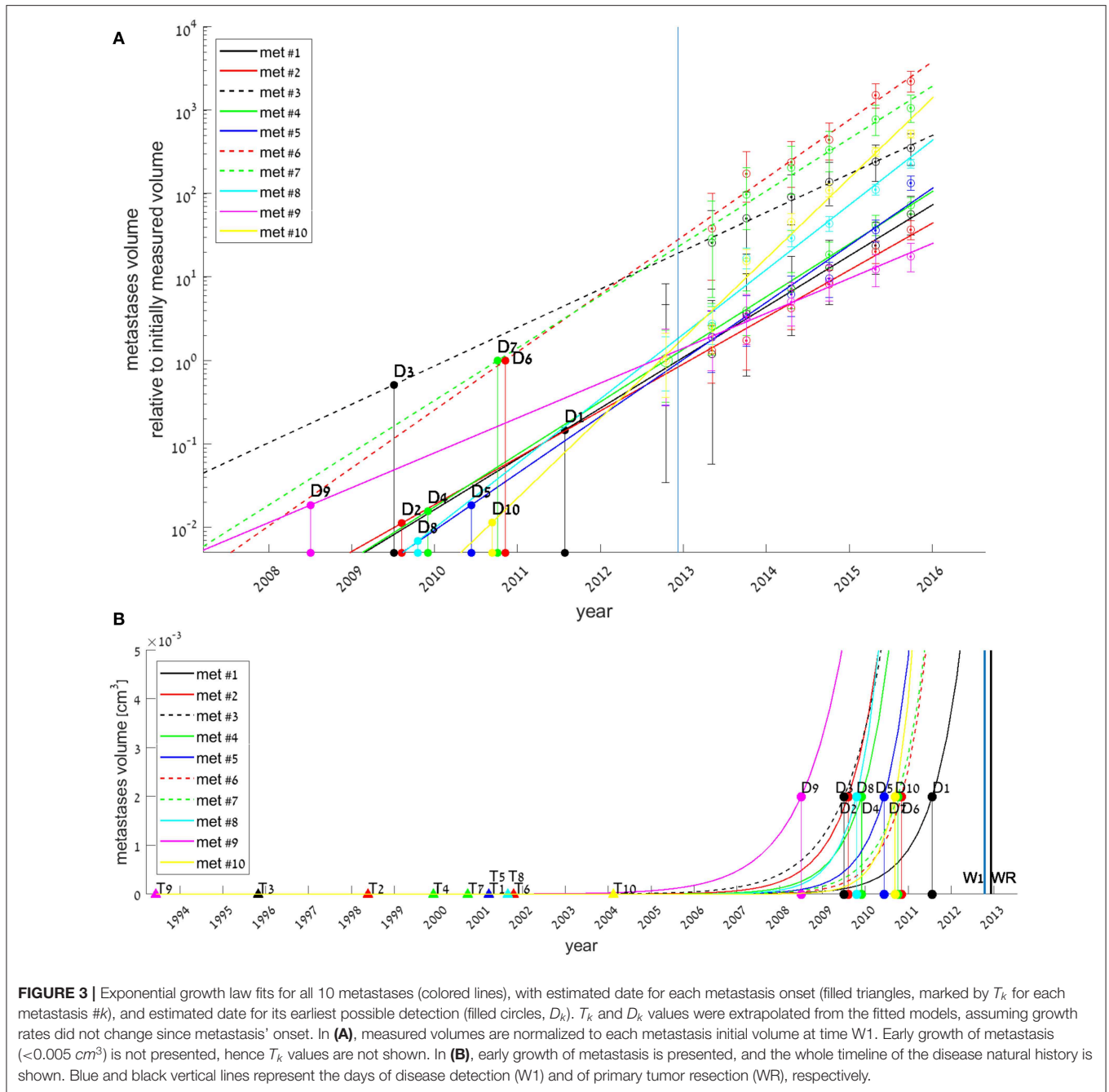
For the period of 6 months up to 2–3 years after surgery, all the metastases could be modeled as exponentially growing. The value of the growth rate parameter λ (see Equation 1) is quite similar for all of them, with a mean value of 1.5 years⁻¹. Based on these results, we can assume that exponential growth assumption is legitimate for most of metastases. Moreover, it is reasonable to assume a single value for the growth rate parameter that would fit all metastases found at the same site, as done in some mathematical models [11, 23]. No relation was found between fitted growth laws or rates to the metastases sizes or locations within the lungs.

A timeline of disease progression was constructed and estimated that onset of metastases occurred 8–19 years before primary tumor was detected (**Figure 3B**), and that they grew slowly and became detectable several years later. This was done by backwards extrapolation of the exponential fitted curves, assuming that growth rates before and after resection were the same. However, if we assume that the growth rate was faster—or at least not slowed—after surgery, then the real inception time was *no later* (and possibly earlier) than at the estimated times shown in **Figure 3B**. This result reinforces the notion that metastases were formed many years before detection of primary tumor [1, 5, 12]. The growth dynamics of metastases before primary tumor resection may be further investigated

by applying a natural history model on this patient's data, developed by Hanin et al. [23]. As validated here, we can use this model under the assumption of exponential growth after resection, with one value for the growth rate parameter for all metastases.

It is well-known that great variability exists between different primary tumors, between different patients with the same primary tumor and even between metastases at different locations in the same patient [24–26]. The main limitation of this work is that it is based on data of a single patient with rectal cancer metastatic to lungs. Different growth patterns might apply to other sites of metastases or to other primary tumors. Also, all measurements are prone to minute deviation errors especially when millimetric lesions are measured on a bidimensional CT scan. It would be interesting to analyze in the same way data of other patients with measurable metastases either in the lungs or other sites, either from CRC or other primary tumors. Another limitation is the fact that metastases were measured along 3 years period only. The dynamics of metastases growth before first diagnosis and more than 3 years after primary tumor resection are lacking. Hence other factors that could influence growth patterns in time are beyond the scope of this case.

In summary, our unique and uncommon data provide firm evidence that exponential growth model demonstrated precise prediction to actual growth measurements of CRC lung metastases, at least for a limited time period, starting half a year after surgery until about 2 years afterwards. In addition, the results imply that growth rate of some metastases might



accelerate shortly after primary tumor surgery, getting more moderated later. These results encourage further research of the suggested mechanisms for metastases growth acceleration caused by short-term effects of surgery, and of the effects of adjuvant treatment in this period of time.

DATA AVAILABILITY STATEMENT

The datasets generated for this study are available on request to the corresponding author.

ETHICS STATEMENT

Written informed consent was not obtained from the individual(s) for the publication of any potentially identifiable images or data included in this article.

AUTHOR CONTRIBUTIONS

SB-M designed research. GH and TH performed research. SB-M and GH contributed analytic tools. SB-M, GH, and ES-S analyzed data. SR measured metastases. ES-S and GH wrote the paper.

SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: <https://www.frontiersin.org/articles/10.3389/fams.2019.00056/full#supplementary-material>

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Conflict of Interest: The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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