



Sleep Disturbances Are Associated With Depressive Symptoms in a Chinese Population: The Rugao Longevity and Aging Cohort

Chunhong Qiao^{1,2}, Hui Zhang^{2,3}, Qi Song³, Xi Wang³, Xiaofeng Wang^{2,3*} and Yin Yao^{1,2*}

¹ Department of Biostatistics and Computational Biology, School of Life Sciences, Fudan University, Shanghai, China,

² Human Phenome Institute, Fudan University, Shanghai, China, ³ State Key Laboratory of Genetic Engineering and Ministry of Education (MOE) Key Laboratory of Contemporary Anthropology, School of Life Sciences, Fudan University, Shanghai, China

Objective: To investigate the cross-sectional and longitudinal relationships between sleep disturbances and depressive symptoms in older Chinese adults.

Methods: This study included baseline and 3.5-year follow-up data of 1,631 Chinese men and women aged 70 years or older from the aging arm of the Rugao Longevity and Aging Study. Depressive symptoms were assessed by the 15-item Geriatric Depression Scale (GDS). Sleep disturbances were assessed by using the Pittsburgh Sleep Quality Index (PSQI). Logistic regression models were used to estimate the odds ratios (ORs) of the associations.

Results: In the cross-sectional analysis, individuals with greater total PSQI scores exhibited significantly higher risk of “depressive symptoms” (OR: 1.31, 95% CI: 1.21–1.41) and “some depressive symptoms” (OR: 1.22, 95% CI: 1.17–1.28). Specifically, higher scores on the sleep efficiency PSQI subscale were associated with greater odds for “depressive symptoms” (OR: 1.54, 95% CI: 1.30–1.84) and “some depressive symptoms” (OR: 1.42, 95% CI: 1.29–1.57). Our longitudinal analyses indicated an association between greater PSQI total scores at baseline and greater odds of having “some depressive symptoms” at follow-up (OR: 1.07, 95% CI: 1.00–1.14). Additionally, higher scores on the sleep efficiency PSQI subscale had an association with higher odds for “some depressive symptoms” (OR: 1.21, 95% CI: 1.04–1.41).

Conclusions: Poor self-reported global sleep quality and sleep efficiency PSQI subscale scores were associated with levels of depressive symptoms in an older Chinese population, indicating that global sleep quality and sleep efficiency may be risk factors for depression and can possibly predict the levels of depressive symptoms.

Keywords: sleep disturbance, depressive symptoms, Chinese older population, risk factor, sleep efficiency

INTRODUCTION

Depression among the elderly population, with an estimated prevalence of 8–16%, is a major public health problem that has attracted worldwide attention (1, 2). Depression presents a heavy disease burden of long-term care on families and society as a whole (3, 4). Therefore, the prevention and treatment of depression have become urgent tasks in the field of public health. Depression has been

OPEN ACCESS

Edited by:

Shuqin Zhan,
Xuanwu Hospital, Capital Medical
University, China

Reviewed by:

Sanford Auerbach,
Boston University, United States
Barbara Galland,
University of Otago, New Zealand

*Correspondence:

Xiaofeng Wang
xiaofengwang71@163.com
Yin Yao
yin_yao@fudan.edu.cn

Specialty section:

This article was submitted to
Sleep Disorders,
a section of the journal
Frontiers in Psychiatry

Received: 27 June 2021

Accepted: 27 October 2021

Published: 08 December 2021

Citation:

Qiao C, Zhang H, Song Q, Wang X,
Wang X and Yao Y (2021) Sleep
Disturbances Are Associated With
Depressive Symptoms in a Chinese
Population: The Rugao Longevity and
Aging Cohort.
Front. Psychiatry 12:731371.
doi: 10.3389/fpsy.2021.731371

recognized as being associated with genetic (5), physical, behavioral, and socioeconomic factors (6). A potential effort to reduce depression levels has been targeting sleep disturbances.

Two previous longitudinal studies revealed that poor self-reported sleep quality was associated with an increased risk of depression. One analysis was conducted in the context of the Study of Osteoporotic Fractures (SOF), and the other was conducted within the prospective Osteoporotic Fractures in Men (MrOS) study. Both studies utilized the Geriatric Depression Scale (GDS) to measure depressive symptoms. Maglione et al. reported that baseline sleep disturbances were associated with a greater chance of worse depressive symptoms 5 years later (7). Paudel et al. revealed that among non-depressed older men, poor self-reported sleep quality was also associated with increased odds of depression 3.4 years later (8).

Before sleep disturbance can be established as a risk factor for depression, more evidence needs to be accumulated in different ethnic and age groups within older populations. In this study, we aimed to explore the relationship between sleep disturbances and depression symptoms at baseline in 1,631 Chinese adult participants aged 70 years or older. We also reassessed depressive symptoms after a 3.5-year follow up, further probing into the question of whether sleep disturbances or its subcomponents at baseline could potentially predict future depressive symptoms.

METHODS

Participants

The data came from the aging arm of the Rugao Longevity and Aging Study (RuLAS), a population-based, observational,

two-arm cohort study conducted in Rugao, Jiangsu Province, China. Approximately 1960 older adults were recruited based on 5-year age and sex strata, equally among 31 villages, in 2014. Our first follow-up was conducted after 1.5 years in the summer of 2016, and the second follow-up was conducted in the winter of 2017 (3 years after baseline). The third follow-up was conducted in the winter of 2019 (5 years after baseline) (9, 10). The current study focused on participants in the second and fourth waves (~1.5 and 5 years after the original assessment) of the RuLAS. In this study, the second wave was recognized as the baseline because data from the Pittsburgh Sleep Quality Index (PSQI) questionnaire were collected starting with the second wave. In the cross-sectional analysis, a total of 1,631 participants were included after excluding participants who appeared to suffer from major diseases (stroke, myocardial infarction, and cancer). Out of 1,631 individuals, 1,279 participants returned and completed both GDS and PSQI questionnaires at baseline. After excluding 459 participants who reported “some depressive symptoms” (GDS 3–5) or “depressive symptoms” (GDS ≥ 6), the remaining 820 participants had reported few depressive symptoms (GDS 0–2). Of these, including older adults with few depressive symptoms, 679 completed GDS questionnaires at the 3-year follow-up. Our longitudinal analyses were conducted on this subset of 679 participants. A schematic of the inclusion and exclusion of older Chinese adults in the longitudinal analysis is shown in **Figure 1**. The Human Ethics Committee of the School of Life Sciences of Fudan University, Shanghai, China, approved this study (No: BE1815). Written informed consent was obtained from all participants prior to the study.

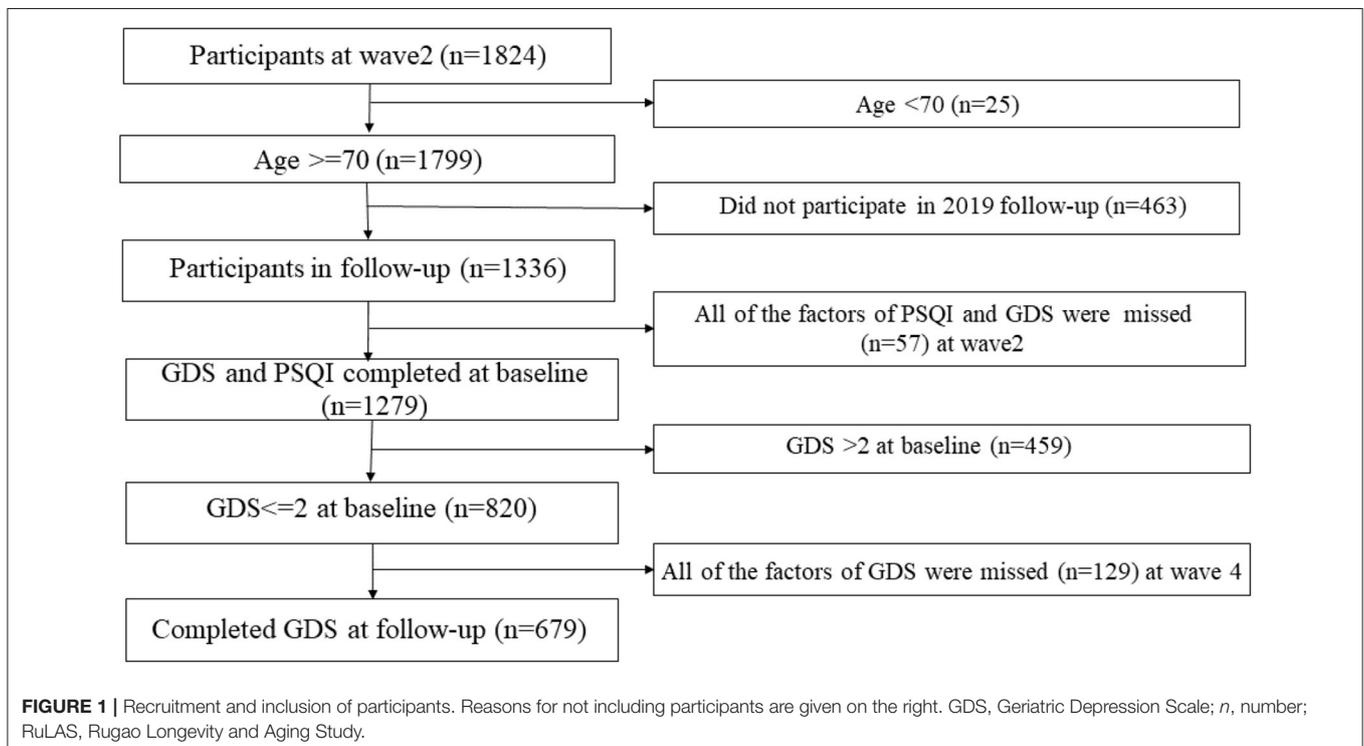


TABLE 1 | Characteristics of participants according to the level of depressive symptoms at baseline.

Characteristics	All (n = 1,631)	GDS ≥ 6 (n = 111)	3 ≤ GDS ≤ 5 (n = 489)	0 ≤ GDS ≤ 2 (n = 1,031)	P-value
Age group, n (%)					0.372
70–74	481 (29.49)	41 (36.94)	128 (26.18)	312 (30.03)	
75–79	687 (42.12)	41 (36.94)	200 (40.90)	446 (42.93)	
80–84	368 (22.56)	18 (16.22)	128 (26.18)	222 (21.37)	
85+	95 (5.82)	11 (9.91)	33(6.75)	51(4.91)	
Sex					0.093
Male	787 (48.25)	58 (52.25)	217 (44.38)	512 (49.28)	
Female	839 (51.44)	53 (47.75)	272 (55.62)	514 (49.47)	
Education					0.597
Illiterate	767 (47.03)	56 (50.45)	258 (52.76)	515 (49.57)	
Literate	829 (50.83)	55 (49.55)	221 (45.19)	491 (47.26)	
Marital status					0.396
Current marital	1,041 (63.83)	70 (63.06)	305 (62.37)	666 (64.10)	
Other	547 (33.54)	41 (36.94)	175 (35.79)	331 (31.86)	
Smoking					<0.001
None	1,425 (87.37)	76 (68.47)	422 (86.30)	902 (86.81)	
Smoker	132 (8.09)	32 (28.83)	46 (9.41)	79 (7.60)	
Drinking					<0.001
None	1,412 (86.57)	62 (55.86)	425 (86.91)	888 (85.47)	
Drinker	134 (8.22)	45 (40.54)	41 (8.38)	85 (8.18)	
BMI category					0.067
<24	907 (55.61)	62 (55.86)	259 (52.97)	586 (56.40)	
24–28	533 (32.68)	43 (38.74)	173 (35.38)	317 (30.51)	
≥28	175 (10.73)	5 (4.50)	53 (10.84)	117 (11.26)	
PSQI score					
≤5	952 (58.37)	43 (38.74)	194 (39.67)	715 (69.35)	
>5	679 (41.63)	68 (61.26)	295 (60.33)	316 (30.65)	

Others in marital status means widowed, divorced, and unmarried. GDS, Geriatric Depression Scale; PSQI, Pittsburgh Sleep Quality Index; SD, standard deviation; BMI, body mass index; Bold means statistically significant.

Depressive Symptoms

Depressive symptoms were assessed using the Chinese version of the 15-item GDS (11). This form of the GDS consisted of 15 self-reported yes-or-no questions that were derived from the GDS-30. The GDS-15 scale has previously been validated for use in community-living older Asian adults (12). A previous review reported that the best performance for the GDS was with a cutoff of 5/6 for the GDS-15 (13). A standard cutoff of ≥ 6 on the GDS revealed 91% sensitivity and 65% specificity when evaluated against diagnostic criteria (14). In this study, we implemented the validated cutoff of 6 to capture no/some and high depressive symptoms. Depressive symptoms were categorized into three groups, based on clinical relevance [0–2 (normal), 3–5 (some depressive symptoms) and ≥ 6 (depressed)] (7). For simplicity and consistency, patients with particular levels of depressive symptoms at follow-up were divided into three groups [0–2 (normal), 3–5 (some depressive symptoms) and ≥ 6 (depressed)]. The changes in GDS score between baseline and follow-up were calculated for each study participant. The participants with an increased GDS score ≥ 2 points were considered to have “worsening depressive symptoms.” This threshold was chosen

based upon the distribution of changes in GDS scores in our samples because the GDS scores used in this study were able to distinguish older Chinese adults with the greatest increase in GDS score (highest quartile) between baseline and 3-year follow-up (7).

Sleep Measures

Global sleep quality was assessed using the PSQI (15). The PSQI is a validated 19-item self-reported measure of sleep disturbances. It has been validated and demonstrated to have good psychometric properties in an ethnically similar population (16). The Chinese version of the PSQI has good overall reliability ($r = 0.82–0.83$) and test-retest reliability ($r = 0.77–0.85$) in community of adults with primary insomnia (16, 17). The PSQI had internal consistency and a reliability coefficient (Cronbach's alpha) of 0.703 for its seven components. The PSQI was divided into seven subcomponent scores: sleep quality, sleep latency, sleep duration, sleep efficiency, sleep disturbance, sleep medication use, and daytime dysfunction. Each subcomponent score ranges from 0 to 3, and global PSQI scores range from 0 to 21, with higher scores reflecting more severe symptoms. Total PSQI scores are

TABLE 2 | Associations between sleep disturbance and depressive symptom levels at baseline.

Sleep variables	Some depressive symptoms OR (95% CI)	P-value	Depressive symptoms OR (95% CI)	P-value
PSQI Total Score				
Base model	1.22 (1.17–1.27)	<0.001	1.30 (1.21–1.40)	<0.001
Multivariable adjusted	1.22 (1.17–1.28)	<0.001	1.31 (1.21–1.41)	<0.001
PSQI >5 PSQI ≤5				
Base model	3.44 (2.75–4.31)	<0.001	3.58 (2.40–5.40)	<0.001
Multivariable adjusted	3.56 (2.77–4.58)	<0.001	3.81 (2.45–6.02)	<0.001
Sleep quality factor				
Base model	1.37 (1.15–1.64)	<0.001	1.69 (1.23–2.33)	0.001
Multivariable adjusted	1.37 (1.10–1.70)	0.004	1.61 (1.10–2.34)	0.013
Sleep latency factor				
Base model	1.40 (1.24–1.59)	<0.001	1.69 (1.35–2.10)	<0.001
Multivariable adjusted	1.35 (1.18–1.54)	<0.001	1.67 (1.33–2.10)	<0.001
Sleep duration factor				
Base model	1.60 (1.35–1.89)	<0.001	1.91 (1.47–2.47)	<0.001
Multivariable adjusted	1.65 (1.38–1.97)	<0.001	1.96 (1.48–2.59)	<0.001
Sleep efficiency factor				
Base model	1.43 (1.31–1.57)	<0.001	1.55 (1.32–1.82)	<0.001
Multivariable adjusted	1.42 (1.29–1.57)	<0.001	1.54 (1.30–1.84)	<0.001
Sleep disturbances factor				
Base model	1.56 (1.24–1.97)	<0.001	1.76 (1.16–2.62)	0.007
Multivariable adjusted	1.46 (1.14–1.87)	0.003	1.61 (1.03–2.49)	0.034
Sleep medication use factor				
Base model	1.95 (1.16–3.65)	0.018	1.01 (0.07–3.24)	0.995
Multivariable adjusted	2.23 (1.16–5.19)	0.029	1.36 (0.09–6.45)	0.74
Daytime dysfunction factor				
Base model	1.35 (1.19–1.53)	<0.001	1.63 (1.33–2.00)	<0.001
Multivariable adjusted	1.33 (1.16–1.52)	<0.001	1.62 (1.31–2.01)	<0.001

Odds ratios and 95% confidence intervals for falling into the “some depressive symptoms” or “depressed” groups are given according to baseline subjective sleep measures. GDS, Geriatric Depression Scale; PSQI, Pittsburgh Sleep Quality Index; CI, confidence interval. The base model p-value adjusted nothing, and the multivariable adjusted p-value adjusted for age, sex, education, marriage, drinking, smoking, and self-reported health status. Bold means statistically significant.

expressed as a continuous variable and as a categorical variable: PSQI > 5 vs. PSQI ≤ 5 (18). Questionnaires were administered by trained physicians. The trained physicians were present during the entire duration of the assessment to provide clarifications if participants had trouble understanding any of the items in the questionnaires. These measures were verbally administered by these physicians in cases where the participants were illiterate.

Covariables

Demographic information included age, sex (male or female), marital status (currently married, others), and education (illiterate, literate). Lifestyle included smoking status (non-smoker, smoker), drinking status (non-drinker, drinker), body mass index (BMI) [normal or underweight (<24.0), overweight (24.0–27.9), and obese (≥28.0)] and self-reported health status.

Statistical Analyses

Continuous variables are reported as the mean ± standard deviation (SD), and categorical variables are reported as percentages. The relationship between sleep disturbance and depressive symptoms were assessed using logistic regression models. The differences in the characteristics of the participants

at baseline and at the 3.5-year follow-up based on the level of depressive symptoms were also assessed using logistic regression models. Covariates known to be associated with levels of depressive symptoms at follow-up or with sleep disturbances were included in multivariable models. Logistic regression models were first used to estimate the odds ratio (ORs) for falling into different depressive symptom level categories (i.e., “some depressive symptoms” (GDS 3–5) or “depressed” (GDS ≥ 6) at follow-up). Logistic regression models were then used to estimate the OR for having a two-point or greater increase in GDS score at follow-up. Models were adjusted for age, sex, education, marriage status, smoking status, drinking status and self-reported health status for multiple variables. P-values < 0.05 were considered statistically significant. All analyses were conducted using R x64 4.0.2 (“<https://www.r-project.org/>”).

RESULTS

Characteristics of the Participants

Table 1 describes the baseline characteristics of the study population categorized into different depressive symptom

TABLE 3 | Associations between sleep disturbance at baseline and depressive symptom level at the 3-year follow-up.

Sleep variables	Some depressive symptoms OR (95% CI)	P-value	Depressed symptoms OR (95% CI)	P-value
PSQI total score				
Base model	1.09 (1.03–1.15)	0.003	1.09 (0.98–1.21)	0.115
Multivariable adjusted	1.07 (1.00–1.14)	0.036	1.03 (0.97–1.16)	0.669
PSQI >5 PSQI ≤5				
Base model	1.44 (1.03–2.02)	0.032	2.03 (1.09–3.82)	0.026
Multivariable adjusted	1.36 (0.94–1.96)	0.104	1.84 (0.94–3.66)	0.077
Sleep quality factor				
Base model	1.18 (0.88–1.56)	0.266	1.22 (0.71–2.09)	0.467
Multivariable adjusted	1.10 (0.76–1.59)	0.601	0.71 (0.35–1.40)	0.337
Sleep latency factor				
Base model	1.25 (1.03–1.51)	0.022	1.31 (0.91–1.85)	0.132
Multivariable adjusted	1.15 (0.94–1.41)	0.173	1.32 (0.89–1.94)	0.16
Sleep duration factor				
Base model	1.16 (0.91–1.46)	0.221	1.05 (0.65–1.62)	0.822
Multivariable adjusted	1.10 (0.86–1.40)	0.455	0.90 (0.52–1.48)	0.705
Sleep efficiency factor				
Base model	1.25 (1.09–1.44)	0.002	1.34 (1.04–1.72)	0.024
Multivariable adjusted	1.21 (1.04–1.41)	0.014	1.21 (0.91–1.61)	0.173
Sleep disturbances factor				
Base model	1.44 (1.00–2.07)	0.0496	1.03 (0.48–2.08)	0.939
Multivariable adjusted	1.29 (0.88–1.89)	0.186	0.76 (0.33–1.65)	0.5
Sleep medication use factor				
Base model	1.58 (0.69–4.35)	0.286	1.46 (0.12–5.52)	0.605
Multivariable adjusted	1.14 (0.43–3.21)	0.776	1.27 (0.11–5.16)	0.741
Daytime dysfunction factor				
Base model	1.08 (0.88–1.31)	0.471	0.98 (0.64–1.43)	0.931
Multivariable adjusted	1.06 (0.86–1.32)	0.575	0.85 (0.52–1.31)	0.497

GDS, Geriatric Depression Scale; PSQI, Pittsburgh Sleep Quality Index; CI, confidence interval. The base model was not adjusted, and the multivariable model was adjusted for age, sex, education, marriage, drinking, smoking, and self-reported health status. Bold means statistically significant.

groups. The mean age was 77.2 ± 4.12 years; 51.44% ($n = 839$) were females; 47.03% ($n = 767$) were illiterate; and 63.83% ($n = 1041$) were currently married. A total of 6.81% ($n = 111$) of the participants had depressive symptoms, and 16.19% ($n = 264$) had sleep disturbance at baseline. Smoking status and drinking status were associated with “some depressive symptoms” and “no or few depressive symptoms.” There was no significant difference among the groups with “depressive symptoms,” “some depressive symptoms,” and “no or few depressive symptoms” in age, sex, marital status, literacy, or BMI. Of the participants without evidence of depression at baseline, 31.66% exhibited “some depressive symptoms” and 6.48% revealed “depressive symptoms” at the 3.5-year follow-up examination.

Associations Between Sleep Disturbances and Depressive Symptoms at Baseline

Table 2 indicates that poor global sleep quality (PSQI > 5) was significantly associated with the follow up “some depressive symptoms” (OR: 3.44, 95% CI: 2.75–4.31) and “depressive symptoms” (OR: 3.58, 95% CI: 2.40–5.40) after adjusting for confounding factors. PSQI subcomponents, including sleep quality (OR: 1.61, 95% CI: 1.10–2.34), sleep latency (OR:

1.67, 95% CI: 1.33–2.10), sleep duration (OR: 1.96, 95% CI: 1.48–2.59), sleep efficiency (OR: 1.54, 95% CI: 1.30–1.84), sleep disturbance (OR: 1.61, 95% CI: 1.03–2.49), and daytime dysfunction (OR: 1.62, 95% CI: 1.31–2.01), were also associated with “depressive symptoms” after adjusting for confounding factors.

Associations Between Baseline Sleep Disturbances and Depressive Symptoms at Follow-Up

Table 3 indicates that global sleep quality was associated with “some depressive symptoms” (OR: 1.09, 95% CI: 1.03–1.15) at follow-up, and this association remained significant after adjusting for confounding factors. Poor global sleep quality (PSQI > 5) was associated with “some depressive symptoms” (OR: 1.44, 95% CI: 1.00–2.02), and no significant association was detected after adjusting for confounding factors. Interestingly, the PSQI subcomponents sleep latency (OR: 1.25, 95% CI: 1.03–1.51), sleep efficiency (OR: 1.25, 95% CI: 1.09–1.44), and sleep disturbance (OR: 1.25, 95% CI: 1.00–2.07) were associated with “some depressive symptoms.” After adjustment, only sleep efficiency retained an association with “some depressive

TABLE 4 | Associations between baseline sleep disturbance and odds of increased depressive symptoms at the 3-year follow-up.

Sleep factors	≥2 Point increase in GDS OR (95% CI)	P-value
PSQI Total Score		
Base model	1.05 (0.99–1.10)	0.086
Multivariable adjusted	1.03 (0.97–1.08)	0.359
PSQI >5 PSQI ≤5		
Base model	1.30 (0.95–1.79)	0.102
Multivariable adjusted	1.14 (0.82–1.59)	0.344
Sleep quality factor		
Base model	1.14 (0.88–1.49)	0.313
Multivariable adjusted	1.12 (0.81–1.57)	0.487
Sleep latency factor		
Base model	0.96 (0.80–1.15)	0.683
Multivariable adjusted	0.95 (0.78–1.14)	0.563
Sleep duration factor		
Base model	1.09 (0.87–1.36)	0.445
Multivariable adjusted	1.03 (0.82–1.31)	0.774
Sleep efficiency factor		
Base model	1.09 (0.95–1.24)	0.216
Multivariable adjusted	1.09 (0.94–1.25)	0.254
Sleep disturbances factor		
Base model	0.98 (0.69–1.39)	0.927
Multivariable adjusted	0.97 (0.67–1.39)	0.857
Sleep medication use factor		
Base model	1.33 (0.58–3.32)	0.482
Multivariable adjusted	1.13 (0.41–2.84)	0.788
Daytime dysfunction factor		
Base model	1.12 (0.93–1.35)	0.234
Multivariable adjusted	1.10 (0.90–1.35)	0.358

Odds ratios and 95% confidence intervals for worsening depressive symptoms (≥2-point increase in GDS score) are given for older adults with sleep disturbances at baseline compared to those without sleep disturbances. Base models were crude models without adjusting any variables. Multivariable models were adjusted for age, sex, education, marriage status, smoking status, drinking status and self-reported health status. CI, confidence interval; GDS, Geriatric Depression Scale; PSQI, Pittsburgh Sleep Quality Index.

symptoms” (OR: 1.21, 95% CI: 1.04–1.41). The category defined by PSQI score > 5 was associated with “depressive symptoms” (OR: 2.03, 95% CI: 1.09–3.82), while no association was revealed after adjusting confounding factors. The PSQI subcomponent sleep efficiency had an association with “depressive symptoms” (OR: 1.34, 95% CI: 1.04–1.72). Interestingly, there was a lack of association between sleep efficiency and depressive symptoms after adjusting for confounding factors (OR: 1.21, 95% CI: 0.91–1.61). The PSQI subcomponents sleep quality, sleep duration, sleep medication use, and daytime dysfunction had no association with “some depressive symptoms” and “depressive symptoms.”

Associations Between Sleep Disturbances and Worsening of Depressive Symptoms at Follow-Up

Table 4 shows that ~38.7% ($n = 263$) of older Chinese adults showed a ≥ 2-point increase in depressive symptoms. However,

we observed no significant association between baseline PSQI total score or PSQI subcomponent scores and worsening of depressive symptoms at follow-up.

DISCUSSION

In this longitudinal analysis of Chinese community-dwelling older people with few or no depressive symptoms at baseline, poorer global sleep quality, and sleep efficiency there appeared to be risk factors for “some depressive symptoms” but not increased risk of worsening of depressive symptoms at follow-up (~3.5 years later). In the cross-sectional analysis, we observed that the level of depressive symptoms was associated with sleep disturbances (self-reported poor global sleep quality and sleep efficiency). Overall, the cross-sectional association between sleep disturbances and depressive symptoms were attenuated in our longitudinal analysis. To our knowledge, we are the first to investigate whether baseline sleep disturbances could increase the risk of worsening depressive symptoms at follow-up in an older Chinese population. More importantly, we found for the first time that those with more reported sleep disturbance at baseline had greater odds of developing “some depressive symptoms.”

One cross-sectional study reported that sleep quality was a risk factor for depression, and the ORs of “depressive symptoms” and “some depressive symptoms” were 3.7- and 2.1-fold, respectively, for those with sleep disturbances at baseline (19). Another study also found that women with “some depressive symptoms” and “depressive symptoms” had greater odds of reporting poor sleep (20). In our cross-sectional analysis, we observed that the ORs for “depressive symptoms” and “some depressive symptoms” were 3.56- and 3.81-fold, respectively, for those with sleep disturbances at baseline which were consistent with the above two studies. A longitudinal study in a large cohort of older men revealed associations between worse sleep quality at baseline and more depressive symptoms at follow-up ~3.4-years later (8). A previous longitudinal study in a large cohort reported that older women with few or no depressive symptoms at baseline who were reported to have sleep disturbances had a greater risk of worse depressive symptoms 5 years later (7). Interestingly, our longitudinal analysis indicated that among non-depressed older adults at baseline, poor self-reported global sleep quality was associated with “some depressive symptoms” but not the risk of worsening depressive symptoms. This discrepancy could have been due to the shorter follow-up period (~3.5 years) in the current study. Another possible reason is that the SOF included only female participants and the MrOS included only male participants. Hence, these findings require verification in more cohorts.

In our cross-sectional analysis, we observed that sleep efficiency and daytime sleep PSQI subcomponents were associated with depressive symptoms regardless of adjustment for confounding factors. In the longitudinal analysis, the sleep efficiency PSQI subcomponent at baseline was associated with “some depressive symptoms” at the 3.5-year follow-up after adjusting for confounding factors. This interesting observation indicated that in older adults without depression, poor sleep

efficiency may increase the risk of “some depressive symptoms.” On the other hand, poor sleep efficiency had no association with “depressive symptoms” at the 3.5-year follow-up after adjusting for confounding factors. The studies mentioned above (7, 8) did not find any relationship between sleep efficiency and depression. This discrepancy could be due to the differences in age, as the latter study consisted of women only, most of whom were older than 80 years, whereas the current study included men and women who were mostly younger than 80 years. A previous study reported that older adults show decreased sleep efficiency over time, with an 18.6% decline observed between 40 and 100 years of age (21). In addition, we found that the daytime sleep PSQI subcomponents were associated with depressive symptoms. This observation was consistent with previous population-based studies, which have examined the association between daytime sleep and depressive symptoms (22). We detected no significant association between PSQI scores at baseline and the odds of worsening depressive symptoms (≥ 2 point increase in GDS). The main difference could be due to the differences in population and sample size.

Marta Jackowska et al. reported that compared to an optimal duration, short (≤ 5 h) but not long (≥ 8 h) sleep hours were linked to elevated depressive symptoms (23). Sun et al. indicated that short sleep duration (< 5 h, 5–6 h) significantly impacted depressive symptoms, while long sleep duration (> 9 h) had no association with depressive symptoms (24). Lai et al. found that long sleep duration (≥ 9 h) was a risk factor for depression (22). Our study did not establish a relationship between sleep duration and the worsening of depressive symptoms. This discrepancy may be because the English Longitudinal Study of Aging (ELSA) included an English cohort, while the China Health and Retirement Longitudinal Study (CHARLS), Yilan Study in Taiwan (YILAN) and RuLAS included Chinese cohorts. Alternatively, the differences may have been influenced by time cutoffs, as Marta Jackowska took 7–8 h as a reference, Sun et al. took 7–8 h as a reference, and Lai et al. took 6–7 h as a reference. In addition, various studies have used different sleep and depression scales. The ELSA used the 8-item Centre for Epidemiological Studies Depression scale (CES-D), the CHARLS used the 10-item version of the Centre for Epidemiological Studies Depression scale (CESD-10), and the YILAN cohort used the Hospital Anxiety and Depression Scale (HADS). Hence, further investigation requires verification in other cohorts.

In summary, the current study assessed global sleep quality among older adults using the PSQI and validated the relationship between baseline sleep disturbances and follow-up depressive symptoms. However, there are several limitations of this analysis. First, our analysis was designed to make use of data that were collected as part of a large study. Hence, it was not designed to address our hypothesis, and the outcome measures were

not predefined. Second, depressive symptoms were assessed by questionnaire rather than standard criteria for depression such as those from the International Classification of Diseases (ICD) or Diagnostic and Statistical Manual of Mental Disorder (DSM). Therefore, conclusions about psychiatric diagnosis cannot be made with certainty. Third, factors such as personal medical issues, antidepressant use, and family stress that may impact both sleep and mental health outcomes were not accounted for in this study. Finally, the generalizability of the study was limited, and more studies are needed to validate these findings in other cohorts.

In conclusion, findings of the present study contribute to the current literature in terms of the relationship between sleep disturbance and depressive symptoms in an older Chinese population. Together with the observations in the aforementioned studies, poor global sleep quality and sleep efficiency may be risk factors for depression and can predict levels of depressive symptoms.

DATA AVAILABILITY STATEMENT

The original contributions presented in the study are included in the article/supplementary material, further inquiries can be directed to the corresponding author/s.

ETHICS STATEMENT

The Human Ethics Committee of the School of Life Sciences of Fudan University, Shanghai, China, approved this research (No: BE1815). Written consent was obtained from all participants prior to participation.

AUTHOR CONTRIBUTIONS

CQ: conceptualization, data analysis, and writing-original draft preparation. YY and XiaW: conceptualization and writing-reviewing and editing. HZ, QS, and XiW: data collection and data cleaning. All authors contributed to the article and approved the submitted version.

FUNDING

This work was supported by the Shanghai Municipal Science and Technology Major Project #1 under Grant No. 2017SHZDZX01 and National Key R&D Program of China #2 under Grant No. 2018YFC2000400 and 2018YFC2000400-3.

ACKNOWLEDGMENTS

We acknowledge all people involved in this study.

REFERENCES

- Jackowska M, Poole L. Sleep problems, short sleep and a combination of both increase the risk of depressive symptoms in older people: a 6-year follow-up investigation from the English longitudinal study of ageing. *Sleep Medicine*. (2017) 37:60–5. doi: 10.1016/j.sleep.2017.02.004
- Malhi GS, Mann JJ. Depression. *Lancet*. (2018) 392:2299–312. doi: 10.1016/S0140-6736(18)31948-2

3. Ouyang P, Sun W. Depression and sleep duration: findings from middle-aged and elderly people in China. *Public Health*. (2019) 166:148–54. doi: 10.1016/j.puhe.2018.10.007
4. Volkert J, Schulz H, Harter M, Włodarczyk O, Andreas S. The prevalence of mental disorders in older people in Western countries - a meta-analysis. *Ageing Res Rev*. (2013) 12:339–53. doi: 10.1016/j.arr.2012.09.004
5. Kendler KS, Gatz M, Gardner CO, Pedersen NL. A Swedish national twin study of lifetime major depression. *Am J Psychiatry*. (2006) 163:109–14. doi: 10.1176/appi.ajp.163.1.109
6. Cole MG, Dendukuri N. Risk factors for depression among elderly community subjects: a systematic review and meta-analysis. *Am J Psychiatry*. (2003) 160:1147–56. doi: 10.1176/appi.ajp.160.6.1147
7. Maglione JE, Ancoli-Israel S, Peters KW, Paudel ML, Yaffe K, Ensrud KE, et al. Subjective and objective sleep disturbance and longitudinal risk of depression in a cohort of older women. *Sleep*. (2014) 37:1179–87. doi: 10.5665/sleep.3834
8. Paudel M, Taylor BC, Ancoli-Israel S, Blackwell T, Maglione JE, Stone K, et al. Sleep disturbances and risk of depression in older men. *Sleep*. (2013) 36:1033–40. doi: 10.5665/sleep.2804
9. Liu ZY, Wang Y, Zhang YC, Chu XF, Wang ZD, Qian DG, et al. Cohort profile: the rugao longevity and ageing study (RuLAS). *Int J Epidemiol*. (2016) 45:1064–73. doi: 10.1093/ije/dyv101
10. Shi GP, Ma T, Zhu YS, Wang ZD, Chu XF, Wang Y, et al. Frailty phenotype, frailty index and risk of mortality in Chinese elderly population-Rugao longevity and ageing study. *Arch Gerontol Geriatr*. (2019) 80:115–9. doi: 10.1016/j.archger.2018.11.001
11. Yesavage JA. Geriatric depression scale. *Psychopharmacol Bull*. (1988) 24:709–11.
12. Nyunt MS, Fones C, Niti M, Ng TP. Criterion-based validity and reliability of the geriatric depression screening scale (GDS-15) in a large validation sample of community-living Asian older adults. *Ageing Ment Health*. (2009) 13:376–82. doi: 10.1080/13607860902861027
13. Dennis M, Kadri A, Coffey J. Depression in older people in the general hospital: a systematic review of screening instruments. *Age Ageing*. (2012) 41:148–54. doi: 10.1093/ageing/afr169
14. Almeida OP, Almeida SA. Short versions of the geriatric depression scale: a study of their validity for the diagnosis of a major depressive episode according to ICD-10 and DSM-IV. *Int J Geriatric Psychiatry*. (1999) 14:858–65. doi: 10.1002/(SICI)1099-1166(199910)14:10<858::AID-GPS35>3.0.CO;2-8
15. Buysse DJ, Reynolds, 3rd CF, Monk TH, Berman SR, Kupfer DJ. The pittsburgh sleep quality index: a new instrument for psychiatric practice and research. *Psychiatry Res*. (1989) 28:193–213. doi: 10.1016/0165-1781(89)90047-4
16. Tsai PS, Wang SY, Wang MY, Su CT, Yang TT, Huang CJ, et al. Psychometric evaluation of the Chinese version of the pittsburgh sleep quality index (CPSQI) in primary insomnia and control subjects. *Qual Life Res*. (2005) 14:1943–52. doi: 10.1007/s11136-005-4346-x
17. Ma T, Shi G, Zhu Y, Wang Y, Chu X, Jiang X, et al. Sleep disturbances and risk of falls in an old Chinese population-rugao longevity and ageing study. *Arch Gerontol Geriatr*. (2017) 73:8–14. doi: 10.1016/j.archger.2017.07.003
18. Hinze A, Glaesmer H, Brahler E, Löffler M, Engel C, Enzenbach C, et al. Sleep quality in the general population: psychometric properties of the pittsburgh sleep quality index, derived from a German community sample of 9284 people. *Sleep Med*. (2017) 30:57–63. doi: 10.1016/j.sleep.2016.03.008
19. Paudel ML, Taylor BC, Diem SJ, Stone KL, Ancoli-Israel S, Redline S, et al. Association between depressive symptoms and sleep disturbances in community-dwelling older men. *J Am Geriatr Soc*. (2008) 56:1228–35. doi: 10.1111/j.1532-5415.2008.01753.x
20. Maglione JE, Ancoli-Israel S, Peters KW, Paudel ML, Yaffe K, Ensrud KE, et al. Depressive symptoms and subjective and objective sleep in community-dwelling older women. *J Am Geriatr Soc*. (2012) 60:635–43. doi: 10.1111/j.1532-5415.2012.03908.x
21. Didikoglu A, Maharani A, Tampubolon G, Canal MM, Payton A, Pendleton N. Longitudinal sleep efficiency in the elderly and its association with health. *J Sleep Res*. (2020) 29:e12898. doi: 10.1111/jsr.12898
22. Lai HC, Hsu NW, Chou P, Chen HC. The associations between various sleep-wake disturbances and depression in community-dwelling older adults- the Yilan study, Taiwan. *Ageing Mental Health*. (2020) 24:717–724. doi: 10.1080/13607863.2019.1582006
23. Poole L, Jackowska M. The epidemiology of depressive symptoms and poor sleep: findings from the english longitudinal study of ageing (ELSA). *Int J Behav Med*. (2018) 25:151–61. doi: 10.1007/s12529-017-9703-y
24. Sun YK, Shi L, Bao YP, Sun Y, Shi J, Lu L. The bidirectional relationship between sleep duration and depression in community-dwelling middle-aged and elderly individuals: evidence from a longitudinal study. *Sleep Med*. (2018) 52:221–9. doi: 10.1016/j.sleep.2018.03.011

Conflict of Interest: The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

Publisher's Note: All claims expressed in this article are solely those of the authors and do not necessarily represent those of their affiliated organizations, or those of the publisher, the editors and the reviewers. Any product that may be evaluated in this article, or claim that may be made by its manufacturer, is not guaranteed or endorsed by the publisher.

Copyright © 2021 Qiao, Zhang, Song, Wang, Wang and Yao. This is an open-access article distributed under the terms of the Creative Commons Attribution License (CC BY). The use, distribution or reproduction in other forums is permitted, provided the original author(s) and the copyright owner(s) are credited and that the original publication in this journal is cited, in accordance with accepted academic practice. No use, distribution or reproduction is permitted which does not comply with these terms.