Check for updates

OPEN ACCESS

EDITED BY Giuseppe Murdaca, University of Genoa, Italy

REVIEWED BY

Koshy Nithin Thomas, Sanjay Gandhi Post Graduate Institute of Medical Sciences (SGPGI), India Masahiro Yoshikawa, Nihon University School of Medicine, Japan

*CORRESPONDENCE Rui Mao 218102100@csu.edu.cn

[†]These authors have contributed equally to this work

RECEIVED 06 April 2023 ACCEPTED 26 June 2023 PUBLISHED 12 July 2023

CITATION

Zhao N, Guo P, Tang M, Yang F, Zhang T and Mao R (2023) Evidence for a causal relationship between psoriasis and cutaneous melanoma: a bidirectional twosample Mendelian randomized study. *Front. Immunol.* 14:1201167. doi: 10.3389/fimmu.2023.1201167

COPYRIGHT

© 2023 Zhao, Guo, Tang, Yang, Zhang and Mao. This is an open-access article distributed under the terms of the Creative Commons Attribution License (CC BY). The use, distribution or reproduction in other forums is permitted, provided the original author(s) and the copyright owner(s) are credited and that the original publication in this journal is cited, in accordance with accepted academic practice. No use, distribution or reproduction is permitted which does not comply with these terms.

Evidence for a causal relationship between psoriasis and cutaneous melanoma: a bidirectional two-sample Mendelian randomized study

Nana Zhao^{1†}, Pengsen Guo^{2†}, Mei Tang^{1†}, Fan Yang³, Tongtong Zhang^{2,4} and Rui Mao^{5*}

¹Department of Operating Room, The Third People's Hospital of Chengdu, Chengdu, China, ²The Center of Gastrointestinal and Minimally Invasive Surgery, The Third People's Hospital of Chengdu, Chengdu, China, ³Emergency Department, Peking University Third Hospital, Peking University School of Medicine, Beijing, China, ⁴Medical Research Center, The Third People's Hospital of Chengdu, The Affiliated Hospital of Southwest Jiaotong University, The Second Chengdu Hospital Affiliated to Chongqing Medical University, Chengdu, Sichuan, China, ⁵Department of Dermatology, Xiangya Hospital, Central South University, Changsha, China

Background and objective: Existing cross-sectional and retrospective studies were unable to establish a causal relationship between psoriasis and cutaneous melanoma (CM). We sought to evaluate the causal role between psoriasis and CM.

Methods: We performed a bidirectional two-sample Mendelian randomization analysis using summary statistics from genome-wide association studies of psoriasis and CM among individuals of predominantly European ancestry. Mendelian randomization–Egger regression, inverse variance weighting, Mendelian Randomization Pleiotropy RESidual Sum and Outlier, weighted mode, and weighted median were used to examine the causal effect between psoriasis and CM.

Results: Genetically predicted psoriasis was a significant risk factor for CM (odds ratio, 1.69; 95% confidence interval, 1.15–2.48; P = 0.025). In contrast, no association was observed between genetically predicted CM and psoriasis.

Conclusion: Our findings corroborated the existence of genetically predicted psoriasis increases risk of CM. Enhanced early screening of cutaneous melanoma in patients with psoriasis may improve clinical burden. However, we did not find evidence for a causal link from CM to psoriasis, so further studies are required to elucidate the effect of CM activity on psoriasis.

KEYWORDS

bidirectional two-sample mendelian randomization, psoriasis, cutaneous melanoma, genome-wide association study, SNP

Introduction

At present, there is controversy regarding whether the risk of CM in patients with psoriasis is higher than that in the general population (1–5). Although the chronic inflammatory state in psoriasis patients may induce carcinogenic effects, psoriasis treatments like systemic therapy, ultraviolet (UV) radiation, and tumor necrosis factor (TNF)- α therapy (6–10) may also put psoriasis patients at greater risk for CM. In addition, studies on psoriasis risk in patients with CM are few in number and inconsistent (11–13).

Traditional observational epidemiological studies face many challenges in discovering disease etiology and inferring causality, such as reverse causal associations, potential confounding factors, minor exposure factors, and multiple tests. In observational and retrospective cohort studies, it is difficult to explore a causal relationship between psoriasis and CM with the existence of treatment or other interfering factors beyond our control. Prospective randomized controlled trials or other research methods that can rule out these interfering factors are urgently needed to establish a causal relationship between psoriasis and CM.

In recent years, with the continuous development of statistical methods, large-sample genome-wide association study (GWAS) data, epigenetics, and various "omics" techniques, the Mendelian randomization (MR) study design has been increasingly widely used in the discussion of the causal association between complex exposure factors and disease outcomes (14-16). The research design of MR follows the Mendelian inheritance law of "parental alleles are randomly assigned to offspring." If the genotype determines the phenotype, the genotype is associated with the disease through the phenotype, so the association between psoriasis and CM can be inferred using the genotype as an instrumental variable. Notably, MR is less susceptible to confounding factors because germline genetic variations are randomly allocated during meiosis and thus can capture exposure without being influenced by reverse causality. As a variant of the MR approach, bidirectional MR can be applied to ascertain the causal direction between two associated phenotypes. In this study, we performed a bidirectional two-sample MR analysis using summary statistics of large GWASs from the FinnGen Consortium to assess the causal association between psoriasis and CM.

Materials and methods

Data sources

The GWAS summary statistics for psoriasis were derived from data published by FinnGen Consortium R9. This study used the "psoriasis" phenotype. The GWAS of psoriasis included 373,338 Finnish adult subjects, including 9,267 cases and 364,071 controls, which excluded subjects with other cancers. Age, sex, top 10 major components, and genotyping batches were corrected during analysis. As a genetic instrumental variable, the GWAS summary data of CM came from the two-stage genome-wide meta-analysis conducted by Matthew et al. (17) that included 13 GWAS datasets from Europe, Australia, the United Kingdom, Athens, and the United States encompassing a total of 15,990 melanoma cases and 26,409 controls.

Instrumental variable

We use the following five criteria to select instrumental variables (IV). First, the SNP-phenotype association level must reach a genome-wide significance threshold (P< 5*10⁻⁸). Second, the linkage disequilibrium between all SNPs is based on the European 1000 Genomes Project reference panel. Third, SNPs with a secondary allele frequency (MAF) of ≤ 0.01 were removed. Fourth, among the SNPs where $R^2 < 0.001$ (clumping window size = 10,000 kb), only those with the lowest P value were retained. Finally, when palindromic SNPs existed, we used the allele frequency information to infer the forward-strand alleles. We used a curated genotype-phenotype database (PhenoScanner) to search for associations between variants used to detect each instrumental variable and other traits that may represent pleiotropic pathways; specific examples are traits associated with hypertension and telomere length, which are recognized risk factors for cutaneous melanoma (18, 19). Variants associated with these and other traits were excluded from sensitivity analysis (using a software default threshold of P< $5 \times 10 - 8$).

Statistical analysis

In the present study, we evaluated the causal relationship between psoriasis and CM using various methods, including inverse variance weighting (IVW), MR-Egger regression, MR Pleiotropy RESidual Sum and Outlier (MR-PRESSO), weighted mode, and weighted median. IVW, also known as the inverse variance weighting method, is characterized by ignoring the existence of an intercept term in regression and using the reciprocal of outcome variance (se²) as the weight for fitting. In the absence of heterogeneity and horizontal pleiotropy, IVW estimates are the most reliable (20). MR-Egger regression is conducted under the assumption of InSIDE (instrument strength is independent of the direct effect), which enables us to assess the presence of poly-efficacity with intercept terms. The intercept term being 0 implies that horizontal pleiotropy is absent, and the result of MR-Egger regression is consistent with IVW (21). In addition, MR-Egger regression is also used to judge whether there is horizontal pleiotropy or not. When more than 50% of the instrumental variables are invalid, the weighted median method can estimate causality more correctly. The weighted model estimation method can better detect the causal effect, has less deviation, and a lower class I error rate than the MR-Egger regression when the InSIDE hypothesis is not met (22). In addition, significant outliers were detected using MR-PRESSO

(23) tests and MR–Egger regression, and horizontal pleiotropic effects were corrected by removing outliers. We used the global test to check if horizontal pleiotropy among all instruments existed (23). Moreover, we applied Cochran's Q-test statistics to further examine the heterogeneity among all SNPs. We identify potentially heterogeneous SNPs by conducting a "leave-one-out" analysis by excluding each instrument SNP in turn. Finally, we also performed reverse MR analysis of psoriasis and CM. We adopted methods and settings that were consistent with the forward MR.

The F statistic was calculated by the formula $F = (\frac{R^2}{1-R^2} \times \frac{(n-k-1)}{k})$ (where R² is the proportion of the variance of the trait explained by the SNP, k is the number of IVs, and N is the sample size of the GWAS of the SNP with the trait), which is used to quantify the strength of the instrument, and a value > 10 is considered sufficient. We used the "TwoSampleMR" R package to calculate the coefficient of determination (R²) of exposure to genetic variants. The R² value was estimated using the formula $R^2 = 2 \times EAF \times (1 - EAF) \times (\beta)^2$ (where EAF is the effect allele frequency [EAF] of SNP, SD is the standard deviation, and β is the estimated effect size of SNP on the trait). Based on the online MR-power calculation tool (https://sb452.shinyapps.io/power/) (24), we calculated the power of the MR estimates.

All analyses were performed in R (version 4.2.2; R Foundation for Statistical Computing, Vienna, Austria), and MR analysis was conducted based on the TwosampleMR (25) (version 0.5.6) and MR-PRESSO (version 1.0) R packages.

Results

Psoriasis to CM

As shown in Table 1, we identified a total of 31 independent SNPs as instrumental variables in psoriasis, explaining 29.22% of the total variation. The F-statistic range for SNPs was 31.9–1018.0, indicating that SNPs explain the potency strength of exposure effectively. Figures 1A, B show the effect of each SNP locus on CM. Furthermore, the results of IVW estimation suggest that psoriasis is an independent risk factor for CM (odds ratio [OR], 1.69; 95% confidence interval [CI], 1.15–2.48; P = 0.025; Figures 1A, C). In addition, the results of the other three estimate methods, including weighted mode (OR, 2.42; 95% CI, 1.16–5.06; P = 0.026), MR–Egger regression (OR, 1.71; 95% CI, 1.06–2.76; P = 0.038), and weighted median (OR, 1.77; 95% CI, 1.09–2.87; P = 0.043), further validated the results of IVW estimation.

Cochran's IVW and MR–Egger Q test showed no significant heterogeneity among these IVs (Table S1). Moreover, no significant change in the estimated causal effect was observed when we excluded each SNP individually during the "leave-one-out" analysis (Figure 1D). Therefore, the estimated effect cannot be explained by any single SNP. In addition, according to the MR– Egger regression intercept analysis (Supplemental Table 2) and MR-PRESSO global test (Supplemental Table 3), there was no significant horizontal pleiotropy. After setting the type I error rate as 0.05 and the outcome variable as binary, we obtained an MR-power value of 100%. The above results prove that the results of our MR analysis are reliable and have high efficiency.

CM to psoriasis

We obtained seven SNPs (P< 5 * 10^{-8} , R²< 0.001) significantly associated with CM from a meta-analysis of large GWASs, which explain 3% of the total variation. Summaries and details of each SNP are provided in Supplementary Table 4. The F-statistic range for SNPs was 30.1–49.5. The individual SNP effects and the combined effects of each method are shown in Figures 2A, B. The results of the four MR estimate methods all suggested that no association was observed between genetically determined CM and psoriasis (Figure 2C, P > 0.05). No notable heterogeneity was detected by Cochran's Q statistics (Supplementary Table 5, P > 0.05). The MR– Egger intercept test (Supplementary Table 6) and MR-PRESSO global test (Supplementary Table 7) suggested no evidence of pleiotropy. Moreover, no significant change in the estimated causal effect was observed when we excluded each SNP individually during the "leaveone-out" analysis (Figure 2D).

Discussion

To our knowledge, this study was the first attempt to explore the causal relationship between psoriasis and CM using the summary statistics of FinnGen Alliance R9 and the large GWAS metaanalysis conducted by Matthew et al. (17). The results of our bidirectional two-sample MR study suggested that a genetic predisposition to psoriasis was associated with a 69% increased risk of CM (hazard ratio, 1.69), while a genetic predisposition to CM was not. In addition, we used a variety of methods to validate the reliability and efficacy of our results.

Most current studies on the effect of psoriasis on CM are observational cross-sectional studies or retrospective cohort studies, with inconsistent findings. For instance, Egeberg et al.'s (1) retrospective cohort study based on a Danish population showed a significantly increased risk of CM in patients with mild psoriasis. Reddy et al.'s (3) retrospective cohort study also confirmed this finding (hazard ratio, 1.53). Nonetheless, a systematic review and meta-analysis of 365 studies found no significant association between psoriasis and CM risk (5). Important confounding factors like phototherapy and systemic antipsoriasis therapy always influence the overall conclusions of the study. However, our study-independent of confounding factors and reverse causality-showed that genetic liability to psoriasis is a risk factor for CM. The chronic inflammatory state of psoriasis may impair immune surveillance and lead to possible tumor development, resulting in a higher risk of CM in psoriasis patients. Many of the transcription factors and cytokines-which include interleukin (IL)-6, TNF-a, and signal transduction and transcription activator 3 (26),-that are thought to play a role in psoriasis may also contribute to tumor development. Signal transduction and transcriptional activator 3 signaling is significantly elevated in keratinocytes in psoriatic lesions (27), and TNF- α has been shown to be critical for skin canceration (28). An intracellular transmitter of inflammatory signals, nuclear factor kappa-light-chain-enhancer of activated B-cells (NF-kB), is activated by pro-inflammatory cytokines such as TNF and IL-1 (29).

NF- κB has pleiotropic properties in the surrounding environment, but three are its basic roles. First, it is involved in

TABLE 1. Instrumental variables used in MR analysis of the association between psoriasis and CM.

exposure	SNP	Effect allele	Other allele	MAF	F	R ² of exposure	Exposure (psoriasis)			Outcome (CM)		
							Beta	SE	P-value	Beta	SE	P-value
psoriasis	rs10193310	А	G	0.248	55.501	0.006	0.127	0.017	9.341E-14	0.054	0.166	0.747
psoriasis	rs12206050	Т	А	0.320	41.025	0.004	0.101	0.016	1.502E-10	-0.006	0.154	0.970
psoriasis	rs1250566	А	G	0.408	33.638	0.004	-0.089	0.015	6.639E-09	-0.111	0.146	0.448
psoriasis	rs13153019	С	Т	0.273	38.114	0.004	0.102	0.016	6.673E-10	-0.033	0.162	0.836
psoriasis	rs139298380	А	G	0.028	56.899	0.005	0.313	0.042	4.589E-14	0.365	0.442	0.409
psoriasis	rs144651842	А	G	0.078	40.923	0.004	0.172	0.027	1.584E-10	-0.070	0.271	0.796
psoriasis	rs1611704	Т	С	0.345	217.349	0.023	0.225	0.015	3.426E-49	-0.120	0.153	0.433
psoriasis	rs16903065	А	С	0.116	34.184	0.004	-0.141	0.024	5.014E-09	-0.317	0.226	0.161
psoriasis	rs181316459	С	G	0.047	102.654	0.010	0.330	0.033	3.992E-24	-0.629	0.354	0.076
psoriasis	rs2021511	Т	С	0.265	40.027	0.005	-0.109	0.017	2.505E-10	-0.219	0.162	0.176
psoriasis	rs2769979	С	Т	0.605	39.928	0.004	-0.096	0.015	2.635E-10	0.025	0.146	0.862
psoriasis	rs28732090	G	С	0.058	1018.007	0.070	0.798	0.025	1.000E- 200	0.735	0.308	0.017
psoriasis	rs28998802	А	G	0.184	73.995	0.008	0.161	0.019	7.834E-18	-0.168	0.185	0.362
psoriasis	rs33980500	Т	С	0.072	78.437	0.008	0.238	0.027	8.260E-19	-0.232	0.278	0.404
psoriasis	rs34536443	С	G	0.030	35.369	0.005	-0.282	0.047	2.728E-09	-0.692	0.418	0.098
psoriasis	rs34955377	А	G	0.116	83.401	0.008	0.202	0.022	6.697E-20	0.159	0.223	0.475
psoriasis	rs4400255	А	Т	0.099	58.398	0.006	0.181	0.024	2.140E-14	-0.016	0.241	0.947
psoriasis	rs60600003	G	Т	0.102	39.937	0.004	0.150	0.024	2.623E-10	-0.167	0.237	0.480
psoriasis	rs653169	G	А	0.570	32.578	0.004	-0.086	0.015	1.145E-08	0.025	0.146	0.862
psoriasis	rs6556423	Т	С	0.642	139.728	0.015	-0.180	0.015	3.054E-32	-0.152	0.150	0.310
psoriasis	rs674451	С	Т	0.344	68.473	0.007	0.128	0.015	1.286E-16	0.158	0.150	0.294
psoriasis	rs7310615	G	С	0.586	32.924	0.004	-0.087	0.015	9.584E-09	-0.119	0.145	0.414
psoriasis	rs74817271	А	G	0.074	70.719	0.007	0.225	0.027	4.119E-17	-0.128	0.272	0.636
psoriasis	rs7542079	С	Т	0.562	39.763	0.004	0.095	0.015	2.867E-10	-0.084	0.144	0.563
psoriasis	rs76741620	G	А	0.058	32.732	0.003	0.174	0.030	1.058E-08	-0.120	0.308	0.696
psoriasis	rs78456138	Т	С	0.020	38.134	0.005	-0.365	0.059	6.606E-10	0.830	0.513	0.106
psoriasis	rs80174646	Т	G	0.046	35.310	0.005	-0.227	0.038	2.812E-09	0.174	0.350	0.619
psoriasis	rs847	С	Т	0.635	64.721	0.007	0.127	0.016	8.628E-16	0.158	0.150	0.292
psoriasis	rs8904	А	G	0.388	66.919	0.008	-0.126	0.015	2.829E-16	0.002	0.147	0.992
psoriasis	rs9273060	Т	А	0.249	174.059	0.018	0.219	0.017	9.618E-40	-0.110	0.168	0.512
psoriasis	rs9346778	Т	С	0.164	31.988	0.004	-0.117	0.021	1.551E-08	-0.030	0.192	0.875

pro-inflammatory responses. In human melanoma, many NF-κBregulated chemokines are expressed at high levels: CXCL8 or IL8, CXCL1 (melanoma growth-stimulating activity), CCL5 (regulatory activation, normal T expression and secretion), and CCL2 (30–32). These NF-κB-regulated chemokines are thought to enhance melanoma progression through autocrine and paracrine loops upon transcriptional activation. In fact, overexpression of CXCL8 leads to metastatic tumor growth in normal primary melanoma cells and is associated with the transition from RGP to VGP in melanoma (33, 34). Second, NF-κB is also a major anti-apoptotic factor that induces the transcription of a variety of anti-apoptotic proteins, such as Bcl-XL, tumor necrosis factor receptor-related factors 1 and 2 (TRAF1 and TRAF2, respectively), and inhibitor-apoptotic (IAP) proteins 1 and 2 (c-IAP1 and c-IAP2, respectively). In advanced melanoma, upregulated NF-κB also enhances above anti-apoptosis molecules (35, 36). Finally, NF-κB promotes cell growth: activated NF-κB enhances the expression of cyclin D1, an important regulator of cell cycle progression. NF-κB



FIGURE 1

MR analysis of the causal relationship between genetically predicted psoriasis and CM. (A) Scatterplots for the causal association between genetically predicted psoriasis and CM. The slope of a straight line indicates the magnitude of causality. Black dots represent genetic instruments included in the main MR analysis. (B) Forest map visualization of the causal impact of each SNP on CM risk. (C) Forest mapping used four methods to visualize the causal effects of psoriasis on CM risk. (D) "Leave-one-out" plots for the causal association between genetically predicted psoriasis and CM.



FIGURE 2

MR analysis of the causal relationship between genetically predicted CM and psoriasis. (A) Scatterplots for the causal association between genetically predicted CM and psoriasis. The slope of a straight line indicates the magnitude of causality. Black dots represent genetic instruments included in the main MR analysis. (B) Forest map visualization of the causal impact of each SNP on psoriasis risk. (C) Forest mapping uses four methods to visualize the causal effects of CM on psoriasis risk. (D) "Leave-one-out" plots for the causal association between genetically predicted CM and psoriasis.

upregulated Myc and cycle regulator proteins, cyclin D1, and cyclin-dependent kinase 2, which further promoted melanoma growth (37–40). Activation of NF-kB enhances tumor cell survival and proliferation and helps to transform tumor-associated macrophages into tumor-promoting phenotypes (41).

In addition, studies on psoriasis risk in patients with CM are few and inconsistent. The retrospective cohort study of Sam et al. (12) suggested that CM patients were at increased risk for psoriasis. However, the prospective pilot study by Matteo et al.11 suggested that CM patients have a reduced risk of psoriasis. Our results were similar to those of the observational cross-sectional study by Elwood et al. (13), where genetic liability to CM was not associated with psoriasis. More clinical and basic studies are needed to explore the risk and mechanism of psoriasis in CM patients.

Strengths and limitations

This study had several advantages. To determine the causal relationship between psoriasis and CM without the interference of confounding factors, we performed bidirectional MR analysis. Bidirectional Mendelian randomization analysis is an extension of standard Mendelian randomization analysis. Compared with the latter, the former can investigate whether there is a two-way causal relationship between exposure variables and outcome variables. In addition, genetic variation in CM was derived from the largest available GWAS meta-analysis, ensuring the strength of instruments in MR analyses. We detected and excluded horizontal poly-tropism using MR-PRESSO and MR–Egger regression intercept tests. Finally, there was no overlap between the GWAS summary data for psoriasis and the meta-analysis GWAS summary data for CM, which led to low heterogeneity, low bias, and high accuracy in our study.

We also acknowledge some limitations. First, most of the data used in the analysis mainly came from individuals with European ancestry, which restricts the applicability of our findings to other ethnic groups. Second, due to the lack of detailed clinical data, we were unable to conduct a subgroup analysis, for example, to distinguish between psoriasis vulgaris, arthritic psoriasis, pustular psoriasis, and erythrodermic psoriasis; explore non-linear relationships; or to sort by the severity of psoriasis. A populationbased study in Denmark suggested a modestly increased risk of melanoma in patients with mild psoriasis and no increased risk in patients with severe psoriasis and psoriatic arthritis (1). Due to the lack of GWAS data on psoriasis severity classification or strictly defined clinical classification to date, we were unable to use MR methods to study the effect of different severities or different clinical subtypes of psoriasis on CM. Notably, each method used in our analysis had advantages and disadvantages. However, the use of many methods with different assumptions may lead to inconsistent or conflicting outcomes and obscure the conclusions.

Conclusion

Our results confirmed genetically predicted psoriasis increases risk of CM. Enhanced early screening of cutaneous melanoma in patients with psoriasis may improve clinical burden. In addition, although the reverse MR estimation does not support the causal relationship from CM to psoriasis, it cannot be ruled out that CM has an effect on the incidence of psoriasis; further research is needed to confirm this.

Data availability statement

The original contributions presented in the study are included in the article/Supplementary Material. Further inquiries can be directed to the corresponding author.

Ethics statement

We did not need further ethical approval for this study because we used data from published studies that had the appropriate ethics committees' approval.

Author contributions

NZ: Conceptualization, Methodology, Software, Investigation, Visualization, Writing an original draft. PG: Methodology, Investigation. MT: Investigation, experiment. FY: Methodology, Investigation. TZ: Data Provision, Methodology. RM: Conceptualization, Writing - original draft, Funding acquisition, Supervision. All authors contributed to the article and approved the submitted version.

Acknowledgments

We thank all the researchers for providing their GWAS data. We also thank the participants and researchers of the FinnGen study.

Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

Publisher's note

All claims expressed in this article are solely those of the authors and do not necessarily represent those of their affiliated organizations, or those of the publisher, the editors and the reviewers. Any product that may be evaluated in this article, or claim that may be made by its manufacturer, is not guaranteed or endorsed by the publisher.

Supplementary material

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/ fimmu.2023.1201167/full#supplementary-material

References

1. Egeberg A, Thyssen JP, Gislason GH, Skov L. Skin cancer in patients with psoriasis. J Eur Acad Dermatol Venereol (2016) 30:1349–53. doi: 10.1111/jdv.13619

2. Margolis D, Bilker W, Hennessy S, Vittorio C, Santanna J, Strom BL. The risk of malignancy associated with psoriasis. *Arch Dermatol* (2001) 137:778–83.

3. Reddy SP, Martires K, Wu JJ. The risk of melanoma and hematologic cancers in patients with psoriasis. *J Am Acad Dermatol* (2017) 76:639–47.e2. doi: 10.1016/j.jaad.2016.09.047

4. Trafford AM, Parisi R, Kontopantelis E, Griffiths CEM, Ashcroft DM. Association of psoriasis with the risk of developing or dying of cancer: a systematic review and meta-analysis. *JAMA Dermatol* (2019) 155:1390-403. doi: 10.1001/jamadermatol.2019.3056

5. Vaengebjerg S, Skov L, Egeberg A, Loft ND. Prevalence, incidence, and risk of cancer in patients with psoriasis and psoriatic arthritis: a systematic review and metaanalysis. *JAMA Dermatol* (2020) 156:421–9. doi: 10.1001/jamadermatol.2020.0024

6. Koo J. Systemic sequential therapy of psoriasis: a new paradigm for improved therapeutic results. J Am Acad Dermatol (1999) 41:S25-28. doi: 10.1016/s0190-9622 (99)70363-4

7. Rs S, Kt N, Lh V. Malignant melanoma in patients treated for psoriasis with methoxsalen (psoralen) and ultraviolet a radiation (PUVA). the PUVA follow-up study. *New Engl J Med* (1997) 336. doi: 10.1056/NEJM199704103361501

8. Yj C, Cy W, Tj C, Jl S, Sy C, Cb W, et al. The risk of cancer in patients with psoriasis: a population-based cohort study in Taiwan. *J Am Acad Dermatol* (2011) 65:84–91. doi: 10.1016/j.jaad.2010.04.046

9. Stern RSPUVA Follow up Study. The risk of melanoma in association with long-term exposure to PUVA. J Am Acad Dermatol (2001) 44:755–61. doi: 10.1067/mjd.2001.114576

10. Rv P, Ln C ML, Jm W. Treatments for psoriasis and the risk of malignancy. J Am Acad Dermatol (2009) 60. doi: 10.1016/j.jaad.2008.12.031

11. Megna M, Napolitano M, Balato N, Scalvenzi M, Ayala F, DI Costanzo L, et al. Psoriasis in melanoma patients: a prospective pilot study. *G Ital Dermatol Venereol* (2017) 152:109–16. doi: 10.23736/S0392-0488.16.05270-6

12. Polesie S, Gillstedt M. Psoriasis and risk of cutaneous melanoma: a retrospective, comparative, registry-based cohort study. *J Am Acad Dermatol* (2022) 86:215–7. doi: 10.1016/j.jaad.2021.01.084

13. Elwood JM, Gallagher RP, Stapleton PJ. No association between malignant melanoma and acne or psoriasis: results from the Western Canada melanoma study. *Br J Dermatol* (1986) 115:573–6. doi: 10.1111/j.1365-2133.1986.tb05767.x

14. Depression and inflammatory bowel disease: a bidirectional two-sample mendelian randomization study. Available at: https://pubmed.ncbi.nlm.nih.gov/34739073/ (Accessed February 18, 2023).

15. Dusingize JC, Olsen CM, An J, Pandeya N, Law MH, Thompson BS, et al. Body mass index and height and risk of cutaneous melanoma: mendelian randomization analyses. *Int J Epidemiol* (2020) 49:1236–45. doi: 10.1093/ije/dyaa009

16. Budu-Aggrey A, Brumpton B, Tyrrell J, Watkins S, Modalsli EH, Celis-Morales C, et al. Evidence of a causal relationship between body mass index and psoriasis: a mendelian randomization study. *PloS Med* (2019) 16:e1002739. doi: 10.1371/journal.pmed.1002739

17. Law MH, Bishop DT, Lee JE, Brossard M, Martin NG, Moses EK, et al. Genomewide meta-analysis identifies five new susceptibility loci for cutaneous malignant melanoma. *Nat Genet* (2015) 47:987–95. doi: 10.1038/ng.3373

18. Telomeres Mendelian Randomization Collaboration, Haycock PC, Burgess S, Nounu A, Zheng J, Okoli GN, et al. Association between telomere length and risk of cancer and non-neoplastic diseases: a mendelian randomization study. *JAMA Oncol* (2017) 3:636–51. doi: 10.1001/jamaoncol.2016.5945

19. Chan II, Kwok MK, Schooling CM. Blood pressure and risk of cancer: a mendelian randomization study. *BMC Cancer* (2021) 21:1338. doi: 10.1186/s12885-021-09067-x

20. Burgess S, Dudbridge F, Thompson SG. Combining information on multiple instrumental variables in mendelian randomization: comparison of allele score and summarized data methods. *Stat Med* (2016) 35:1880–906. doi: 10.1002/sim.6835

21. Bowden J, Davey Smith G, Burgess S. Mendelian randomization with invalid instruments: effect estimation and bias detection through egger regression. *Int J Epidemiol* (2015) 44:512–25. doi: 10.1093/ije/dyv080

22. Hartwig FP, Davey Smith G, Bowden J. Robust inference in summary data mendelian randomization via the zero modal pleiotropy assumption. *Int J Epidemiol* (2017) 46:1985–98. doi: 10.1093/ije/dyx102

23. Verbanck M, Chen C-Y, Neale B, Do R. Detection of widespread horizontal pleiotropy in causal relationships inferred from mendelian randomization between complex traits and diseases. *Nat Genet* (2018) 50:693–8. doi: 10.1038/s41588-018-0099-7

24. Burgess S. Sample size and power calculations in mendelian randomization with a single instrumental variable and a binary outcome. *Int J Epidemiol* (2014) 43:922–9. doi: 10.1093/ije/dyu005

25. Hemani G, Tilling K, Davey Smith G. Orienting the causal relationship between imprecisely measured traits using GWAS summary data. *PLoS Genet* (2017) 13: e1007081. doi: 10.1371/journal.pgen.1007081

26. Mantovani A, Garlanda C, Allavena P. Molecular pathways and targets in cancer-related inflammation. *Ann Med* (2010) 42:161-70. doi: 10.3109/07853890903405753

27. Chan KS, Sano S, Kiguchi K, Anders J, Komazawa N, Takeda J, et al. Disruption of Stat3 reveals a critical role in both the initiation and the promotion stages of epithelial carcinogenesis. *J Clin Invest* (2004) 114:720–8. doi: 10.1172/JCI21032

28. Moore RJ, Owens DM, Stamp G, Arnott C, Burke F, East N, et al. Mice deficient in tumor necrosis factor-alpha are resistant to skin carcinogenesis. *Nat Med* (1999) 5:828–31. doi: 10.1038/10552

29. Beyaert R, Beaugerie L, Van Assche G, Brochez L, Renauld J-C, Viguier M, et al. Cancer risk in immune-mediated inflammatory diseases (IMID). *Mol Cancer* (2013) 12:98. doi: 10.1186/1476-4598-12-98

30. Singh RK, Gutman M, Radinsky R, Bucana CD, Fidler IJ. Expression of interleukin 8 correlates with the metastatic potential of human melanoma cells in nude mice. *Cancer Res* (1994) 54:3242–7.

31. Ghosh S, Karin M. Missing pieces in the NF-kappaB puzzle. *Cell* (2002) 109 Suppl:S81-96. doi: 10.1016/s0092-8674(02)00703-1

32. Karin M, Delhase M. The I kappa b kinase (IKK) and NF-kappa b: key elements of proinflammatory signalling. *Semin Immunol* (2000) 12:85–98. doi: 10.1006/smim.2000.0210

33. Huang S, DeGuzman A, Bucana CD, Fidler IJ. Nuclear factor-kappaB activity correlates with growth, angiogenesis, and metastasis of human melanoma cells in nude mice. *Clin Cancer Res* (2000) 6:2573–81.

34. Bobrovnikova-Marjon EV, Marjon PL, Barbash O, Vander Jagt DL, Abcouwer SF. Expression of angiogenic factors vascular endothelial growth factor and interleukin-8/CXCL8 is highly responsive to ambient glutamine availability: role of nuclear factorkappaB and activating protein-1. *Cancer Res* (2004) 64:4858–69. doi: 10.1158/0008-5472.CAN-04-0682

35. Kashani-Sabet M, Shaikh L, Miller JR, Nosrati M, Ferreira CMM, Debs RJ, et al. NF-kappa b in the vascular progression of melanoma. *J Clin Oncol* (2004) 22:617–23. doi: 10.1200/JCO.2004.06.047

36. Devergne O, Hatzivassiliou E, Izumi KM, Kaye KM, Kleijnen MF, Kieff E, et al. Association of TRAF1, TRAF2, and TRAF3 with an Epstein-Barr virus LMP1 domain important for b-lymphocyte transformation: role in NF-kappaB activation. *Mol Cell Biol* (1996) 16:7098–108. doi: 10.1128/MCB.16.12.7098

37. Senftleben U, Cao Y, Xiao G, Greten FR, Krähn G, Bonizzi G, et al. Activation by IKKalpha of a second, evolutionary conserved, NF-kappa b signaling pathway. *Science* (2001) 293:1495–9. doi: 10.1126/science.1062677

38. Karin M, Ben-Neriah Y. Phosphorylation meets ubiquitination: the control of NF-κB activity. *Annu Rev Immunol* (2000) 18:621-63. doi: 10.1146/ annurev.immunol.18.1.621

39. Guttridge DC, Albanese C, Reuther JY, Pestell RG, Baldwin AS. NF-kappaB controls cell growth and differentiation through transcriptional regulation of cyclin D1. *Mol Cell Biol* (1999) 19:5785–99. doi: 10.1128/MCB.19.8.5785

40. Hinz M, Krappmann D, Eichten A, Heder A, Scheidereit C, Strauss M. NFkappaB function in growth control: regulation of cyclin D1 expression and G0/G1to-S-phase transition. *Mol Cell Biol* (1999) 19:2690–8. doi: 10.1128/ MCB.19.4.2690

41. Ben-Neriah Y, Karin M. Inflammation meets cancer, with NF- κB as the matchmaker. Nat Immunol (2011) 12:715–23. doi: 10.1038/ni.2060